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RAIL-HIGHWAY CROSSING HAZARD PREDICTION RESEARCH RESULTS

Peter Mengert

U.S. DEPARTMENT OF TRANSPORTATION RESEARCH AND SPECIAL PROGRAMS ADMINISTRATION Transportation Systems Center Cambridge MA 02142



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into an absolute hazard inde	x are given. An	introductory d	iscussion is	provided on
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PREFACE

The author wishes to acknowledge the contribution of James Guarente of Kentron Hawaii Limited who did the programming and data handling for this project. Mr. Guarente also was responsible for many day-to-day decisions on the course of the analysis, and assisted in determining the direction of this study. Because of his direct involvement with the actual analysis and experimentation, Mr. Guarente has written one of the sections of this report entitled "Course of Experimentation."

The author is also indebted to B. George and R. Snow of the Federal Railroad Administration for their support and guidance, and to J. Hitz, E. Farr, and R. Hinckley of TSC for useful suggestions throughout the course of this work.

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SUMMARY

The Federal Railroad Administration (FRA), in accordance with the Federal Railroad Safety Act of 1970, is investigating problems of railroad crossing safety improvement. In pursuit of the related studies, the FRA sought the services of the Transportation Systems Center (TSC) in selecting, evaluating, and developing hazard indexes, formulas used to estimate from available quantified information the hazards, or relative hazards, of train/vehicle accidents at railroad crossings. The TSC, currently engaged in a Grade Crossing Funding Allocation Project which also requires state-ofthe-art hazard indexes of the highest selective and predictive capabilities, complied with the FRA request. Thus, it has provided this document report on a study of hazard indexes as evaluated and constructed on the basis of FRA data. The report distinguishes between, develops, and evaluates the following hazard indexes:

- <u>Relative</u> hazard indexes, for <u>ranking</u> crossings according to relative hazard.
- <u>Absolute</u> hazard indexes, for providing an estimate equal to, or at least proportional to, expected accident frequency at the individual crossings.

Comparisons of several previously developed hazard indexes are given. Of these, the New Hampshire and Peabody-Dimmick are widely used. Selected Coleman-Stewart formulas for three specific warning device classes (crossbucks, flashing lights, and automatic gates) have also been evaluated.

New hazard indexes with improved prediction capability have been developed, and are reported on. The performance of these new indexes is compared in detail with the previously proposed formulas. A number of techniques for constructing hazard indexes have been explored, and a particularly effective technique employing nonlinear logistic discriminant techniques was selected for the final models reported on. This method is described in detail, and is suggested as the tool to form the basis of further analysis or development.

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The major results of the study include:

- 1. <u>Techniques</u> and <u>methodology</u> for <u>producing</u>, <u>comparing</u>, and evaluating hazard indexes
- 2. <u>New hazard indexes</u> for three warning device classes (crossbucks, flashing lights, and gates)

Detailed comparisons of the <u>performance</u> of hazard indexes.
Out of these have come specific results:

a. Volume factors (average daily vehicle volume and average daily train volume) account for 90-95 percent of the predictive power obtainable from the factors studied, excluding accident history at present. (See below.)

b. The simple New Hampshire formula (relative hazard proportional to vehicular volume times train volume) is nearly as effective as other volume-only formulas for <u>relative</u> hazard. A procedure and formula are given for converting this to an absolute hazard index (proportional to expected accident frequency). The New Hampshire formula is useful for its combination of power and simplicity.

c. For some uses, and in certain respects described in the report, the TSC formulas exhibited greater selectivity of hazardous crossings (performance as a relative hazard index) than other formulas tested. This is evidenced, for example, by comparing the ten percent most hazardous crossings selected by the TSC formula with the ten percent most hazardous set selected by the New Hampshire formula (crossbuck case). The TSC ten percent set, as determined from the FRA data bases, contains three percent more of the <u>total</u> accidents than the New Hampshire ten percent set. This is statistically significant.

The TSC formulas developed and reported on here may be useful when an absolute hazard index or expected frequency of accidents is needed, as in the funding allocation work. For this purpose, both comprehensive and volume-only formulas are given. The performance of absolute hazard indexes is exhibited in special plots.

vi

d. Because of the large amount of experimentation done and the relatively small improvements in power factors (PF) obtainable, it would appear that the ultimately attainable power factors are not far from those obtained in this study. (The power factor at X% multiplied by X% gives the percent of accidents at the X% most hazardous crossings according to the given hazard index. Thus, if the 5% power factor is 4, then 5% of the crossings have 20% of the accidents.) The following power factors are quoted to illustrate the performance measures attained in a few instances taken as examples.

% Crossings	PF New Hampshire	PF TSC
1	6.80	7.86
2	6.17	5.90
6	4.76	4.92
10	3.83	4.10
20	2.88	3.01
40	2.03	2.03

Crossbucks Power Factors

Thus, this table says that according to the New Hampshire formula, the 10% most hazardous crossings had about 38% of the accidents, while the 10% most hazardous crossings according to the TSC formula had about 41% (the 3% difference was alluded to above) of the accidents, all figured on the FRA data bases (1975 accidents). More complete information and similar information for other warning device classes is presented in the body of this report.

Suggestions for further work are also given. In particular, in one of the appendixes (Appendix II) a proposed means is developed for incorporating accident history at an individual crossing along with crossing characteristics into a hazard index, and some preliminary results are given. The techniques of this appendix are currently being used in the FY79 effort at TSC to produce accident history dependent hazard indexes.

vii

Addendum to Summary

Since this report was written, some power factors have been run using the same best TSC and New Hampshire models discussed in this report, but with the 1976 accident data and the inventory data of May 1978 (about 9 months later than the date of the inventory data used in this report). In order not to delay publication, results are presented only in this summary for this report. In general, the results were quite comparable to the results reported on in detail for the earlier date presented in this report, the comparability holding for all three warning classes. Partial results corresponding to the table given just above are given in the following table:

Crossbucks	Power	Factors
(1976	accider	nts)

0 .	PF	ΡF
Crossings	New Hampshire	TSC
1	7.11	7.73
2	6.11	6.57
6	4.43	4.71
10	3.72	3.92
20	2.78	2.91
4 0	1.98	2.03

The observed results confirm the stability of relative performance for the TSC models when used on accident data for a different year from that of the data used in their construction. The power factors for the other two warning device classes similarly confirmed this stability.

CONTENTS

Section

-

÷

Page

	GLOS	SARY		xii
1.	INTR	ODUCTIC	N	1-1
	$1.1 \\ 1.2 \\ 1.3$	Hazard The FR Projec	A Data Bases t Objectives	1 - 1 1 - 2 1 - 3
			Grade Crossing Funding Allocation Project	1-5
	1.4	Notes	on Structure, Content, and Conventions	1-5
		1.4.1 1.4.2 1.4.3	Definitions and Terms Appendixes Notes on Suggested Order of Reading	1-5 1-6 1-6
2.	METH	ODOLOGY		2-1
	2.1	The Em	pirical Operating Characteristics (EOC).	2-1
		$2.1.1 \\ 2.1.2$	The Power Factor The EOC Described	2 - 2 2 - 3
	2.2	Hazard	Index Construction	2 - 7
		2.2.1 2.2.2	The Linear Regression Approach Iterated Weighted Logistic Discriminants	2 - 7 2 - 8
		2.2.3	Hazard Function Development and Vari- able Selection by Synthesis of Regression and EOC Techniques	n 2-11
			2.2.3.1 Linear Case 2.2.3.2 Logistic Case	2-11 2-16
	2.3	Expect Indexe	ed Accident Frequency (Absolute Hazard s)	2 - 19
		2.3.1	Expected Accident Frequency from Logistic Discriminants (TSC) Nonlinear Hazard Indexes	2-19
		2.3.2	Absolute Hazard Indexes Based on Power Factor Information	2-21
3.	COURS	SE OF EX	(PERIMENTATION	3-1
4.	RESUI	LTS		4-1
	4.1	EOCs an	nd Power Factors	4 - 1

CONTENTS (CONTINUED)

Section				Page
		$4.1.1 \\ 4.1.2$	Introductory Comment The Crossbucks Case	4 - 1 4 - 2
			Analytic Power Factor Curves	4 - 5
		4.1.3 4.1.4 4.1.5	Flashing Lights Case Automatic Gates Case An EOC over the Full Flashing Lights	4 - 6 4 - 8
		4.1.6	Data Base Evaluation of Other Previously Proposed Hazard Indexes	4-9 4-10
	4.2	Discus: Volume	sion of TSC Comprehensive Models and Models	4 - 1 2
		$\begin{array}{c} 4.2.1 \\ 4.2.2 \\ 4.2.3 \end{array}$	Crossbucks Flashing Lights Automatic Gates	4-12 4-14 4-15
	4.3	Expect Summary	ed Accident Frequency Plots y of Results	4 - 1 8 4 - 2 3
5.	SOME	PRACTIO	CAL CONSIDERATIONS	5-1
	5.1 5.2	Select: Use of	ing a Hazard Index the Hazard Index	5 - 1 5 - 3
			Relative or Absolute Hazard Index	5 - 3
APPEND	IX A.	THE LO HAZARI	DGISTIC DISCRIMINANT APPROACH TO D INDEX CONSTRUCTION	A-1
APPEND	IX B.	HAZARI REFERI	D INDEX FORMULAS - DEFINITIONS FOR	B-1
APPEND	IX C.	EOC AN	ND POWER FACTOR TABLES	C-1
APPEND	IX D.	REGRES	SSIONS	D-1
APPEND	IX E.	EXPECT EOC PI	TED ACCIDENT FREQUENCY PLOTS AND LOTS	E - 1
APPENDI	IX F.	DATA E	BASES	F-1
APPENDI	IX G.	NONLIN CONSTR	NEAR (LOGISTIC) VERSUS LINEAR RUCTION	G-1
APPENDI	[х н.	HAZARE	INDEXES BASED ON ACCIDENT HISTORY	H-1
REFERENCE	ES			R-1

LIST OF ILLUSTRATIONS

Figure		Page
2 - 1	HAZARD FUNCTION CONSTRUCTION LOGISTIC DISCRIMINANT & EOC APPROACH	2-12
	LIST OF TABLES	
Table		Page
2 - 1	FACTORS FOR ABSOLUTE EXPECTED ACCIDENT FREQUENCY	2-20
2 - 2	EMPIRICAL POWER FACTOR FORMULA FOR THE NEW HAMPSHIRE MODEL	2-24
4 - 1	ESTIMATES OF POWER FACTORS AND "ACHIEVABLE" POWER FACTORS (CROSSBUCKS)	4 - 3
4 - 2	POWER FACTORS (FLASHING LIGHTS)	4 - 7
4 - 3	ESTIMATED POWER FACTORS (GATES)	4 - 9

ABSOLUTE HAZARD INDEX -- A hazard index which is also proportional to expected number of accidents per year. (See HAZARD INDEX, PROBABILITY OF ACCIDENT, and EXPECTED FREQUENCY OF ACCI-DENTS.) (See Sections 1, 2.3, and 4.3.)

<u>ACCIDENT</u> -- "A public grade crossing accident/incident is any impact between railroad on-track equipment and an automobile, bus, truck, motorcycle, bicycle, farm vehicle, or pedestrian, regardless of whether it resulted in any casualties or damage." (See Reference 3.) The ratio of fatalities to accidents in the year 1975 was approximately 0.105. (See Section 1 and Appendix F.)

<u>COLEMAN-STEWART MODEL</u> -- One of several specific hazard indexes. (See Appendix B and Reference 2.)

EMPIRICAL OPERATING CHARACTERISTICS (EOC) -- A table giving power factors, cumulative accidents at various percentages of hazardous crossings, etc. Also a graph of percent accidents versus percent crossings. (See also POWER FACTOR.) (See Sections 2 and 4 and Appendix C.)

EOC -- See EMPIRICAL OPERATING CHARACTERISTIC.

EXPECTED FREQUENCY OF ACCIDENTS (or expected number of accidents) -- For a given value of the hazard index, the expected number of accidents to occur at a given crossing in a given year. Related approximately to probability of accident, p, by f = p/1-p. (See PROBABILITY OF ACCIDENT.) (See Sections 2.3 and 4.3.)

HAZARD INDEX (or hazard function, hazard model). A formula relating relative hazard of accident to quantifiable crossing characteristics. The higher the hazard index the higher the probability of accidents (if the hazard index holds good). If the hazard index is also proportional to probability of accident, then it is an absolute hazard index. (See also ABSOLUTE HAZARD INDEX.) (See Section 1.)

ITERATED WEIGHTED LOGISTIC REGRESSION -- Each of the component parts of this expression has a common meaning in statistical analysis. They are combined in this project to produce a technique especially adapted for producing hazard models. (See Section 2.2 and Appendix A.)

<u>NEW HAMPSHIRE MODEL</u> -- A very simple hazard index (often given other names) which states that for a given warning device class the (relative) hazard increases with the product of the average vehicular volume and the average train volume. This gives a good relative hazard index, but not a good absolute hazard index (except by modification). (See Section 4.1 and Appendix B.)

<u>NON-VOLUME VARIABLES</u> -- All crossing characteristics not derived only from volume variables, e.g., number of tracks, train speed, number of night trains, etc. (See also VOLUME VARIABLES.) (See Sections 1 and 3.)

<u>PEABODY-DIMMICK MODEL</u> -- A hazard index developed many years ago depending only on vehicular volume and train volume for a crossing of a given warning device class. (See Section 4.1, Appendix B, and Reference 11.)

<u>POWER FACTOR</u> -- The fraction of accidents occurring at a given fraction of the most hazardous crossings. If the 5 percent factor is 4, then the 5 percent most hazardous (according to a given hazard index) crossings have 20 percent of the accidents. (See Section 2 and Appendix C.)

<u>POWER FACTOR FUNCTION</u> -- An analytic representation of the EOC by means of a function which fits the observed functional relation of power factor to fraction of crosisngs. (It is used for a given warning device class and hazard index.) It is usually of the form log $\rho = a(\log \lambda)^b$, where ρ is the λ x 100% power factor, and λ is a given fraction of the crossings. (See also POWER FACTOR.) (See Sections 2.3 and 4.1 and Appendix C.)

<u>PROBABILITY OF ACCIDENT</u> -- For a given value of a given hazard index the probability, p, of a crossing (with this value for the hazard index) experiencing an accident in a given year. (See EXPECTED FREQUENCY OF ACCIDENTS.)

xiii

<u>TSC COMPREHENSIVE MODFL</u> -- A hazard index, for each warning device class constructed in this project, which uses the TSC volume model as a base, and which includes non-volume variables as well. The best comprehensive models for each warning device class are given in Appendix B. The best comprehensive models are of logistic construction, and easily yield an absolute hazard index.

<u>TSC VOLUME MODEL</u> -- A hazard index for each warning device class, constructed in this project, using only <u>volume variables</u>. Best volume models are of logistic construction, and yield absolute hazard as well. (See Sections 4.1 and 4.4 and Appendix B, etc.)

<u>VOLUME REGRESSION</u> -- Any hazard index optimized as a function of volume variables only, whether constructed by linear regression, logistic discriminant analysis, etc. It has been observed that apparently 90-95 percent of the predictive power of any hazard index is accounted for by its volume dependence. (See also VOLUME VARIABLES.) (See Sections 2 and 3.)

<u>VOLUME VARIABLES</u> -- Those variables and functions, entering any hazard index, which depend only on C, the average total daily vehicular volume, and T, the average total daily train volume. Only total traffic of either kind is included. By this definition, any breakdown such as day/night, through/switch, car/truck involves non-volume variables. (See also NON-VOLUME VARIABLES.)

<u>WARNING DEVICE CLASS</u> -- A class of crossings determined by the warning devices for highway vehicles. The most effective device present at each crossing in the class gives its name to the class. Thus, each crossing in the warning device class designated "crossbucks" has no active warning equipment, but has the standard crossbucks to warn highway traffic. The separate and disjoint warning device classes "crossbucks," "flashing lights," and "automatic gates" together encompass about 90 percent of all public grade crossings accounting for over 90 percent of all grade crossing accidents, and are the focus of this report. All the analyses and constructions in this report are disaggregated by

xiv

warning device class. The term "warning device class" refers to essentially the same classes of crossings as does the term "protection class" used in some previous literature. (See Table F-2, LOC 47.)

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1. INTRODUCTION

1.1 HAZARD INDEXES



There has been long and continued interest in objective formulas for comparing individual railroad grade crossings with respect to accident hazard. These formulas are usually relatively simple functions of easily quantifiable characteristics of the grade crossing, and are called "hazard indexes." (See References 1, 2, 5, 8, and 12.)

An example of a hazard index in common use is the so-called New Hampshire formula:

 $H = K_{D} \cdot CT$

where C is the average daily vehicular traffic volume at the crossing, and T is the average number of trains per day. K_p is constant, differing for each warning device class. As will be shown in this report, the New Hampshire formula can be of value for comparing crossings of the same warning device class with respect to relative hazard. The means of comparing formulas in their ability to predict hazard will be shown and these methods will be used with the comprehensive data (all to be described presently) to make relative assessments of the <u>hazard ranking</u> efficiency of various formulas. (The New Hampshire formula is not the best in this regard, but is a good example because it is surprisingly efficient, given its simplicity.

A hazard index, as referred to above, gives a <u>relative</u> indication of hazard. An <u>absolute</u> indication of hazard is a quantity which is proportional to expected frequency of accidents per year (at a crossing with the characteristics represented in the formula). The New Hampshire formula, as stated above, is good as a relative hazard index only. Other formulas, which are good as absolute hazard indexes, will be covered, as will the method of obtaining an absolute hazard index from a relative hazard index.

Relative hazard indexes are used for <u>ranking</u> and <u>comparing</u> grade crossings as to their hazard level. One practical use would be in preliminary selection of a group of crossings (out of some

population) for closer examination in order to select from this group a smaller group for improvement, i.e., upgrading of warning device class.

For certain more analytic applications, especially those carried out on a large scale, it may not be enough to have a <u>rela-</u> <u>tive</u> hazard index; rather, an absolute hazard index is needed. For example, in calculating benefit/cost ratios on a per crossing basis, the benefits may be based on the expected (predicted) accident frequency at the particular crossing, and this is given only by an absolute hazard index.

Construction of a superior relative hazard index is the more difficult part. "Shaping" the hazard function to an absolute hazard index is easier and more straightforward.

Mathematically, the term "relative hazard index" can be defined in terms of "absolute hazard index" (even though the calculation may go the other way). An <u>absolute</u> hazard index is any quantity directly proportional to expected frequency of accident. A <u>relative</u> hazard index is any monotonic (always increasing) function of an absolute hazard index. From almost every intuitive, computational, and practical point of view, however, the concept of relative hazard index may be thought of as prior. The relative hazard index indicates which crossings are more hazardous, but not by how much; the absolute hazard index answers the latter question.

1.2 THE FRA DATA BASES

The FRA has compiled a comprehensive data base containing data on a large number of qualitative and quantitative characteristics of all public roadway-railway grade crossings and all private grade crossings in the United States. This crossing inventory is briefly described in Appendix F (see also Section 3), and is also the subject of an earlier report (Reference 4). It contains, in quantified fixed-format records, information on a great many factors, of which total average daily vehicle volume, total average daily train volume, and maximum warning device class are just three (derived) quantities. There are other quantities related to

vehicle volume, train volume, and crossing warning equipment included, and many quantities not related to these, such as estimates of typical train speeds, functional class of road, type of development of the area, etc. In the work reported on here, only those records which refer to <u>public</u> grade crossings are used (219,162 in number). This data base is referred to as "the crossing inventory."

In addition, the FRA has been keeping a complete file on the grade crossing accidents which occur at these crossings; this data base is described briefly in Appendix F (see also Section 3), and also reported on in Reference 3. For the year 1975, a total of 8,028 accidents are represented. This data base (1975 only) will be called the "accident file." The accident file (since 1975 inclusive) is keyed to the crossing inventory by a crossing identification number, uniquely associated with each crossing. This number is included in the record associated with that crossing in the crossing inventory, and included in the record(s) for all accidents which occurred at that crossing. (A certain number of accidents are not linked to crossing because of technical difficulties. See Subsection 2.3.1.)

1.3 PROJECT OBJECTIVES

The FRA asked TSC to investigate accident prediction using the data bases just referred to in order to construct an efficient hazard index whose overall performance on public crossings in the United States would be as good as possible.

The overall goal of this project was to construct and test hazard indexes with the intent of attaining or estimating the ultimate attainable prediction power. In pursuing this goal, it was endeavored to:

- a. Rate hazard indexes (previously proposed or arising in this project) on their ability to predict relative and absolute hazard of accidents.
- b. Construct hazard indexes which are better in performance than previously proposed or previously used hazard indexes.

c. Attempt to define the limits of power achievable by accident hazard prediction functions, i.e., hazard indexes.

In all three of these endeavors, available information is confined to that in our data bases. In particular, the hazard indexes which are constructed, compared, etc., are all based on the data items which pertain to each crossing in the crossing inventory. Hazard indexes which are based on data items not included in the crossing inventory (for example, "unobstructed sight distance") cannot be compared, evaluated, or constructed in this manner. The data bases at hand are about the most comprehensive of this type ever gathered; therefore, the results reported here should be representative of overall U.S. experience.

The methodology, as described in the next section, is based on the assessment of the hypothetical performance of candidate hazard indexes if they had been used to predict the accidents which have been observed (as recorded in the accident file).

The key elements are:

- 1. A good representation of performance quality.
- 2. A means of assessing the sample variability and the capacity for generalization of our measures.
- 3. A good means of setting up a family of hazard functions which can easily be optimized with respect to an appropriate criterion.

Although goal c above, "define the limits of power achievable," is not a feasible task in so complex a situation, the results may still be quite helpful in this regard.

Note that the goals and methods all pertain to <u>prediction</u>. There has been no attempt to isolate factors which are causally related to accidents. There is a connection between the two endeavors, but since the efforts here are directed solely at predictive capability, i.e., hazard estimation, the results will not necessarily be readily interpretable from a causal point of view.

In more explicit terms, the hazard indexes are for the purposes of identifying hazardous crossings, but if the form of the hazard index formula suggests the direction and magnitude of the influence of a certain factor, this aspect is incidental, and could be misinterpreted. This and related questions will be dealt with in Section 4.

Grade Crossing Funding Allocation Project

An absolute hazard index will be used in the Grade Crossing Funding Allocation Project (FRA-TSC-RR833) currently underway at TSC. In that project, strategies for allocating funds for warning device class improvements among groups of crossings are being worked out and incorporated into computer programs. The marginal benefit/ cost ratios for individual crossings which these strategies are based on are proportional to the expected frequency of accidents at the individual crossings; hence, the need for an absolute hazard index. It is clear that absolute, and not relative, hazard indexes are necessary for input to the funding allocation algorithms. One of the goals of this project was to supply such hazard indexes for the funding allocation project.

1.4 NOTES ON STRUCTURE, CONTENT, AND CONVENTIONS

1.4.1 Definitions and Terms

In this report, several expressions will be used as synonyms for "hazard index": "hazard function," "hazard model," "hazard," "discriminant function," and "probability function." The synonyms will clearly refer to hazard index, but may connote interest in a special aspect in certain contexts.

A list of selected terms used in this report will be found in the front section titled "Glossary." This is to provide emphasis and clarification of key concepts, especially when they are discussed in several sections. A brief definition is given in the glossary, and section references are given to key passages dealing

with the item. It is, consequently, suggested that the glossary be read straight through by the interested reader, as it will aid in developing the desired perspective.

1.4.2 Appendixes

The appendixes are an integral part of this report. Nearly all the substantive data are in the appendixes. The appendixes allow results to be found easily for reference at any time. They also allow lengthy parenthetical comments without interrupting the discussion.

1.4.3 Notes on Suggested Order of Reading

Section 3, entitled "Course of Experimentation," is, from a slightly different perspective, an overview of the whole project. Some readers have found this section a helpful introduction as well as a general description of what was done. The separate perspective provided by this section is useful for the purpose of helping to communicate a general review of a complex program.

Section 2 (Methodology) can be read over quickly at first. However, the part regarding empirical operating characteristics and power factors is a prerequisite to understanding the results. The rest of Section 2 may be primarily of interest to someone interested in doing further work in hazard index construction.

Section 4 (Results) contains the primary material on what was discovered about hazard indexes and their performance.

It is suggested that for a first reading of this report the easiest sequence to follow would be: Summary; Section 1, Introduction; Section 3, Course of Experimentation; Sections 2.1 and 2.3; then Section 4 (with all its cross references).

Appendix H treats the problem of hazard indexes based on accident histroy. This important subject is placed in an appendix because the treatment here was a late development in the project, with the empirical results being preliminary; further development is anticipated.

2. METHODOLOGY

This section outlines and discusses, from a practical point of view, the methods used to pursue the three goals listed in Section 1.3. The latter section also provides the key elements involved in the methods adopted. A parallel report on the methods used, including more details and theoretical considerations, is under preparation. This will be referred to as "Comprehensive Methodology Report" (CMR).*

2.1 THE EMPIRICAL OPERATING CHARACTERISTICS (EOC)

The primary tool for the comparison of relative hazard indexes used in this study is the empirical operating characteristic (EOC). This term refers to a set of derived data to be used for comparing the performance of two or more hazard indexes on a given data base. The EOC is a set of data derived from an accident and crossing data base which has been ordered according to some hazard index.

In verbal discussions the power factors are often referred to in contexts where reference would be made to the EOC in a more formal discussion. The power factor is closely related to the EOC, and is easy to motivate and to define; hence its currency in verbal discussions without access to lengthy tables or graphical presentations necessary to communicate the entire EOC. The power factor is defined first.

The present report is meant to stand alone regarding support for the conclusions. However, there are a number of innovative techniques used here which are discussed more fully in the CMR. A complete discussion is therefore not warranted here. In addition, the CMR contains many techniques which would be applicable to this overall project if time had permitted and is being prepared as a companion report to this document. The methodology covered is applicable to analysis of accident data of various types and in general situations where predictive discriminant analysis is to be used. The Comprehensive Methodology Report does not contain information specific to grade crossing analysis, which this report contains.

2.1.1 The Power Factor

The power factor (PF) is defined as follows: The 10 percent power factor, also written PF(10%), is the percent of accidents which occur at the 10 percent most hazardous crossings (as determined by the given hazard index) divided by 10%. The same sort of definition holds for the 5 percent power factor PF(5%), etc. Thus, if PF(5%) = 3.0, then 5 percent of the crossings account for 15 percent (3x5% = 15%) of the accidents (when the 5% referred to is the 5 percent most hazardous according to the hazard index in question).

The power factor can be seen as a direct primary measure of the efficacy of a hazard index for relative ranking of crossings. Thus, suppose 10 percent of a certain group of crossings is to be selected for improvement, and assume that one wishes to select the most hazardous crossings for this purpose. Then, if a given hazard index is used, the 10 percent most hazardous crossings will be selected according to that hazard index. The number of accidents that may be expected at these selected crossings in any period of time is proportional to the power factor for the given hazard index. The greater the proportion of the total accidents that would occur at the crossings selected as most hazardous, the more effective is the hazard index as evidenced by the power factor; in fact, for some purposes, the "payoff," or benefit, will be proportional to the number (or proportion) of accidents that would occur at the selected crossings, as these accidents may be partially or totally prevented. Consequently, when the hazard index is to be used for selecting the 10 percent most hazardous crossings, the 10 percent power factor seems to be the most direct measure of its effectiveness. The same would hold for the 20 percent power factor if 20 percent of the crossings were to be selected, etc. The complete set of power factors computed at each percentile of hazard (with percentile of hazard defined as the percent more hazardous, and hence, with small order percentiles indicating higher hazard) will give the same information as the EOC. (EOC is, as has been implied, to be considered a more comprehensive term.)

The EOC contains the power factor and other related information. The power factors and the EOC are always computed on a specified data base containing a certain number of accident crossings and a certain number of inventory crossings. The data base information on which the EOC is computed thus actually comprises two data bases: the "accident" data base, which is a random sampling of accident crossings (repeated as many times as accidents occurred at the crossing in 1975); and the "non-accident" data base, which is a random sampling of <u>all</u> crossings (each repeated only once, whether or not it had an accident). Section 3 and Appendix F provide further descriptions of the primary data bases and subsampled data bases.

Appendix F describes the various sampled data bases used in this project. Subsamples of the total data base were used for two reasons:

a. The total number of crossings compared with the number of accident crossings was so large that no appreciable increase in accuracy could be achieved by using all non-accident crossings versus a fractional subsample.

b. The total number of accidents was small compared with the total number of crossings, and therefore all accidents must be used for the purposes of this analysis. Nevertheless, the number of accidents was sufficient to justify dividing them into two groups such that one could be used for hazard index construction, and the other for validation. All subdata bases were further broken down by warning device class for all model development and testing.

2.1.2 The EOC Described

The EOC refers to a large derived data set. A number of EOCs are given in Appendix C. Based on Table C-1 of Appendix C, the information contained in an EOC is described. The first six columns give EOC information pertaining to a given hazard index (labeled there as the TSC model, but numerous TSC models are represented in various EOCs.) Tables C-2 through C-13 have the same format as

that of Table C-1. In this section the format referring to Table C-1 is described; however, the most essential information in Table C-2 is presented graphically in Figure E-9. Thus, the horizontal axis of Figure E-9 corresponds to column 1 of Table C-2, and the vertical axis corresponds to column 4 of Table C-2. Further inspection of Figure E-9 will enable one to understand the EOC. For example, Figure E-9 shows that when 20 percent of the crossings are selected as most hazardous by the TSC model, over 50 percent of the accidents occur at them. The tables show the same information (and other information) more accurately.

Table C-1 will now be considered. The table's first column refers to a given percentage of the "non-accident" crossings. This was a straight sampling from the inventory, and included both accidents and non-accidents. Consequently the first column is labeled "% Crossings." The sixth column (labeled "Hazard Index") gives the value of a relative hazard index for the least hazardous nonaccident crossing in the group. Thus, column 1 gives the percentage of the "non-accident" crossings whose hazard index equaled or exceeded the value given in column 6. Similarly, column 4 gives the percentage of accidents whose (crossing) hazard index equals or exceeds the same value. Column 5 gives PF(X), where X percent is the value in column 1. Thus, column 5 is the ratio of column 4 to column 1. Column 3 gives the actual number of accidents on which the percentage in column 4 is based (column 3 and 4 are proportional) and column 2 indicates the increments in column 3 (first differences). The next five columns give the same EOC information on another hazard index -- the New Hampshire formula (based on the same data base). The rest of the columns give information for comparing the two hazard functions. Moving ahead to the 15th column, entitled "CUMMTCH" for cumulative match, one gets a very important number. It tells how many of the accidents counted in columns 3 and 8 (both labeled "CUM#ACC") are identical, i.e., how many matches there are. Thus, the 128 for the second entry in column 15 means that of the 144 accidents selected by the "TSC model" and the 135 accidents selected by the New Hampshire model, 128 were identical, i.e., included in both groups. (All this

refers to crossings selected with the 1 percent most hazardous "non-accident" crossings -- for each model in turn.) The cumulative match is important, as explained in the CMR, because it can be used to construct statistical tests for the significance of the observed difference between the two models.

The next-to-the-last column (column 20) will now be considered. That number is the difference between the percent accidents for the two models, i.e., the difference between column 4 and column 8. The last column (column 21) is for testing the statistical significance of the given observed difference. It is called the "t value," but is properly referred to a normal distribution. Thus, a t value of about 2 means "significant at the 5% level," and a t value of 3 means "significant at the 0.5 percent level." (Of course, a t value of 3.5 or 4 would be extremely significant.)

The key consideration here is that the significance refers to each row in the table separately, and does not apply if the row with the maximum value is selected by searching for it. However, when the t value exceeds 3.5, it is always significant. The reason why TVAL = 3.5 is statistically significant even if it is the largest TVAL at any of the 100 half-percentiles is that the probability of getting a standard normal deviate greater than 3.5 is 0.00033, which is less than 0.05 even when multiplied by 100 (a very conservative requirement).

The formula for the quantity "TVAL" is

 $C_{21} = (C_3 - C_8) / \sqrt{C_3 + C_8 - 2C_{15}}$

where C_{i} denotes the value in the ith column (see the CMR for derivation).

The formula for TVAL is derived informally as follows: C_3 (column 3) gives the number of accidents selected by the first formula (TSC), while C_8 gives the number of accidents selected by the second formula (New Hampshire). Now C_{15} gives the overlap. Thus, $C_3 = C_3 - C_{15}$ gives the accidents selected by TSC over and above the common accidents, while $C_8 = C_8 - C_{15}$ gives the same for New Hampshire. The variance in C_3 is approximately C_3 , and that

in C_8 is approximately C_8' (as Poisson variables). The variance in their difference is the variance of the difference of independent random variables (since the overlap has been subtracted out), and so the variance of $C_3 - C_8 = C_3' - C_8'$ can be approximated by $C_3 + C_8 - 2C_{15}$. Thus, the test of significance for the comparison of C_3 and C_8 is based on $C_3 - C_8/\sqrt{C_3 + C_8 - 2C_{15}}$.

The other columns (12, 13, 14, 16, 17, 18, and 19) are described as follows:

- Column 12: same as column 20
- <u>Column 13</u>: same as column 21 TVAL, except that C_{15} is set to 0
- Column 14: first differences of column 15
- Column 16: column 3 minus column 15
- <u>Column 17</u>: <u>column 16</u> expressed as a percentage of all accidents in the data base (for the particular warning device class)
- <u>Column 18</u>}: similar to columns 16 and 17, but for New <u>Column 19</u>}: Hampshire instead of for TSC.

Note: Columns 12, 13, 14, 16, 17, 18, and 19 will not be referred to further in this report.

Once again attention should be called to the fact that Table C-1 (like the other EOC tables) was computed on a specific sampled data base ("Test Data Base -- Crossbucks" of Table F-5). Certain columns (the key columns), however, are referable to the entire data base of all accidents and all crossings (crossbucks) for 1975. The universal columns (estimates for any sampled data base) are 1, 4, 5, 6, 9, 10, and 11. Thus, from columns 1 and 9 we see that the 15% power factor for the New Hampshire formula is 3.32 in the crossbucks case. This means that if one chooses the 15 percent most hazardous public crossbucks crossings (according to the New Hampshire formula) throughout the United States, one may expect that in a given period of time 49.8 percent of the (crossbucks) accidents will occur at these crossings. In particular,

the power factors and the basic EOC information -- percent accidents versus percent crossings as well as the relevant hazard index values -- are referable to the entire 1975 data base.

2.2 HAZARD INDEX CONSTRUCTION

2.2.1 The Linear Regression Approach

With the goal of constructing an "optimal" hazard index, this study is, in many ways, similar to a regression problem of fitting an equation in several "independent" variables to observed past concomitant values of the "dependent" variable, and then using the resultant equation to predict future values of the dependent variable when only the independent variables are known. Such a technique is used, for example, in realty tax assessment to estimate what a house would sell for if it were on the market, based on the cost of similar houses which have been sold recently. (Such systems have been used and are being adopted by communities in various states for determining tax valuation based on "fair market" value).

As used here, the dependent variable, as observed in the past, is the occurrence (or number of occurrences) of an accident in a given time period. The prediction is on the relative likelihood of occurrence of an accident in a future time period. One of the mathematical techniques used is identical in some respects to that used in the tax assessment problem. However, although the technique is in part identical to ordinary linear regression, and therefore permits the use of a standard linear regression package, the dependent variable is related to a "yes-no" situation, i.e., is it or is it not an accident? It is not so widely known that such a technique yields an indicator of the probability of an accident for the given values of independent variables. The theoretical considerations will not be covered here (see the CMR), but it is worth mentioning that the particular way the ordinary regression approach was used is equivalent to the Fisher linear discriminant technique. Specifically, classical linear discriminant functions

27

were generated, and the regression package was used for convenience. The precise form that the regression problem takes is:

$$\hat{Y}_{i} = \sum_{k=1}^{K} X_{i,k} b_{k}$$

minimizing

$$\sum_{i=1}^{N} (\hat{Y}_{i} - Y_{i})^{2}$$

where $X_{i,k}$ represents the numerical value of the kth variable evaluated for the ith crossing. K is the number of variables considered, and N the number of crossings in the sample. Y takes the value of 0 if there was no accident (in the time considered) at the ith crossing, and it takes the value 1 if there was an accident. The purpose is to determine the b_ks [for the interval (k = 1,K)], which specify the hazard function Y. \hat{Y} is an estimate of the accident probability at the ith crossing. Since \hat{Y} can be less than zero and greater than one, and is not a very good estimate of the probability (it may be transformed into a much better estimate of the probability), we consider it a relative hazard function, indicating only relative probabilities. As stated previously, it is a classical discriminant function.

2.2.2 Iterated Weighted Logistic Discriminants

Besides the ordinary regression approach to hazard index construction (which, as noted above, is also the Fisher discriminant function approach), other approaches were used. The most important of these techniques, which were used to construct the most valid and useful models, was a particular iterative weighted regression approach. Since it fit a logistic function to the probability of accident, it was called the "logistic discriminant approach." Iterative weighted regression has become the subject of much interest in recent years, but, for the connection to logistic discriminants (as used in this study), the reader is referred to the CMR, as one cannot do justice to the subject here. Logistic discriminants have been available for many years. The classical approach will be found in Cox (Reference 6). However, the approach used here is better described by reference to robust and iterative weighted regression techniques; these considerations are well described in Mosteller and Tukey (Reference 7). The particular approach used here and its justification can, as far as can be determined, be found only in the CMR. The salient features which distinguish it from ordinary logistic discriminant analysis become, in this context:

a. The ability to put the major emphasis on correctly identifying the high hazard crossings.

b. "Robustness" and "resistance" -- technical terms for important qualities in regressions. In this case, the benefit is that the logistic model doesn't have to hold exactly for the estimates of the hazard function to be valid, and also errant data points, i.e., those in strong disagreement with the others on the model parameters, have small effect. (These points are dealt with more completely in the CMR.) The version of logistic discriminant analysis used here is especially suited to this problem.

What then is logistic discriminant analysis? It has been noted in the statistical literature that the logistic function is a good model for the probability of an event when expressed as a function of a number of variables in a fairly wide variety of cases. Such an argument is often preferred by some statisticians, even when it is very unlikely that the "wide variety of cases" covers the case at hand. The simple fact remains that the logistic function, or logit function, has some useful properties. Logistic discriminant functions are discriminant functions, i.e., hazard indexes, which are logistic functions of linear combinations of the independent variables. The logistic function is simply:

$$H(h) = \frac{1}{2} + \frac{1}{2} \tanh(h) = \frac{1}{1 + e^{-2h}}$$

Logistic discrimination is the seeking of coefficients \mathbf{b}_k such that:

$$h_i = \sum_k b_k X_{i,k}$$

where $X_{i,k}$ is, as in ordinary regression, the kth characteristic of the ith crossing. The coefficients are to be chosen such that H(h) is a good absolute hazard index, i.e., $H(h_i)$ accurately estimates the probability of accident for the ith crossing. If the probability of accident is considerably less than 1, the probability of accident is equal to the expected number of accidents during the same time period.

Since H(h) is a good absolute hazard function and H is a strictly increasing function, this means that h is a good relative hazard function. The statement "h is a good relative hazard function" ("hazard function" is synonymous with "hazard index") means that if $h_i > h_j$, then crossing i has greater accident hazard than crossing j, or the expected number of accidents is higher at crossing i than at crossing j, or the probability of an accident is higher at crossing i than at crossing j.

It is important to note that to get a good <u>relative</u> hazard index it is necessary to construct a good <u>absolute</u> hazard index, i.e., predicted accident frequency. This fact leads to the use of the logistic discriminant analysis, since the logistic function has properties which make it a suitable foundation for an absolute hazard function. The chief properties which make it a reasonable function to "shape" a linear hazard index into a function which gives probability of accident are:

- a. It is strictly increasing.
- b. It does not go below zero or above one.
- c. In the "tails," i.e., for very large or very small values of the hazard index, it approaches its limit (either 0 or 1) exponentially.

The actual construction of a logistic discriminant type hazard function or hazard index is described in Appendix A.



2.2.3 <u>Hazard Function Development and Variable Selection by</u> Synthesis of Regression and EOC Techniques

The basic principles used in hazard function development and variable selection are about the same, whether the development is based on an ordinary linear regression or on the iterative weighted version described in subsection 2.2.2 and Appendix A. The procedure for the ordinary linear regression case is considered first. Figure 2-1 will assist the reader in following the discussion below; however, it must be kept in mind that the figure describes the process using the logistic approach and not straight linear regression, which will be described first. The full process, including the logistic approach, will be outlined immediately after that.

2.2.3.1 Linear Case -- The fundamental unit of search or research is a stepwise regression followed by an empirical operating characteristic (or power factor) calculation. A "variable pool" of from 2 to about 35 raw and derived variables is supplied to the stepwise regression. This is a set of characteristics or variables as quantified in the crossing inventory and perhaps transformed by some function. For example, log C and log T have been previously mentioned as possible variables. (Recall C = AADT = average daily vehicle volume, and T = average daily train volume.) "Is the highway paved?" yields a variable which is 0 for unpaved, 1 for paved. "Population" is another variable (see Appendixes B and D for definition), as is "functional class of road," and "number of highway lines," "number of main tracks," "number of switch trains," etc. Derived variables include log C, log T, log C x log T, "highway paved" (0,1) times "nearby intersecting highway" (0,1). The last variable is determined by the product of a variable which is 0 or 1 depending on whether the highway is paved, times another variable which is 0 or 1 depending on whether there is a nearby intersecting highway. If 1 represents "yes" in both cases, the result will be 0 in all cases except the one in which both the highway is paved and there is a nearby intersecting highway. Clearly, an infinite number of derived variables can be generated. The





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variables which can be generated from functions of C (average daily vehicular traffic) and T (average daily train traffic) are what are called volume variables. Thus, C, T, log C, log T, and log C \cdot log T, etc., are all volume variables. As is noted below, volume variables were found to be the chief determinants of hazard functions. Thus, it might be hoped that simple functions of the other variables would be sufficient when combined with optimal volume functions. Basically, if a non-volume variable were needed in the hazard index, the raw form should suffice when an optimal volume function was already contained in the regression. Thus, the stepwise regressions were primarily of two types.

- 1. A regression containing different volume functions in order to find an optimum volume function.
- 2. A regression in which a "pre-optimized" (volume only) function was included in the variable pool as well as non-volume variables (volume variables were also included to test the optimality of the volume only function).

A stepwise regression selects the variables one at a time from the variable pool according to how much each variable adds to the "goodness of fit" of the regression to that point. Therefore, the stepwise regression was run for several steps, adding more and more variables into the regression. Later, the results were examined to see at which steps the variables entering the regression made a "significant" contribution according to their "t values." The t values, as printed out by the regression package, are not directly interpretable in terms of prediction capability. Therefore, an empirical operating characteristic (EOC) had to be calculated for selected steps of the stepwise regression. For example, after running 15 steps of a stepwise regression (after which 15 variables had been included into the regression), an EOC might be run for each of steps 8, 9, 10, 11, and 12, since after step 12 the t values indicated that the regression was not being contributed to significantly.

If the t value indicates no significant contribution to the regression, for example t < 3, then it may safely be assumed that the further variables will add nothing to the accident prediction capability, since the t values are based on the regression criterion (i.e., least squares). If the t values are small, the regression isn't being helped (in terms of minimizing the square error) and since the regression criterion isn't being helped, its use in another connection (accident prediction) won't be helped either. However, the converse is not, in general, true: a variable can contribute significantly to the regression, i.e., to lower the sum of the squares of the residuals, without contributing to the predictive power of the resulting hazard index. For a measure of the latter feature the EOC is needed. Thus, from the EOCs at a number of steps, the best step is selected, and the result is a hazard function for further analysis, comparison, or even for use in further constructions. Thus, for either volume or non-volume (comprehensive) regressions, the regression must be followed by one or more EOCs (power factor table or plot).

In using EOCs for model development, the EOCs for selected steps in a regression or for the best steps from a number of regressions are compared. The criterion is the number of accidents included in the highest hazard groups, which include 5 percent to 25 percent of the non-accident crossings. In other words, the 5%, 5.5%, 6% ... 24%, 24.5%, 25% power factors are compared (PF1 or PF2). If two hazard functions from two different steps of one regression or from two different regressions intertwine their EOCs, i.e., if they alternate several times in which one has the higher power factor over this whole range of 5% to 25%, then they may be considered roughly equivalent. But if one has higher power factors for most or all this range (and substantially higher at some points), then that one is to be considered tentatively identified as a superior hazard function. Notice that in general the regression statistics are not sufficient to make this distinction. In other words, the t values, multiple correlation coefficients, and F values do not point to the better hazard function except in a general way. In a gross manner they do -- otherwise, regression,

2 - 1 4

and especially stepwise regression, would be useless -- but the regression statistics do not reliably tell the whole comparison story. If, of two regressions, one has a much better multiple correlation coefficient than another, then the one with the higher correlation coefficient will probably have the better hazard function, but only the EOCs can enable one to make the final decision.

The result of this is that many regressions must be run and their EOCs checked at a number of steps. Since the EOCs are more time-consuming and costly to produce than regressions, this leads to a more costly and time-consuming process than if all information were contained in the regression statistics.

Of the various ways of ensuring external validity of the models (hazard functions) produced, one of the simplest has been chosen. Two separate data bases have been developed disjoint and created under statistically identical conditions. This was done by dividing the accident crossing set into two equal parts and adjoining a separate fractional sample of the non-accident crossings* to each. The details are described in Appendix F. Although this was done in different ways at different stages of study, the discussion is simplified by referring to the two separate data bases as data base A and data base B which were separate, disjoint, independent, but identically created from a statistical point of view. The idea is to create hazard functions on one data base, and test or validate them on the other. The original plan was to do both regression and initial selection on one data base and final selection on the other. But since the EOCs are costly to run, it became apparent that the testing had to be speeded up; the eventual procedure, then, was to run all regressions on one data base (for example, data base A) and all EOCs on the other data base (data base B). Thus, every power factor is, in effect, a "validation." This weakens the ultimate validation of the models used here, because some of the selection process was carried out

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A random crossing from the crossing data base is consistently referred to as a "non-accident crossing." Whether or not it experienced an accident is not determined. See also Section 3.

on the "validation data base." The amount of weakening is, however, far less than if the regressions were also done on that data base. Since the data bases are quite large, the weakening should be slight. If time permits, the crucial steps of final regression and testing should be repeated, reversing the roles of the two data bases.

2.2.3.2 Logistic Case -- The procedures of stepwise regression alternated with EOCs (power factor tables or plots) when ordinary linear regression is used has been described above. When iterated weighted regression, i.e., logistic discriminant function construction, is used, the procedure is very similar except the stepwise regressions must be iterated several times. With reference again to Figure 2-1, the whole process is discussed with the logistic, iterative regressions used. Figure 2-1 represents something of a simplification, because the process is not quite so formal as depicted there. The steps there were carried out many times, and much effort was spent in trying to find better functions and combinations. This is indicated to a degree by the dotted flow arrows. As this process is now discussed, Figure 2-1 should be referred to frequently.

Typically, a certain small group of volume variables -- e.g., log(T+1), log(C+1), log(C+1)X log(T+1), $[log(C+1)]^2$, $[log(T+1)]^2$, see Section 3 for details -- is run through a complete iterated regression. Any terms with small t values are dropped, and the iteration continued until convergence is achieved. The resultant hazard function has an EOC run on it, and its performance as compared to the New Hampshire formula is tested. (One warning device class is worked on at a time, i.e., the whole procedure, as described in this section, is done separately on each warning device class and repeated three times for three warning device classes.)

At that point, it is determined whether any of the other simple volume terms used will enter a stepwise weighted regression with significant values. It is important to note several key points:

- a. The hazard function at this point could be called 90 95 percent complete in terms of its performance.
- b. Because of point (a), the t values for small groups (1,2, etc.) of variables which were then entered into the "selection" regression (with the primary volume hazard function as one variable) were indications of the true t values of these variables.

The use of t values to guide the choice of new variables for the regression (which is how stepwise regression works) is problematic in ordinary linear regression, and even more so in the iterative weighted (non-linear) case, since the t values are not even correct estimates of the uncertainty in the coefficients any more. However, as just noted, they are somewhat indicative of the precision and statistical significance of the corresponding coefficients b_{ν} . They will, in general, be overestimates of the true t values, i.e., the t values as output by the linear regression routine will in general overestimate the true t value which would be obtained if the estimated coefficient were divided by a good estimate of its standard error. Using more sophisticated methods (cf "jacknifing," Reference 7), one can calculate an estimate of the true covariance matrix, and thus, the true standard errors of the b_k s (regression coefficients) can be calculated in the iterated case. This step has been bypassed in order to focus the available time and resources on a wide exploration of regression equations and tests of these equations. The assumption is that the "linear" t values can serve as a crude guide, letting the EOC be the final arbiter of which variables add to the predictive power of the hazard function.

In contradistinction to the above less-than-wished-for state of affairs, it should be noted that what are called "selection regressions" are run with a nearly optimal hazard index determining the weights. Furthermore, they are run only once, and are thus "semi-linearized." It is possible to run such a regression that the regression statistics (t values, etc.) are essentially true conditional (intermediate) values. Since the selection regression is used to select variables, it can be run in the stepwise mode,

and the near validity of t values and other regression statistics is especially fortunate. The point is that the t values in the selection regressions have very close to true validity (conditioned on the volume hazard index), and this is just the sort of thing desired for the selection regression.

As has been noted, a "best volume function" is found first using the iterative procedure described above. An additional iteration can be run using the resulting h in a polynomial. This step will be described in Section 3. When the volume function has been improved, a new weighted stepwise regression is run in which nonvolume terms are allowed to enter (selection regression). Some volume terms are included in the variable pool, so that if the volume function is not completely optimized, they will be picked up early. They may be picked up at later steps to compensate or adjust for the effects of non-volume variables which have already entered. Each of the non-volume variables used is based on a single variable in the crossing inventory. Since the volume part of the function is so important, it was felt that non-volume terms, which add little to the function, could be expected to make their contribution as single variables, i.e., no cross products. Sections 3 and 4 show the justification for this.

The stepwise regression selects non-volume variables from a large variable pool. They are selected, of course, by the t values. These t values may be expected to be rough to good indications of which variables to select as indicated above. The key to the procedure is that the stepwise procedure is cut off at a very high t value, and the variables selected at that step are then to be used in a final iterative regression. The stepwise regression just described is of the weighted type with U(h) and V(h) (see Appendix A) determined by the h for the best volume regression called V(C,T). From this point, the final iterations use the volume function V(C,T) as one of the variables along with the non-volume variables (and any additional volume variables) selected in the stepwise selection regression just run.

Several sets of variables, as chosen at different steps of the selection regression, are run through the iterated regression. Each is run through several steps until convergence of a comprehensive (volume and non-volume) hazard index is achieved. EOCs are run on each, and the best is selected as the final hazard index for the given warning device class.

2.3 EXPECTED ACCIDENT FREQUENCY (ABSOLUTE HAZARD INDEXES)

2.3.1 Expected Accident Frequency from Logistic Discriminants (TSC) Nonlinear Hazard Indexes

As noted earlier, the primary interest in comparing hazard indexes is to determine their relative ability to select hazardous crossings -- that is, to compare them as relative hazard indexes. Once a good relative hazard index is determined, it can be converted to an absolute hazard index. Presently it will be shown how to get an absolute hazard index from the analytic expression for the power factor curve. In addition, the logistic discriminant procedure produces hazard indexes which are immediately interpretable as absolute hazard indexes. This is because H(h) = $1/1+e^{-2h}$ gives the probability of the crossing being an accident crossing. (One needs to recall that h is any of the HIs of paragraph B.5 of Appendix B. It is a relative hazard index. The value 10,000h is the value indicated by the numbers in column 6 of Tables C1-C6, the HAZARD INDEX column for the TSC model.) Because the "non-accident data" base was really a straight sample of the inventory, and since the "accident crossings" file had each crossing repeated for each accident occurring at the crossing (proportional representation) this results in the estimate of the frequency of accidents per year at a crossing as:

 $C_1 H$ $T_1 H$ or $C_1 e^{2h}$

The quantity h x 10^4 is the quantity tabulated under "TSC MODEL-HAZARD INDEX" in Table C-1, etc. The constant C₁ is dependent on how many accident crossings (accidents) and how many crossings

were selected for the sub data base versus the same ratio in the total data base. If, for the given warning device class, there were M total accidents (in the 1975 accident file) and N total crossings while in the sub data base used for creating the hazard index, there were m total "accident crossings" and n total sample from the inventory i.e., "non-accident crossings," then:

 $C_1 = \frac{M}{N} \frac{n}{m} r.$

The quantity r is a scale factor which takes into account that not all accidents occurring at public grade crossings in 1975 were represented in the accident data base used here. The reason for this lack of total representation is that some of the accident records could not be linked to the crossing records because of missing or invalid crossing i.d. information. With 8,028 accidents represented in the data base, and with a total of 11,350 accidents at all public grade crossings in 1975, one can scale C_1 by the factor r = 11,350/8,028 to arrive at a final scale factor for converting $e^{2h}TSC$ into expected accident frequency in 1975. Table 2-1 gives the values of M, N, m, n, and C_1 for each warning device class considered.

TABLE 2-1. FACTORS FOR ABSOLUTE EXPECTED ACCIDENT FREQUENCY

Class	М	Ν	m	n	r	C ₁	
Crossbucks	3,969	141,477	1,985	20,188	1.414	0.403	
Flashing Lights	2,650	33,969	1,326	13,250	1.414	1.10	
Automatic Gates	707	11,983	354	3,535	1.414	0.833	

Warning Device

Thus, hazard index h gives rise to the absolute hazard index $C_1 e^{2h}$ as an estimate of the expected yearly accident frequency at the crossing. The statistical measures of goodness of a probability estimate are not always enlightening, and so one seeks to construct a direct (if crude) estimate of this probability

function against which to check the hazard index. One such estimate is to base it on the relative numbers of incremental accidents in each 1-percent interval. Figures E-2, etc., show plots of frequency of accident, $f(h) = C_1 e^{2h}$, versus the above crude empirical estimate for the TSC comprehensive models for crossbucks, flashing lights, and gates. The scatter of the empirical estimates is to be expected, and does not reflect a fluctuation of the true value. The true value should cut through the center of the scatter, and so a rather good fit is observed. Section 4.3 will discuss these plots in detail.

2.3.2 Absolute Hazard Indexes Based on Power Factor Information

It has been noted that the EOC and power factor information is useful for comparing hazard indexes on a relative basis, i.e., to assess their efficiency in ranking crossings for relative hazard. Every absolute hazard index is a relative hazard index, and so absolute and relative hazard indexes can be compared with each other, all on a relative basis. In this section the new hazard index construction techniques employed have been discussed, and the ultimate techniques result in absolute hazard indexes (the TSC models yield an estimate of expected accident frequency per year) at each crossing. Section 4.3 exhibits the results of such estimates, and gives an indication of how accurate they are.

It is a property of most hazard indexes previously given in the literature that they are poor estimators of absolute hazard, although they may not be poor estimators of relative hazard. (Some of the Coleman-Stewart models are possibly exceptions. See Section 4.3.) Therefore, a technique for deriving absolute hazard indexes from relative ones has been developed. As noted above, the new hazard index construction technique used in this project yields <u>directly</u> an absolute hazard index. Consequently, the method given in this subsection is primarily to convert <u>other</u> relative hazard indexes into absolute hazard indexes.

If the power factor as a function of percentage of crossings is denoted by $\rho(\lambda)$, where λ is expressed as a fraction rather than

as a percent, i.e., λ = percent crossings more hazardous/100%, and $\rho(\lambda)$ is the power factor at that percent (or fraction), then an analytic expression may sometimes be found which approximates this relation. Thus (see Section 4), to some degree of approximation, log $\rho = \alpha (\log \lambda)^{\beta}$ (where α and β are constants) can be used to represent the power factor function. The key fact is that if an analytic representation for $\rho(\lambda)$ can be found, then, whatever the form, an expression for the expected accident frequency, f, can be found. One needs to remember that the expected accident frequency (number per year) f is what is called an absolute hazard index. If $\rho(\lambda)$ is the power factor as a function of proportion of most hazardous crossings, then:

$$\mathbf{f} = C_2 \ \rho \left(1 + \frac{d \ \log \ \rho}{d \ \log \ \lambda} \right) = C_2 \ \rho \left(1 + \frac{\log \ \rho}{\log \ \lambda} \frac{d \ \log \ \log \ \rho}{d \ \log \ \log \ \lambda} \right)$$
$$C_2 = \frac{M}{N} \mathbf{r}$$

where,

M = total number of accidents in data base which ρ represents N = total number of crossings in the data base

r = scale factor (see Section 2.3.1 and Table 2-1).

This expression is exact and will give an exact absolute hazard index f unless the expression $\rho(\lambda)$ does not adequately express the relation of power factor to percent (proportion) of crossings. If log $\rho \approx \alpha(\log \lambda)^{\beta}$, then d log log ρ/d log log $\lambda \approx \beta$, and so

 $f \approx C_2 \rho [1 + (\log \rho / \log \lambda) \beta].$ ** Section 4 shows that an expression of this form (with β constant) for f gives satisfactory estimates of accident frequency (absolute hazard index) in cases of interest.

The expression Δ log log ρ/Δ log log λ can be evaluated over short ranges in the EOC table, so that, in effect, β is not a constant. This possibility is mentioned for later analysis, since, although the behavior of Δ log log ρ/Δ log log λ for 1/2-percent intervals for one data base is exhibited (see Table C-7), the form (**) with β constant is adequate for the present purposes.

(Note: Up to this point the methodology has been explained; for the results, refer to Section 4.)

The use of the expression (**) to determine f requires an EOC table in which f can be calculated as a function of raw (relative) hazard index. As has been noted, for each 1/2 percent increment in crossings the hazard index (New Hampshire) h has been tabulated (in Table C-1 and other tables of that format) in column 11; in column 10 the power factor, ρ ; and in column 1 the (cumulative) percentage of more hazardous crossings, λX 100. Thus, for a given value of h (New Hampshire), h is first located in column 11, then the corresponding value of ρ has to be found from column 10 and λ from column 1 (dividing the latter by 100), and this substituted in expression ****** to determine f(h). (See Table 2-2.)

This procedure need not be carried out if one of the new hazard indexes (produced in this report) is used, as they directly yield a value for f. However, the technique is valid for converting relative hazard indexes of any sort to absolute hazard indexes.

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TABLE 2-2.EMPIRICAL POWER FACTOR FORMULA FOR
THE NEW HAMPSHIRE MODEL

$$\rho \approx \alpha (\log \lambda)^{\beta} \leftrightarrow f \approx \frac{M}{N} r \left(1 + \frac{\log \rho}{\log \lambda} \beta\right) \cdot \rho$$

for the New Hampshire formula:

	β	<u>M/N</u>	α
Crossbucks	.7 ± .1	.028	.76
Flashing Lights	.75 ± .1	.078	.65
Gates	$.6 \pm .1$.059	.62

Thus, for $\lambda = .005$ (.5%) to $\lambda = .5$ (.50%)

$$\beta = \frac{\Delta \log \log \rho}{\Delta \log \log \lambda} \qquad \alpha = \frac{\log \rho}{(\log \lambda)^{\beta}} \qquad r = 1.41$$

 $\log = natural \log arithm (\log 2.71828 = 1).$

Information for calculating best α and β values for all cases in EOC Tables C-1 to C-7. (See Sections 4.1.1, 4.1.2 and 4.1.3.)

Note: The second term $[(\log \rho/\log \lambda) \cdot \beta]$ within parentheses in the above equation for f is to be considered negative.

COURSE OF EXPERIMENTATION

Over the years, several models have been developed with the purpose of creating a relative ranking of accident potential (hazard index) attributed to railroad crossings. Many of the models were constructed using local (and in some cases specialized) data, and were not representative of other areas of the country. Moreover, the small amounts of data used to construct the models challenges the accuracy of the results.

The creation of the inventory data base meant that, for the first time, large amounts of standardized inventory data were available for the construction of an accident prediction model. Table F-1 of Appendix F describes the format of the inventory data base.

The data available for this study were molded into two forms:

- The inventory characteristics of <u>all</u> public railroad crossings (219,162 crossings).
- The inventory characteristics of public railroad crossings which had an accident in 1975 (8,028 crossings, with 943 duplications for multiple accidents).

These two sets of data became known as the "non-accident" data base (inventory characteristics of railroad crossings) and the "accident" data base (inventory characteristics of railroad crossings which had an accident in 1975).

The "non-accident" data base contained each crossing whether or not it had an accident and each crossing occurred in the "accident" data base as many times as it had accidents. As a result, there was a certain amount of overlap (redundancy) between the two data bases as well as within the accident data base. Since the accident base consists of the inventory characteristics of the public crossings which had accidents in 1975, and the non-accident data base consists of the inventory characteristics of all public crossings as of 1975, by definition the characteristics of all crossings with accidents are contained in both data bases. This

precipitates an overlap of 3.66 percent. Furthermore, on investigation of the crossing identification numbers for the accident data base, 943 crossings out of 8,028 were found to have multiaccidents. Table F-4 of Appendix F contains a breakdown of these multi-accident crossings.

It should be pointed out here that the non-accident and accident data bases were the complete sets of data available. Variable subsets of data were extracted from both data bases to provide input for the particular mathematical tool being implemented. Subsets were used because of the economics of time and money which could be saved instead of using the complete sets of data.

In most cases, these subsets consisted of two pairs of accident/non-accident data bases; one pair was used for developing a model, and the other pair for validating the model developed. These pairs were, for the most part, statistically and numerically identical while, at the same time, disjoint. Table F-5 provides the breakdown of the non-accident and accident data bases by warning device class. (Table F-5 refers to the data bases used for the nonlinear logistic case only. All results reported on in Section 4 are for this case.)

As a starting point for the construction of an accident prediction model, both data bases were reduced. These data originally contained 84 fields of information, but for the purpose of this study, only 51 fields were extracted (see Table F-2). Some fields were eliminated because they were descriptive in nature and hard to quantify.

The 51 fields describing railroad crossing inventory characteristics were examined. One particular report (Reference 8) was very useful in providing areas of concentrated and isolated inventory characteristics.

After familiarization with the data available and extracting only the fields of interest, several existing models which provide a hazard index were examined. The purpose of this examination was two-fold in nature. First, by examining inventory characteristics

used in previous models, insight was gained in finding those characteristics which seemed to be more predictable. Secondly, it was intended to use these previous models as benchmarks to which the newly constructed models might be compared.

Of these models examined, only three were found to be applicable to the data bases: the New Hampshire formula, the Peabody-Dimmick formula, and several formulas calculated by Coleman and Stewart. These equations appear in Appendix B. The remaining models were deemed unusable, because their formula called for variables not included in the operating data bases. Some of the variables were accident experience (or accident probability) and a sight distance rating. (Note that accident history may and should be used in further studies with the FRA data bases.)

By use of a linear regression technique (the first technique that was tried), variables (inventory characteristics) were examined further to determine their relative rank of predictability for accidents. Several observations were noted at this point. They will be discussed later.

From previous studies it was decided that the best approach to constructing an accident prediction model hinged upon the segregation of the data by warning device class (see Table F-3). Because they contained the largest amounts of data, crossbucks were examined first. This particular theme of examining only those crossings with crossbucks was carried throughout the model building until an adequate crossbucks model was obtained. At that point the other warning device classes were examined separately.

From the total reserves of data, 15,654 "non-accident" crossings (actually without regard to whether or not accident occurred, and which were selected randomly) and 1,246 accident crossings, all of which had crossbucks, were extracted for testing purposes. A linear regression was to be used. Later it was shown that a nonlinear regression worked better.

First to be described here is some extensive experimentation with linear regression because of some basic principles brought out

in the course of that work. However, the chief results reported in Section 4 came from the nonlinear regressions which are also reported on in this section.

The linear regression was constructed as follows. There were 27 independent variables and one dependent variable. The dependent variable was either a zero (0) for the crossings from the non-accident subset, or a one (1) for those crossings from the accident subset. Theoretically, each dependent variable approximated the probability for an accident at each crossing. Thus, the crossings that had an accident were given a probability of 1, while those crossings that did not have an accident were given a probability of 0. The values for the independent variables were the inventory characteristics themselves as they appeared in the subsets. No distinction between accident and non-accident data was made for the independent variables. The results of this regression appear in Table D-1 of Appendix D.

The conclusion drawn from this initial examination was simplistic: volume variables (train and vehicle movements) contributed more to the predictability of the regression than did the non-volume variables. This conclusion is drawn from two pieces of evidence. First, the correlation matrix indicates that the volume variables consistently have higher coefficients (when crossed with the dependent variable) than the non-volume variables. Second, three of the first four terms selected by the regression were volume terms.

It was decided that the best approach to understanding the volume variables was to construct and run regressions which contained only volume terms. In order to meet this end, many regressions were run which contained many varied functional forms of all volume variables plus combinations of volume variables. Those variables and functional forms which showed promise were sifted out for further examination and optimization.

One of the problems faced was the determination of which regression model was the better predictor. Not only were there numerous regressions to choose from, but each regression had many

steps (a stepwise regression model was utilized) in which a model could be constructed. There were well over 100 possible models from which to choose.

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Power factors and "empirical operating characteristic" (EOC) curves* were utilized to determine the best models. After each regression was run, power factors were computed for several different steps of that regression. These power factors could then be examined to determine the "best step" for that regression. Each "best step" for a regression could then be compared with other "best steps" from other regressions to find the optimal model.

During these proceedings, several trends were observed. First, in any given regression, colinearity usually destroyed the power factors when 10 or more variables composed the model. This manifested itself in the regression through low t and F values, and in many cases, the signs of the regression coefficients were reversed. Second, the LOG_{10} functional form was a better predictor than other functional forms.

With these discoveries in mind, a final regression was constructed to obtain an "optimum volume model" The independent variables for this regression are shown in Table D-2 of Appendix D. As usual, the dependent variable was 0 or 1. Several steps of this regression were compared against each other using power factors. The "optimum volume model" was derived from step 5. This formula appears in Appendix B.4.1.

Having obtained the best volume model, the next logical step was to piggyback non-volume variables onto the volume model. To obtain this effect, the volume model was made an independent variable and fed into the pool of independent variables along with non-volume variables and combinations of non-volume variables and volume variables.

[&]quot;The terms "power factor" and EOC are defined in the Glossary and Index of Selected Terms and Expressions and also in Section 2.1. The power factors and EOCs are the primary measures of predictive power used throughout the study. They were calculated on separate data bases from those on which the regressions were run.

The purpose of inserting additional volume variables as possible selections for a comprehensive regression is twofold. First, it serves as a check on the volume equation to insure that various volume data are best represented, and secondly, combining them with non-volume variables enables the possibility of several cross-product variables being selected which would not have been chosen by themselves.

Several regressions were processed using this approach. Two examples of these comprehensive linear regressions are shown in Tables D-3 and D-4 of Appendix D. Again, several steps from each regression were compared; power factors were used to determine the "best step" for each regression, and "best steps" from all the regressions were compared to determine the "best model." From the many steps and regressions examined, two models were selected as the "best models." These appear in paragraphs B.4.2 and B.4.3 of Appendix B. (They are referred to as "best linear models." The best nonlinear models, to be discussed shortly, were better.)

The power factors and EOCs showed these "best" linear models to be superior to the New Hampshire and Coleman-Stewart models. However, the data base for testing the models was inadvertently not disjoint from the data base for constructing them. When this was discovered, some preliminary testing indicated that the improvement was not statistically significant. Much later, a special experiment was run to compare the best linear models with the best nonlinear (logistic) models developed in this project. This experiment is reported in Appendix G.

The construction of a comprehensive "best model" (linear) concluded the work in which a linear regression technique was used. The observations from this segment are summarized below.

- Volume variables (train and vehicle movements) account for approximately 90 percent of the predictive powers of the regression.
- Regressions with more than eight variables frequently produced poor power factors due to colinearity. The signs of colinearity included low t values, low F values,

misdirected signs of variables, and high cross-correlation coefficients as seen in the correlation matrix.

- 3. The New Hampshire formula and the Coleman-Stewart formula (for urban/rural crossbucks) were similar to each other in that they produced almost equivalent power factors. Furthermore, the ranking capabilities of both these models were good. In contrast, the Peabody-Dimmick formula was inferior to both the New Hampshire and Coleman-Stewart formulas.
- 4. The best TSC model was at least as good as both the New Hampshire and Coleman-Stewart formulas.
- 5. The best TSC model appeared to be not quite statistically significantly better than the New Hampshire and Coleman-Stewart models.
- 6. Linear regression techniques might not be providing an adequate tool for producing optimal or near optimal models. Many regressions containing hundreds of different combinations and functional forms were tested, evaluated, and validated, yet none appeared to significantly improve the ranking capabilities over either the New Hampshire formula or the Colemen-Stewart formula.

Because of the less than satisfactory results using linear regressions techniques, other approaches were tried. The next approach required the calculation of cross-tabulation tables for both the non-accident and accident data bases.

Having obtained the cross-tabs ratios (ratios of the number of accidents in a particular cell to the number of non-accidents in the corresponding cell), these ratios were used in conjunction with a stepwise linear regression to construct an accident prediction model. Several regressions were run using various functional forms of these cross-tabs ratios. It soon became apparent that many regressions which could introduce various functional forms and combinations were necessary to obtain a significant model. Since the approach was time-consuming compared to

other approaches, it was abandoned before any conclusive results could be found.

At this point a different approach was tried. The following serves as a simplistic explanation of the problem to be solved and the new directions available.

If one were to represent the accident and non-accident data as probabilities, the point and line curve would look like the sketch below.



If the data are from an accident crossing, its probability for an accident is 1; thus, the points along P(ACC) = 1. If the data are from a non-accident crossing, its probability for an accident is 0; thus, the points along P(ACC) = 0.

If one were to fit these points using a linear regression model, the approximation would be a straight line as shown below. Note that this approximation is weak where P(ACC) = 0 and P(ACC) =1, the places where good fits are necessary.



Not only is the linear approximation a crude fit, but it also produces an overshoot where P(ACC) = 1, and an undershoot where P(ACC) = 0. This appears to be a satisfactory explanation of the linear regression not being able to produce a significant accident prediction model.

Another approach to the problem considered the use of piecewise linear fits. The theory behind this approach is to approximate the hypothetical probability curve by several linear curves. Theoretically, this approach is shown below.



However this approach did not yield any conclusive results. In fact, the problem of overshoot still existed. It is possible that this approach would produce significant results if more time were allotted for testing.

However, it was hopeful that a third approach, which considered the nonlinear shape of the hypothetical probability curve, would render a quicker solution. The shape of the hypothetical probability curve resembles the hyperbolic tangent curve. Whereas the hypothetical probability curve takes on values between 0 and +1, the hyperbolic tangent curve ranged from -1 to +1. It was decided to incorporate the hyperbolic tangent curve into the linear regression package to produce a nonlinear regression. The mechanics of this method are described in Section 2 and Appendix A.

The execution procedure for this nonlinear regression becomes iterative in nature. That is, the regression coefficients that result from the first regression are used as parameters in the second regression; the coefficients from the second regression are used as parameters for the third regression, etc.

The independent variables were the same variables used for the linear regression technique. However, the dependent variable took on a value of either -1 or +1 (for non-accident and accident crossings, respectively) whereas the linear regression had a value of either 0 or 1. The change in the dependent variable only required a simple transformation to produce the probability of an accident.

Since the iterative process is innovative, several behavioral observations should be noted at this point. The number of independent variables should be limited to about six, since a greater number could introduce noise into the regression. As a starting point, the initial values for the regression coefficients for each independent variable may be set to 0. This allows each variable to reflect its original value. During each iteration, the values of the coefficients will change. Some coefficients will migrate and converge toward their final values, while others will migrate in an unsystematic fashion and not converge. Those variables which do not converge on a final value (steady state) are unstable, and should be dropped from the pool of independent variables. When the coefficients approach their final values, a tolerance can be imposed for judging when the iterative process is complete. The tolerance used in this study was based on the standard error of estimate for each coefficient. When the change in coefficient (from one iteration to the next) for all variables was less than 1/10th of their respective standard error of estimates, the coefficients were said to be converged to their steady state and the iterative process complete. One final word of caution: the t values and F values of the regression become meaningless due to the nonlinearity of the regression. Therefore, the acceptance of a variable into the equation is based on the systematic migration and convergence of its coefficient to a final value. Most of the coefficients migrate at a fairly consistent rate of change.

The data base subsets used for the nonlinear regression were constructed quite differently than the subsets for the linear regressions. First, it was decided to construct two subsets of accident/non-accident data; one pair of accident/non-accident data for experimental purposes, and the other pair for validation purposes. These data bases were named the "TEST" data base and the "VALIDATION" data base, but since they were constructed to be statistically and numerically identical, they could be interchanged for testing and validation purposes. A table of the breakdown of these subsets appears in Table F-5 (Appendix F).

It was decided to construct an accident prediction model for each warning device class. It was also decided to construct each model in three phases. The first phase was to consider only volume variables and produce a "best volume" model. The second phase was intended to refine the "best volume" model, shaping it into a polynomial of up to the third degree. The third phase incorporates non-volume variables into an equation with the refined volume model (a polynomial) to produce the accident prediction model or a "comprehensive" model. These phases may be expressed in general terms as shown below.

Phase 1 (Best Volume Model)

 $H = a_0 + a_1 \ \log_{10}(T + 1) + a_2 \ \log_{10}(C + 1) + a_3 \ [LOG_{10}(T + 1)]^2$ $+ a_4 \ [LOG_{10}(C + 1)]^2 + a_5 \ LOG_{10}(T + 1)*LOG_{10}(C + 1)$

Phase 2 (Refined Volume Model)

 $HR = b_0 + b_1 H + b_2 H^2 + b_3 H^3$

Phase 3 (Accident Prediction Model)

 $HI = c_0 + c_1 HR + c_2 X_1 + c_3 X_2 + c_4 X_3 + \dots + c_n X_{n-1}$

where

a_i, b_i and c_i are regression coefficients
X_i are non-volume variables
T is train movements
C is vehicle movements
HI is the resultant hazard index. (Referred to as "h" in
Sections 2 and 4.)

Crossbucks were the first warning device class to be modeled. The series of regression printouts in Table D-5 of Appendix D illustrates how the regression coefficients migrate. This series of iterations is from the "refined volume" model described in phase 2. The iteration was terminated, since the changes in the coefficients are all less than 1/10th of their respective standard error of estimates. It should be noted that the best volume model contained only four terms: the intercept, $LOG_{10}(T+1)$, $LOG_{10}(C+1)$, and $[LOG_{10}(T+1)]^2$. The other two terms, $[LOG_{10}(C+1)]^2$ and $LOG_{10}(T+1) + LOG_{10}(C+1)$, were dropped because their coefficients did not systematically migrate. Thus, for the crossbucks "best volume" model, a_4 and a_5 were zero (from the general equation in phase 2). The equations for the crossbucks "best volume" model and "refined volume" model are shown in Appendix B.5.1. The power factors appear in Table C-1 of Appendix C.

Once the volume model for crossbucks was found, the next step (phase 3) called for the incorporation of non-volume variables. A "selection" regression was developed and utilized to introduce the most effective non-volume variables. The methodology behind the "selection" regression is discussed in Section 2. The "selection" regression is a technique that was developed to indicate which non-volume variables would be most effective at directly predicting accidents in the nonlinear environment. In contrast to the iterative nonlinear regressions, the t values provided by the "selection" regression are true measurements. The selection regressions for crossbucks is shown in Table D-6 of Appendix D. Volume variables appear with non-volume variables in the "selection regression" pool in order to provide different combinations for the comprehensive model as well as to strengthen the volume model if necessary.

After choosing the non-volume terms to combine with the volume model, the regression was iterated until a steady-state solution for the regression coefficients was found. This equation, the comprehensive accident prediction model, is shown in Appendix B.5.1.3.

Power factors were calculated from this comprehensive model to provide an indication of success when measured against the

benchmark models (New Hampshire and Coleman-Stewart). The power factors for the crossbucks model are shown in Tables C-2 and C-8 of Appendix C. In order to better illustrate the success (or failure) of the accident prediction model for crossbucks, a plot of the power factors derived from the equation was contrasted against the New Hampshire and the Coleman-Stewart formula. This plot appears in Figures E-9, E-10, and E-11 of Appendix E.

Besides providing a hazard index, one of the features of the accident prediction model is its ability to supply the probability of accidents. Thus, the hazard index may be transformed into a frequency of accident curve. A plot of this transformation for crossbucks is shown in Figures E-1 through E-8 of Appendix E. The TSC model is contrasted against the Colemen-Stewart model for crossbucks. The solid curve represents the performance of the model, while the points, denoted by x's, represent the empirical data.

After the crossbucks model was completed, the three processing phases were repeated for crossings with flashing lights, and again for crossings with gates.

In both the models for flashing lights and gates, phase 2 (where the refined volume model is constructed in the form of a polynomial) is not used. They are constructed this way because the coefficients for the polynomial in phase 2 would not migrate in a systematic manner and converge to a steady state. Therefore, it was deemed that the polynomial was unstable and caused by colinearity. Thus, for both the flashing lights model and the gates model, b_0 , b_2 , b_3 in the general equation for phase 2 were equal to zero, while b_1 was equal to one.

After the volume models for both flashing lights and gates were obtained, "selection" regressions were run for each warning device class to determine which non-volume variables should be incorporated into the comprehensive equation. The selection regressions for flashing lights and gates are depicted respectively in Tables D-7 and D-8 of Appendix D. After the comprehensive model was iterated, power factors were calculated as well as frequency of accident curves and EOC curves.

The power factors for flashing lights appear in Tables C-3, C-4, and C-9 of Appendix C. Those for gates are in Tables C-5, C-6, and C-10. The EOC plot for flashing lights appears in Figure E-10 of Appendix E and the EOC plot for gates appears in Figure E-11. Frequency of accident plots for flashing lights appear in Figures E-4, E-6, E-7, and E-8 and for gates in Figure E-5.

The last four iterations of the "best volume" model for flashing lights are depicted in Table D-9 of Appendix D. The last four iterations of the comprehensive model (phase 3) are shown in Table D-10. These examples illustrate how the coefficients of the variables migrate toward a steady state. Also shown are the last four iterations of phase 1 and phase 3 for gates. These appear in Tables D-11 and D-12, respectively.

The "best volume" equation and the comprehensive equation for flashing lights are given in Appendix B.5.2. Those for gates are shown in Appendix B.5.3.

As a feature of this study, all three classical models were examined. Tables C-8 and C-11 of Appendix C indicate the power factors of the Coleman-Stewart formula versus the New Hampshire formula and the Peabody-Dimmick formula versus the New Hampshire formula for crossbucks. Coleman-Stewart versus New Hampshire and Peabody-Dimmick versus New Hampshire for flashing lights are shown in Table C-9 and C-12. Tables C-10 and C-13 show both comparisons for gates.

As noted in Section 1, the chief objective of this study was to produce a better hazard index. According to the rather rigorous criteria used, this objective was achieved. The magnitude of the improvement may be judged by the reader as to its importance.

Some rather comprehensive displays and comparisons of TSC's new models and previous models have also been produced. The third objective, which was to estimate the limits achievable in prediction models (hazard indexes) for railroad grade crossings, is, of course, almost a philosophical contradiction. There is no way to say that a formula cannot be found which will reliably predict accidents with any arbitrary precision. Nevertheless, the experience and results of this study strongly suggest fairly sharp limits on the power factors achievable with any hazard index. The reason for this is that similar results have been obtained with numerous attempts to find the best hazard index.

Note that in the subsections of this section, frequent reference is made to the appendixes. In particular, the definitions of the various hazard index formulas discussed appear in Appendix B. As each of these formulas is first referred to in this section, the number of the precise formula (or paragraph) in Appendix B is given, e.g., Appendix B.5.1.3. Also, EOC tables and curves and accident frequency curves will be referred to in a similar manner in Appendixes C and E. The reader is urged to turn to each such reference as it occurs, for the first time at least. Appendixes B, C, and E provide the primary empirical result information, and might well be marked by index tabs for ready reference.

4.1 EOCs AND POWER FACTORS

4.1.1 Introductory Comment

As noted numerous times in Sections 2 and 3, the bulk of this project was exploratory in nature: an attempt to find what worked and what techniques yielded the best hazard functions. Therefore,

most of the work was concentrated on a single warning device class. Crossbucks were chosen because of the large number of crossings involved. Flashing lights might have been profitably chosen, since that warning device class involved only about one-third fewer accidents and would, therefore, have yielded nearly as precise estimates.

Nevertheless, the objective of this study was to do the whole study for all warning device classes. Preferred methods have been applied, developed from extensive testing on the crossbucks case to the other two warning device classes, flashing lights and automatic gates. The results are shown in the power factors in Tables C-3 to C-7 of Appendix C.

4.1.2 The Crossbucks Case

Table C-2 (see also Figure E-9 of Appendix E) gives the EOC for the best TSC crossbucks model (comprehensive) compared to the New Hampshire (crossbucks) model. (The best TSC models will be discussed below in Section 4.2. They are defined in Appendix B.5.1.3). Remember that the best models are all of logistic construction. The power factors for each percent level may be seen. Thus, if 2 percent of the crossings are chosen, a power factor of about 6 is obtained (PF(.02) = 5.9). If 6 percent of the crossings are chosen, the power factor becomes about 5 (PF(.06) = 4.92, and 6 percent of the crossings are chosen, the power factor is about 4 (PF(.105) = 4.0). At 20 percent of the crossings, the power factor becomes about 2.

To get some idea of how close these are to the best achievable, they may be compared to the New Hampshire power factors. Denoting the power factors for the TSC model by PFTSC and the New Hampshire power factors by PFNH, one notes the data of Table 4-1:

Crossing	PF TSC	PF NH	PF ''X''	±
1	7.86	6.80	8.8	10%
2	5.90	6.17	6.4	7 %
6	4.92	4.76	5.2	5 %
10	4.10	3.83	4.4	4 %
20	3.01	2.88	3.3	3%
40	2.03	2.03	2.1	2.5%

TABLE 4-1. ESTIMATES OF POWER FACTORS AND "ACHIEVABLE" POWER FACTORS (CROSSBUCKS)

The t values (Table C-2) show that the TSC model is significantly better at certain percentage levels, most notably the 9 percent level, where the t value is 4.4. The reader will recall that, as indicated in Section 2.1, a t value of 3.5 is to be considered <u>more</u> than significant (statistically) at the 0.05 level (even though, as in this case, the largest of 100 t values has been selected). Nevertheless, the improvement is small. The improvement is best measured by the percent of the total accidents the TSC model selects over that which the New Hampshire model does. It is seen that this is a maximum in the 8-10 percent range, where it is 3 percent of the accidents (out of about 40%, which are picked up by both models). Therefore, at the point of most favorable comparison, the TSC model is relatively about 7-8 percent better.

It may be expected that any other model would be better than the TSC model by no more than the TSC model is better than the New Hampshire model. This statement cannot be proven, and is only approximate. It may overestimate or underestimate the maximum achievable. The column labeled PFX would represent this crudely estimated bound on the ultimately attainable power factors. The reasons for surmising that PFX may represent the best attainable are:

a. The large amount of experimentation and testing done leads one to suspect that the TSC model is near the best.

b. The improvement achieved over the simplest good model (New Hampshire) is relatively small, and, therefore, would be expected to estimate in order of magnitude an upper bound on what is left for improvement.

In practical terms the improvement achieved -- about 3 percent more of the accidents identified (out of 40 percent total for both models) for 10 percent total crossings chosen may or may not be considered pratically significant. If the hazard index were being used only at this level, then upwards of 3 percent of all accidents (at crossbucks) could be anticipated (by the new model over and above the New Hampshire model) and thus avoided by upgrading the warning equipment at the crossing. This could amount to many accidents over many years. However, in general, other information may be available than is used by these formulas, and, in this case, they may be expected to be used mostly for preliminary screening, so that accuracy would not be such a large factor. Incidentally, as noted in the column + (i.e., absolute percentage errors in an estimated power factor) in Table 4-1 above, the power factors as estimated here are about as uncertain in the absolute magnitude* as is the difference between what is estimated to be the power factor of the best model and the power factor of the best attainable model.

In cases where the hazard models are to be used for final selection of crossings for improvement expenditures for whatever reasons (such as objective formula required by law), then the roughly 3 percent more accidents which may be "pre-identified," i.e., anticipated or predicted by the TSC model over and above the New Hampshire model, could be of considerable interest. As further aids in interpreting and understanding the power factor information in the EOC, Appendix E contains plots of EOC curves (which give percent accidents vs. percent crossings) for various models and warning device classes as indicated. The relatively small

This quantity, which appears in the column headed "+", is $1/\sqrt{\text{CUMACC}}$ expressed as a percent. CUMACC is in column 3 of the C tables. Thus, $\pm = 1/\sqrt{\text{col } 3} \times 100\%$. This number has been rounded up for Table 4-1 to be conservative.

but noticeable improvement of the best TSC models over the previous models can be observed.

Analytic Power Factor Curves

Tables C-1 to C-6 also present data on single parameter representations of entire power factor curves. Since power factor curves are empirically determined, they are inherently noisy (see Figures E-9 to E-11). An analytic expression can remove this noisiness to facilitate power factor estimates and comparisons. Such expressions can be used to estimate power factors without having the entire tables at hand, and also to interpolate and extrapolate. They are also used to obtain accident frequency estimates below.

The numbers labeled CON1, CON2, etc., in Tables C-1 through C-6 are "constants" appropriate for four empirical formulas for power factors. Let λ be the percent level expressed as a fraction, and ρ the power factor at the level. Thus, $\rho(.1) = 2.1$ means "the 10% power factor is 2.1." The four functions of ρ , λ are:

- A) $\frac{\rho}{\log \lambda} = C1$
- B) $\frac{1}{\log \rho} + \frac{1}{\log \lambda} = \frac{1}{C2}$
- $\frac{1 \log \rho}{(\log \lambda)^{1.5}} = C3$

D)
$$\frac{\log \rho}{(\log \lambda) \cdot 75} = C4$$

These functions are chosen on a rationale that $\log \rho$ and $\log \lambda$ are simply related. This result can be derived theoretically under certain assumptions. Since ρ is a function of λ , then Cl, C2, C3 and C4 are also functions of λ . They are tabulated in Tables C-1, etc. as "CON1, etc." Insofar as any of the functions are approximately constant, then the appropriate function gives ρ as an implicit function of λ . Thus, if C4 is observed to hold rather constant

(in a given EOC) for a whole range of λ , then log $\rho = C4(\log \lambda)^{.75}$ represents the EOC or power factor curve. Notice that for each "constant" C1, C2, C3, C4, $\partial \rho / \partial C1 > 0$, $\partial \rho / \partial C2 > 0$, etc., so that the larger the value the "better" the power factor curve. Thus, if two hazard indexes can be given the same representation with different values of C1, C2, C3, etc., then the hazard index with the larger value (of C1, etc.) is better. It can be seen that the forms log $\rho = \alpha (\log \lambda)^{\beta}$ give the best fits. Later it will be shown what values of α and β to use, and further uses for the power factor formula. (See Subsection 2.3.2.)

4.1.3 Flashing Lights Case

Table C-3 shows a comparison of the power factors for volumeonly hazard index for the flashing light warning device class versus the New Hampshire hazard index. (See Appendix B.5.2.2 for the definition of the TSC flashing lights volume-only model.) It should be remembered that all best TSC models reported in Section 4 are of logistic construction.

The volume model is roughly no better and no worse in ranking crossings than the New Hampshire formula. In fact, they come quite close to giving the same ranking. The practical distinction between the two formulas is that the TSC hazard index h yields the absolute hazard index C_1e^{2h} , as noted above. This makes the TSC volume model suitable for adding on non-volume terms. (This surprising assertion is discussed in Section 2 and in the CMR.)

Even with the non-volume terms added in, the TSC model is only slightly better than the New Hampshire model, as can be seen in Table C-4 (a maximum of 2% absolute or, in other words, "2 percentage points" difference at the 16% to 18% level).

The conclusion here, as with crossbucks, is that a slight improvement over the New Hampshire formula is possible; a twopercent reduction in accidents could certainly be well worth considering. However, these formulas are very likely not as efficient as professional judgement with an on-site inspection by local authorities. Thus, a small difference in efficiency (as

measured by the power factor), although important absolutely on a national basis, could be insignificant in the use of the formula for a pre-screening process.

The scantiness of the improvement again gives evidence that the TSC model is near the ultimate. Table 4-2 lists the power factors (they should be considered in the PF1 sense for the 1975 accidents) for the TSC model, the New Hampshire model, and, as a guess at an upper ground, for the ultimate attainable:

% Crossings	PF TSC	P F NH	PFX Ultimate (guess)
1 %	7.55	6.72	8.5
2 %	5.74	5.44	6.1
5%	4.35	4.12	4.6
10%	3.48	3.32	3.7
20%	2.62	2.59	2.7
30%	2.09	2.11	2.2

TABLE 4-2. POWER FACTORS (FLASHING LIGHTS)

Again, what the "ultimate power factor" might be is an "educated guess." At the same time, the experience of many fruitful (and fruitless) attempts would have one believe in the stability of performance of good reasonable models on these data; it would be remiss not to impart a sense of the knowledge which was gained from this experience.

A comparison of Table C-4 with Table C-2 shows that power factors for flashing lights are smaller than those for crossbucks. This says that the dependence of <u>relative</u> probability of accident on the factors in the crossing inventory is stronger for crossbucks than for flashing lights. Roughly speaking, high car and train volumes have a larger <u>relative</u> effect (multiplicative) on the probability of accident at crossbucks than at flashing lights. The absolute probability of accidents is considerably smaller at crossbuck than at flashing lights (with similar volumes), but the volume dependence is more pronounced for crossbucks.

4.1.4 Automatic Gates Case

The last warning device class for which results are given in this report is automatic gates. Here the sample size is so small that the results must be treated cautiously. There were only slightly over 700 accidents at gates, giving about 350 for both data bases. The power factors comparing the TSC volume model with the New Hampshire model are shown in Table C-5. The fact that at some places the TSC model gets up to 6 percent more of the total accidents (at 50 percent of the crossings, TSC gets 80.5 percent of the accidents, while the New Hampshire model gets 74.9 percent of the accidents) should be treated cautiously. The statistical significance is not as high as might be desired. The evidence for this is that New Hampshire catches up to TSCat at 32.5 percent, and even goes ahead by nearly a full percentage point after dropping over 4 percent of the total accidents behind at 17 percent of the crossings. However, there is evidence that the TSC formula performs better.

Adding in the non-volume variables makes the case stronger at the lower percentages (below 35% of the crossings), where the formula is expected to be of more use; this improvement is at the expense of poorer performance above 35 percent. (Compared to TSC volume only, it remains better than New Hampshire.) Table C-6 gives power factors for the TSC comprehensive models for gates. Although it gives up to 6.5 percent improvement over New Hampshire at some levels, the variability of quality of performance makes it appear that the amount of improvement in quality is very hard to estimate. The TSC model is evidently about equal to the New Hampshire model in the region of 6% to 15%, a conceivably very important region in application. It is not certain that statistically significant improvements have been obtained; in other words, the 6-percent improvement over New Hampshire in the gate category is much less significant than the 3-percent improvement in the crossbuck category.

In spite of all the variability and imprecision of the gates power factors (because of the small number of accidents), Table 4-3 lists the following estimated power factors.

% Crossings	TSC Power Factor	NH Power Factor	Ultimate PFX
. 2	3.8-4.8	3.4-3.6	5.5
5	3.4	2.8	4.0
10	2.7	2.7	3.0
20	2.2	2.1	2.5
30	1.9	1.9	2.0
	<u> </u>		

TABLE 4-3. ESTIMATED POWER FACTORS (GATES)

4.1.5 An EOC over the Full Flashing Lights Data Base

All the power factor tables (EOCs) presented in Appendix C were calculated on random subsamples of the data base (crossing inventory + accident file). One EOC was calculated on the basis of all relevant data. This was primarily to exhibit the consistency of the sampled data EOCs. Table C-7 shows an EOC for the New Hampshire model for flashing lights. It is calculated on the basis of all public flashing light crossings and all 1975 public flashing light crossing accidents. Table C-3 gives the same information for a sampled data base containing 50 percent of the accidents and a sampling of the crossings. The agreement is quite satisfactory. Since Table C-7 is based on 2,650 accidents and Table C-3 on 1,324 accidents, one can see that Table C-7 should have approximately twice as many (plus maybe 1 or 2) accidents at each level, as has Table C-3. Table C-3 has 198 accidents (for New Hampshire) at the 3-percent level, while Table C-7 has 392 accidents (2 x 198 = 396). The difference of 4 accidents out of 392is quite small.

At the 10-percent level, 893 accidents in Table C-7 compares with 2 x 440 = 880 accidents in Table C-3. The difference of 13 is small compared to 893. In general, the difference between the accidents in Table C-7 and twice those in Table C-3 could reach as high as 60 without there being a significant difference. The difference doesn't reach this size, and in general there is no evidence for any statistical difference between the two.

In Table C-7, extra data are given for interpreting equation

 $\log \omega = \omega (1 + \frac{\lambda}{2})^3$.

Values of $\alpha = \frac{\log \rho}{(\log \lambda)^{\beta}}$ for the values of β :

 β = .65, .70, .75 and .80 are given in the columns headed **.65, **.70, etc.

The next-to-last column is $\Delta \log \log \rho / \Delta \log \log \lambda$ figured over 1/2 percent intervals in λ , while the last column gives averages over groups of seven of the latter quantity. It can be seen that for the New Hampshire flashing lights case, β ranges from a stable 0.7 to a stable 0.8. Therefore, using $\beta = 0.7$ (as in Figure E-7) can be expected to underestimate f for values of $\lambda \ge 100\% > 25\%$, but should otherwise do rather well. (See Table 2-2.)

4.1.6 Evaluation of Other Previously Proposed Hazard Indexes

Continuing with the comparison of relative hazard indexes, the classical Peabody-Dimmick formula and the Coleman-Stewart formulas will be considered. As with the comparison of the TSC formulas with the New Hampshire model in Subsections 4.1.2-4.1.4, the comparisons in this sections will all be of the EOC type; thus, the models will be compared as relative hazard indexes, i.e., with respect to their ability to rank crossings according to hazard. Tables C-8 to C-13 present comparisons between Coleman-Stewart models and the New Hampshire formula, and between the Peabody-Dimmick formula and the New Hampshire formula for the three warning device classes considered. Comparisons with the TSC models are omitted, as the TSC comprehensive models outperform the Coleman-Stewart models tested here (those from Reference 2 for the three warning device classes) and the Peabody-Dimmick formula as well as the New Hampshire formula for all three warning device classes. Further, the New Hampshire formula is, for the most part, somewhat better than the other two models as measured on these data and exhibited in Tables C-8 to C-13.
The reason why the New Hampshire model is superior to the Peabody-Dimmick is probably that the train volume appears to a lower power than car volume in the Peabody-Dimmick formula, while the evidence indicates that the train volume should appear to a higher power.

The superiority of the New Hampshire formula over the Coleman-Stewart models is difficult to explain. It may be related to Coleman-Stewart's use of four distinct formulas to produce one relative hazard index. Thus, if the normalization, i.e., scaling factor, is slightly off for one formula relative to another, the performance as a relative hazard index is thrown off. To put it differently, since a different formula is used for each state of the urban/rural and single/multiple track variable, there is a need for the formulas to perform accurately as absolute hazard index as a relative hazard index. Accuracy of an absolute hazard index is difficult to achieve (see Section 4.3), and so the effectiveness as a relative hazard index is impaired.

Earlier in the project, when some models were tested which were constructed by Coleman-Stewart and which contained the urban/rural and single/multiple track variables directly in the formula, the hazard indexes appeared to perform somewhat better than the New Hampshire model, although somewhat poorer than the TSC models. Such results cannot be reproduced and included here because of time constraints. The models were, as has been indicated, no better than the TSC models, and their form aided in suggesting forms for the optimum TSC models.

In summary, Tables C-8 through C-13 indicate that the New Hampshire formula provides a relative hazard index superior to that of the other two families of previous models. It follows, then, that the TSC models, since they have already been described as superior in accuracy to the New Hampshire model, are likewise superior to the other two models.

Figures E-9 through E-11 of Appendix E show graphically the effectiveness of the TSC, Coleman-Stewart, and New Hampshire

models in terms of relative hazard indexes, which, in turn, are manifested as respective EOCs in the figures.

The following points are pertinent to the definition of the formulas:

- 1. The Peabody-Dimmick formula, as it occurs in the original article, amounts to a complex function of $C^{\cdot 17} \cdot T^{\cdot 15}$; but since this complex function is strictly increasing (monotonic), $C^{\cdot 17} \cdot T^{\cdot 15}$ is equivalent to the complete Peabody-Dimmick formula for use as a <u>relative</u> hazard index (such as the comparison in this subsection relates to).
- The New Hampshire formula has a warning device class factor, but this can be ignored here, as the comparisons are limited to single warning device classes.
- 3. The Coleman-Stewart formulas are, strictly speaking, undefined if either C or T is zero. This is resolved here by taking C = 1/2 and T = 1/2 respectively in the zero cases. (Actually, T = 0 represents the case when the inventory form had a zero for each of items 24, 26, 28, and 30 of Table F-2.)

4.2 DISCUSSION OF TSC COMPREHENSIVE MODELS AND VOLUME MODELS

4.2.1 Crossbucks

The volume model requires little discussion (see Appendix B.5.1.1 for definition). It provides a ranking of the crossings which differs little from the New Hampshire model, and in the effectiveness or predictive value of the rankings, differs by still less. The chief purpose of the complexity of the volume model is to ensure that the hazard index becomes an absolute hazard index. It is noted in Section 2 that this is a necessary property for discriminants to ensure their optimal construction. In other words, unless a good estimate of the actual probability can be worked out, it is not possible to construct the best discriminant (hazard index) even from the relative point of view.

Since some of the terms in the hazard index have negative coefficients, it is good to check that the hazard index is an increasing function of the car and train volume over a sufficiently wide range. In this connection, $\partial h_v / \partial c < 0$, for all values of c, but $\partial h_v / \partial T > 0$ only when 1.158 - 2. * .2212 $\log_{10}(T+1) < 0$, or T > 414. So, if there are more than 414 trains per day, the hazard index starts to decrease with increasing train volume, but not before. The decrease is very small; for example, doubling the number of trains (to 828) decreases the probability (frequency) estimates by less than 1 percent of its value, i.e., the relative decrease is 1%. Of course, the estimates are not valid at these very high train volumes. A separate analysis would be required to estimate the dependence of hazard on very high train volumes. The point of this discussion is to verify that the negative coefficient of $\left[\log_{10}(T+1)\right]^2$ does not result in a decrease in estimated hazard at ordinary train volumes. The non-volume terms that were selected by the selection regression (and not dropped later due to shrinking contribution) are more notable.

Besides the volume term there are (see B.5.1.3):

- The logarithm of the number of day through-trains (plus 1): (positive coefficient).
- 2. The number of main tracks: (positive coefficient).
- 3. Is the highway paved? (positive coefficient).
- 4. Population (see Appendixes B and D for definition of the variable): (positive coefficient).
- 5. Function class of the road (see Appendix F for definition of this variable): (negative coefficient).

As stated in Section 1, the goal of the study was predictive effectiveness and not identification of the causes of accidents. Therefore, caution must be exercised in interpreting these terms as <u>causally</u> connected with accidents. However, the terms are all intuitively reasonable as contributors to accident probability (taking into account the sign of the coefficient).

4 - 1 3

The positive connection of accidents to day through-trains is perhaps the most obvious. Since "volume" variables are based on <u>total</u> average vehicular traffic and <u>total</u> average train traffic, it does not take into account extra exposure of day trains to higher day traffic; therefore, an extra term for day trains is to be expected.

The positive term for number of main tracks is, likewise, evident, especially since many accidents occur involving slow or stalled vehicles, and some involve vehicles not seeing trains masked by other trains.

The positive correlation with the highway being paved would seem harder to explain. Perhaps this ends up being a proxy variable for vehicle speed. Because of the volume optimization and presence of track and lane variables in the variable pool, it should be no more than a partial proxy for them (the same holds for functional class). However, a road that is not paved would very likely carry cognizant local traffic at lower speeds than would a paved road with the same vehicular volume.

The population variable is also tricky -- the higher the population, the higher the risk. This is likely tied up with the correlation of sight distance with population. One finds that population is a better predictor than urban/rural, so population is probably a proxy for urban/rural, even though urban/rural was part of the variable pool available to the stepwise regression. Population is probably also a proxy for percent familiar drivers. The demographic implications of this variable could provide endless speculation. Functional class is perhaps a proxy for lanes, vehicle speed, percent familiar drivers, etc.

4.2.2 Flashing Lights

The flashing lights comprehensive model is given in B.5.2.2. It involves exactly the same non-volume variables as crossbucks and with the same signs for the coefficients, except for the added term "number of highway lanes" replacing the "highway paved" variable. This is very reasonable on two counts:

- a. The more lanes, the more likely the driver is not to notice the flashing lights.
- b. In the case of flashing lights, the variable "highway paved" is not very useful, since it almost always has the same value. Number of highway lanes then becomes a reasonable proxy for the same set of conditions.

4.2.3 Automatic Gates

The automatic gates comprehensive model is contained in B.5.3.2. In this case the smaller sample size supports the incorporation of fewer variables into the model. The selection regression used to select the variables for the automatic gate model is given in Appendix D, Table D-13. It is interesting to examine the regression at step 4, step 6, and again at step 10. At step 4 it can be seen that the first variable chosen (at step 1) was "presence of railroad advanced warning signs." This enters with a negative sign. In other words, if the advance warning signs are present, the crossing is more dangerous. This class of phenomenon may be called the warning device level paradox (or perhaps the Coleman-Stewart paradox).

This paradoxical phenomenon is fairly widespread in safety analysis of this type, and bears some discussion. The characteristic of interest (in this instance and in certain others) is that the "higher warning device level" crossing is more hazardous than the "lower warning device level" crossing, even accounting for other factors. The warning device level variable has become a proxy for "local judgment." In other words, if the higher warning device level is present, this reflects a judgment on the part of someone that the crossing warrants the warning equipment, perhaps based on factors not in the data base used here, including past accident history. The other element necessary to produce the paradoxical situation is that, on the average, the effectiveness of the extra warning device does not offset the increased hazard implied by its presence. In the case of the "RR advance warning sign," this is especially likely to be true, since it presumably is of low effectiveness but is also relatively inexpensive.

4 - 1 5

Coleman-Stewart noted the paradox (if not this explanation) in connection with flashing lights and gates under certain conditions and a similar anomaly was noted in the California report (Reference 1). The variable "RR advance warning signs present" is not included here mostly because of the odd nature of its promised contribution. If more time had permitted, it would have been included in a final model, since it does add to predictive power, and that is what is desired. In general, a paradoxical variable is usually a proxy for some other variable, so that including it will improve predictability; however, it can also either improve or reduce one's ability to give a causal interpretation of the model.

Step 6 of Table D-13 shows that a number of variables have entered the regression. "Number of lanes" and "number of main tracks" were chosen because of their relatively large t values. The variable $\left[\log(C+1)\right]^2$ could have been included, but since it is a volume variable and, therefore, only adjusting for the addition of the other variables (not all of which would be kept), and since the sample size is small (supporting few variables), it was, decided to keep only the tracks and lanes variables and the "variable" 1, i.e., a constant. The constant has already been chosen by the volume regression, but with only two other non-volume variables, the inclusion of a constant term is surely advisable. (Many "blind alleys" and abandoned attempts have shown the advantages and disadvantages of including the constant term.) It is probably always advisable to include it on theoretical grounds, and it was left out of the models only when the power factor showed it did not help. Looking at step 10, one sees that the constant "1" entered only on the 8th step. Some of the t values of step 10 are so small that interpretation is problematical.

The positive dependence on night trains need not be interpreted, even though it is contrary to the positive dependence on day trains for crossbucks and flashing lights. All variables which entered those models had much larger t values in the selection regression.

Next to be considered is what did not show up in the models. The gate selection regression variable pool (Table D-13) shows the

most important ones. Considerable experience had led to the dropping of other variables previously as especially unpromising. "Crossing angle" is a notable one. Nearly every selection regression contained that variable to no avail. The variable was just not pertinent to any of them. However, it should have been included in the variable pool for gates in spite of its lack of usefulness for crossbucks. But this oversight is unlikely to be of any significance.

Of the variables shown in Table D-13, "maximum train speed" (variable 10) was an example of a variable that was carried in virtually every selection regression to practically no avail. It never entered the gate selection regression shown in the table except at a very late step with completely insignificant t value. It entered selection regressions for flashing lights and crossbucks, but at late steps, with too small t values to justify keeping the variable for prediction purposes, but probably large enough to prove its significance. It entered with a positive coefficient. Apparently, "max speed," which is better than the other two variables representing train speed, is of ambiguous value for determining accident hazard. On the one hand, the faster the train, the more likely for the train to strike the car because of lack of warning (of more importance for crossbucks and flashing lights). On the other hand, the slower the train, the more likely the car is to strike the train. (Notice also that the train speed is a proxy for number of cars in the train.) The two effects evidently cancel in the case of gates and partially cancel in the case of lower warning device classes, with slight tipping of the balance to higher hazard for higher speed. If accident severity or just fatal accidents had been considered instead of all accidents, the finding might have been completely contrary.

It should be noted that the variable "urban/rural" was forced out of the regression in Table D-13. This variable was always passed up in favor of "population" in the case of selection regressions for crossbucks and gates. Various indications led to <u>suppressing</u> it from this regression since it had entered another gate selection regression (see Section 3), and since the other

variables seemed to perform better in its absence. This kind of judgment sometimes led to choices that were not dictated by the selection regression, but as far as possible the EOC was, as noted in Section 2, the final arbiter.

4.3 EXPECTED ACCIDENT FREQUENCY PLOTS

It has been noted that there is need in some contexts for an <u>absolute</u> as well as a relative hazard index. An absolute hazard index is one which is proportional to the expected frequency of accidents each year at the given crossing. It has been indicated that the expected accident frequency can be estimated in a number of ways. In Appendix E plots are presented which show a comparison of the expected accident frequency as computed four ways:

a.
$$f_1 = r \frac{M}{N} \rho \left(1 + \frac{\log \rho}{\log \lambda} \beta \right)$$

The second term within the parentheses is considered negative. (See Table 2-2 and Section 2.3.)

b.
$$f_2 = \frac{M}{N} \frac{\Delta \frac{6}{6}}{\Delta \frac{6}{6}} \frac{Accidents}{Crossings}$$

where \triangle % accidents is the percentage of all accidents occurring in a 1-percent interval of crossings:

$$\triangle$$
 % Crossings = 1%
c. $f_3 = C_1 e^{2h}TSC$

where h_{TSC} is any one of the TSC hazard indexes. The formula for f_3 is given in Section 2.3.1. The value for h_{TSC} can be obtained from any of the following formulas, depending on warning device class and on whether the non-volume variables are to be used:

B.5.1.2	B.5.2.1	B.5.3.1
B.5.1.3	B.5.2.2	B.5.3.2

The constant C_1 is determined by normalization, so that the cumulative number of (predicted) accidents over the 50 percent most hazardous crossings agrees with observation. Alternatively, the values of C_1 given in Section 2.3.1 can be used.

d.
$$f_4 = k e^{h_{cs}}$$

where h_{cs} is the appropriate Coleman-Stewart model (see Appendix B) and k is determined as in c. above by normalization.

The quantity f_2 (see b. above) is the observed empirical frequency of accidents in a given interval (of ranked crossings). Consequently, the other formulas are especially to be compared to this one. When the estimate provided by f_2 jumps around a lot, the other frequency estimates, which are strictly decreasing, are accurate if the "empirical" estimates average around the "theoretical" estimates. (By empirical f_2 is meant, and by theoretical f_1 , f_3 , and f_4 are referred to.) Also, when any two of the estimates tend to agree, this is confirmation for both estimates, since they are derived and calculated by quite different methods.

For example, Figure E-1 of Appendix E compares the frequency estimates for the TSC comprehensive crossbucks model, as given by formulas f_1 and f_3 . The solid line represents f_3 and the X's represent f_1 . It is not really possible to say which gives the better estimate (indicated below will be how this can be resolved with a little more analysis). It appears that the agreement is rather good, especially in the 5 percent to 30 percent region.

Figure E-2 is for the same hazard index (TSC comprehensive cross-bucks). The formulas represented are: f_3 by the solid line and f_2 by the X's. Of course, f_2 estimates the "true" values in the sense of unbiased but it is quite noisy.* Given that f_2 is unbiased, it seems that f_3 is not underestimating out near 50 percent and is as good as the eye can distinguish for low percentage

[&]quot;'Noisy" = "has large random component." This quality is evident in the figure.

(high hazard) crossings. It may be assumed that f_3 gives rather good probability estimates.

Figure E-4 provides the same comparison for the TSC comprehensive flashing lights model. The agreement is as good as the eye can suggest for a strictly decreasing smooth function (except possibly below 5%).

Figure E-5 provides the same comparison $(f_2 vs f_3)$ for the TSC comprehensive automatic gates model. The large scatter is consistent with the findings of large variability of estimates in the automatic gates case. The agreement is difficult to assess.

Figure E-7 is the f_1 , f_2 comparison for the New Hampshire model in the flashing lights case. In earlier work it was shown that the New Hampshire formula H=CT does not provide an absolute hazard index directly, so something like an f_1 estimate is necessary. In Figure E-7, f_1 and f_2 are represented by the solid line and X's respectively. The agreement is about as good as the eye can suggest (β was taken to be 0.7).

Figure E-8 is also for the flashing light case, but compares f_1 and f_2 applied to the TSC model. (In f_1 , β was set to 0.7 again.) The agreement is as good as the eye can tell. Referring to Figure E-4, it can be seen that f_3 perhaps overestimates for low percentages. Figure E-8 shows that f_1 estimates smaller values in the same region. In this case, f_1 may be a better frequency estimator than f_3 .

Figure E-3 is for the Coleman-Stewart crossbucks model, and Figure E-6 represents the Coleman-Stewart flashing lights model. Note that in the crossbucks model the Coleman-Stewart formula, normalized to this data base, fits the empirical frequency points relatively well.

It might appear that the Coleman-Stewart crossbucks formula is as good an absolute hazard index as the TSC (comprehensive) crossbuck formula. However, the closeness with which the curve matches the empirical frequencies is not the only test. The other consideration is how much of a differential spread in

frequencies the formula creates. Thus, a formula that simply assigned the same average accident frequency to all crossings would match the empirical value almost exactly, but would be useless as an absolute hazard index, since it contains almost zero information. The TSC crossbucks formula results in more of a spread, as evidenced by a more sharply rising curve (compare Figure E-2 with Figure E-3), and thus is apparently more useful as an absolute hazard index. The sharper frequency curve is reflected directly in the fact that the TSC formula has higher power factors.

The Coleman-Stewart flashing lights model (Figure E-6) has a high spread of empirical frequencies about the formula curve. This is probably partially due to a mismatched combination of the four formulas which make up the Coleman-Stewart model for this warning device class*, and thus gives added evidence that separate formulas may be giving trouble when used to compare crossings. The use of separate (TSC) formulas for separate warning device classes is all right under either of two circumstances:

- 1. No cross comparison between warning device classes is made.
- The crossings are compared over the whole population on which the hazard indexes are constructed and calibrated (the entire United States).

In other cases, a single combined formula for all warning device classes might be considered (with warning device class as a variable). Such a formula has not yet been constructed (using our technique).

The Coleman-Stewart flashing lights formula also results in a less sharply increasing curve than the TSC flashing lights formula.

In general, the frequency estimating formulas f_1 and f_3 are fairly accurate. Formula f_3 is probably good within \pm 10 percent for the flashing lights and crossbuck cases. It is difficult to evaluate these frequency estimates. A careful analysis using the

^{*}The four formulas involved for each warning device class are shown in Appendix B, Section B.1. The fact that distinct formulas are used to compare distinct crossings is the issue in point.

exact formula

$$f = \frac{M}{N} \rho \left(1 + \frac{\log \rho}{\log \lambda} \frac{d \log \log \rho}{d \log \log \lambda} \right)$$

(see Subsection 2.3.2) could be undertaken with the data available. The assumption above has been that

ê

$$\beta = \frac{d \log \log \rho}{d \log \log \lambda}$$

is a constant. The other definition of β is the value that makes

$$\frac{\log \rho}{(\log \lambda)^{\beta}} = \alpha$$

stay constant. In general, β will not be constant, but a function of λ . β may be determined by best fit locally of Δ log log $\rho =$ $\beta \Delta$ log log λ or, according to an alternative hypothesis, Δ log log $\rho = r \Delta$ log log $\lambda + s \Delta$ log λ for which $\beta = r + s \log \lambda$. β or r and s can be fit locally over several percent (to smooth out the noise). Although there was not time to do this analysis for this report, it can nevertheless be surmised that the estimated accident frequencies already reported will be useful even if estimated to within about 10 percent.

It seems almost unreasonable to predict these rare events more accurately. Furthermore, it might be suggested that most selection processes, whether they be real or hypothetical for analysis purposes, can be based on power factors rather than on probabilities (frequencies). It should be remembered that when power factor ρ is associated with a hazard index value h (assuming λ is known as well, i.e., h selects λ x 100% of the crossings), then one has the answer to the question of the number of accidents that are to be expected in the next year at the λ x N crossings with this or greater hazard index (in the given warning device class). On the other hand, if "f," is known one has the answer to the question of the number of accidents that can be expected at a crossing with this value of "h" in one year. It is difficult to know what the comparison is of the errors in ρ with those in f; the error in ρ may have been overestimated or underestimated in f.

In general, f estimates the derivative of a random function and, therefore, cannot be known as well as ρ , which is based on cumulative quantities only.

4.4 SUMMARY OF RESULTS

With a data base of unusually large size and degree of completeness, and with the statistical tools for constructing, testing, evaluating, and exhibiting properties of hazard indexes, a number of things have been quantitatively demonstrated about hazard indexes for railroad crossing accidents that could not be done under other circumstances.

Extensive experimentation has been done, on the basis of which three comprehensive models are offered; these models are also : tested and compared in performance to other hazard indexes on the overall United States accident experience for 1975. Extensive tests and analyses have been performed to provide clear and specific information on what these hazard indexes indicate; in this regard, some similar analyses were performed on earlier simpler hazard indexes.

It has been shown by extensive experimentation and tests that simple volume-dependent formulas appear to have 90-95 percent of the predictive power of more complex formulas. Of volume-only formulas, the New Hampshire hazard index for a given warning device class gives nearly as good a ranking as an optimized formula. However, the straight formula H = k CT should not be used to estimate accident frequency. In this regard, Table 2-2 and the discussion in Subsection 2.3.2 show how to convert a value of CT (average daily vehicular volume times average daily train volume) into a power factor and also an expected accident frequency.

The new formulas that are derived are, in certain respects as described above, more selective than the New Hampshire formula, even though the difference, while statistically significant, may not be considered large enough in certain applications to forgo the simplicity of the New Hampshire formula. If the TSC comprehensive formulas are used, they too may be converted into expected frequency of accident by simply forming C_1e^{2h} . The value of

h x 10,000 is listed under the column HAZARD INDEX, TSC in the EOCs (Appendix C); the New Hampshire hazard index value given is simply CT.

Finally, it has been shown explicitly how the power factor and expected frequency of accidents change and what values they take (for the three warning device classes -- crossbucks, flashing lights, and gates) when a given percentage of the most hazardous crossings is chosen. This information is in the EOC tables and in the fitted formulas relating f and ρ to λ .

The techniques used have been exceptionally effective in illuminating the quantitative aspects of hazard as predicted by simple quantitative characteristics. However, there has not been time to realize all of the intended applications of these constructions and tests. Indeed, a very major product of this project consists of the tools and techniques and procedures which were found useful. Further use of these tools and similar techniques will no doubt sharpen the picture that has been produced.

In the near future hazard indexes that use accident history of the individual crossing to help determine hazard are expected to be developed. (See Appendix H for details.) Preliminary results indicate that accident history can be of great predictive value when combined with the other factors considered in this report.

Finally, the reader or the user of information in this report must be cautioned that the work presented here is more of the nature of "experimentation" than of "production." A great deal of care has gone into ensuring the accuracy of the results shown, but all the data and formulas should be considered subject to refinement as more experience and data are gained.

5. SOME PRACTICAL CONSIDERATIONS

The practical worker is faced with two questions:

a. Which hazard index to use?

b. How is it to be used?

5.1 SELECTING A HAZARD INDEX

The informal recommendations given in this section will not substitute for a more thorough appraisal based on a reading of the whole report, but may help get the worker started on the task.

For choosing a hazard index, one needs to ask the following questions:

- 1. What information is available for the construction of the hazard index?
- 2. How much complexity in the formula is feasible?
- 3. How important is accuracy?
- 4. Is an absolute hazard index (frequency of accidents per year) necessary, or is a relative hazard index sufficient (for ranking crossings by relative hazard)?

Now one needs to focus on a particular warning device class (crossbucks, flashing lights, or gates). Suppose that only volume information is known, i.e., average daily number of vehicles over the road and average daily number of train movements. The simple New Hampshire formula is available for relative ranking within a warning device class, but if an absolute hazard index is needed, or if the moderate complexity of the TSC volume formulas is not a sufficient drawback, then a TSC volume formula is to be used. So, with volume-only information, if (a) an absolute hazard index is indispensable, such as for an absolute comparison of hazard between warning device classes or for use in a cost/benefit ratio, or (b) the complexity of the TSC formulas is not considered a serious drawback, i.e., the calculations can be done with ease, then it is necessary to use one of the TSC best volume formulas as follows: for crossbucks one uses B.5.1.1 (Appendix B); for

flashing lights one uses B.5.2.1; and for automatic gates one uses B.5.3.1. If, on the other hand, a simple formula is desired, and accuracy differences within a small percentage range are tolerable, and an absolute hazard index is in no way necessary, then one uses the simple New Hampshire formula.

Now, suppose other information is available besides vehicular volume and train volume. In this case, one has to check the relevant TSC best comprehensive formulas: for crossbucks, B.5.1.3; for flashing lights, B.5.2.2; and for automatic gates, B.5.3.2. One sees if all the information required for these formulas is available; for example, if the warning device class is gates, then the number of highway lanes must be known for each crossing. This information resides in the FRA data base (see Sections 1 and 3 and Appendix F), and will be available from that source if from no other. If all the information required by these formulas is available, then use of these foumulas is suggested. If not, then one needs to decide whether a volume-only formula is satisfactory; for this decision, the discussion in Section 4 concerning the EOC tables of Appendix C may be helpful. In general, use of the comprehensive formulas when one or more data items are missing for each crossing, or for most crossings, is not to be recommended. At the same time, there is no evidence that a reasonable attempt along these lines is sure to be unsuccessful.

If there are some data items (quantified crossing features) available on the crossings to which the hazard index is to be applied but not represented in the FRA data base, e.g., "clear sight distance down track," then the question as to how these items should be treated has, of course, not been directly answered. But after reading this report, the user may be reluctant to assign a large effect to any additional non-volume variables (except accident history, considered below). It should be remembered also that any additional variable will, in general, already have been partially accounted for by other variables acting as proxy.

This report has stressed that accident history at particular crossings is probably of great importance if available. We are

not yet ready to report on hazard indexes involving accident history, but Appendix H shows how they can be constructed by use of the same data base of accidents and crossings.

5.2 USE OF THE HAZARD INDEX

Next to be considered is the practical use of the hazard index. It seems that this always involves carrying out some form of the following basic procedure:

- 1. Rank all the crossings under consideration according to the value of the hazard index. (This ranking is probably most reliable if all the crossings of only one warning device class are considered at one time. If crossings of two or more warning device classes are to be ranked together, i.e., interspersed, then it is essential that an absolute hazard index proportional to expected accident frequency be used.)
- Select from this ordered list of crossings under consideration a specific number or a specific percentage (proportion) for some action, e.g., improvement of crossing warning equipment.

Relative or Absolute Hazard Index

Now, to convert a relative hazard index to an expected frequency of crossing (or an absolute hazard index), it is necessary to employ some transformation, and this is simplest and probably most accurate when one of the TSC hazard indexes is used. In this case, one simply forms $C_1 e^{2h}$, where h = HI, the hazard index given by the desired formula chosen from the following group in Appendix B:

- 1. Crossbucks-comprehensive: B.5.1.3.
- 2. Crossbucks-volume-only: B.5.1.2.
- 3. Flashing lights-comprehensive: B.5.2.2.
- 4. Flashing lights-volume-only: B.5.2.1.
- 5. Automatic gates-comprehensive: B.5.3.2.

6. Automatic gates-volume-only: B.5.2.1. If an absolute hazard function involving the New Hampshire formula is desired, the transformation is discussed in Subsection 2.3.2, but, as already noted, when an absolute hazard function is to be used, a TSC form is probably best.

Once a portion of the crossings has been chosen as the most hazardous, some measure of the expected effectiveness of the procedure may be desired. As noted throughout this report (see especially Section 2.1), the EOC and the power factor are the recommended measures of performance. Thus, if 15 percent of the crossings have been chosen using the TSC comprehensive crossbucks formula, than, as has been seen, on a national basis (see Table C-2) 51.3 percent of the accidents will occur at these crossings, giving a power factor of 3.42.

Now, if the formula is used in a certain locality, say a certain state, it is to be expected that various relationships that are observed nationally may, to some extent, not hold true. Logically, one would expect that the most invariant quality of a hazard index would be its general goodness as a relative hazard index -- not necessarily the specific measures of this quality of performance, but the fact that it is a good relative hazard index when compared with the performance of other hazard indexes. Next in variability would be its quality as an absolute hazard index, which might be a little more variable. Following in variability the plain absolute hazard index, which is, by definition, only proportional to expected accident frequency, would be the expected accident frequency itself, which might require further scaling. Last, and most variable, would be the EOC curve itself -- the power factors, etc. This would be more variable, because it would reflect local variability in the crossing characteristics themselves, not just the relation of these characteristics to accident frequency. Thus, if 15 percent of a local population of crossings is chosen according to the TSC comprehensive crossbucks model, it can be expected that a statistical deviation from the national average, 51.3 percent of the accidents in that local population, will occur at those crossings.

If one refers to Table C-2, one sees that the value of HI for the 15 percent most hazardous crossings nationally is greater than, or equal to, -0.8585, which represents the HI of the least hazardous crossings of the 15-percent chosen as the most hazardous. The least hazardous crossing of the 15 percent of the crossings chosen locally could be greater than, or less than, this value. It is suggested here that the nationally determined power factor for the actual hazard index of the least hazardous crossing in the set chosen should be used for reference to the EOC tables rather than the percentage of the local population selected for the hazardous crossings group.

By way of illustration, it is assumed that 15 percent of the crossbuck crossings have been selected using the TSC crossbucks comprehensive formula, and that the hazard index of the least hazardous crossing in the 15-percent group is HI = - 079 (from equation B.5.1.3). This HI is then referred to Table C-2 (first multiplying by 10,000 as noted in paragraph C.3 of Appendix C), and it is seen that -7900 corresponds to the national 12.5 percent point, that is, to a power factor of 3.67 rather than to the 15-percent power factor of 3.4. This small difference in power factor is not very significant, and statistical fluctuation will probably be larger. When the local estimate is close to the national average, confidence in calculation based on national statistics may have been increased.

The practical significance of a power factor of 3.6 for 15 percent of the crossings selected can be explained as follows. If all accidents at those selected crossbuck crossings could be prevented, in this hypothetical case 15 percent of all local crossbuck crossings, 54 percent of all local crossbuck crossing accidents could be prevented.

When the formulas and tables of this report are used on a national basis, they are most reliable. It is in just such circumstances that expected accident frequency based on objective crossing characteristics may be most useful in large-scale cost/benefit analyses. .

APPENDIX A

THE LOGISTIC DISCRIMINANT APPROACH TO HAZARD INDEX CONSTRUCTION

First, an ordinary linear regression is run, identical to the one described in Section 2.2 except for the trivial difference that an accident is represented by +1 while a non-accident crossing is represented by -1. (In Section 2 the non-accident crossing was represented by 0.) The effect on the hazard index generated is to multiply the previous linear regression hazard index by 2 and subtract 1. That is, the present procedure produces the same hazard function except for this multiplication and subtraction. This difference is immaterial, as the hazard index so generated at this point is good as a relative hazard index only. An interative procedure is then used to find the b_k 's as follows:

One should recall that $X_{i,k}$ is the kth characteristic for the ith crossing. Thus, letting C = AADT or average daily vehicle traffic, $X_{i,k} = \log (C_i)$ would be a possibility. Equation $X_{i,k} = \log T_i$, where T_i is the average number of trains at crossing i, would be another possibility.

Several variables of this type are carried in the regression. The iteration is of the so called "fixed point" or "implicit equation" type, with each step in the iteration being the equivalent of an ordinary regression. At each step, s, of the iteration, there is an estimate of b_k ; k = 1, K denoted by $b_k^{(s)}$. From $b_k^{(s)}$, $h_i^{(s)}$ is defined as:

$$h_{i}^{(s)} = \sum_{k} b_{k}^{(s)} X_{i,k}.$$

The variables $X_{i,k}$ and Y_i undergo a nonlinear transformation before the next regression is performed:

$$x_{i,k}^{(s)} = X_{i,k} \sqrt{U(h_i)^{(s)}}$$

 $y_i^{(s)} = Y_i \sqrt{V(h_i)^{(s)}}$

where:

$$U(h) = \operatorname{sech}^{2}(h) \left(\frac{\tanh(h)}{h}\right)$$
$$V(h) = \operatorname{sech}^{2}(h) \left(\frac{h}{\tanh(h)}\right)$$

The regression seeks an ordinary least squares solution to:

$$\hat{y}_{i}^{(s)} = \sum_{k=1}^{k} b_{k}^{(s+1)} x_{i,k}^{(s)}$$

The least squares solution $b_k^{(s+1)}$ is the one which minimizes $\sum_i (\hat{y}_i^{(s)} - y_i^{(s)})^2$. The regression is carried out by an ordinary

linear least squares package (the IBM SSP stepwise regression package was used on a DEC PDP-10) which has been modified slightly (one or two statements) so that it does not automatically correct for the mean. This is because for this type of iterated nonlinear procedure the constant term must be handled separately and explicitly; thus, $X_{i,1} - 1$ is always used. One of the basic variables must be a constant 1 (unity) with no loss of generality. (See CMR for further details.)

In this manner, $b_k^{(s)}$ leads to $b_k^{(s+1)}$; as noted already, $b_k^{(1)}$ is obtained from straight regression (no weights) which is, in effect, the same as taking $b_i^{(0)} = 0$. This will result in $h_i^{(0)} = 0$, $U[h_i^{(0)}] = 1$, and $V[h_i^{(0)}] = 1$ for all i and therefore, as stated an ordinary untransformed, unweighted linear regression.

Classical logistic discriminant analysis (Reference 6) can be shown (see CMR) to be equivalent to this procedure with U(h)and V(h) replaced by

$$U(h) = h/tanh (h)$$

and

V(h) = tanh (h)/h.

The functions $[sech(h)]^2$.h/tanh(h), $[sech(h)]^2$.tan(h)/h, h/tanh(h), and tanh(h)/h are all plotted in Figures A-1 and A-2 with

 $H(h) = \frac{1}{2} + \frac{1}{2} \tanh(h) = \frac{1}{1 + e^{-2h}}$

superimposed for reference. The classical procedure may be seen to give exceptionally heavy weight (by the factor h/tanh (h) to accident crossings for which a very low accident probability is estimated $(h\rightarrow\infty)$. In this case all low probability crossings are weighed low [by the factor $\operatorname{sech}^2(h)$] and only those crossings for which H(T) takes intermediate values are weighed high. (The intermediate values are, relativley speaking, the high hazard values.)

It is arranged that all accident probabilities fall below 1/2. This is not required of all crossings, but of all but the 1/2 percent most hazardous. This is achieved by balancing the sample with the correct number of accidents and non-accidents, and by the use of a separate constant weighing factor for accidents in some warning device classes. The reason for having all accident probabilities fall below 1/2 (they are later scaled to their real values) is that the weighting factor sech² (h) will then always be smaller for smaller H(h), i.e., smaller accident probability. This is because H(h) = 1/2 and sech² (h) takes its maximum when h = 0, while H(h) is less than 1/2 for negative values of h. Later, when the EOC curves for specific hazard functions are presented, the effect of the factor sech^2 (h) will be shown; this distinguishes the TSC technique from classical logistic discriminant analysis. As just noted, the effect is to emphasize the

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4.5

FIGURE A-2. tanh (x) and tanh $\frac{(x)}{x}$ vs. x

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performance of the hazard index for high hazard crossings; just how much will be shown when the results are presented.

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APPENDIX B

HAZARD INDEX FORMULAS-DEFINITIONS FOR REFERENCE

B.1 COLEMAN-STEWART MODEL

$$\log_{10} A = C_0 + C_1 \log_{10} C + C_2 \log_{10} T + C_3 (\log_{10} T)^2$$

where

- Λ = number of accidents (proportional to f)
- T = average daily train movements. (If T=0, use 1/2 instead for this model <u>only</u>.)

	С	Cl	C_{2}	C ₃
Category	Ŭ	1	2	5
Single-track urban				
Automatic gates	-2.17	0.16	0.96	-0.35
Flashing lights	-2.85	0.37	1.16	-0.42
Crossbucks	-2.38	0.26	0.78	-0.18
Single-track rural				
Automatic gates	-1.42	0.08	-0.15	-0.25
Flashing líghts	-3.56	0.62	0.92	-0.38
Crossbucks	-2.77	0.40	0.89	-0.29
Multiple-track_urban				
Automatic gates	-2.58	0.23	1.30	-0.42
Flashing lights	-2.50	0.36	0.68	-0.09
Crossbucks	-2.49	0.32	0.63	-0.02
Multiple-track rural				
Automatic gates	-1.63	0.22	-0.17	0.05
Flashing lights	- 2 - 7 5	0.38	1.02	-0.36
Crossbucks	-2.39	0.46	-0.50	0.53

B.2 NEW HAMPSHIRE MODEL*

 $HI = T \times C$ for Crossbucks, Flashing Lights, Gates Where: T = train movements C = vehicle movements HI = hazard indexB.3 PEABODY-DIMMICK MODEL* $HI = C^{170} \cdot T^{151}$ for Crossbucks, Flashing Lights, Gates Where: T = train movements C = vehicle movementsHI = hazard indexB.4 TSC LINEAR MODELS (CROSSBUCKS ONLY) B.4.1 Car-Train Equation (Linear) HI = -0.02022+0.01509 LOG₁₀ (C+1) LOG₁₀ (T+1) +0.01391 $[LOG_{10}^{10} (C+1)]^2$ +0.06330 LOG₁₀ (C+1) LOG₁₀ (DT+1) $-0.11039 \text{ LOG}_{10}^{-0} (\text{DT}+1)$ +0.03907 LOG₁₀ (NITE+1) Where: T = train novements C = vehicle movementsDT = day thru trains NITE = night trains HI = hazard index

*Note: Original versions of these models had additive and/or multiplicative terms, depending on warning device class. Since these formulas are used only as <u>relative</u> hazard indexes and on only one warning device class at a time, the extra factors and terms are irrelevant. Also, the Peabody-Dimmick formula as originally given was a complex function of the formula given here, but if it is to be evaluated (or to be used) as a relative hazard index, the functional transformation becomes irrelevant (see Subsection 4.1.5).

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B.4.2 Model 8C

HI = 0.11074 +1.43432 (VOL) -0.48848 (VOL x NRBY XING HWY) +0.07906 (VOL x POP) +0.01996 (MAIN TRACKS) -0.00001 (AADT/LANES+1) -0.01349 (FC) -0.01283 (NRBY XING HWY) -0.01232 LOC₁₀ (AADT+1)

Where:

VOL	=	Volume equation (HI from B.4.1)
NRBY XING HWY	=	nearby crossing highway?
POP	=	population; the tens digit of functional
		classification of road over crossing
MAIN TRACKS	=	number of main tracks
AADT	=	vehicle movements
LANES	=	number of traffic lanes
FC	=	the units digit of functional classification
		of road over crossing
ΗI	=	hazard index

B.4.3 <u>Model 8D</u>
HI = 3.67821
+0.75952 (VOL/NRBY XING HWY)
+0.06678 (VOL x POP)
-0.00194 (AADT ² /LANES)
-0.01327 (FC)
+6.80342 LOG_{10} (VOL ²)
-6.63985 (VOL) ²
-0.01282 (NRBY XING HWY)
-0.11629 (VOL x NRBY XING HWY)

.

B - 3

B.5 TSC BEST NONLINEAR (LOGISTIC)

- B.5.1 Crossbucks
- B.5.1.1 Best Volume Model-Crossbucks h = -3.0264 $+1.1580 \ LOG_{10} \ (T+1)$ $+0.48654 \ LOG_{10} \ (C+1)$ $-0.22122 \ [LOG_{10} \ (T+1)]^{2}$ $T = train movements \ (Sum of inventory items 24-30, Table F-2)$ $C = vehicle movements \ (Inventory item 81, Table F-2)$
- B.5.1.2 Refined Volume Model-Crossbucks

HI = -0.13711+0.38069 h -0.66800 h² -0.19171 h³

Where:

h = best volume model for crossbucks given in B.5.1.1
HI = hazard index

```
B.5.1.3 Comprehensive Model-Crossbucks
```

```
HI = 0.74982 HVOL
+0.19474 LOG<sub>10</sub> (DT+1)
+0.17491 MAIN TRACKS
+0.17780 HWY PAVED
+0.045405 POP
-0.13139 FC
```

Where:

- HVOL = the refined volume equation for crossbucks given in B.5.1.2
 - DT = number of day thru trains
- MAIN TRACKS = number of main tracks
 - HWY PAVED = is highway paved? (No = 0, Yes = 1, Note difference in coding from item 67, Table F-2
 - POP = population; the ten digit of functional classification of road over crossing
 - FC = the units digit of functional classification of road over crossing
 - HI = hazard index -- this is "h" in Sections 2 and 4. $C_1 e^{2h}$ is an absolute hazard index, see Subsection 2.3.1. In the tables of Appendix C, HI is multiplied by 10⁴.

B.5.2 Flashing Lights

B.5.2.1 Best Volume Model-Flashing Lights

HI = -2.8395 +0.75477 LOG_{10} (T+1) +0.083292 $[LOG_{10}(C+1)]^2$ Where: T = train movements C = vehicle movements

HI = hazard index

B.5.2.2 Comprehensive Model-Flashing Lights

HI = 1.0422HVOL +0.13737 MAIN TRACKS $-0.097584 [LOG_{10}(T+1)]^2$ +0.018064 LANES -0.036259 FC +0.12137 LOG, (DT+1) +0.018944 POP Where: HVOL = the best volume equation for flashing lights as given in B.5.2.1 MAIN TRACKS = number of main tracks T = train movements LANES = number of traffic lanes FC = the units digit of functional classification of road over crossing DT = number of day thru trains POP = population; the tens digit of functional classification of road over crossing HI = hazard index -- "h" (see B.5.1.3)

B.5.3 Automatic Gates

B.5.3.1 Best Volume Model-Gates HI = -1.9674 $+0.18621 \ LOG_{10} \ (T+1) \ LOG_{10} \ (C+1)$ Where: T = train movements C = vehicle movementsHI = hazard index

B.5.3.2 Comprehensive Model-Gates

HI = -0.83656 +0.74849 HVOL +0.19139 TRACKS +0.093829 LANES

where

HVOL = best volume model for gates as given in B.5.3.1
TRACKS = number of main tracks
LANES = number of traffic lanes
HI = hazard index -- "h" (see B.5.1.3)

B.6 TSC GRADE CROSSING HAZARD MODELS (Consolidated for Easy Reference)

The formulas in this section are essentially those in B.5 repeated. They are presented in a form more convenient to use. In addition, the overall factor for each formula has been changed slightly, reflecting the normalization appropriate for using the formulas for <u>all</u> warning device classes. (This overall factor was referred to as C_1 in B.5 and Section 2.3.1, and is now changed to the values indicated below in the expression for H, i.e., 0.389, 1.084, and 0.820. Note that the hazard indexes here are of the form H=ce^{2h}, where h is an HI from B.5.)

The models to be used for Warning Device Classes 1,2,3 and 4, are: Comprehensive Model: $H=0.389 \text{ EXP}^{2X_1}$

2HVOL₁ Volume Model: H=0.389 EXP

where

$$\begin{split} & X_1 = 0.74982 \text{HVOL}_1 + 0.19474 \ \text{LOG}_{10} (\text{DT+1}) + 0.17491 \ \text{MAIN TRACKS} \\ & + 0.17780 \ \text{HWY PAVED} + 0.045405 \ \text{POP} - 0.13139 \ \text{FC} \\ & \text{HVOL}_1 = -0.13711 + 0.38069 \text{h}_1 - 0.66800 \text{h}_1^2 - 0.19171 \text{h}_1^3 \\ & \text{h}_1 = -3.0264 + 1.1580 \ \text{LOG}_{10} (\text{T+1}) + 0.48654 \ \text{LOG}_{10} (\text{C+1}) - 0.22122 \\ & \left[\text{LOG}_{10} (\text{T+1}) \right]^2. \end{split}$$

The models to be used for Warning Device Classes 5, 6 and 7 are: 2X Comprehensive Model: H=1.084 EXP 2HVOL Volume Model: H=1.084 EXP

where

 $X_2 = 1.0422HVOL_2 + 0.13737$ MAIN TRACKS $-0.097584 [LOG_{10}(T+1]^2 + 0.018064$ LANES -0.036259 FC +0.12137 $LOG_{10}(DT+1) + 0.018944$ POP HVOL_2 = -2.8395 + 0.75477 $LOG_{10}(T+1) + 0.083292 [LOG_{10}(C+1)]^2$. The models to be used for Warning Device Class 8 are:

Comprehensive Model: H=0.820 EXP 2HVOL₃ Volume Model: H=0.820 EXP

where

 $X_3 = -0.83656 + 0.74849 \text{ HVOL}_3 + 0.19139 \text{ MAIN TRACKS}$ +0.093829 LANES HVOL₃ = -1.9674 +0.18621 LOG₁₀(T+1) LOG₁₀(C+1).

Explanation of symbols:

H = Expected number of accidents per year
T = Number of trains per day
C = Number of cars per day
DT = Number of day thru trains per day
MAIN TRACKS = Number of main tracks
HWY Paved = 1 if highway paved, 0 if not paved
POP = Population. This is the tens digit of functional classification of road over crossing.
FC = The units digit of functional classification of road over crossing
LANES = Number of traffic lanes
EXP = 2.71828.... Warning Device Class 8 = Automatic Gates

- 7 = Flashing LIght
- 6 = Highway Signals, Wigwags, or Bells
- 5 = Special Protection
- 4 = Crossbucks
- 3 = Stop Signs
- 2 = Other Signs
- 1 = None

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APPENDIX C

EOC AND POWER FACTOR TABLES

C.1 BRIEF EXPLANATION OF EOC TABLES

Most of the tables in this appendix are in the format of Tuble C-1. The information on the first page of the table us also described in Section 2.1. The columns are numbered along the bottem of the first page. Column 1 is labeled "%Xings", and should be interpreted as a percentage of the total crossings selected by the hazard index. Column 4 is labeled "% ACC", and gives the percent of accidents accumulated with the percent of crossings (for the TSC model), while column 3 labeled "CUM#ACC", gives the actual number of accidents with hazard index as great or greater than the value in column 6 ("HAZARD INDEX") in the data base used. Column 5, "POWER FACTOR", is the ratio of column 4 to column 1, i.e., the ratio of the cumulative percent of accidents to the cumulative percent of crossings. Thus, columns 1, 4, 5, and 6 can be interpreted without reference to a particular data base, and although they are calculated in a specific data base, they estimate the corresponding quantities with reference to all (crossbuck) crossings and all (crossbuck) accidents (at public crossings in the entire U.S.). Columns 1, 9, 10, and 11 give the same information for the New Hampshire formula, while column 20 gives the difference in percent accidents for the two formulas. The other columns (described in Section 2.1) refer mostly to the particular data base used, and can be used for calculating the accuracy and statistical significance of the results. Since, for the TSC model, 10 percent of the crossings correspond to 769 accidents, the 10 percent power factor has a relative standard error of $\sqrt{769}$ 769 = .036 or 3.6 percent. This is reflected by the 4 percent in Table 4-1 opposite 10 percent of the crossings. Other accuracy and significance information can be derived from these columns -especially from the t value in the last column (see Section 2.1).

C-1

C.2 LIST OF TABLES IN APPENDIX C (EOC'S AND POWER FACTORS)

<u>Table</u> :	Models Represented
C- 1	TSC (Volume only) vs. New Hampshire (NH) crossbucks
C- 2	TSC (Comprehensive) vs. NH crossbucks
C- 3 ·	TSC (Volume only) vs. NH flashing lights
C- 4	TSC (Comprehensive) vs. NH flashing lights
C- 5	TSC (Volume) vs. NH automatic gates
C- 6	TSC (Comprehensive) vs. NH automatic gates
C- 7	Special EOC and Power Factors for New Hampshire
	Flashing Lights case, Full data base
C- 8	Coleman-Stewart vs. NH crossbucks
C- 9	Coleman-Stewart vs. NH flashing lights
C-10	Coleman-Stewart vs. NH gates
C-11	Peabody-Dimmick vs. NH crossbucks
C-12	Peabody-Dimmick vs. NH flashing lights
C - 1 3	Peabody-Dimmick vs. NH automatic gates

C.3 LEGEND

This section identifies formulas from Appendix B used to provide data listed under the column heading HAZARD INDEX, column 6 of Tables C-1 through C-6. In each table, HAZARD INDEX is determined by the expression HIx10,000, where HI is given in the designated subsections of Appendix B as listed below for the respective tables.

Table	TSC Subsection
C-1	B.5.1.1
C-2	B.5.1.3
C-3	B.5.2.1
C - 4	B.5.2.2
C-5	B.5.3.1
C-6	B.5.3.2

TABLE C-1. TSC (VOLUME ONLY) VERSUS NEW HAMPSHIRE CROSSBUCKS

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														-			- X E	K / 📜	DAT		
0/			TSC NO	זפה כ	+		N I		90 T U 20	СТ						~					
10	INC	CU 3		POWER	HAZARD	LNC	CUN		PONER	 HA73RD	¥779 8	Атсн	TNC	спи	1955	L. BARCH	1 800		*		
Xing	ACC	#ACC	ACC	PACTR	INDEX	ACC	FACC	S ACC	FACTR	INDEX	SDIPP	TVAL	ATCH	BTCH	\$17C	S ACC	61CC	S LCC		ATCH	
							-									<i>x</i>		~	NULFE	1146	
0.50	86	86	4.33	8.67	-2274	83	83	4.18	8.37	72000	0.15	0.2	81	81	5	0.25	2	0.10	0.15	1.1	
1.00	50	144	7.26	7.26	-2941	52	135	6.80	6.80	47300	0.45	0.5	47	128	16	0.81	7	0.35	0.45	1.9	
1.50	45	189	9.53	6.35	-3404	54	189	9.53	6.35	34500	0.00	0.0	46	174	15	0.76	15	0.76	0.00	0.0	
2.00	65	254	12.80	6.40	-3864	56	245	12.35	6.17	27200	0.45	0.4	48	222	32	1.61	23	1.16	0.45	1.2	
2.50	47	301	15.17	6.07	-4221	56	301	15.17	6.07	22800	0.00	0.0	57	279	22	1.11	22	1.11	0.00	0.0	
3.00	47	348	17,54	5,85	-4561	40	341	17.19	5.73	19200	0.35	0.3	38	317	31	1.56	24	1.21	0.35	0.9	
3.50	50	3.88	20.06	5.73	-4843	34	375	18.90	5.40	16800	1.16	0.8	37	354	44	2.22	21	1.06	1.16	2.9	
4.00	31	429	21.62	5.41	-5096	41	416	20.97	5.24	14800	0.66	0.4	32	386	43	2.17	30	1.51	0.66	1.5	
4.30	4.)	470	23.99	5.33	-53.30	45	461	23.24	5.16	13000	0.76	0.5	45	431	45	2.27	30	1.51	0.76	1.7	
5.50	20	514	23.71	5.18	-0046	10	494	24.90	4.98	12000	1.01	0.6	38	469	4.5	2.27	25	1.26	1.01	2.4	
5.30	25	507	20.07	0.10	-5/6/	1 6	525	20.40	4.81	11000	1.61	1.0	31	500	57	2.87	25	1.26	1.61	3.5	
6.50	31	613	20.60	4.03	-6126	42	507	20.08	4.76	10000	0.76	0.4	33	533	49	2-47	34	1.71	0.76	1.6	
7-00	26	619	32.21	4 60	-6735	24	615	11 00	4.58	9150	1 21	0.6	35	568	45	2.27	23	1.16	1.11	2.7	
7.50	12	651	32.81	4.00	-6436	23	610	37 16	4.43	8000	1.21	0.7	10	595	44	2.22	20	1.01	1.21	3.0	
8-00	21	672	37.87	4.23	-6603	27	665	32.10	4.27	3000	0.00	0.4	21	610	41	2.07	28	1.41	0.66	1.6	
4.50	24	696	35.08	4 1 3	-67.65	21	6.88	77.75	4.12	6900	0,00	0.2	10	631	41	2.07	34	1.71	0.35	0.8	
9.00	36	732	36.90	4.10	-6906	22	710	15 70	3 00	6400	1 11	n 4	10	476	49	7.4/	41	2.07	0.40	0.8	
9.50	20	752	37.90	3.99	-7067	26	736	37.10	3.90	6000	0.01	0.0	19	601	50	2.01	30	1.70	1.11	2.3	
10.00	17	769	38.76	3.80	-7198	23	759	38.26	1.81	5600	0.50	0.3	22	716	53	2.72	42	2.12	0.01	1.0	
10.50	27	796	40.12	3.82	-7368	31	790	39.82	3.79	5250	0.30	0.2	26	742	54	2 77	5	2.17	0.30	0.4	
11.00	21	817	41.18	3.74	-7490	27	817	41, 18	1.74	5000	0.00	0.0	29	770	1 1 1 1 7	2 17	40	2.42	0.30	0.0	
11.50	28	845	42.55	3.70	-7647	22	839	42.79	3.68	4800	0.30	0.1	25	795	50	2 52	47	2 2 2 2	0.00	0.0	
12.00	28	873	44.00	3.67	-7764	30	869	43.80	3.65	4400	0.20	0.1	29	824	70	2.47	45	2 2 2 2	0.30	О.0 Л л	
12.50	25	898	45.26	3.62	-7886	22	891	44.91	3.59	4200	0.35	0.2	25	849	49	2.47	42	2.12	0 35	0.7	
13.00	13	916	46.17	3.55	-8001	36	927	46.72	3.59	4000	-0.55	-0.3	28	877	39	1.97	50	2.52	-0.55	-1.2	
13.50	25	941	47.43	3.51	-8183	13	940	47.38	3.51	380 0	0.05	0.0	17	894	47	2.37	46	2.32	0.05	0.1	
14.00	6	947	47.73	3.41	-8274	17	957	48.24	3.45	3600	-0.50	-0.2	9	903	44	2.22	54	2.72	-0.50	-1-0	
14.50	22	969	48_84	3.37	-8395	8	965	48.64	3.35	3500	0.20	0.1	19	922	47	2.37	43	2.17	0.20	0.4	
15.00	14	983	49.55	3.30	-8460	23	988	49.80	3.32	3240	-0.25	-0.1	15	937	46	2.32	51	2.57	-0.25	-0.5	
15.50	21	1004	50.60	3.26	-86 28	15	1003	50.55	3.26	3070	0.05	0.0	19	956	48	2.42	47	2.37	0.05	0.1	
16.00	15	1019	51.36	3.21	-8723	31	1034	52.12	3.26	3000	-0.76	-0.3	29	985	34	1.71	49	2.47	-0.76	-1.6	
16.50	10	1629	51.86	3.14	-8758	1	1035	52.17	3.16	2948	-0.30	-0.1	1	986	43	2.17	49	2.47	-0.30	-0.6	
17.00	24	1053	53.07	3.12	-8906	19	1054	53.13	3.13	2800	-0.05	-0.0	24	1010	43	2.17	44	2,22	-0.05	-0.1	
17.50	12	1065	53.68	3.07	-9000	21	1075	54.18	3.10	2600	-0.50	-0.2	10	1020	45	2.27	55	2.77	-0,50	-1.0	
18.00	26	1091	54.99	3.05	-9101	15	1090	54.94	3.05	2500	0.05	0.0	26	1046	45	2.27	44	2.22	0.05	0.1	
18.50	24	1115	56.20	3.04	-9250	23	1113	56.10	3.03	2400	0.10	0.0	20	1066	49	2.47	47	2,37	0.10	0.2	
19.00	19	1134	57.10	3.01	-9329	0	1113	56.10	2.95	2400	1.06	0.4	7	1073	61	3.07	4 D	2.02	1.06	2.1	
20.00	14	1140	54 67	2.00	-9432	17	1130	50.96	2.92	2240	0.81	0.3	16	1089	57	2.87	41	2.07	0.81	1.6	
20.50	20	1104	50.07	2.93	-9511	12	1142	57.56	2.88	2130	1.11	0.5	11	1100	64	3.23	42	2.12	1.11	2.1	
21.00	20	1104	57.00	2.91	-9033	45	1107	59.83	2.92	2000	-0.15	-0.1	40	1140	44	2.22	47	2.37	-0.15	-0.3	
21.50	11	1002	60.00	2.80	-9727	0	1187	59.83	2.85	2000	0.25	0.1	5	1145	47	2.37	42	2.12	0.25	0.5	
22 00	13	1216	61 29	2.02	-7003	U A	1105/	27.03	2.78	2000	V.81	0.3	R R	1153	50	2.52	34	1.71	0.81	1.7	
22.50	16	1210	62.10	2.19	-10017		1717	61 30	2.14	1900	1.06	0.4	7	1160	56	2.82	35	1.76	1.06	2.2	
23.00	.5	1270	62.45	2.70	- 100.00	1,	1270	61 00	2.13	1750	0.75	0.5	41	1181	51	2.57	30	1.81	0_75	1.5	
23.50	11	1250	63,00	2.68	- 10237	14	1242	62.60	2.09	1600	0.05	0.2	10	1797	42	2.12	11	1.56		1.1	
24.00	10	1260	63.51	2.65	- 10127	18	1760	63.51	2.00	1600	. 0.40	0.4	10	1213	.10 72	1 0 -	27	1.30	4.40	0.0	
24.50	22	1282	64.62	2.64	- 10 384	10	1270	64.01	2 6 1	1520	0.00	0.0	87	1220	51	7 67	ינ	1.00	0.00	1 2	
25.00	15	1297	65.37	2.61	- 10503	25	1295	65.27	2,61	1500	0_10	0.0	י רכ	1250		2.02	40 (L)	2.02	0.00	0.2	
													<u>د </u>								
1	2	て	4	5	6	7	8	q	1.0	11	12	13	1/	15	16	17	18	10	20	21	Coll
-	-	5	т	5	v	/	0	5	тU	· • •	16	тJ	7.4	r T J	10	т,	тo	1.5	20		

* See LEGEND, Section C.3 (Hazard Index = hx10,000)

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TABLE C-1. (cont.)

																	X I	3KVL -	DA	Т
%			TSC N	ODEL	4	-	N	EV HAND	SHIRE	СТ					TS	c	1	(H		
V =	INC	CON	_	POWER	HAZARD "	INC	CUN		POWER	HAZ ARD	VITH P	ATCH	INC	CŪM	LESS	BATCH	LESS	HATCH	LESS 8	INTCH
Aing	ACC	#ACC	% ACC	PACTR	INDEX	ACC	#A CC	🗶 YCC	FACTR	INDEX	SDIPP'	TVAL	NTCH	NTCH	#AtC	X ACC	₿ ACC	S ACC	SDI PP	TVAL
25 50	0	1306	65 03	7 60	106 07									_						
26.00	21	1307	66 99	2.50	-10607	16	1290	66.00	2.50	1450	0.50	0.2	3	1256	50	2.52	40	2.02	0.50	1.1
26 50	2	1376	67 34	2.5/	- 10665	13	1311	66.08	2.54	1400	0.01	0.3	19	1275	52	2.62	36	1.61	0.81	1.7
27.00	10	1346	67.84	2.51	+ 108 19	9	13.26	66.03	2.30	1350	0.96	0.4		1286	50	2.52	31	1,56	0.96	2.1
27.50	7	1351	68.20	2.48	-10918	, ,	1334	67 24	2.40	1200	0.06	0.4	10	1296	50	2.52	30	1.51	1.01	2.2
28.00	13	1366	68.85	2.46	- 10963	30	1364	68 75	2.45	1220	0.10	0.4	26	1220	51	2.5/	. 32	1.61	0.96	2.1
29.50	10	1376	69.35	2.43	- 11059	ŏ	1364	68.75	2.41	1200	0.60	0.0	20	1320	30	1.92	36	1.81	0.10	0.2
29.00	10	1386	69.86	2.41	-11173	13	1377	69.41	2.39	1120	0.45	0.2	6	1339	2 P 11 J	2.17	16	1 02	0.60	1.4
29.50	13	1399	70.51	2.39	-11267	13	1390	70.06	2.37	1080	0.45	0.2	11	1352	47 47	2 37	19	1 92	0.45	1.0
30.00	9	1408	70.97	2.37	-11316	11	1401	70.61	2.35	1020	0.35	0.1	14	1366	47	2.12	35	1.76	0.95	1.0
30.50	7	1415	71.32	2.34	-11390	37	1438	72.48	2.38	1000	~1.16	-0.4	14	1380	35	1.76	58	2.92	-1 16	-2 #
31.00	12	1427	71.93	2.32	- 11536	0	1438	72.48	2.34	1000	-0.55	-0.2	4	1384	43	2.17	54	2.72	-0.55	-1 1
31.50	5	1432	72.18	2.29	- 11615	0	1438	72.48	2.30	1000	-0.30	-0.1	1	1385	47	2.37	53	2.67	-0.30	-0.6
32.00	16	1448	72.98	2.28	-11649	19	1457	73.44	2,29	925	-0.45	-0.2	10	1395	53	2.67	62	3.13	-0.45	-0.8
32.50	19	1467	73.94	2.28	-11754	19	1476	74.40	2.29	900	-0.45	-0.2	27	1422	45	2.27	54	2.72	-0.45	-0.9
33.00	0	1467	73.94	2.24	- 11754	1	1477	74.45	2.26	890	-0.50	-0.2	0	1422	45	2.27	55	2,77	-0.50	-1.0
33.50		1476	74.40	2.22	-11851	12	1489	75.05	2.24	840	-0.66	-0.2	8	1430	46	2.32	59	2.97	-0.66	-1.3
34.00	11	1487	74.95	2.20	- 11964	27	1516	76.41	2.25	800	-1_46	-0.5	18	1448	39	1.97	68	3.43	-1.46	-2.8
34.50	14	1545	12.00	2.19	-12081	0	1516	76.41	2.21	800	-0.76	-0.3	6	1454	47	2.37	62	3.13	-0.76	-1.4
35 50	13	1520	70.30	2.10	- 12 158	ر	1519	10.56	2.19	776	-0.20	-0.1	9	1463	52	2.62	56	2.82	-0.20	-0.4
36.00	32	15.0	77 62	2.1/	- 12260		1528	77.02	2.17	750	0.00	0.0	12	1475	53	2.67	53	2.67	0.00	0.0
36.50	17	1557	78 64	2.10	- 12297	2	1560	70 77	2.10	720	-0.25	-0.1	24	1499	41	2.07	46	2.32	-0.25	-0.5
37.00	6	1563	78.79	2.13	- 12 3 0 9		1557	70.23	2.14	700	0.25	0.1	18	1517	40	2.02	35	1.76	0.25	0.6
37.50	10	1573	79.28	2.11	- 12587	7	1560	78 93	2.12	600	0.30	0.1		1521	42	2.12	36	1.81	0.30	0.7
33.00	6	1579	79.59	2.09	-12599	37	1601	80.70	2.12	600	-1 11	-0 4		1524	49	2.4/	40	2.02	0.45	1.0
39.50	10	1589	80.09	2.08	- 126 79	ō	1601	80.70	2.10	600	-0.60	-0.4	6	1551	20	1 07	50	2.92	-1.11	- 2. 3
39.00	9	1598	80.54	2.07	- 12763	ō	16.01	80.70	2.07	600	-0.15	-0.1	a a	1550	30	1 07	20	2. 32	-0.60	-1.3
39.50	6	1604	80.85	2.05	- 12871	0	1601	80.70	2.04	594	0,15	0.1	1	1560	55	2.27	4 Z U 1	2.07	-0.75	-V.J
40.00	6	1610	81.15	2.03	- 129 11	11	1612	81.25	2.03	550	-0.10	-0.0	10	1570	u 0	2.02	47	2 12	-0 10	-0.2
40.50	5	1615	81.40	2.01	-12986	6	1618	81.55	2.01	522	-0.15	-0.1	7	1577	3.8	1.92	41	2.07	-0.15	-0.3
41.00	4	1619	81.60	1.99	- 13048	23	1641	82.71	2.02	500	-1-11	-0.4	10	1587	32	1.61	54	2.72	-1.11	-2.4
41.50	6	1625	81.91	1.97	-13160	0	1641	82.71	1.99	500	-0.81	-0.3	2	1589	36	1.81	52	2.62	-0.81	-1.7
42.00	9	1634	82.36	1.96	- 13235	0	1641	82.71	1.97	500	-0.35	-0.1	5	1594	40	2.02	47	2.37	-0.35	-0.8
42.50	4	1638	82.56	1.94	- 13305	0	1641	82.71	1.95	500	-0.15	-0.1	2	1596	42	2.12	45	2.27	-0.15	-0.3
43.00	0	1638	82.56	1.92	-13347	7	1648	83.06	1.93	480	-0.50	-0.2	4	1600	38	1.92	48	2.42	-0.50	-1.1
43.50	9	1647	83.01	1.91	- 13453	4	1652	83.27	1.91	460	-0.25	-0.1	8	1608	39	1.97	44	2,22	-0.25	-0.5
hu 50	11	1652	03.27	1.89	- 13555	9	1661	83.72	1.90	450	-0.45	-0.2	7	1615	37	1.86	46	2.32	-0.45	-1.0
44.00	11	1667	03.62	1.00	-13607		1668	84.07	1.89	420	-0.25	-0.1	14	1629	34	1.71	- 39	1.97	-0.25	-0.6
45.50	- A	1675	84 UZ	1.07	-13721	23	1691	85.23	1.89	400	-1.21	-0.4	5	1634	33	1.66	57	2.87	-1.21	-2.5
46.00	Š	1680	84 68	1 84	-13861	ň	14.01	03.23	1.07	400	-0-61	-0.3	6	1640	35	1.76	51	2.57	-0.61	-1.7
46.50	17	1697	85.51	1.84	- 13966	Ň	1691	85 23	1 0 2	400	-0.55	-0.2		1642	9 E	1.92	49	2.47	-0.55	-1.2
47.00	Ó	1697	85.53	1.82	- 13966	1	1692	AS. 78	1.81	376	0.30	0.1	16	1656	19	1.97	33	1.66	0.30	0.7
47.50	13	1710	86.19	1.01	- 140 44	÷	1699	85.64	1.80	360	0.55	0.1	17	1675	29	1.9/		1./1	0.25	0.6
48.00	0	1710	86.19	1-80	- 14044	, D	1699	85 64	1.78	250	0.55	0.2		1676	30	1 76	24	1 21	0.55	1.4
48.50	1	1711	86.24	1.78	- 14 1 1 3	Š	1704	85.89	1.77	326	0.35	0.1	2	1679	دد د ۲	1 66	24	1 2 1	V.))	1.4
49.00	8	1719	86.64	1.77	- 14216	ŝ	1709	86.14	1.76	320	0.50	0.2	5	1684	33	1.76	20	1.31	0.33	1 2
49.50	8	1727	87.05	1.76	- 14250	14	1723	86.84	1.75	300	0.20	0.1	10	1694	33	1.66	29	1.46	0.20	0.5
50.00	3	1730	87.20	1.74	- 14268	0	1723	86.84	1.74	300	0.35	0.1	2	1696	34	1.71	27	1.36	0.35	0.9

*See LEGEND, Section C.3.

TABLE C-1* (cont.)

(TSC)

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%						%					
Xing	P P i	cos 1 c	ON 2	CON 3	CON 4	Xing	PF	CO8 1	CON 2	CON 3	CON 4
0.50	8,67	3.65	1.64	0.94	0.62	25.50	2.58	3.10	1.89	0.81	0.75
1.00	7.26	3,48	1.58	0.92	0.63	26.00	2.57	3.16	1.91	0.81	0.76
1.50	6.40	3.30	1.64	0.90	0.67	26.50	2.54	3-13	1.91	0_81	0.75
2.50	6.07	3.53	1.65	0.94	0.68	27.50	2.48	3.06	1.92	0.80	1.75
3.00	5.05	3.56	1.67	0.94	0.69	23.00	2.46	3.07	1.93	0.80	0.75
3.50	5.73	3.64	1.71	0.95	0.70	28.50	2.43	3.05	1.94	0.79	0.75
4,00	5 33	3.64	1.72	0.95	0.70	29.00	2.41	3.05	1.95	0.79	0.75
5.00	5.18	3.65	1.73	0.95	0.72	30.00	2.37	3.02	1,96	0.78	0.75
5.50	5.10	3.72	1.76	0.96	0.73	30.50	2.34	2.98	1.97	0.78	.75
6.00	4.89	3.64	1.74	0.95	0.73	31.00	2.32	2.99	1.98	0.78	1.75
6.50 7 00	4.60	3.53	1.73	0.94	0.73	31.50	2.29	2.94	2.00	0.77	0.75
7.50	4.38	3.43	1.69	0.92	0.72	32.50	2.28	3.06	2.02	0.78	3.75
8.00	4.23	3.37	1.68	0.91	0.72	33.00	2.24	2.96	2.02	0.77	0.75
8.50	4.13	3.34	1.67	0.90	0.72	33.50	2.22	2.95	2.03	0.76	0.75
9.00	3,99	3.41	1.69	0.90	0.73	34,60 30 50	2.20	2.96	2.04	0.76	0.75
10.00	3.80	3.29	1.69	0.89	0.72	35.00	2.18	3.04	2,08	0.76	0.75
10.50	3.82	3.31	1.70	0.39	0.73	35.50	2.17	3.07	2.09	0.76	0.75
11.00	3.74	3.28	1.70	0.89	0.73	36.00	2.16	3.10	2.11	0.76	0.76
11.50	3.70	3.32	1.73	0.89	0.74	30.50	2.10	3 15	2.13	0.76	0.76
12.50	3.62	3.38	1.74	0.89	0.74	37.50	2.13	3.16	2.16	0.76	0.76
00.د1	3.55	3.25	1.74	0.89	C.74	34.00	2,09	3.13	2.16	0.75	0.76
13.50	3.51	3.37	1.75	0.89	0.75	38.50	2.08	3.15	2.18	0.75	0.76
14_00	3-41	3.20	1.74	0.87	0.74	39.00	2.07	1.15 7.13	2.19	0.75	0.76
15.00	3.30	3.23	1.74	0.87	0.74	40.00	2.03	3.10	2.21	0.74	0.76
15.50	3.26	3.24	1.75	0.87	0.74	40.50	2.01	3.07	2.22	0.73	0.75
16.00	3.21	3.21	1.75	0.86	0.74	41.00	1.99	3.02	2.23	0.73	0.75
16.50	3.14	J.14 3 18	1.74	0.85	0.74	41.50	1.97	3.00	2.24	0.72	0.75
17.50	3.07	3.14	1.76	0.85	0.74	42.50	1.94	2.96	2.27	0.72	0.75
18.00	3.05	3.20	1.78	0.85	0.75	43.00	1.92	2.87	2.27	0.71	0.74
18.50	3.04	3.25	1.80	0.86	0.75	43.50	1.51	2.59	2.29	0.71	0.74
19.00	3.01	3.27	1.81	0.85	0.75	44.0C 44.5C	1.89	2.86	2.31	0.70	0.74
26 66	2.90	3.25	1.82	0.85	0.75	45.00	1.87	2.36	2.34	0.70	0.74
20.50	2.91	3.28	1.84	0.85	0.76	45.50	1.86	2.88	2.36	0.70	0.74
21.00	2.86	3.22	1.83	0.84	0.75	46.00	1.84	2.35	2.37	0.69	0.74
21.50	2.82	3.19	1.83	0.84	0.75	46.50	1.84	2.99	2.40	0.70	0.74
22.50	2.76	3.18	1_65	0.83	č.75	47.50	1.81	2.98	2.44	0.69	0.74
23.00	2.72	3.12	1.85	0.82	0.75	48.00	1.80	2.89	2.45	0.68	0.74
23.50	2.68	3.09	1.85	C. 82	0.75	46.50	1.78	2.81	2.46	0.68	0.73
24.00	2.65	3.06	1.85	0.87	0.75	49.00	1.76	2.84	2.48	0.67	0.74
25.00	2.61	3.13	1.89	0.62	0.75	50.00	1.74	2.81	2.52	0.67	0.73

*See Subsection 4.1.2 for discussion of CON 1-CON 4.

TABLE C-1 (cont.) (NH)

%						%					
Xing	PFI	CON 1	CON 2	CON 3	CON 4	Xing	P F	CON 1	CON 2	CON 3	CON 4
0.50	4.37	3.55	1.53	0.92	0.61	25.50	2.56	3.02	1.87	0.80	0.74
1 60	6.80	3.29	1.48	0.89	0.61	26.00	2.54	3.03	1.89	0.80	0.75
1.50	6.35	3.30	1.51	0.90	0.63	26.50	2.50	2.98	1.89	0.80	0:74
2.00	6.17	3.40	1.58	0.92	0.65	27.00	2.48	2.95	1.89	0.79	0.74
2.50	6.07	3.53	1.65	0.94	0.68	27.50	2.45	2.91	1.89	0.79	0.74
3.00	5.73	3.48	1.63	0.93	0.68	23.00	2.46	3.05	1.93	0.80	0.75
3.50	5.40	3.39	1.61	0.92	0.68	28.50	2.41	2.95	1.92	0.79	C. 74
4.00	5.24	3.41	1.63	0.92	0.69	29.00	2.39	2.96	1.93	0.78	0.74
4.50	5.16	3.49	1.67	0.93	0.70	29.50	2.37	2.97	1.95	0.78	0.74
5.00	4.98	3.46	1.66	0.93	0.71	30.00	2.35	2.96	1.96	0.78	0.74
5.50	4.81	3,43	1.66	0.92	0.71	30.50	2.38	3.19	2.00	0.79	0.76
6.00	4.76	3.51	1.69	0.93	0.72	31.30	2.34	3.09	2.00	0.78	0.75
6.50	4.58	3.44	1.68	0.92	0.72	31.50	2.30	2.99	1.99	0.78	0.75
7.00	4.43	3,38	1.67	0.91	0.71	32.00	2.29	3.07	2.01	0.78	0.75
7.50	4.29	3.32	1.65	0.90	0.71	32,50	2.29	3.15	2.04	0.78	0.76
6.00	4_19	3.31	1.66	0_90	0.72	33.03	2.26	20.L	2.03	0.77	0.75
8.50	4.08	3.27	1.65	0,90	0.71	33.50	2.24	3.07	2.00	0 70	0.75
9.00	3.98	3.23	1_65	0.89	0.71	34.00	2.20	3.23	2.03	0.77	0.76
9.50	3.90	3.23	1.66	0.89	0.72	34.50	2.21	3.13	2.00	0.74	0.76
10.00	3.83	3.22	1.66	0.88	0.72	35,00	2.19	3.00	2.00	0.76	0.75
10.50	3.79	3.26	1.68	0.89	0.72	35,30	2.17	3 15	2.07	0.76	0.76
11.00	3.14	J. 20	1.70	0.09	0.73	36.00	2.10	3 13	2 1 2	0.76	0.76
11.50	3.68	3.21	1.70	0.09	0.73	37 00	2 12	3 08	2 13	0.75	0.76
12.00	3.65	56.C	1 72	0.89	0 74	37.50	2 10	3.06	2.14	0.75	0.75
12.50	3.59	3.34	1 76	0.07	0.75	18 00	2 12	3 40	2 19	0.77	0.77
13.00	3.59	3.43	1 75	0.50	0 75	34.50	2.10	1.29	2.20	0.76	0.77
13.50	3.31	3.34	1.75	0.88	0.75	33.00	2.07	3, 19	2.20	0.75	0.76
14 50	3, 15	3.24	1.74	0.87	0.74	39,50	2.04	3.09	2.20	0.74	0.76
15.00	3, 32	3.27	1.75	0.87	0.74	40.00	2.03	3.13	2.22	0.74	0.76
15 50	3 26	3, 23	1.75	0 87	0.74	40.50	2.01	3.10	2.23	0.74	0.76
16 00	3.26	3, 32	1.78	0.87	0.75	41.00	2.02	3.30	2.26	0.74	0.76
16.50	3.16	3.19	1.75	0.86	0.74	41.50	1.99	3.20	2.27	0.74	0.76
17.00	3.13	3.19	1.76	0.86	0.74	42.00	1.97	3.10	2.27	0.73	0.75
17.50	3.10	3,21	1.78	0.86	0.75	42.50	1.95	3.00	2.27	0.72	0.75
18.00	3.05	3.19	1.78	0.85	0.74	43.00	1.93	2.99	2.29	0.72	0.75
19.50	3.03	3,24	1.80	0.85	0.75	43.SC	1.91	2.95	2.30	0.71	0.75
19.00	2.95	3.11	1.78	0.84	0.74	44.00	1.90	2.97	2.32	0.71	0.75
19.50	2.92	3.11	1.79	0.84	0.74	44.50	1.89	2.97	2.33	0.71	0.75
20.00	2.88	3.08	1.79	0.83	0.74	45.00	1.89	3.19	2.37	0.71	0.76
20.50	2.92	3.30	1.84	0.85	0.76	45.50	1.87	3.09	2.38	0.71	0.75
21.00	2.85	3.18	1.83	0.84	0.75	46.00	1.85	3.00	2.39	0.70	0.75
21.50	2.78	3.06	1.81	0.83	0.74	46.50	1.83	2.90	2.39	0.69	0.74
22.00	2.74	3.01	1.81	C.82	0.74	47.00	1.81	2.83	2.40	0.69	C • 74
22,50	2.73	3.06	1.83	0.82	0.74	47.50	1.00	2.83	2.42	0.68	0.74
23.00	2.69	3.03	1.83	0.82	0.74	48.00	1 77	2.74	2.43	0.67	0.73
23.50	2.66	3.03	1.84	0.81	0.74	40.5U 40.5U	1 76	2.74	2.40	0.0/	0.73
24.00	2.65	3.06	1.85	0.81	0.75	49.00	1 75	2.80	2.4D 7 40	0.67	0.73
24.50	2.61	5.03	1.86	0.81	0.74	50 00	1 7/1	2.00	2.47	0.67	0.73
25.00	2.61	3.12	1.88	0.82	0.75	30.00	1.4.7.9	2. 1	4.21	0.00	0.75

*See Subsection 4.1.2 for discussion of CON 1- CON 4.

TABLE C-2. TSC (COMPREHENSIVE) VERSUS NEW HAMPSHIRE CROSSBUCKS

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0/			TSC NO	זיזמנ			vi			СТ					_		_		77777	'DAT
10	TNC	CIL	19C 11	PINER	HA7 ARD *	TNC	COM			U17100		11924	THO	0.0.0	T			H		
Xinq	ACC	A ACC	S ACC	FACTE	INDEX	ACC	BACC.	\$ 100	PACTR	TNDPY	50100	TATCU	ALC R	HTCH	6622	S ACC	LESS	SATCH .	LESS	ATCH
- -	u ș e							<i>#</i> NCC			AVAL	1107	aich	arch	*ACC	A ACC	WALL	N ACC	XDIFF	TVAL
0.50	96	96	4.84	9.68	-1095	83	83	4.18	8.37	72000	0.66	1_0	39	19	57	2.87	44	2 22	0 66	
1.00	60	156	7.86	7.86	-2279	52	135	6.80	6.80	47300	1.06	1.2	3.6	77	79	3.98	5.8	2.22	1 06	1 9
1.50	38	194	9.78	6.52	-2896	54	189	9.53	6.35	34500	0.25	0.3	37	114	RÓ	4.03	75	1 78	0.25	0 1
2.00	40	234	11.79	5.90	-3326	56	245	12.35	6.17	27200	-0.55	-0.5	35	149	85	4.28	96	4.84	-0.55	-0.8
2.50	49	283	14.26	5.71	-3800	56	301	15.17	6.07	22800	-0.91	-0.7	52	201	82	4.13	100	5.04	-0.91	-1.3
3.00	70	353	17.79	5.93	-4220	40	341	17.19	5.73	19200	0.60	0.5	46	247	106	5.34	94	4.74	0.60	0.8
3.50	41	394	19.86	5.67	-4579	34	375	18.90	5.40	16800	0.96	0.7	26	273	121	6.10	102	5, 14	0.96	1.3
4.00	41	435	21.93	5.48	-4920	41	4 16	20.97	5 24	14800	0.96	0.7	39	312	123	6.20	104	5.24	0.96	1.3
4.50	46	481	24.24	5.39	-5212	45	461	23.24	5.16	13000	1.01	0.7	41	353	128	6.45	108	5.44	1.01	1.3
5.00	36	517	26.06	5.21	-5468	33	494	24,90	4.98	12000	1.16	0.7	32	385	132	6,65	109	5.49	1.16	1.5
5.50	34	551	27.77	5.05	-5700	31	525	26.46	4.81	11000	1.31	0.8	30	415	136	6.85	110	5.54	1.31	1.7
6.00	12	586	29.54	4.92	-5928	42	567	28.58	4.76	10000	0.96	0.6	40	455	131	6.60	112	5.65	0.96	1.2
7 00	37	623	31.40	4.83	-6152	24	591	29.79	4.58	9150	1.61	0.9	33	488	135	6.80	103	5.19	1.61	2.1
7.00	20	649	34.71	4.0/	-6312	24	615	31.00	4.43	8500	1.71	1.0	33	521	128	6.45	94	4.74	1.71	2.3
4.00	31	724	34.73	4.00	-6500	23	676	32.10	4.29	3000	2.17	1.5	35	556	137	6.91	82	4.13	2.77	3.7
8 50	2 μ	74.9	30.43	4.50	-66/0	27	200	33.32	4.17	7484	2.97	1.0	30	586	138	6.96	79	3.98	2.97	4.0
9-00	27	775	19.06	4.74	-6985	23	710	34.00	3 00	6900	3.02	1.0	21	607	141	7.11	61	. 4. 08	3.02	4.0
9.50	23	798	40.22	4.21	-7117	26	716	37 10	3.70	6000	3.20	1.6	23	012	14.3	7.21	78	3.93	3.28	4.4
10.00	15	813	40.98	4.10	-7258	23	759	38.26	3.83	5600	2.72	1 4	27	683	121	/ UT		3.88	1.1	4.2
10,50	20	833	41.99	4.00	-7391	31	290	19.82	1.79	5250	2.17	1 1	23	705	124	6 UV	00	J.00 // 70	2.72	3.7
11.00	19	852	42.94	3.90	-7540	27	817	41.18	3.74	5000	1.76	0.9	19	700	129	6 45	03	4.20	4.17	2.7
11.50	22	874	44.05	3.83	-7681	22	839	42.29	3.68	4800	1.76	0.8	21	745	129	6.50	91	4.03	1 76	2.4
12-00	19	893	45.01	3.75	-7820	30	869	43.80	3.65	4400	1-21	0.6	29	774	119	6.00	95	4 79	1 21	1 6
12.50	18	911	45.92	3.67	-7956	22	891	44.91	3.59	9200	1.01	0.5	23	797	114	5.75	94	4.70	1.01	1.4
13.00	21	932	46.98	3.61	-8068	36	927	46.72	3.59	4000	0.25	0.1	26	823	109	5.49	104	5.24	0.25	0.3
13.50	21	953	48.03	3.56	-8209	13	940	47.38	3.51	3800	0.66	0.3	16	839	114	5.75	101	5.09	0.66	0.9
14.00	20	973	49.04	3.50	-8339	17	957	48.24	3.45	3600	0.81	0.4	27	866	107	5.39	91	4.59	0.81	1.1
14.50	23	996	50.2 0	3.46	-8458	8	965	48.64	Э.Э5	3500	1.56	0.7	12	878	118	5.95	87	4.39	1.56	2.2
15.00	21	1017	51.26	3.42	-8585	23	988	49.80	3.32	3240	1.46	0.6	28	906	111	5.59	82	4.13	1.46	2.1
15.50	21	1038	52.32	3.38	-8724	15	1003	50.55	3.26	3070	1.76	0.8	16	922	116	5.85	81	4,08	1.76	2.5
16.00	18	1056	53.23	3.33	-8827	31	1034	52.12	3.26	3000	1.11	0.5	24	946	110	5.54	88	4,44	1.11	1.6
17.00	15	1071	53.98	3.27	-8937	1	1035	52.17	3.16	2948	1.81	0.8	12	958	113	5.70	77	3.88	1.81	2.6
17.00	14	1107	54.99 56 BA	3.23	-9055	19	1054	53.13	3.13	2600	1.86	0.8	21	979	112	5.65	75	3.78	1.86	2.7
18 00	16	1122	55.80	3 19	-9154	21	1075	54.18	3.10	2600	1.61	0.7	17	996	111	5.59	79	3.98	1.61	2.3
18 50	23	1146	57 76	3.19	-9230	10	1090	54.94	3.05	2500	1.66	0.7	14	1015	113	5.70	80	4.03	1.66	2.4
19.00	23	1151	58.11	3.12	-9343	23	1113	56 10	3.03	2400	1.00	0.7	16	1026	120	6.05	87	4.39	1.66	2.3
19.50	21	1176	59.27	3.04	-9548	17	1130	56 96	2.77	2400	2.02	1.0	24	1050	123	6.20	83	4.19	2.02	2.8
20.00	20	1196	60.28	3.01	-9650	12	1142	57 56	2.72	2120	2.32	1.0	20	1050	120	6.05	74	3.13	2.32	3.3
20.50	13	1209	60.94	2.97	-97 15	45	1187	59 81	2 97	2130	1 11	0.0	21	1103	127	5.40	/ 3	J.68	2.12	3.8
21.00	13	1222	61.59	2.93	-9817	0	1187	59.87	2.85	2000	1.76	0.7	ננ ר	1102	107	5 70	85 7 P	4.28	1.11	1.0
21.50	11	1233	62.15	2.89	-9916	õ	1187	59.83	2.78	2000	2.32	0.9	, ם	1110	115	5 80	10	3.7J	7 2)	2.5
22.00	8	1241	62.55	2.84	- 10021	8	1195	60.23	2.74	1900	2.32	0.9	, ,	1:25	115	5.85	70	3.53	2.32	1.8
22.50	18	1259	63.46	2.82	- 10120	22	1217	61.34	2.73	1800	2.12	0.0	18	1143	116	5.85	74	3.71	2.12	3 - 0
23.00	13	1272	64.11	2.79	- 10242	11	1228	61.90	2.69	1750	2.22	0.9	15	1158	114	5.75	70	3.5	2.22	3
23.50	19	1 29 1	65.07	2.77	- 10350	14	1242	62.60	2.66	1620	2.47	1.0	19	1177	114	5.75	65	3.28	2.47	3.7
24.00	6	1297	65.37	2.72	- 10443	18	1260	63.51	2.65	1600	1.86	0.7	11	1188	109	5.49	72	3.63	1.86	2.8
24.50	15	1312	66.13	2.70	- 10538	10	1270	64.01	2.61	1520	2.12	0.8	12	1200	112	5.65	70	3.53	2,12	3.1
25.00	19	1 3 3 1	67.09	2.68	- 10641	25	1295	65.27	2.61	1500	1.81	0.7	21	1221	110	5.54	74	3.73	1.81	2.7

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* See LEGEND, Section C.3.

C - 7

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TABLE C-2. (cont.)

0/	TSC 80	INPI				СТ								77777 DAT
<i>/</i> 0	INC CUM	POWER HAZAI	ס או לס		POWER	HAZARD		4 TNC	CTH	TS	5C		1	LELLE DAT
Xing	ACC #ACC % ACC	PACTE INDI	X ACC #/	CC S ACI	C PACTR	INDEX		. 190	10 A	6 L C S S	MATCH S	LESS N	ATCH	LESS MATCH
•							PD1 (1 174)		uten	VACC	A AUC	WALL X	a ALC	NULPP TVAL
25.50	11 1342 67.64	2.65 -107	2 1 12	96 65.3	2 2.56	1460	2.32 0.9	96	1227	115	5.80	69	3_48	2.32 3.4
26.00	7 1349 67.99	2.62 - 1082	8 15 13	111 66.0	8 2.54	1400	1.92 0.1	7 8	1235	114	5.75	76	3.83	1.92 2.8
20.50	8 1357 68.40	2.58 -1092	7 61.	117 66.3	8 2.50	1350	2.02 0.0	87	1242	115	5.80	75	3.78	2.02 2.9
27.00	10 1367 68,90	2.55 -1101	4 9 1.	26 66.8	3 2.48	1280	2.07 0.1	89	1251	116	5.85	75	3.78	2.07 3.0
27.50	11 1378 69,46	2.53 - 111	5 8 1.	134 67.2	4 2.45	1220	2.22 0.1	89	1260	118	5.95	74	3.73	2.22 3.2
28.00	12 1390 70.06	2.50 -1120	9 30 1	64 68.7	5 2.46	1200	1.31 0.5	5 27	1287	103	5.19	77	3.80	1.31 1.9
20.50	7 1010 71 07	2.48 - 112	8 01.	64 68.7	5 2.41	1200	1.97 0.	7 7	1294	109	5.49	70	3.53	1.97 2.9
20.00	1 1410 71.07	2,45 -1137	4 13 13	11 69.4	1 2.39	1120	1.66 0.0	59	1303	107	5.39	74	3.73	1.66 2.5
30 00	10 1010 72 20	2.43 - 1148	i 131. // 111.	90 70.0	6 2.37	1080	1.71 0.0	5 13	1316	108	5.44	74	3.73	1.71 2.5
10.50	21 1455 73 34	2.41 ~115:	יוויי	30 70.6	1 2.35	1020	1.66 0.0	5 14	1330	104	5.24	71	3.58	1.66 2.5
31.00	6 1461 73 64	2.40 -117		30 72.44	0 4.30	1000	0.86 0.	3 28	1358	97	4.89	80	4.03	0.86 1.3
31.50	7 1468 73.99	2.35 - 1181	μ 0 14	30 12.4	0 2.34	1000	1.16 0.4		1361	100	5.04	77	3.88	1.16 1.7
32.00	16 1434 74.80	2-34 -1190	ม 1914	57 73 4	u 7 .30	975	1.31 0.0	3 J	1304	104	5.24	/4	3.13	1.51 2.2
32,50	9 1493 75.25	2.32 -1196	6 19 14	76 74 4		900	· 0.96 0.1	1 15 1 15	1309	20	4 / / 9	08	3.43	1.36 2.1
33.00	6 1499 75.55	2.29 - 120	0 1 14	77 74.4	5 2.26	890	1.11 0.4		1404	60	4,49	72 60	3.03	0.86 1.3
33.50	9 1508 76.01	2.27 -1214	0 12 14	89 75.0	5 2.24	840	0.96 0.	, , , ,	1017	91	4.37	77	3.40	
34.00	14 1522 76.71	2.26 - 1224	1 27 1	16 76.4	1 2.25	800	0.30 0.	í 17	14.14	88	4.57	97	5.03	0.96 1.5
34.50	7 1529 77.07	2.23 -1230	0 0 15	16 76.4	1 2.21	800	0.66 0.2	 7 u	1418	91	4 59	78	3 6 3	0.50 0.5
35.00	8 1537 77.47	2.21 -1239	8 3 19	19 76.5	6 2.19	776	0.91 0.	3 6	1444	93	4.69	75	3.75	0.00 1.0
35.50	12 1549 78.07	2.20 -1247	6 9 19	28 77.0	2 2.17	750	1.06 0.0	a 13	1457	92	464	71	3.58	1.06 1.6
36.00	14 1563 78.78	2,19 -1256	5 17 15	45 77.8	7 2.16	720	0.91 0.3	3 21	1476	85	4.28	67	3.38	0-91 1-5
30.50	6 1569 79.08	2.17 - 1263	2 7 19	52 78.2	3 2.14	700	0.86 0.3	37	1485	84	4.23	67	3.38	0.86 1.4
37.00	3 1572 79.23	2.14 - 1270	1 5 1 5	57 78.41	8 2.12	660	0.76 0.3	36	1491	81	4.08	66	3.33	0.76 1.2
37.50	4 1576 75.44	2.12 - 1276	2 7 1	64 78.8	3 2.10	640	0.60 0.3	29	1500	76	3.83	64	3.23	0.60 1.0
18.00	9 1242 79.89	2.10 -1285	0 37 16	01 80.7	0 2.12	600	-0.81 -0.1	3 25	1525	60	3.02	76	3.83	-0.81 -1.4
18.50	4 1589 80.09	2.08 -1292	8 0 16	01 80.7	0 2.10	600	-0.60 -0.2	2 2	1527	62	3, 13	74	3.73	-0.60 -1.0
39.00	8 1597 80.49	2.06 - 1300	3 0 16	01 80.7	0 2.07	600	-0.20 -0.1	1 43	1531	66	3.33	70	3.53	-0.20 -0.3
40.00		2.05 -1308	4 0 16	01 80.7	0 2.04	594	0.10 0.0) 3	1534	69	3.40	67	3.38	0.10 0.2
40.50	11 1621 81 70	2.03 ~131;	0 11 10	12 81.2	5 2.03	550	-0.10 -0.0) 5	1539	71	3,58	73	3.68	-0.10 -0.2
41.00	5 1626 81 96	$2 \cdot 02 = 1321$	/ 010 7 1214	18 81.5	5 2.01	522	0.15 0.1	1 15	1554	67	· 3. 38	64	3.23	0.15 0.3
41.50	9 1635 82.41	1 99 - 132		41 0 <u>7</u> ,7		500	-0.76 -0.2	8 8	1562	64	3.23	79	3.98	-0.76 -1.3
42.00	3 1630 82.56	1 97 - 1343	1 0 16	41 02.7	1 1.99	500	-0.30 -0.	16	1568	67	3,38	73	3.68	-0.30 -0.5
42.50	5 1643 82.81	1.95 -1351	μ 0.16	41 87 7	1 1.97	500	-0.15 -0.	د ۱	1571	67	3.38	70	3.53	-0.15 -0.3
43.00	5 1648 83.06	1.93 -1361	7 7 16	48 83.0	6 1.93	480		J 4	15/5	08	3.45	66	3.33	0.10 0.2
43.50	9 1657 83.52	1.92 -1369	3 4 16	52 83.2	7 1.91	460	0.00 0.0	, 0 , 7	1500	C 0 7 4	3.28	60	3.28	0.00 0.0
44.00	9 1666 83.97	1.91 - 1378	6 9 16	61 83.7	2 1.90	450	0.25 0	' / 1 11	1601	67	3,30	62	3.13	0.25 0.4
44.50	9 1675 84.43	1.90 -1381	3 7 16	68 84 0	7 1.89	420	0.15 0.1	, i 1 e	1610	65	3 20	59	3.02	0.25 0.4
45.00	7 1682 84.78	1.88 - 1392	6 23 16	91 85.2	1.89	400	-0.45 -0.3	, , 7 18	1678	50	2,20	63	2.76	
45.50	5 1687 85.03	1.87 -1398	2 0 10	91 85.2	3 1.87	400	-0.20 -0.	1 5	1633	54	2.77	58	7 97	-0.20 -0.8
46.00	5 1692 85.28	1.85 -1405	6 0 16	91 85.2	3 1.85	400	0.05 0.0		16 38	54	2.72	53	2.67	0.05 0.1
46.50	10 1702 85,79	1.84 - 1414	3 0 16	91 85.2	3 1.83	400	0.55 0.2	2 6	1644	58	2.92	47	2. 17	0.55 1.1
47.00	3 1705 85.94	1.83 -1422	4 1 16	92 85.2	8 1.61	376	0.66 0.3	2 2	1646	59	2.97	46	Z. 32	0.66 1.3
47.50	2 1707 86.04	1.81 - 1429	4 716	99 85.6	4 1.60	360	0.40 0.	1 5	1651	56	2.82	48	2.42	0.40 0.8
48.00	7 1714 86.39	1.80 -1435	8 0 16	99 85.6	4 1.78	350	0.76 0.3	3 3	1654	60	3.02	45	2.27	0.76 1.5
48.50	6 1/20 86.69	1.79 -1441	7 513	04 85.8	9 1.77	326	0.81 0.3	35	1659	61	3.07	45	2.27	0.81 1.6
49.00 Ng 50	13 1733 87.35	1.78 - 1449	6 5 1 7	09 86.1	4 1.76	320	1.21 0.4	i 13	1672	61	3.07	37	1.86	1.21 2.9
47.JU 50 00	5 1/38 87.60 1 1730 07 /r	1.77 - 1456	3 14 11	23 86.8	1.75	300	0.76 0.3	9 9	1681	57	2.87	42	2.12	0.76 1.5
30.00	1 1132 01.02	1./5 -1464	J 017	23 86.8	4 1.74	300	0.81 0.3	31	1682	57	2.87	41	2.07	0.81 1.6

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* See LEGEND, Section C.3.

TABLE C-2*(cont.) (TSC)

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Xing	PP	CON 1	CON 2	CON 3	CON 4	Xing	P F	CON 1	CON 2	CON 3	CON 4
0.50	9.68	3.97	1.63	0.99	0.65	25.50	2.65	3.41	1.94	0.83	0-77
1.00	7.86	3.73	1.71	0.96	0.66	26.00	2,62	3.36	1.94	0.83	0.77
1.50	6.52	3.39	1.55	0.91	0.64	26.50	2.58	3.32	1 99	0.82	0.77
2.00	5.90	3.25	1.51	0.90	0.64	27.00	2.55	3.28	1.96	0.82	0.76
2.50	5.71	3.30	1.55	0.91	0.65	27.30	2.50	3.28	1.97	0_81	0.77
3.00	5.93	3.02	1.69	0.95	0.09	28.00	2.48	3.29	1.98	0.81	0.77
3.50	5.48	3.60	1.70	0.95	0.71	29.00	2.45	3.25	1.98	0.81	0.76
4.50	5.39	3.69	1.74	0.96	0.72	29.50	2.43	3.27	1.99	0.80	0.77
5.00	5.21	3.68	1.74	0.95	0.73	30.00	2.41	3.26	2.00	0.80	0.77
5.50	5.05	3.67	1.74	0.95	0.73	30.50	2.40	3.36	2.02	0.81	0.77
6.00	4.92	3.68	1.75	0.95	0.73	31.00	2.38	3.31	2.03	0.80	0.77
6.50	4.83	3.72	1.77	0.95	0.74	31.50	2.35	3.28	2,03	0.79	0.77
7.00	4.67	3.67	1_76	0.95	0.74	32.00	2.34	3.33	2.05	0.80	0.77
7.50	4.66	3.79	1.80	0.96	0.75	32.50	2.32	3.32	2.00	0.79	0.77
8.00	4.56	3.80	1.01	0.95	0.76	33.00	2 27	3.20	2.07	0.78	0.77
8.50	4.44	3.76	1.80	0.95	0.76	33.00	2 26	3.31	2.09	0.78	0.77
9.00	4.34	3.70	1.80	0.95	0.76	34.50	2.23	3.28	2.10	0.78	0.77
9.50	4.23	3.73	1.00	A 01	0.70	35.00	2.21	3.27	2.11	0.78	0.77
10.00	4.10	3.04	1 77	0.93	0.75	35.50	2.20	3.30	2.12	0.77	0.77
11.00	3 90	3.56	1.77	0.92	0.75	36.00	2.19	3.35	2.14	0.77	0.77
11. 50	3.83	3.54	1.77	0.91	0.75	36.50	2.17	3.32	2.15	0.77	0.77
12.00	3.75	3.51	1.77	0.91	0.75	37,00	2.14.	3.25	2.15	0.76	0.76
12.50	3.67	3.48	1.77	0.90	0.75	37,50	2.12	3.20	2.16	0.76	0.76
13.00	3.61	3.47	1.77	0.90	0.75	38.00	2.10	3.20	2.17	0.76	0.76
13.50	3.56	3.47	1.78	0.90	0.75	38.50	2.08	3, 15	2.18	0.75	0.76
14.00	3.50	3.46	1.78	0.89	0.76	39.00	2.06	3.14	2.19	0.75	0.76
14.50	3.46	3.48	1.79	0.89	0.76	39.50	2.05	3.12	2.20	0.74	0.76
15.00	3,42	3_49	1_80	0.89	0.76	40.00	2.03	3,10	2.21	0.74	0.76
15.50	3.38	3.50	1.01	0.89	0.76	40.50	2.02	3.14	2 2 2 4	0.73	0.75
16.00	3.33	3.49	1.02	0.89	0.76	41.00	1 99	3, 12	2.26	0.73	0.76
16.50	3.2/	3.40	1.82	0.88	0.76	42.00	1.97	3.06	2.27	0.73	0.75
17.00	2 10	3.40	1 63	0.00 0.00	0.76	42.50	1.95	3.03	2.28	0.72	0.75
18 00	3.14	3.45	1.83	0.87	0.76	43.00	1.93	2.99	2,29	0.72	0.75
18.50	3.12	3,50	1.85	0.88	0.77	43.50	1,92	3.01	2.31	0.71	0.75
19.00	3.06	3.42	1.84	0.87	0.76	44.00	1.91	3.04	2.32	0.71	0.75
19.50	3.04	3.48	1,86	0.87	0.77	44.50	1.90	3.06	2.34	0.71	0.75
20.00	3.01	3,51	1.87	0.87	0.77	45.00	1,68	3.06	2.36	0.71	0.75
20.50	2.97	3.49	1.88	0.87	0.77	45.50	1.8/	3.04	2.37	0.70	0.75
21.00	2.93	3.47	1.88	0.86	0.77	46.00	1.85	3.01	2.39	0 70	0.75
21.50	2.89	3.43	1,88	0.86	0.77	46.50	1,84	3.00	2 4 1	0.70	0.75
22.00	2.84	3.37	1.88	0.85	0.77	47.00	1 91	2.01	2.42	0.69	0.74
22.50	2.82	3.40	1.89	0.85	0.77	47.00	1.01	2.95	2.45	0.69	0.74
23.00	2.79	7.72	1.90	0.83	0.77	48.50	1.79	2.94	2.47	0.68	0.74
23.30	2.1/	3.43	1.01	0.80	0,//	49.00	1.78	3.05	2.50	0.68	0.74
24.00	2 70	5 2 5 0 1 ° 1 0	1 0 7	0.04	0.//	49.50	1,77	3.03	2.52	0.68	0.74
25.00	2.68	3.43	1.94	0.84	0.77	50.00	1,75	2.95	2.53	0.67	0.74
23100		3443	1	5.01							
* See	Subs	ectio	on 4	.1.2	for	discus	ssion	of (CON 1	1 - (CON 4

C-9

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TABLE C-2* (cont.)

(NH)

%						0/ /0					
Xing	PP	CON 1	CON 2	CON 3	CON 4	Xing	2 2	CON 1	CON 2	CON 3	COM 4
0.50	8-37	3,55	1.58	0.92	0.61	25.50	2.56	3.02	1.87	0.80	0.74
1.00	6.60	3.29	1.48	0.89	C.61	26.00	2.54	3.03	1.89	0.80	0.75
1.50	6.35	3.30	1.51	0.90	0.63	26.50	2.50	2.98	1.89	0.80	0.74
2.00	6.17	3.40	1.58	0.92	0.65	27.00	2.48	2.95	1.89	0.79	0.74
2.50	6.07	3.53	1.65	0.94	0.68	27.50	2.45	2.91	1.89	0.79	0.74
3.00	5.73	3.48	1.63	0.93	0.68	28.00	2.46	3.05	1.93	0.80	0.75
3.50	5.40	3.39	1.61	0.92	0.68	28.50	2.41	2.95	1.92	0.79	0.74
4.00	5.24	3.41	1.63	0.92	0.69	29.00	2.39	2.96	1.93	0.78	0.74
4.50	5.16	3.49	1.67	0.93	0.70	29,50	2.37	2.97	1.95	0.78	0.74
5.00	4.98	3.46	1.66	0.93	0.71	30.00	2.35	2.90	1.96	0.78	0.74
5.50	4.81	3.43	1.66	0,92	0.71	30.50	2.38	3.17	2.00	0.79	0.70
6.00	4.76	3.51	1.69	0.93	0.72	31.00	2.34	3.09	1 00	0.70	0.75
6.50	4.58	J. 44	1.00	0.92	0.72	31.00	2.30	3 07	2 01	0.78	0 75
7.00	4.43	3.38	1.0/	0.00	0.71	32.50	2 2 2 9	3 15	2.04	0.78	0.75
7.50	4 - 29	<u>د د</u>	1.00	0.90	0.77	32.30	2 2 2 5	3 06	2.03	0.77	0.75
8.00	4.19	3.31	1,00	0.90	0 71	33.50	2 74	3.07	2.05	0.77	0.75
8.00	3 68	3.27	1.65	0.89	0.71	34.00	2.25	3.25	2.08	0.78	0.76
9.00	3 00	3 23	1.66	0.89	0.72	34.50	2.21	3, 15	2.08	0.77	0.76
10 00	3 83	3. 22	1.66	0.88	0.72	35.00	2.19	3.08	2.08	0.76	0.75
10.50	1.79	3.26	1.68	0.89	0.72	35.50	2.17	3.07	2.09	0.76	0.75
11.00	3.74	3,28	1.70	0.89	0.73	36.00	2,16	3.15	2.12	0.76	0.76
11.50	3.68	3.27	1.70	0.89	0.73	36.50	2.14	3.13	2.13	0.76	0.76
12.00	3.65	3.33	1.72	0.89	0.74	37.00	2.12	3.08	2.13	0.75	0_76
12.50	3.59	3.32	1.73	0.89	0.74	37.50	2.10	3.06	2.14	0.75	0.75
13.00	3.59	3.43	1.76	0.90	0.75	38.00	2.12	3.40	2.19	0_77	0.77
13.50	3.51	3,37	1.75	C.89	0.75	38.50	2,10	3.29	2.20	0.76	0.77
14.00	3.45	3.34	1.75	0.88	0.75	39.00	2.07	3.19	2.20	0.75	0.76
14.50	3.35	3.24	1.74	0.87	0.74	39.50	2.04	3.09	2.20	0.74	0.76
15.00	3.32	3.27	1.75	0.87	0.74	40.00	2.03	3.13	2.22	0.74	0.76
15.50	3.26	3.23	1.75	0.87	0.74	40.50	2.01	J. 10	2.23	0.74	0.76
16.00	3.26	3.32	1.78	C.87	0.75	41.00	2.02	1.10	2.20	0.74	0.76
16.50	3.16	3.19	1.75	0.86	0.74	41.50	1.99	3,20	2.21	0.79	0.75
17.00	3.13	3.19	1.70	0.00	0.74	42.00	1 05	3 00	2.11	0.7J	0 75
17.50	3.10	3.21	1 74	0.00	0.75	42.00	1 93	2 90	2.27	0 72	0.75
10.00	3.03	3.17	1 80	0.05	0.75	43.50	1.91	2.95	2.30	0.71	0.75
16.50	2 95	3 11	1 78	0.84	0.74	44.00	1.90	2.97	2.32	0.71	0.75
19.50	2.92	3, 11	1.79	0.84	0.74	44.50	1.89	2.97	2.33	0.71	0.75
20.00	2.88	3.08	1.79	0.83	0.74	45.00	1.89	3.19	2.37	0.71	0.76
20.50	2.92	3.30	1.84	0.85	0.76	45.50	1.87	3.09	2.38	0.71	0.75
21.00	2.85	3.18	1.83	0.84	0.75	46.00	1.85	3.00	2.39	0.70	0.75
21.50	2.78	3.06	1.81	0.83	0.74	46.50	1.83	2.90	2.39	0.69	0.74
22.00	2.74	3.01	1.81	0.82	0.74	47.00	1.81	2.83	2.40	0.69	0.74
22.50	2.73	3.06	1.83	0.82	0.74	47.50	1.80	2.å3	2.42	0.68	0.74
23,60	2.69	3.03	1.83	0.82	0,74	48.00	1.78	2.74	2.43	0.68.	0.73
23.50	2.66	3.03	1.84	0.81	0.74	48.50	1.77	2.72	2.45	0.67	0.73
24.00	2.65	3.06	1.85	0.81	0.75	49.00	1.76	2.70	2.46	0.67	0.73
24.50	2.61	3.03	1.86	0.81	0.74	49.50	1.75	2.80	2.49	0.44	0.73
25.00	2.61	3.12	1.88	U.82	0.15	20.00	1.74	4.1	2.51	V.00	U • / 3
* 500	Sub	secti	on -	4.1.	2 for	discus	ssion	of (ON 1	(CON 4.
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TABLE C-3. TSC (VOLUME ONLY) VERSUS NEW HAMPSHIRE FLASHING LIGHTS

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9		ጥሮ	с жа	DEL			N RU	НАЛР	SHIRP	CI				TS	с	18	н	
10	TNO	CU #		DOUED	83718D ¥	TNC	сли		DOWPD	RATARD	9798 81764	THC	СЦВ	1 25 5	BATCH	TREE		87948 2291
Xina	TAC	1.0.3		PUEL DA	186AND **	100		100		THODY		100	10.00	4100		5100		
	ACC	BACC %	ACC	FACTE	IN DEL	ALC	UALC X	ACC .	PACTR	TNDET	BOTLE IAT	arca	8104	UNCC	B ALC	GALL	กลเเ	SDIFF LANT
							• • • •							· .				
0.50	56	56 4	1.23	8.46	-2310	55	55 4	5.15	8,31	487800	0.08 0.1	55	55	1	0,08	0	0.00	0.08 1.0
1.00	40	96 7	1.25	7.25	3472	34	89 6	-72	6.72	337790	0.53 0.5	34	89	7	0.53	0	0.00	0.53 2.6
1.50	19	115 8	3.69	5.79	-4051	26	115 8	3.69	5.79	268380	0.00 0.0	22	111	4	0.30	4	0.30	0.00 0.0
2.00	27	142 10	.73	5.36	-4521	29	144 10	.68	5.44	234000	-0.15 -0.1	27	130	4	0.30	6	0.45	-0.15 -0.6
2 50	24	166 12	5.5.0	5.02	-4847	25	169 12	2.76	5.11	210000	-0.23 -0.2	26	164	2	0.15	5	0.38	-0.23 -1.1
2.00	27	103 14	5.0	J 86	-5218	29	198 14	95	4 Q.A	183820	-0.38 -0.3	27	191	2	0.15	7	0.53	-0.38 -1.7
3.00	50	313 14	0.00	4 CO	-51.60	10	217 16	10	n 49	169#00	-0.10 -0.2	20	211	2	A 15	È.	0 45	-0.30 -1.4
3.00	20	213 10		4.00	-5409	4 1	221 12	1 46	1 36	162090	-0.00 -0.2	15	226		0.10	Š	8 39	-0.09 -0.3
4.00		230 17		4.34	3780	1.0	231 17		N 20	133000		10	240		0.30	5	0,00	
4.50	17	247 18	5.00	4.15	-6075	13	250 18		4.20	133000	-0.23 -0.1	19	240	4	0.15		0.30	-0.23 - 1.1
5.00	19	266 20	.09	4.02	-6273	23	2/3 20	1.62	4.12	128000	-0.53 -0.3	18	263	5	0.23	10	0.75	-0-53 -1-9
5.50	22	289 21	1.75	3.95	-6468	23	296 22	2.36	4.06	120800	-0.60 -0.3	20	X83	5	0,38	13	0.98	-0.60 -1.9
6.00	28	316 23	.87	3.98	-6633	26	322 24	1.32	4.05	112720	-0.45 -0.2	23	306	10	0.76	16	1.21	-0.45 -1.2
6.50	21	337 25	5.45	3.92	-6832	18	340 25	68	3.95	106981	-0.23 -0.1	24	330	7	0.53	10	0.76	-0.23 -0.7
7.00	19	356 26	. 89	3.84	~7003	24	364 27	.49	3.93	100000	-0.60 -0.3	21	351	5	0,38	13	0.98	-0.60 ~1.9
7.50	20	376 28	40	3.79	-7160	15	379 28	.63	3.82	94560	-0.23 -0.1	20	371	5	0.38	6	0,60	-0.23 -0.8
8 00	15	391 29	5.5	3.69	-7309	18	197 29	. 98	3.75	90000	-0.45 -0.2	14	385	6	0.85	12	0.91	-0.45 -1.4
0.00		300 30	10	3 55	-74.22	10	407 20	70	3 67	95.000	-0 60 -0 3		300	5	0,38	13	0 00	-0.50 -1.9
0.00		377 30		7.17	-7423	10	407 30	07	3 64	01700			507	5	A 60	16	1 13	-0 57 -1 5
9.00	16	415 31	. 34	3.40		15	422 31		3.34	81700	2°0- C °0- 7	12	~U/			15	1.13	
9.50	12	427 34		7-14	-7689		428 32		3.40	78880	~0.08 ~0.0		418	9	0.08	10	0.76	-0.08 -0.2
10.00	10	437 33	.01	3.30	-7788	15	440 33	.23	3.32	75600	-0.23 -0.1	10	428	y	0.68	12	0.91	~0.23 ~0.7
10.50	10	447 33	3.76	3.22	~7900	14	454 34	.29	3.27	72000	-0.53 -0.2	10	438	9	0,68	16	1.21	~0.53 •1.4
11.00	21	468 35	5.35	3.21	8027	18	472 35	i.65	3.24	69225	-0.30 -0.1	14	452	16	1,21	20	1.51	-0.30 -0.7
11.50	19	487 36	5.78	3.20	-8140	20	492 37	1.16	3.23	66000	-0.38 -0.2	20	472	15	1.13	20	1.51	-0.38 -0.8
12.00	18	505 38	3.14	3.18	-8257	12	504 38	.07	3.17	64000	0.08 0.0	18	490	15	1.13	14	1.06	0.08 0.2
12.50	12	517 39	.05	3.12	-8366	11	515 38	1.90	3.11	61088	0.15 0.1	12	502	15	1.13	13	0.98	0.15 0.4
13.00	11	528 39	. 88	3.07	-8473	8	523 39	50	3.04	59800	0.38 0.2	11	513	15	1.13	10	0.75	0.38 1.0
13 50	ġ	537 40	56	3_00	-8573	Ā	531 80	. 11	2.97	56480	0.45 0.2	10	523	14	1.06	ค	0.60	0.45 1.3
10 00	10	547 11	31	2 92	- 96 70	16	547 49	31	2 95	54,000		10	533	14	1.06	14	1.06	B-00 0-0
14.00		560 00		201	-0070	9.5	547 41		2077	51600			565	17	4 70	79	1 50	~0.20 ~0.6
14.50	- 11	226 42	. 15	2.91	-0/00	12	202 42		6.73	51500	······	6	241	16	1 36	21	3 90	-0 53 -1 1
15.00	6	564 42	- 60	2.84	-8678		5/1 43	1.13	2.80	4994U	-0.23 -0.2		200	10	0,00	20	1.07	
15.50	14	578 43	- 60	2.82	-9015	12	283 44	LUA	2.84	48000	~0.38 ~0.1	13	222	19	1.44	~~	1.01	~0.30 ~0.0
16.00	14	592 44	.71	2.79	-9091	5	588 44	.41	2.78	46554	0.30 0.1	13	570	22	1.00	18	1, 30	0.30 0.6
16.50	14	606 45	. 77	2.77	-9192	22	610 46	ia 07	2.79	45000	-0.30 -0.1	12	562	24	1.81	28	2.17	~0.30 ~0.6
17.00	11	617 46	.60	2.74	-9262	6	618 46	n 53	2.74	43260	0.08 0.0	11	593	24	1.01	23	1.74	0.08 0.1
17.50	16	633 47	.81	2.73	-9361	7	623 47	.05	2.69	41990	0.76 0.3	13	606	27	2.04	17	1.28	0.76 1.5
18.00	14	647 4B	.87	2.71	94 50	17	640 48	1.34	2.69	40390	0.53 0.2	15	621	26	1.96	19	1.44	0.53 1.0
18.50	10	657 49	.62	2.68	-9524	8	648 48	. 94	2.65	40000	0.68 0.2	10	631	26	1.96	17	1,28	0.69 1.4
19.00	ů.	661 49	92	2.61	-9613	Ř	656 49	55	2.61	38800	0.38 0.1	8	639	22	1.66	17	1.28	0.38 0.8
19 50	7	668 50	45	2.50	-9714	17	669 50	51	2,50	37500	-0.08 -0 0	10	649	10	1. BA	20	1.51	-0.08 -0.2
20.00	12	101 101	22	2,23	-0701	17	692 64	Q.4	2 50	36000	-0 15 -0 1	17	666	10	1 24	20	1.51	-0.15 -0.1
20.00	10	004 01	.00	2.00	~7/01	12	000 00		2.09	30000			470	10	10,30	20	1 51	
20.50		691 52	. 19	2.35	9845	0	092 52		4.35	35000	-0-08 -0-0	0	014	19	1.95	20	1. J. I 	~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~
21.00	5	696 52	.57	2.50	-9914	8	700 52	. 87	2.52	34000	-0.10 -0.1	8	680	16	1.21	20	1.21	00030 000/ 000 - 0 5
21.50	10	706 53	2 د .	2.48	- 10007	9	/09 53	1.55	2.49	33000	-0.23 -0.1	6	666	16	1.30	21	1.27	~V°T C°O
22.00	14	720 54	-38	2.47	- 10089	6	715 54	.00	2,45	32000	0.38 0.1	9	697	23	1.76	18	1. 36	0.38 0.8
22.50	7	727 54	.91	2.44	- 10157	7	722 54	. 53	2.42	30940	0.38 0.1	8	705	22	1.66	17	1.28	0.38 0.8
23.00	6	733 55	.36	2.41	10229	9	731 55	5.21	2.40	30000	0.15 0.1	6	711	22	1.66	20	1.51	0.15 0.3
23.50	6	739 55	.82	2.38	- 10305	4	735 55	.51	2.36	29700	0.30 0.1	6	717	22	۱.66	18	1.36	0.30 0.6
24.00	10	749 56	57	2.36	- 10369	ę	744 56	19	2 34	28800	0.38 0.1	11	728	21	1.59	16	1.21	0.38 0.8
24 50	<u>и</u>	751 54	.87	2. 12	- 10443	á	753 54	. 87	2.12	28000	0.00 0.0	5	777	20	1.51	20	1.51	0.00 0.0
25 00		762 50	55	2 30	-105 19	, 0	762 57	55	2 20	27000		á	74.2	20	1.51	20	1.51	0.00 0.0
2 J o VV	,	102 31		6000	- 103 17	7	792 J1		4 0 J V	2 1 U U U		,		20		~ V		
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* See LEGEND, Section C.3.

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TABLE C-3. (cont.)

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%			TSC M	DDEL	-		N I	SH HAN	PSHIRE	U I					T	5C		B		
Vina	INC	CUN		POWER	HAZARD 🛪	INC	CUN		POWER	HAZARD	BITH B	ATCH	INC	CON	L ES S	BATCH	LESS	BATCH	LBSS	BATCH
Aing	ACC	# ∆ CC	S ACC	FACTR	INDEX	ACC	#ACC	% ACC	FACTE	LNDEK	%DIFP	TVAL	BICH	BICH	\$ACC	% ACC	PACC	S ACC	%DIFI	TTAL
25.50	5	767	57.93	2.27	- 10608	7	769	58.08	2.28	26270	-0.15	-0.1	7	789	18	1.36	20	1.51	-0.19	5 -0.3
26.00	, Š	776	58-61	2.25	- 106 57	13	7 82	59.06	2.27	25600	-0.45	-0.2	12	761	15	1 11	21	1 59	-0 85	-10
26 50	9	785	59 29	2.24	- 10714		789	59.59	2 25	25000	-0.30	_0 1	5	766	19	1 44	23	1 74		
27 00	11	796	60 12	2 23	- 10789	6	705	60 05	5 23	20000	0.00	_0.1	2	775	24	1 50	20	9 6 9	~0.30	, -v.a
27.50	6	002	60 57	2 20	- 10961	5	900	60 82	3 30	24230	V. UO A 1E	n 0	×	794	21	1.57	20	1.51	0.00	, v.z
27. 30	0	002	61 10	2 10	- 10001	5	000	40 00	2.20	20000	0.15	0.0		701	21	1.37	19	1.04	U. 1	> 0.3
20.00	10	010	61.10	2.10	*10917		040	C 4 70	2.11	23200	0.38	0.1	0	787	£ ∑	1.74	18	1.30	0.3	9 0.8
20.00	10	820	01.37	2.1/	~ 10900	21	010	010/0	4.1/	22500	0.15	0.0	16	799	21	1.59	19	1.44	0.15	0.3
29.00		823	62.10	2.14	~ 110 60		845	02.31	2.15	22000	~0.15	-0.0	4	903	20	1.51	22	1.65	~0.1	» =0.3
29.50	10	833	62.92	4.13	- 11121	8	833	62.92	2.13	21440	0.00	0.0	/	810	23	1.74	23	1.74	0.00) 0.0
30.00	8	841	63.52	2.12	-11184	4	837	63.22	2.11	20900	0.30	0.1	9	819	22	1.66	18	. 1. 36	0.30) 0.6
30.50	5	846	63,90	2.09	~ 11234	4	841	63.52	2.08	20400	0.38	0.1	9	823	23	1.74	18	1.36	0.30	3 0.0
31.00	10	856	64.65	2.09	- 11325	5	846	63.90	2.06	20000	0.76	0.2	8	831	25	1.69	15	1.13	0,76	i 1.6
31.50	10	866	65.41	Z.08	- 11387	1	847	63.97	2.03	19680	1.44	0.5	5	836	30	2.27	11	0.83	1.80	J 3.0
32.00	6	872	65.86	2.06	- 11438	8	855	64.58	2.02	19200	1.28	0_4	7	843	29	2.19	12	0.91	1.2	3 2.7
32.50	4	876	66.16	2.04	- 11487	5	860	64.95	2.00	18792	1.21	0.4	5	848	28	2.11	12	0.91	1.2	1 2.5
33.00	7	883	66.69	2.02	-11545	12	872	65.86	2.00	18040	0.83	0-3	Ÿ	855	28	2.11	17	1.28	0.83	3 1.6
33.50	2	885	66.84	2.00	- 11589	6	878	66.31	1.98	18000	0.53	0.2	2	857	28	2.11	21	1.59	0.53	1.0
34.00	4	889	67.15	1.97	- 11645	8	886	66.92	1.97	17325	0.23	0.1	6	863	26	1.96	23	1.79	0.2	1 0.9
34.50	4	893	67.45	1.95	-11704	6	892	67.37	1.95	16848	0.08	0.0	5	868	25	1.89	24	1.81	0.0	a 0.1
35.00	8	901	68.05	194	- 11762	11	903	68.20	1.95	16380	~0.15	-0.0	R	876	25	1.89	27	2.00	o0.15	5 -0. T
35.50	9	910	68.73	1.94	-11823	4	907	68.50	1.93	16000	0.23	0.1	, e	885	25	1.49	22	1.66	0 2	1 0.4
36.00	8	918	69.34	1.93	-11897	4	911	68.81	1.91	15600	0.53	0.2	ś	890	28	2 11	21	1.59	0 5	1 1 6
36-50	6	924	69.79	1.91	-11945	ò	911	68.81	1.89	15230	n 98	0.3	Ē	895	29	2 10	16	1 21	0.90	1 1 0
37.00	ā	932	70.39	1,90	- 12003	Ř	919	69.41	1.88	15000	0.98	ñ. 1	á	000	20	2 11	15	1 17	0.90	3 2 0
37 50	ă	041	71 07	1.90	- 120 73	2	922	69 64	1 94	14560	4 4 4 4	0.5		044	20	2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2		0 0 0	4 4 4	, 1.0
38 00	<u>,</u>	945	71 17	1 88	- 12075	5	922	20 02	1 94	14 220	1 74	0.4		017	30	7 19		0.03	4 74	·
39 50		0/10	71 60	1 84	- 12220		0.25	70.02	1.03	10000	1.30	0.4		917	20	2 40	10	4 4 3	1.30	207
30.00	°	057	73 30	1.00	- 12220		333	70.02	1000	14000	1.00	0.3	4.	920	2 Y	2017	15	1.13	1.00	
30 50	~	057	70 00	1.00	- 12204	1	740	71.93	1.03	13000	0.03	0.3	8	920	29	K . 19	16	1.20	0.82	
33-20		937	73 61	1.03	42200		733	71.90	1.04	13200	0.30	0.1	3	931	20	1.90	22	1.00	0.3	/ 0.0
NO 50		300	72.04	1.01	~ 12380	2	928	12.30	1.81	12800	0.15	0.0	4	533	21	8.04	25	1.89	0.1	LaU (
40.50	4	964	12.01	1.80	*12445		964	72.61	1.80	12500	0.00	0.0	9	939	25	1.09	25	1.89	0.0	1 0.0
41.00	6	970	73.20	1. 19	· 12495	צו	983	74.24	1.81	12000	-0.98	-0.1		947	23	1.74	36	2.72	-0.90	5 -1.7
41.50	2	973	/3.04		- 12576	0	983	14.20	1.79	12000	-0.60	-0.2	3	950	25	1.89	33	2.49	-0.60) -1.1
42.00	د	978	13.81	1.76	-12630	1	984	74.32	1.77	11792	-0.45	-0-1	3	953	25	1.89	31	2.34	-0.4	5 -0-8
42.50	в	986	74.47	1.75	-12683	4	988	74.62	1.76	11388	-0.15	-0.0	5	958	28	2.11	30	2.27	-0.1	j ~0₀3
41.00	10	996	75.23	1.75	- 12750	1	989	74.70	1.74	11025	0.53	0.2	6	964	32	2.42	25	1.89	0.53	3 0.9
43.50	9	1005	75.91	1.74	- 12814	4	993	75.00	1.72	10800	0.91	0.3	11	975	30	2.27	18	1.36	0.9	1 1.7
44.00	8	1013	76.51	1.74	12885	8	1001	75.60	1.72	10500	0.91	0.3	6	991	32	2.42	20	1. 51	D_9'	1 1.7
44.50	8	1021	77.11	1.73	- 12939	8	1009	76.21	1.71	10200	0.91	0.3	11	99 Z	29	2。19	17	1.20	0.9	1 1.8
45.00	4	1025	77.42	1.72	- 12996	10	1019	76.96	1.71	10000	0.45	0.1	4	996	29	S° 18	23	1.74	0.4	5 0.8
45.50	4	1029	77.72	1.71	-13047	1	10 20	77.04	1.69	9840	0.68	0.2	2	998	31	2.34	22	1.66	0.6	3 1.2
46.00	5	1034	78.10	1.70	- 130 87	3	1023	77.27	1.68	9600	0.83	0.2	5	1003	31	2.34	20	1.51	0.8	3 1.5
46.50	Э	1037	78,32	1.68	- 13140	2	1025	77.42	1.66	9324	0.91	0.3	3	1006	31	2.34	19	1.44	0.9	1 1.7
47.00	6	1043	78.78	1.68	- 13198	6	1031	77.87	1.66	9180	0.91	0.3	5	1011	32	2.42	20	1.51	0.9	1 1.7
47.50	4	1047	79.08	1.66	- 13269	13	1044	78.85	1.66	9000	0.23	0.1	11	1022	25	1.89	22	1.66	0.2	3 0.4
48.00	6	1053	79.53	1.66	- 13322	4	1048	79.15	1.65	8800	0.38	0.1	u	1026	27	2.04	22	1.66	0.3	8 0.7
48.50	5	1058	79.91	1.65	- 13378	6	1054	79.61	1.64	8500	0.30	0.1	à	1035	21	1.74	19	1.90	0,3	0.6
49.00	4	1062	80.21	1.64	-13404	Ř	10.62	80.21	1.64	8250	0.00	0.0	7	104.2	20	1.51	20	1.51	0.0	0.0
49.50	5	1067	80.59	1.63	- 134 72	ă	1071	80.89	1.67	8000	-0.30	-0.1	2	1045	22	1.66	26	1.96	-0.3	0.6
50.00	6	1073	81.04	1.62	~ 13535	á	1071	80.89	1.62	8000	0.15	0_0	2	1047	26	1.96	24	1.81	0.1	5 0.1
	-									0000		~ • •	4		20	10.70	- 4			

* See LEGEND, Section C.3.

FPF01 DAT

TABLE C-4. TSC (COMPREHENSIVE) VERSUS NEW HAMPSHIRE FLASHING LIGHTS

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FLCN1 DAT

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%		7	rsç au	DORUGO	4 4 7 8 7 6 4	TNC	R.	CO NAB.	POURD		utma u		110	CT []	1966		1 900	8		
Xina	INC	LON CON	1	PUCER	TND 97	100	0100	a	PLCTD PUT DUY			8108 8761	100	Deca	6633	d rcc	0100	darte d icc	KDYPP	54 IVA T
Aring	ALL	VALC)	o ACC	FACIA	AN DEA	466	VACC	D ACC	FACIA	LOVEA	PDIEF	1442	BICH	arca	UACC	ARCC	UNCC	A ACC	PDILL	1 447
0.50	63	63	4.76	9.52	-1801	55	55	4.15	8.31	#87800	0.60	0.7	43	63	20	1.51	12	0.91	0.60	1.8
1 00	37	100	7.55	7.55	-2932	34	89	6.72	6.72	317790	0.83	0.8	27	70	30	2.27	19	1.48	0.83	1.6
1.50	27	127	9.59	6.39	-3649	26	115	0.69	5.79	268380	0.91	0.8	22	9 Z	35	2.64	23	1.78	0.91	1.6
2,00	25	152 1	11.48	5.74	-4050	29	144	10.88	5.44	234000	0.60	0.5	32	124	28	2.11	20	1.51	0.60	1.2
2.50	18	170	12.84	5.14	~4537	25	169	12.76	5.11	210000	0.08	0.1	22	146	29	1.81	23	1. 74	0.08	0.1
3.00	22	192 1	4.50	4.83	~4868	29	198	14.95	4.98	183820	-0.45	-0.3	19	165	27	2.04	33	2.49	-0.45	-0.8
3.50	30	222	16.77	4.79	-5212	19	217	16.39	4.68	169400	0.3B	0.2	30	195	27	2.04	22	1.66	0.38	0.7
4.00	19	241 1	18.20	4.55	5479	14	231	17.45	4.36	153088	0.76	0.5	11	206	35	2.64	25	1.89	0.76	1.3
4.50	25	266 2	20.09	4.46	~5751	19	250	18.88	4.20	139000	1.21	0.7	22	228	38	2.87	22	1.66	1.21	2.1
5.00	22	288 2	21,75	4.35	-5972	23	273	20.62	4.12	128000	1.13	0.6	1 9	247	41	3.10	26	1.96	1, 13	1.8
5.50	18	306 2	23.11	4.20	-6 2 19	23	296	22.36	4.06	120800	0.76	0.4	18	265	41	3.10	31	2.34	0.76	1.2
6.00	24	330 2	24.92	4.15	-6392	26	322	24.32	4.05	112720	0.60	0.3	26	291	39	2.95	31	2.34	0.60	1.0
6.50	18	348	26.28	4.04	-6580	18	340	25.68	3.95	105981	0.60	F.0	14	305	43	3.25	35	2.64	0.60	D-9
7.00	19	367 4	<i>(1, 12</i>)	3.90	-0/30	24	204	21.09	3.93	100000	0.23	0.1	22	321	40	3.02	37	2.19	0.23	0.3
7.50	18	- 206 -	(9°09)	3.88	~0915	10	3/9	20.00	1.04	9438U	0.45	V. 2	13	340	45	3.40	39	2. 95	0.45	0.7
6.00	13	398 3		3.70	-7025	10	397	27.70	3.75	90000	0.08	0.0	10	333	43	3.43	42	3.11	0.08	0.1
8.50	20	410 3	30 <u>5</u> 77	7.04	-7117	10	407	30.7%	3.04	81700	0.23	0.1	13	300	92 NG	3.70	23	2.77	0 40	0.3
9.00	17	450 -	22.40	3 55	-7459	6	4.2B	27 27	1 " " "	78980	1 4 0	0.1	19	307	9.5 6.6	5.7V	26	5 75	1 44	2 0
10 00	14	461	34 82	3 4 8	-7598	12	420	12.00	1 32	75600	1.59	0.7	17	409	52	101	31	2. 74	1 59	2.1
10.50	14	u 75 3	15.88	3.42	-7775	10	858	10.29	3.27	72000	1.59	0.7	17	476	ม์จ	3,70	28	2.11	1.59	2.0
11.00	13	488	36.86	3.35	~7872	1.8	472	35.65	3,24	69225	1.21	0.5	19	445	43	3,25	27	2.05	1,21	1.9
11.50	11	499 3	37.69	3,28	-7998	20	492	37.16	3.23	66000	0.53	0.2	12	457	42	3.17	35	2.64	0.53	0.8
12.00	19	518 3	39, 12	3.26	-8116	12	504	38.07	3.17	64000	1.06	0.4	16	473	45	3,40	31	2,34	1.06	1.6
12.50	15	533 4	0.26	3.22	-8214	11	515	38.90	3.11	61088	1.36	0.6	13	486	47	3.55	29	2.19	1.36	2.1
13.00	12	545 4	1.16	3.17	8319	8	523	39.50	3.04	59800	1.66	0.7	7	893	52	3.93	30	2.27	3.66	2.4
13.50	10	555 4	\$1.92	3.11	8447	8	531	40.11	2.97	56480	1.81	0.7	9	502	53	4.00	29	2.19	1.01	2.7
14.00	16	571 4	43.13	3.08	-8553	16	547	41.31	2.95	54000	1.81	0.7	14	516	55	4.15	31	2.34	1.81	2.6
14.50	8	579 4	3.73	3.02	~8654	15	562	42.45	2.93	51500	1.28	0.5	. 16	532	47	3.55	30	2.27	1.28	1.9
15.00	11	590 4	4.56	2.97	-8746	9	571	43.13	2.88	49940	1.44	0 - 6	9	541	49	3.70	30	2.27	1.44	2.1
15.50	12	602 4	15.47	2.93	-8880	12	583	44.03	2.84	48000	1.48	0.6	14	555	47	3.55	28	2,11	1.60	2.2
16.00	13	615 4	46.45	2.90	-8995	5	588	66.61	2.78	46554	2.04	0.8	7	562	53	4.00	26	1.96	2.04	3.0
16.50	11	626 4	1.28	2.8/	-9095	22	610	46.07	2.79	45000	1.21	0.5	20	582	44	3.32	28	2.11	121	1.9
17.00	10	636 4	48.04	2.03	-9188	<u>0</u>	616	40.55	2.74	43260	1.51	0.0	8	590	46	3.67	26	1,90	1,51	2.4
19 00	14	450 4	0 70	2-01	-9290	.,	640	10 20	2.07	01990 00200	2.00	V.0	13	603	57	3.33	2V 20	1.01	1 74	2,3
10.00	10	460 4	1757U	20/0	-0865		2/10	40.J9 N0 0h	2.09	*0330	1.30	0.5	1.3	6 0 10 6 2 M	9 <u>6</u> bii	2 2 2 2	213	1 01	1 51	2.2
19 00	10	678 9	1 21	2 70	-9903	9	656	40.JA	2.05	38800	1.51	0.0	10	630	4-5 11 (1)	1 12	29	1 66	1 66	2.7
19 50	5	683 9	1 59	2.65	-9500	12	669	50 51	7 50	37500	1 06	0.0	14	648	75	7 KA	21	1 59	1 06	1.9
20.00	11	694 9	52.42	2.62	-9749	17	686	51.81	2.59	36000	0.60	0.2	13	661	33	2.49	25	1.89	0.60	1.1
20.50	5	699 5	52.79	2.58		6	692	52.27	2.55	35000	0.53	0.2		668	31	2.34	24	1.81	0.53	0.9
21.00	5	704 5	53.17	2.53	≙9938	Ă	700	52.87	2.52	34000	0.10	0_1	, 6	678	30	2.27	26	1.96	0.30	0.5
21.50	6	710	53.63	2.49	- 100 11	9	709	53.55	2,49	33000	0,08	0.0	ŭ	678	32	2.42	31	2.34	0.08	0.1
22.00	13	723	54.61	2.48	- 10090	6	715	54.00	2.45	32000	0.60	0.2	10 [°]	688	35	2.64	27	2.04	0.60	1.0
22.50	7	730 9	55.14	2.45	- 10166	7	722	54.53	2.42	30940	0.60	0.2	6	694	36	2.72	28	2.11	0.60	1.0
23.00	10	740 5	5.89	2.43	- 10250	9	731	55.21	2.40	30000	0.68	0.2	9	703	37	2.79	28	2.11	0.68	1.1
23.50	7	747 5	56.42	2.40	- 10327	4	735	55.51	2.36	29700	0.91	0.3	6	709	38	2.07	26	1.96	0.91	1.5
24.00	10	757 9	57.18	2.38	- 104 18	9	744	56.19	2.34	28800	0.98	0.3	9	7 19	39	2.95	26	1.96	0.98	1.6
24.50	9	766	57,85	2.36	- 105 16	9	753	56.87	2.32	28000	0.98	0.3	3	721	45	3.90	32	2.42	0.98	1.5
25.00	3	769	58.08	2.32	- 10595	9	762	57.55	2.30	27000	0.53	0.2	8	729	4 0	3.02	33	2.49	0.53	0.8

* See LEGEND, Section C.3.

C-13

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TABLE C-4 (cont.)

															LCI	IT DAT
% Xing	INC COM ACC DACC	TSC N I S ACC	ODEL POWER HAZARD [*] PACTR INDEX	ÎNC Acc	NEN HI Cun Oacc % Ag	AMPSHIBE Pover CC Pacte	C I HAZABD INDEX	BITH H Sdiff	IATCH TVAL	ISC HTCH	COA Atca	TS Less Oacc	SC HATCH % ACC	LESS Ørcc	НН ВАТСВ % асс	LESS MATCH Sdiff tval
			2 20 10/76	-			7/ 774				776		3 30	30	2 67	
25.50	5 112	308+J1 Co 40	2.29 - 100/5		707 58.0	18 2.28	26270	-0.23	-0.1	0	735	37	2.19	34	2.31	0.23 0.4
20.00	5 777	50.07	2.20 - 10/39	13	702 39-0		22000	-0-30	-0.1	10	743	26	2 42	37	4.73	
20.00	0 700	50 07	2.24 0 100 32		709 59.	>> <u>∠</u> ₀∠⊃	20000	-0.23	•U•I	10	700	31	2.34		2	-0.23 -0.4
27.00	5 794	27.71	2.22 - 10900	0	795 60.0	15 2.22	24290	-0.08	-0.0	6	763	31	2.34	22	2.42	
27.50	5 /59	00.35	2.19 - 10971	2	800 60.	32 2.20	24000	~0.08	~0.0	6	769	20	2.2/	31	2.34	-0.00 -0.1
28.00	3 602	00.3/	2.10 -11037		010 60.0	30 <u>2</u> .17	23200	-0.23	-0.1	2	700	20	2 11	31	2,34	
28.00	0 010	01.00	2.15 - 11099	13	010 01.	70 2.17	22300	-0.60	- 0 J	0	70%	20	2	25	2.012	
29.00	6 017	6 1 . 7 16	2,13 -11102		023 020	2013	21000	-0.00	-0-2	c c	790	27	2.04	10	2.0%	-0.76 -1.0
29.30	7 923	62.10	2 0 0 0 11200		23 D 403	94 4.JJ	20940	-0.53	-0.2	2	801	20	2 10	36	2.0/	-0.53 -0.8
30.00	7 630	4 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	2.07 -11310			2 2011	20 900		-0.3	0 /	805	22	2 2	36	2.72	-0.30 -0.5
31.00	12 050	61 20	2.07 - 110.5	- -	041030.	12 2°00	20400	0.10	0 1		005	36	2 72	30	2.12	0.30 0.5
21.50	13 030	64 BO	2.07 - 11440	1	040 000	7 7 7 00	19690	0.30	0.1	2	817	20	3 10	30	2 3 7 7	0.90 0.0
31.50	010 0	65 11	2.00 - 11518		A22 64	1 2 VJ	19000	0.53	0.3	2	822	40	3.10		2.21	0.53 0.9
32.00	4 502	65 71	2.03 - 11607	6	950 64 9	35 2 DA	19707	A 76	0.2	2	820	4.1	3 10	31	2 34	0.76 1.7
32.30	0 870	66 70	2.02 - 11720	12	877 65 1	16 2 00 16 2 00	18040	0.53	Å 2	11	840	20	2 95	27	2.37	0.53 0.8
33.00	10 999	67 15	2.01 - 11720	6	979 66 1	1 1 98	18000	0 63	Λ J	''	847	47	3 17	31	2 34	0 43 1 3
37.00	5 99/1	67 52	1 99 11857		996 66 9	1 1 9 7	17325	0.03	0.3		954		3 02	12	2 4 2	0 60 0 9
34.50	7 901	68 05	1 97 - 11915	6	892 67.	17 1 95	16948	0.68	0.2	'n	858	6.2	3 25	30	2.57	0.68 1.0
35 00	5 904	68 43	1 96 - 11982	11	903 68 3	0 1 95	16390	0.00	0.1	ě	867	30	2 95	36	2.72	0 23 0 3
35 50	6 912	68.88	1.94 - 12048	 u	907 68.9	10 1 93	16000	0 38	D. 1	,	874	18	2 97	77	2.09	0.38 0.6
36.00	9 921	69.56	1.97 -12114	ц Ц	911 68.1	1 1.91	15600	0.76	0.2		880	41	3 10	31	2. 34	0.76 1.7
36.50	2 921	69.71	1.91 .12185	õ	911 68.8	1 1.89	15230	0.91	0.3	1	881	42	3.17	30	2.27	0.91 1.4
37 00	6 929	70 17	1.90 = 12754	Â	919 69.1	1 1.88	15000	0 76	0.2		887	82	3.17	32	2.42	0.76 1.2
37 50	A 917	70.17	1 89 - 12235	ž	972 69.1	ia 1.86	14560	1 13	0.1	ž	AGT	43	3 25	28	2.11	1, 13 1 8
38.00	10 947	71.51	1.88 - 12410	5	927 70 0	2 1.84	14,220	1.51	0.5	, ,	901	46	3.47	26	1.96	1.51 2.4
38 50	7 954	72.05	1.87 - 12470	Â	935 70.0	1 1 81	14000	1.14	0.1		907	47	3.55	28	2,11	1.44 2.2
39.00	8 962	72_66	1.86 -12537	11	946 71-1	1.83	13600	1.21	0.4	11	919	u u	3.32	28	2.11	1.21 1.9
39.50	5 967	73.04	1.85 - 12598	7	953 71-	8 1.82	13200	1.06	0.3	4	922	45	3.40	31	2.34	1.06 1.6
40.00	4 971	73.34	1.83 - 12652	ś	958 72.	16 1.81	12800	0.98	0.3	6	928	83	3.25	30	2.27	0.98 1.5
40.50	7 978	71.87	1.82 - 127 11	6	964 72.1	81 1.80	12500	1.06	0.3	7	935	43	3.25	29	2.19	1.06 1.6
41.00	9 987	74.55	1.82 - 12772	19	983 74	24 1.81	12000	0.30	0.1	14	949	38	2 . 87	34	2.57	0.30 0.5
41.50	3 990	74.77	1.80 - 12832	Ő	983 74.2	24 1.79	12000	0.53	0.2	i	950	40	3.02	33	2.49	0.53 0.8
42.00	3 993	75.00	1.79 - 12896	1	984 74.	32 1.77	11792	0.68	0.2	2	952	41	3,10	32	2.42	0.68 1.1
42.50	8 1001	75.60	1.78 - 12981	4	968 74.1	52 1.76	11368	0.98	0.3	4	956	45	3.40	32	2.42	0.98 1.5
43.00	5 1006	75.98	1.77 -13041	i	989 74.	70 1.74	11025	1.26	0.4	4	960	46	3.47	29	2.19	1.28 2.0
43.50	4 1010	76.28	1.75 ~13085	4	993 75.0	0 1.72	10800	1.28	0.4	4	964	46	3.47	29	2.19	1.28 2.0
44.00	5 1015	76.66	1.74 -13146	8	1001 75.0	50 1.72	10500	1.06	0.3	8	972	43	3.25	29	2.19	1.06 1.6
44.50	4 1019	76.96	1.73 - 13203	8	1009 76.3	21 1.71	10200	0.76	0.2	8	980	39	2.95	29	2.19	0.76 1.2
45.00	6 1025	77.42	1.72 -13270	10	1019 76.1	96 1.71	10000	0,45	0.1	9	989	36	2.72	30	2.27	0.45 0.7
45.50	5 1030	77.79	1.71 - 13345	1	1020 77.0	9 1.69	9840	0.76	0.2	1	990	40	3.02	30	2.27	0.76 1.2
46.00	8 1038	78.40	1.70 -13411	3	1023 77.	27 1.68	9600	1.13	0.3	8	998	40	3.02	25	1.89	1.13 1.9
46.50	6 1044	78.85	1.70 - 13477	2	1025 77	12 1.66	9324	1.44	0.4	Ű	1032	42	3.17	23	1,74	1.44 2.4
47.00	5 1049	79.23	1.69 - 13520	6	1031 77.	87 1.66	9180	1.36	0.4	7	1009	40	3.02	22	1.66	1.36 2.3
47.50	5 1054	79.61	1.68 -13566	13	1044 78.	95 1.66	9000	0.76	0_2	11	1020	34	2.57	24	1.81	0.76 1.3
48.00	5 1059	79.98	1.67 -13646	4	1048 79.	15 1.65	8800	0.83	0.2	4	1024	35	2.64	24	1.81	0.83 1.4
48.50	5 1064	80.36	1.66 ~13704	6	1054 79.0	51 1.64	8500	0.76	0.2	5	1029	35	2.64	25	1.89	0.76 1.3
49.00	7 1071	80.89	1.65 -13777	8	1062 80.3	21 1.64	8250	0.68	0.2	11	1040	31	2.34	22	1.66	0.68 1.2
49.50	1 1072	80.97	1.64 -13832	9	1071 80.	89 1.63	8000	0.08	0.0	1	1041	31	2.34	30	2.27	0.08 0.1
50.00	7 1079	81.50	1.63 -13899	0	1071 80.	89 1.62	8000	0.60	0.2	5	1046	33	2.49	25	1.89	0.60 1.1

* See LEGEND, Section C.3.

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FLCN1 DAT

TABLE C-5. TSC (VOLUME ONLY) VERSUS NEW HAMPSHIRE AUTOMATIC GATES

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· GPF01 DAT

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										СТ										
%		T	SC NO	DEL			N E	ER HARI	PSHIRE						T	5C	1	8		
V #	INC	CUN		POSER	HAZARD	INC	CUN	. .	POREB	HAZ AR D	WITH P	JATCH	INC	CUE	LESS	NATCH	LESS	HATCH	LESS E	TC
xing	VCC	acc 3	ACC	PACTR	INDEX	ACC	ØACC	X ACC	PACTR	INDEX	SDIPP	TVAL	RTCH	ATCA	PACC	X VCC	ØACC	S ACC	RDIPP	TVAL
	-	-		3 05	6.0 h h	-	-								-		-			<u> </u>
0.50		15	1.90	7.20	~4844	4		1.78	3.95	020020	0.00	0.0	4	0 • •	5	0.05	1	0.85	0.00	0.0
1.50	11	24	4.24 7 7h		-5075	ć	10	5 37	2 60	919360	0.20	1 0		16		1,13	3	0.00	0.28	0.4
2.00	'''	20	0.30	4 66	-5525	5	20	6 79	3.30	748000	1.30	1.0	çı n	10	11	3.11	с С	1,11	1.98	1.0
2.50	<u></u>	36 1	0 17	4 07	-6418	7	27	7 63	3 05	677660	20J7 250	1 1	و. 1	22	14	3,95	5	3 4 1	2.34	2.1
3.00	5	38 1	0_73	3.58	-67 10	, A	35	9.89	3,30	593400	0.85	0_4	7	29	9	2.54	5	1 69	A. 54 D. 85	1. I 1. D
3.50	10	48 1	3.56	3.87	-6968	š	44	12.43	3.55	540000	1,13	0.1	10	19	ģ	2.54	Š	1.07	1.13	1.1
4.00	2	50 1	4.12	3,53	-72 08	2	46	12.99	3.25	488400	1, 13	0.4	1	ŰĎ	10	2.82	6	1.69	1,13	1.0
4.50	4	54 1	5.25	3.39	-7353	3	49	13,84	3.08	458080	1.41	0.5	1	41	13	3.67	ā	2.26	1.01	1.1
5.00	6	60 1	6.95	3.39	-7507	0	49	13.84	2.77	425000	3.11	1.1	0	61	19	5.37	8	2.25	3, 11	2.1
5.50	4	64 1	8.08	3.29	-7719	5	54	15.25	2.77	400320	2.82	0.9	3	44	20	5.65	10	2.82	2.82	٩.0
6.00	0	64 1	8.08	3.01	7902	11	65	18.36	3.06	373800	-0.28	-0,1	6	50	14	3.95	15	4.24	-0.28	-0.2
6.50	5	69 1	9.49	3.00	-8036	7	72	20.34	3.13	350000	~0.85	-0.3	7	57	12	3.39	15	4.24	-D.85	-0.6
7.00	4	73 2	0.62	2.95	-8168	5	77	21.75	3.11	334400	-1.13	-0.3	2	59	14	3.95	18	5.08	-1.13	-0.7
7.50	5	782	2.03	2.94	-82 55	2	79	22.32	2,98	321600	-0.20	-0.1	6	65	13	3.67	10	3.95	~0.28	-0.2
8.00	5	83 2	3.45	2.93	-8370	4	83	23.45	2.93	308100	0.00	0.0	3	68	15	4,24	15	4.24	0.00	0.0
8.50	4	87 2	4.58	2.89	-8443	6	89	25.14	2.96	297600	r0.56	~0.2	6	72	15	4.24	17	4.80	-0.56	-0.4
9.00	T	88 2	4,86	2.76	-8532	1	90	25.42	2.82	285200	-0.56	-0.1	1	73	15	4_24	17	4.80	-0.56	-0.4
9.50	4	92 2	5.99	2.74	-8603	4	94	26.55	2.80	270100	-0.56	-0.1	5	78	14	3.95	16	4.52	-9.56	-0.4
10.00	3	95 2	6.84	2.68	-8680	2	96	27.12	2.71	261000	-0.28	-0.1	1	79	16	A. 52	17	4.80	-0.28	-0.2
10.50	2	97 2	7.40	2.61	-8781	0	96	27.12	2.58	254400	0.28	0.1	1	80	17	4.00	16	4.52	0.28	0.2
11.00	6	10.3 2	9.10	2.65	-8858	0	96	27.12	2.47	246400	1.98	0.5	D	80	23	6.50	16	4.52	1.98	1.1
11.00	1	104 2	9.38	2.00	-8921	4	100	28.25	2.40	237000	1.13	0.3	3	83	21	2.91	17	4.80	1.13	0.6
12.00	1	105 2	9.00	2.4/	-9005	4	104	29.30	2.43	228600	0.28	-0.1	4	85	20	5,65	19	5.37	0.28	0.2
13 00	11	110 2	1 07	2.00	-9038	2	111	31 03	2.01	210000	- 1, 41	-0.3	2	50	19	5 08	21	2.33	- 1.41	-0.0
13 50	6	114 3	2 20	2 39	-9098	2	113	31 02	2 36	202800	~0.05	0.1	1	92	21	5,V0 6 03	20	3.7J 6 46	0 20	~0.5
14.00	5	119 1	3.62	2.40		2	115	32 49	2.50	197145	1 13	1 7		95	21	5.50	10	5.07	1 1 1	7 2
14-50	ű.	123 3	4.75	7.40	-9317	ົ້	117	33.05	2.28	190400	1 69	D. A	2	90	23	6 7A	16	5 08	1 69	n e
15.00	ŝ	126 3	5.59	2.37	-9380	รื	122	34.46	2.30	183720	1,13	0_3	ž	101	25	7.06	21	5,93	1.13	0.6
15.50	4	130 3	6.72	2.37	-94 14	3	125	35.31	2.28	173600	1.41	0.3	5	107	23	6.50	16	5.08	1.41	0.6
16.00	3	133 3	7.57	2.35	-9488	2	127	35.00	2.24	170000	1.69	0.4	3	110	23	6.50	17	4.80	1.69	0.9
16.50	10	143 4	0.40	2.45	-9572	2	129	36.44	2.21	168000	3.95	0.8	8	118	29	8.19	15	4.24	3.95	2.1
17.00	4	147 4	1.53	2.44	-9642	3	132	37.29	2.19	161500	4.24	0.9	2	116	31	8.76	16	8.52	8.20	2.2
17.50	0	147 4	1.53	2.37	-9687	4	136	38.42	2.20	155800	3.11	0.7	1	117	30	6.07	19	5.37	3. 11	1.6
18.00	4	151 4	2.66	2.37	-9761	3	139	39.27	2.19	150000	3.39	0.7	2	119	32	9.04	20	5.65	3.39	1.7
18.50	2	153 4	3.22	2.34	-9804	2	141	39.83	2.15	147400	3.39	0.7	2	121	32	9.04	20	5.65	3.39	1.7
19.00	1	154 4	3.50	2.29	~9842	0	141	39.83	2.10	144000	3.67	0.8	1	122	32	9.04	19	5.37	3.67	1.0
19.50	I	155 4	3.79	2.25	9911	6	147	41.53	2.13	13995 0	2.26	0.5	4	126	29	6.19	21	5.93	2.26	1.1
20.00	4	159 4	4.92	2.25	~9962	4	151	42.66	2.13	135000	2.26	0.5	9	130	29	8.19	21	5.93	2.26	1.1
20.50	4	163 4	6.05	2.25	~ 100 17	1	152	42.94	2.09	131810	3.11	0.6	3	133	30	8.47	19	5.37	3.11	1.6
21.06	1	164 4	6.33	2.21	- 10041	2	154	43.50	2.07	127500	2.82	0.6	2	135	29	B. 19	19	5.37	2.82	1.4
21.50	۲ ۲	167 4	1.18	2.79	-10091	4	158	44.63	2.08	123760	2.54	0.5	5	741	26	7.34	17	9.80	2.54	1.4
22.00	2	109 4	9 5 0	2.1/	- 1014/	1	159	44.92	2.09	110500	2.62	0.6	2	103	26	1.59	10	4.52	2.82	1.5
22.00	ב "	172 4	0.77	2.10	~ 10193		100	VJ.20	2.01	110500	5.59	0.7	2	145	27	1.03	15	4.24	5. 59	1 3
23.00	4	170 4	7.12	2 14	- 10241	2	107	10.01 17 10	2.01	114308	۱۱ ، د	V.0	6	151	25	1.00	14	3.95	3,11	1.0
24.00	2	180 5	0.20	2 12	- 10294	2	160	17 70	1 00	100120	2 11	0.0	G) Q	155	23	6 70	12	3°17 27	3 94	1 0
24,00	1	181 5	1.12	2.00	- 10333	2	177	4//4 10 50	1 0 0	105000	J. J. 54	0.0 A E		150	24	7 64	16	5.07	J 6 1 1	1 8
25.00	- -	184 5	1.98	2.09	- 104.08	0	173	48.59	1 04	103000	4°74 V	0 ° C	U 2	120	20	מניך	10	3 65	2+34 3 20	1.0
23.00		.04)	• • 7 9	6400	- 10-100	U	1.12	70.39	1.74	102400	ער יר	Vab	2	138	20	1. 34	14	J. 7J	2.37	107

* See LEGEND, Section E.3.

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TABLE C-5 (cont.)

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%			TSC MC	DEL	بد		NEM HVE	PSHIRE	U I					19	SC .		8		
Xina	INC	CUM		POVER	HAZARD ~	INC	CDM	POVER	HAZARD	WITH P	ATCH	180	C 0 8	L ES S	AVLCH	L ESS	BATCH	LESS 8	ATCH
Aring	ACC (¢λCC ¦	% ACC	FACTR	INDEX	ACC	#ACC % ACC	PACTR	INDER	X DIYY	TVAL	atca	HTCH	OACC	X ACC	Ø Å CC	% ACC	SDIFF	TVAL
25.50	1	185	52.26	2.05	- 104 48	Э	175 49.44	1.94	100000	2.82	0.5	0	150	27	7.63	17	4.80	2.62	1.5
26.00	3	188	53.11	2.04	- 10487	4	179 50.56	1.94	98064	2.54	0.5	1	159	29	8.19	20	5.65	2.54	1.3
26.50	2	190 !	53.67	2.03	- 105 19	0	179 50.56	1.91	96000	3.11	0.6	1	160	30	8.47	19	5.37	3.11	1.6
27.00	2	192 9	54.24	2.01	~10555	0	179 50.56	1.87	93500	3.67	0.7	1	161	31	8.76	18	5.00	3.67	1.9
27.50	0	192	54.24	1.97	- 10597	6	185 52,26	1.90	90000	1.98	0.9	6	167	25	7.06	18	5.08	1.98	1.1
28,00	5	197	55.65	1.99	- 10650	3	180 53.11	1.90	87696	2.54	0.5	3	170	27	7.63	18	5.09	2.54	1.3
23.50	1	198	55.93	1.96	~ 10695	1	189 53.39	1.87	86350	2.54	0.5	1	171	27	7.63	18	5.08	2.54	1.3
29.00	0	198	55.93	1.93	- 10741	5	194 54,80	1.89	84000	1.13	0.2	3	174	24	6.78	20	5.65	1,13	0.6
29.50	3	201	56.78	1.92	~ 10786	0	194 54.80	1-86	82 0 8 0	1.98	0.4	2	176	25	7.06	14	5.08	1.98	1.1
30.00	3	204	57.63	1.92	- 10832	ů.	198 55.93	1.86	80000	1-69	0.3	ā	179	25	7.06	19	5 37	1.69	6.9
30.50	1	205	57.91	1,90	- 10856	ò	198 55.91	1.43	79200	1.98	0.3	ñ	170	26	7.38	19	5.37	1 00	1 6
31 00	i	206 9	58.19	1.88	- 108.86	2	200 56.50	1.82	78200	9 69	6 3	1	180	26	7 34	20	5 65	1 60	8 9
31 50	;	200 .	58 17	1 96	- 100 00	ñ	200 56 50	1 70	75400	1 04	6.0		180	20	7.50	20	5,05	5 00	107
37.00	'n	207	58 47	1 87	- 10952 - 10967		204 57 67	1 60	73100	16 70 A AE	0.3	2	19.2	21	7.00	20	2.03	1.50	1.V 6 4
33 50	1	200	50 72	1 0 1	- 100.07	г	204 57805	4 03	71000		-04		104	23	1.00	44	9 64		000
32.50	2	200 .	CA 17	1.01	~ 10 7 02		211 37.00	1.03	71000	··v. 65	-0.1		100	42	0.21	25	7.06	~V.03	- V.aq
33.00	2	213		1.07	110 17		213 60.17	1.02	70000	0.00	0.0	3	189	24	0./8	24	0.70	0.00	0.0
33.00		219 0	01.00	1.00	~ 110 52	~	215 60.73	1.81	56000	1-13	0.2	4	131	26	7.91	24	0.78	1.13	0.6
34.00	1	220 0	62.15	1.83	~ 11103	0	215 60.73	1.79	65009	1.41	0.2	1	192	20	7.91	23	6.50	1.41	0.7
34.50	2	223 0	62.99	1.83	-11137	1	216 61.02	1.77	64880	1.98	0.3	3	195	26	7.91	21	5.93	1.98	1.0
35.00	5	228 0	64.41	1.84	- 11193	2	218 61.58	Y. 76	63250	2.82	0.5	1	196	32	9.04	22	6.21	2.82	1.4
35.50	1	229 (64.69	1.82	- 11218	3	221 62.43	1.76	62020	2.26	0.0	3	199	30	0.47	22	6.21	2,26	1.1
36.00	2	231 (65.25	1.81	- 11253	1	222 62.71	1.74	60600	2.54	0.4	1	200	31	8.76	22	6.21	2.54	1.2
30,.50	1	232 (65.54	1.80	- 11303	0	222 62.71	1.72	60000	2.82	0.5	t	201	31	8.76	21	5.93	2.82	1_4
37.00	3	235 (66.38	1.79	- 11358	1	223 62.99	1.70	58104	3.39	0.6	2	233	32	9.08	20	5.65	3.39	1.7
37.50	5	240	67.80	1.81	~ 11397	6	229 64.69	1.73	56240	3.11	0.5	9	212	28	7.91	17	6,80	3.11	1.6
38.00	З	243 (68.64	1.81	-11443	Э	232 65.54	1.72	55000	3.11	0.5	3	215	28	7.91	17	4.80	3, 11	1.6
38.50	2	245 6	69.21	1.80	- 11477	4	236 66.67	1.73	54000	2.54	0.4	2	217	28	7.91	19	5.37	2.54	1.3
39.00	6	251	70.90	1.82	- 11532	0	236 66.67	1.71	52 500	4.24	0.7	2	219	32	9.04	17	4,80	4.24	2.1
39.50	1	252	71.19	1.80	-11556	Ō	236 66.67	1.69	51320	6.52	0.7	0	219	33	9.32	17	8.80	4.52	2.3
40.00	1	253	71.47	1.79	~ 11587	2	218 67.23	1.68	50040	4.24	0.7	ĩ	220	33	9.32	18	5.08	4.26	2.1
40.50	4	257	72.60	1.79	-11633	1	239 67.51	1.67	#900D	5.08	0.8	2	222	35	9,89	17	B. 80	5.08	2.5
41.00	2	259	73.16	1.78	~ 11676	1	240 67.80	1.65	48000	5.17	0.9	1	223	36	10.17	17	8,80	5.37	2.6
41,50	1	260	73.45	1.77	~ 11706	i	241 68.08	1.64	87100	5.37	0.0	9	226	36	10.17	17	A . AO	5.17	2.6
42.00	1	261	73.73	1.76	-11755	ò	241 68.08	1-62	#6000	5.65	0.9		224	17	10.45	17	A . AO	5.65	2 7
42.50	1	262	74 01	1.74	a 11779	ň	200 68.91	1.62	45000	5.09	n e		226	36	10 17	10	5 00	5 0.9	7 a
42,00	ò	262	74.01	1 72		5	244 60.03	1 60	40000	5 08	0°0	ົ	220	26	10.17	10	5.00	5 00	2
43.50	ž	2611	74 50	1 79	-11977	Ň	244 000 31	1 50	44000	5,00	0.0	2	220	20	0 90	16	5,00 h 50	5 37	2.0
41.00	5	204	74 50	1 20	- 11809	-	293 03061	4 60	41700	J.J.J/ M. 00	0.0	د	229	30	7007	10	0 a J2	3.37	2.01
44.50	2	204	74.30	1.07	- 110/6	2	24/ 03.//	1033	% 1 2 0 U	7.0V	0.0	-	230	30	9.00		4.00	9.0U	2.9
44.30	2	201	75 00	1.03	* 11343	6	249 70.39	1.38	00000	2.08	0.0	4	232	22	3.09	11	0,00	5.08	2.3
43.00	2	209	76 66	1.09	- 11772	0	249 70.34	1000	80000	3.65	0.9	ų į	226	31	10.43	1/	4.00	3.65	1.1
40.00	~	271	70.00	1.00	• 120 30	2	251 70.90	1,30	39080	5.65	0.9	6	230	31	10.05	17	4.00	5.05	8.1
40,00	v v	2/1	10.55	1.66	~ 12048	4	255 12.03	1.57	37800	4.52	0.7	4	∡38	13	9 J2	17	4.00	9.52	Z.3
46.50	4	273	11.12	1.66	~ 12087	1	256 72.32	1.56	37200	4.80	0.7	3	241	32	9.08	15	8.24	4.80	2.5
47.00	2	275	//.68	1.65	~ 12110	4	260 73.45	1.56	36160	9.29	0.6	0	245	30	8.47	15	6.Z4	4.24	2.2
47.50	3	278	78.53	1.65	- 12158	1	261 73.73	1.55	36000	4.80	0.7	1	286	32	9.00	15	4,24	4.80	2.5
48.00	1	279	78.81	1.64	-12188	્ 1	262 74.01	1.54	34420	4.80	0.7	1	207	32	9.04	15	4.24	4.60	2.5
48.50	2	281 '	79.38	1.64	- 12204	5	264 74.58	1.54	33600	4.80	0.7	3	2 5 0	31	0.76	14	3.95	4.80	2.5
49.00	4	285	80.51	1.64	- 12240	0	264 74.50	1.52	32500	5.93	0.9	2	252	33	9.32	12	3,39	5.93	3, 1
49.50	0	285	80.51	1.63	- 12276	0	264 74.58	1.51	32000	5.93	0.9	0	252	33	9.32	12	3.39	5.93	3.1
50.00	0	285	80.51	1.61	- 12312	1	265 74.86	1.50	30800	5.65	0.9	D	252	33	9.32	13	3.67	5.65	2.9

*See LEGEND, Section C.3.

GPF01 DAT

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TABLE C-6. TSC (COMPREHENSIVE) VERSUS NEW HAMPSHIRE AUTOMATIC GATES

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GPF10 DAT

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9			75 7 8 0	JDEL			N RU HA	APSHIRE	CT					79		19	A		
<i>/</i> 0	INC	CUM	100 11	POWER	HAZARD*	INC	CUA	POVER	HAZABD	WITH H	ATCH	INC	CDA	LESS	SATCE	LZSS	BATCH	LESS B	ATCH
Xing	ACC	BACC S	S ACC	FACTR	INDEX	ACC	OACC S AC	C PACTE	INDEX	SDIPP	TVAL	STCH	ATCH	ØACC	S ACC	ACC	S ACC	SDIPP	TVAL
-																			
0.50	10	10	2.82	5.65	-3964	7	7 1.9	8 3,95	1272000	0.85	0.7	б	6	5	1.13	1	0.20	0.85	1.3
1.00	5	15	4.24	4.24	-4745	7	14 3.9	5 3.95	939420	0.28	0.2	3	9	6	1.69	5	1.41	0.20	0.3
1.50	3	18	5.08	3.39	-5173	5	19 5.3	7 3,50	819360	-0.28	-0.2	2	11	7	1.98	8	2.26	-0.28	-0.3
2.00	10	28	7.91	3.95	-5728	5	24 6.7	8 3.39	748000	1.13	0.6	2	13	15	4.24	11	3.11	1.13	0.0
2.50	14	42	11-86	4.75	~6079	3	27 7.6	3 3.05	677660	4.24	1.8	0	17	25	7.06	10	2.82	9.24	2.5
3.00	2	44	12.43	4.14	~6324	8	35 9-8	9 3.30	593400	2.54	1.0	6	23	21	5.93	12	3.39	2.50	1.6
3.50	5	47	13.28	3.79	-6551	2	44 12.4	5 3.55	540000	0.85	0.3	Þ	29	18	5.08	15	4.24	0.85	0,5
4.00	2	50	15 50	3, 23	~6983	ź	40 14.59	00 C N	400400	1.60	0.4	2	29	21	3,93	10	4.00	1.13	0.0
N. JU	2	50	12.24	2,27	-0902	2	47 13.0	יסס כי יי דיד רי וו	436000	2 9 2	3 0	2	30	29	7 04	101	J . UO A Jh	1.07	0.9
5.00	5	59	10.07	3,33	-7101	š	47 13.0 Ea 15 7	\$ <u>6</u> ,11 \$ 7 77	400325	2 8 2	^ a	3	10	23	6 70	10	3 05	3.04	1.0
5.00	5	55	10.00	3 11		11	55 10 2	5 3 06	377000	0.26	0.j	š	36	20	C 66	10	5,33	4.94	0.0
6 50	2	68	19.71	2.96	-7594		77 76 1	1 3,13	350000	-1.13	-0.1	2	NA NA	20	5.65	20	6.7A	-111	-0.6
7 00	ű.	72	20 34	2.91	-7731	5	77 21.7	5 1.11	334400	-1.41	-0.0	A	52	20	5.65	25	7.06	-1.81	-0.7
7.50	7	79	22.32	2.98	-7868	ž	79 22.3	2 2.98	321600	0,00	0.0	0	56	23	6.50	23	6.50	0.00	0.0
8.00	ų	83	23.45	2.93	-8009	ū	83 23.4	5 2.93	308100	0.00	0,0	2	58	25	7.06	25	7.06	0,00	0.0
8.50	3	86	24.29	2.86	-8138	6	89 25.1	¥ 2.96	297600	-0.85	-0.2	4	62	24	6.78	27	7.63	-0.05	-0.4
9.00	3	89 2	25.14	2.79	-8280	1	90 25.4	2 2.82	285200	-0.28	-0.1	8	66	23	6.50	24	6.78	-0.28	-0.1
9.50	5	94	26.55	2.80	-8381	4	94 26.5	5 2.80	270100	0.00	0.0	5	71	23	6.50	23	6.50	0.00	0_0
10.00	1	95 2	26.84	2.68	8506	2	96 27.1	2 2.71	261000	-0.28	-0.1	1	73	23	6.50	24	6.78	-0.28	~0.1
10.50	1	96 .	27.12	2.58	-86 01	0	96 27.1	2 2,50	254400	0.00	0.0	D	72	24	6.78	24	6.78	0.00	0.0
11.00	2	98 2	27.68	2.52	-8676	Û	96 27.1	2 2,47	246400	0.56	0.1	1	73	25	7.06	23	6.50	0.56	0.3
11,50	6	104	29.38	2.55	-8778	3	100 28.2	5 2.46	237000	1.13	0.3	6	79	25	7.06	21	5.93	1.13	0.5
12.00	2	106 2	29.94	2.50	-8660	4	104 29.3	B 2.45	228600	0.56	0.1	3	01	25	7.06	23	6.50	0.56	0.3
12.50	3	109	30.79	2.46	-8942	7	111 31.3	6 2.51	220000	-0.56	-0.1	1	82	27	7.63	29	8.19	-0.56	-0.3
13.00	E	112	31.64	2.43	-8994	2	113 31.9	2 2.46	210000	-0.28	-0.1	6	86	26	7.30	27	7.63	-0.28	-0.1
13.50	ź	114	32.20	2.39	-9076	0	113 31.9	2 2.36	202800	0.28	0.1	2	68	26	1.30	25	7.05	0.20	0.1
14.00	2	119	33.62	2.40	-9120	4	115 32.4	9 2.32	197145	1.13	6.0	5	91	28	7.91	24	0.75	1.19	0.8
14.00		121.	34.10	2.30	~7220	<u>د</u>	111 33.0	2 <u>%</u> 8%0 K 2 20	190400	ندارد مع ۲	V.J	6	y 3 0 0	20	0 10	69 72	00/0 K EN	1.13	0.0
15.00		120	30°10'	2.41	~93 18	2	122 34.4	0 2.JV	103/20	1.09	0,4	5	עע פרמים	27	7 67	23	100 JU 4 ED	1.07	0.0
16 00	-	130	76 77	2 30	-9476	2	122 22.3	1 2.20 9 7 70	170000	0.95	0.3	5	102	27	7 91	25	9.00 9.06	0.85	0.0
16.50		133	37.57	2.28	-9420	2	120 16.4	ሀ ፈሪፈማ ዘ ጋ ጋዓ	168000	1 1 7	n 2	1	10.3	10	8.47	25	7.38	1.17	0.5
17.00	า้	136	38 42	2.26	-9556	า	132 37.2	ຊີ້ງຳຊ	161500	1,11	ň 2	,	105	31	8.76	27	7.63	1.11	0.5
17.50	6	142	40-11	2.29	-9610	ū	136 38-8	2 2.20	155800	1.69	0.4	ŝ	110	32	9,00	26	7.30	1-69	0.8
18.00	3	145 /	40.96	2.28	~9690	3	139 39.2	7 2.18	150000	1.69	0.4	Š	115	30	8.47	25	6.78	1.69	0,8
18.50	6	151	42.66	2.31	-9763	ž	141 39.8	3 2.15	147400	2.82	0.6	Š	120	31	0.76	21	5.93	2.62	1.8
19.00	5	156	44.07	2.32	-9865	0	141 39.8	3 2.10	144 000	4.24	0.9	2	122	34	9,60	19	5.37	9,29	2.1
19.50	0	156	44.07	2.26	~9926	6	147 41.5	3 2.13	139950	2.54	0.5	1	123	33	9.32	24	6.78	2,54	1.2
20.00	5	16 1	45.48	2.27	-9982	4	151 42.6	6 2.13	135000	2.82	0.6	5	128	33	9.32	23	6.50	2.82	1.3
20.50	2	163	46.05	2.25	-100.34	1	152 42.9	8 2.09	131810	3.11	0.6	3	131	32	9.04	21	5.93	3.11	1.5
21.00	2	165 /	46.61	2.22	- 10080	2	154 43.5	0 2,07	127500	3.11	0.6	1	132	33	9.32	22	6.21	3.11	1.5
21.50	7	172	48.59	2.26	- 10 1 29	4	158 44.6	3 2.08	123760	3.95	0.8	6	130	34	9.60	20	5.65	3.95	1.9
22.00	3	175	49.44	2.25	- 10161	1	159 44.9	2 2.04	120000	4.52	0.9	2	140	35	9.89	19	5.37	9.52	2.2
22.50	2	177 9	50.00	2.22	- 10217	1	160 45.Z	0 2.01	118500	4_80	0.9	3	103	34	9.60	17	4.90	4.80	2.4
23.00	4	181	51.13	2.22	- 10283	5	165 46.6	2.03	114504	4.52	0.9	5	140	33	9.32	17	8.80	8 .52	2.3
ZJ. 50	Ł	184	51.98	2.21	- 10321	2	167 47.1	8 2.01	111650	4.80	0.9	3	151	33	9.32	16	5.52	4.00	2.6
24.00	1	100	32°76 23'7'	2.18	10395	2	109 07.7	0 1.99	108120	8.52	0.9	Z	153	52	3.04	10	9.JZ	4 .32 E 02	2.3 2 F
24.JU	Š	190 3	53.0/	2.19	- 10449	1	1/2 48.5	y 1.98	105000	5.08	0.9	2 Z	155	55	y.09	47	4.0U 5.70	5.08	2,3 74
¥3.VV		וצו	27°22	2.10	- 10432	a a	1/2 40.5	y 1.94	102400	2.51	1.0	Q.	105	96	10.17		4.00	3.3/	2.0

* See LEGEND, Section C.3.

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TABLE C-6 (cont.)

									_	СТ									-	
%			TSC M	DDEL			n	EN HAR	PSHIRE	01					T:	5C		R		
Vána	INC	CUM		POWER	HAZARD 🕊	INC	CON		POWER	HAZARD	WITH M	SATCH	INC	COH	L ES S	NATCH	LESS	HATCH	LESS A	A TC H
xing	ACC	# ACC	% ACC	PACTR	INDEX	ACC	0 ACC	X ACC	¥ ACT R	INDEI	% DIFF	I AVT	OTCE	RTCB	ÛACC	Z VCC	ØACC	S ACC	NDIFF	TVAL
25.50	3	194	54.80	2.15	- 10541	Э	175	49.44	1.94	100000	5.37	1_0	0	155	39	11.02	20	5.65	5.37	2.5
26.00	з	197	55,65	2.14	-10574	ų	179	50,56	1.94	9806%	5.08	0.9	3	158	39	11.02	21	5.93	5.08	2.3
26.50	0	197	55.65	2.10	- 10643	0	179	50,56	1.91	96000	5.08	0.9	0	158	39	11.02	21	5.93	5.08	2.3
27.00	2	199	56.21	2.08	- 106 87	0	179	50.56	1.87	93500	5.65	1.0	0	158	41	11.58	21	5.93	5.65	2.5
27.50	2	20.1	56.78	2.06	- 10738	6	185	52.26	1.90	90000	4.52	0.8	5	163	38	10.73	22	6.21	4.52	2.1
28.00	1	202	57,06	2.04	- 10786	3	186	53.11	1.90	87696	3,95	0.7	2	165	37	10.45	23	6.50	3, 95	1.8
28.50	3	205	57.91	2.03	~ 10845	1	189	53.39	1.07	86350	6.52	0.8	2	167	38	10.73	22	6.21	8.52	2.1
29.00	1	206	58.19	2.01	- 10877	Ś	194	54.80	1.89	B&000	3.39	0.6	ă	171	35	9.89	23	6.50	3.39	1.6
29.50	3	20.9	59.04	2.00	-10898	Ō	194	54.80	1.86	82080	6.78	0.7	1	172	37	10.05	22	6.21	N. 20	2.0
30 00	6	215	60.73	2.02	- 10944	ũ	198	55.93	1.86	80000	8.80	0.8	, 0	172	43	12.15	26	7.38	a . 80	2.0
30.50	Ň	215	60 73	1.99	- 10986	õ	199	55.93	1.83	79200	4. BO	0.8	ň	172	63	12.15	26	7.30	A_ 60	2.0
31 00	č	221	67 113	2 01	-11048	ž	200	56 50	1 82	78200	5 9 1	1 0	2	175	46	12 00	25	7 06	5 0 1	2 5
31.50	ິ້	221	67 99	2 00		ñ	200	56 50	1 79	75400	6 50	1 1	4	176	л7 Д7	13.78	20	% 70	6 SN	2 7
32.00	1	223	61 39	1 99	-1100/	н	200	57 61	1.00	77100	5 6 5	1 0	9	177	A7	17 78	27	7 63	5 65	2 2
32.00	2	229	63.20	1 06	~111.34	ž	211	50 60	1 03	71000	5.03	0 7	0	105	4 I	11 50	26	7.05	, J. V.J.	1 0
32.00		220	4 h h h	1 05	- 1124	5	293	50.00	1 03	70000	4.24	87		103		11.50	30	7.39	4 20	1.0
33.00	2	220	24.41	1.75	- 11200	2	213	60.17	9 64	60000	9.49	0.1	<u>د</u>	107	41	11.30	20	7030	4.24	1.0
22.20		2 2 2 2	45 03	1.00	11230	4	213	60,73	\$ 70	66000	4.UV C 04	0.0		107	45	12.71	20	7 85	5.00	2.0
34.00		233	66 30	1 00	- 11323	ÿ	213	61 02	4 77	66000	5.08	0.7		107	903 h E	16077	20	7.91	2°08 5°18	2.1
34.50	2	233	00.30	1.72	14507		210	01.02	1.11	60000	2.37	0.7	2	190	63	12.00	20	1.30	5.5/	2.2
35.00	<u>د</u>	238	07.23	1.92	- 11423	4	218	01.30	1.70	63450	5.03	0.9		192	40	12.99	20	7.34	2002	5.7
35.50	2	240	67.80	1.31	-11460	5	221	02.43	10/0	62020	2,11	0.9		190	40	12.53	25	7.00).J/	2.3
36.00	0	240	67.80	1.88	-11499		666	02.71	1.74	60600	5.08	0.0	, u	196	44	12. 53	26	1.30	5.08	~ ~ ~
36.50	1	241	68.08	1.87	-11553	0	222	62.71	1.72	60000	5.37	0.9		197	99	12.01	25	1.06	5.3/	<u> </u>
. 37.00	1	242	46.50	1.85	-11596	1	223	02.99	1.70	58104	5.37	0.9	0	197	45	12.71	25	7.54	5.37	2.3
37.50	1	243	68.64	1.83	- 1164/	b	229	04.09	1.73	56240	3.42	0.0		204	39	11.02	25	7.06	7.92	1.0
38.00	1	244	98"23	1.81	-11672	اد	232	65.54	1.72	55000	3.39	0.6	2	206	30	10.73	26	/. 39	3. 19	1.5
38.50	0	244	68.93	1.79	~11687	4	2 36	66.67	1.73	54000	Z.26	0.4	2	206	36	10.17	28	7.91	2.26	1.0
39.00	1	245	69.21	1.77	~ 11725	0	236	66.67	1.71	52500	2.58	0.4	0	208	37	10.45	28	7.91	2.54	1.1
39.50	3	248	70.06	1.77	-11771	0	236	66.67	1.69	51320	3.39	0.5	1	209	39	11.02	27	7.63	3.39	1.5
40.00	0	248	70.06	1.75	- 11793	2	238	67.23	1.68	50040	2.02	0.5	1	210	36	10.73	20	7.91	Z. 82	1.2
40.50	1	249	70.34	1.74	- 11823	ĩ	239	67.51	1.67	49000	2.82	0.5	2	212	37	10.45	27	7.63	2.02	1.3
41.00	1	250	70.62	1.72	- 11863	1	240	67.80	1.65	48000	2.82	0.5	1	213	37	10.45	27	7.63	2.02	1.3
41,50	0	250	70.62	1.70	-11906	1	241	60,08	1.64	47100	2.54	0.4	1	215	36	10.17	27	7.63	2.54	1.1
42.00	0	250	70.62	1.68	- 119 39	0	241	68.08	1.62	46000	2.54	0.4	0	218	36	10.17	27	7.63	2.54	1.1
42,50	1	251	70.90	1.67	~ 11969	3	244	68.93	1.62	45000	1.98	0.3	1	215	36	10.17	29	0.19	1.98	0.9
43.00	1	252	71.19	1.66	-12000	0	244	68.93	1.60	44000	2.26	0.4	3	215	37	10.45	29	8.19	2.26	1.0
43.50	1	253	71.47	1.64	- 12046	1	245	69.21	1.59	42650	2.26	0.4	8	216	37	10.45	29	0.19	2.26	1.0
44.00	1	254	71.75	1.63	- 12077	2	247	69.77	1.59	41280	1.98	0.3	2	218	36	10, 17	29	8.19	1.98	0 .9
44.50	0	254	71.75	1.61	-12120	2	249	70.34	1.50	40000	1.41	0.2	9	218	36	10.17	31	8.76	1.41	0.0
45.00	0	254	71,75	1.59	- 12142	0	249	70.34	1.56	40000	3.41	0.2	0	216	36	10.17	31	8.76	1.41	0.6
45.50	0	254	71.75	1.58	- 12181	2	251	70.90	1.56	39060	0.85	0.1	1	219	35	9.89	32	9.04	0.05	0.9
46.00	1	255	72.03	1.57	- 12220	4	2 5 5	72.03	1.57	37800	0.00	0.0	3	221	34	9.60	34	9.60	0.00	0.0
46.50	5	260	73.45	1.58	∝ 12261	1	256	72.32	1.56	37200	1,13	0.2	8	225	35	9.89	31	B.76	1.13	0.5
47.00	2	262	74.01	1.57	-12306	4	260	73.45	1.56	36160	0.56	0.1	3	220	34	9.60	32	9.04	0.56	0.2
47.50	1	263	74.29	1.56	<u>- 12339</u>	1	261	73.73	1.55	36000	0.56	0.1	1	229	34	9.60	32	9.04	0.56	0.2
48.00	1	264	74.58	1.55	-12364	1	262	74.01	1.54	34450	0,56	0.1	1	230	39	9.60	32	9.04	0.56	0.2
48.50	1	265	74.86	1.54	~ 124 17	2	264	74.58	1.54	33600	0,28	0.0	2	232	33	9.32	32	9.00	0.28	0.1
49.00	1	266	75.14	1.53	- 12468	0	264	74.58	1.52	32500	0.56	0.1	1	233	33	9.32	31	8.76	0.56	0.3
49.50	2	268	75.71	1.53	- 12502	0	264	74.58	1.51	32000	1.13	0.2	2	235	33	9.32	29	8.19	1.13	0.5
50.00	2	270	76.27	1.53	- 12548	1	265	74,86	1.50	30800	1.41	0.2	2	237	33	9.32	28	7.91	1.41	0.6

* See LEGEND, Section C.3.

GPF10 DAT

TABLE C-7^{*} SPECIAL EOC AND POWER FACTORS FOR NEW HAMPSHIRE--FLASHING LIGHTS CASE FULL DATA BASE

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%		-	CT									L	NLN (DAT
Xing	# ACC	% АСС	HAZ IND	PP	**.65	** .70	**.75	**.80	LOG LOG	LOSLOG	DELLN	DELLO	DELTA	7-1VE
0.50	120	4.53	464000	9.06	0, 75	0.69	0.63	0.58	0-790	1-667	0.000	0 000	0 000	n 00n
1.00	183	6.91	332000	6.91	0.72	0.66	0-61	0.57	0.659	1.527	0.111	D 140	0 9 37	0.000
1.50	233	8.79	267820	5.86	0.70	0.65	0.60	0.56	0.570	1.435	0.089	0 092	0 962	0 200
2.00	292	11.02	230400	5.51	0.70	0.66	0.61	0.57	0.534	1.364	0.036	0.071	0.503	0.000
2 50	347	13.09	205200	5.24	0_71	0.66	0.62	0.58	0.504	1. 305	0.030	0.059	0.512	0 748
3.00	392	14.79	181300	4.91	0.71	0.66	0.62	0.58	0.467	1.255	0.037	0.051	0.711	0 715
3.50	438	16.53	165600	4 77	0 71	0.67	0.63	0.59	0 040	1 2 10	0.027	0 045	0 6 10	0 663
4_00	472	17-81	152688	4.45	0.70	0.66	0.62	0.59	0 401	1 169	0.039	0 04 1	0 950	0 607
4.50	510	19.25	140480	4.28	0.70	0.66	0.62	0.59	0 374	1.117	3.027	0 037	0.735	0.609
5.00	550	20.75	129540	4.15	0.70	0.66	0.63	0.59	0.353	1.097	0.021	0.035	0.601	0.591
550	602	22.72	121200	0.13	0.71	067	0.64	0.61	0.350	1.065	0.003	0.032	0.108	0 589
6.00	642	24.23	114520	4.04	0.71	0.68	0.64	0.61	0.333	1.034	0.016	0.030	0.529	0.502
6.50	679	25.62	108000	3.94	0.71	0.68	0.65	0.61	0.316	1.006	2.017	0.029	0.601	0 517
7.00	715	26.98	101500	3.85	0.71	0.68	0.65	0.62	0.300	0.976	0_016	0.027	0.600	0.548
7.50	757	28.57	96000	3.81	0-72	0.69	0.65	0.62	0-291	0-952	0.009	0.026	0.337	0.640
8.00	785	29.62	91040	3.70	0.72	0.68	0.65	0.62	0.269	0.927	0.021	0.025	0.845	0.687
8.50	813	30.68	86560	3.61	0.71	0.68	0.65	0.62	0.250	0.907	0.020	0.024	0.812	0.716
9.00	841	31.74	82500	3,53	0.71	0.68	0.65	0.62	0.231	0.879	0_018	0.023	0.781	0.760
9.50	867	32.72	79200	3_44	0.71	0.68	0.65	0.62	0.212	0.856	0.019	0.023	0.833	0.813
10.00	893	33.70	75400	3.37	0.71	0.68	0.65	0.62	0.195	0.834	0.018	0.022	0.805	0.734
10.50	916	34-57	72000	3.29	0.70	0.67	0.65	D-62	0.175	0.813	0.019	0.021	0.907	0.745
11.00	943	35.58	69749	3.23	0.70	0.67	0.65	0.62	0.160	0.792	0.015	0.021	0.708	0.748
11.50	979	36.94	66450	3,21	0.71	0.68	0.65	0.63	0.154	0.771	5,006	0.020	0.293	0.751
12.00	1001	37.77	64000	3.15	0.70	0.68	0.65	0.63	0.137	0.752	3.018	0.070	0.885	0.747
12.50	1025	34.68	61180	3.09	0.70	0.68	0.65	0.63	0.122	0.732	0.015	0.019	0.774	0.726
13.00	1046	39.47	60000	3.04	0.70	0.67	0.65	0.63	0.105	0.713	0.017	0.019	0.888	0.710
13.50	1069	40.34	56968	2.99	0.70	0.67	0.65	0.63	0.090	0.694	2.015	0.019	0.777	0.785
14.00	1092	41.21	54880	2.94	0.70	0.67	0.65	0.63	0.077	0-676	0.014	0.018	0.757	0.731
14.50	1118	42.19	52800	2.91	0.70	0.67	0.65	0.63	0.066	0.658	0.011	0.018	0.598	0.773
15.00	1139	42.98	50656	2.87	0.69	0.67	0.65	0.63	0.051	0.640	D.014	0.018	0.014	0.797
15.50	1166	44.00	49200	2.84	0.70	0.67	0.65	0.63	0.042	0.623	0.009	D.017	0.512	0.695
16.00	1181	44.57	48000	2.79	0.69	0.67	0.65	0.63	0.024	0.606	0.018	0.017	1.068	0.629
16.50	1156	45.13	46000	2.74	0.69	0.67	0.65	0.63	0.006	3.589	0.018	D.017	1.056	0.703
17.00	1231	46.45	44800	2.73	0.69	0.67	0.65	0.64	0.005	0.572	0.001	0.017	0.060	0.644
17.50	1248	47.09	43050	2.69	0.69	0.67	0.65	0.63	0.010	0.556	0.005	0.016	0.296	0.704
18.00	1261	47.58	41720	2.64	0.68	0.67	0.65	0.63	0.028	0.539	0.018	0.016	1.114	0.684
10.50	1288	48.60	40170	2.63	0.69	0.67	0.65	0.64	0.035	0.523	0.006	0.016	0.398	0.627
19.ÒO	1304	49.21	39520	2.59	0.68	0.67	0.65	0.63	0.050	0.507	0.015	0.016	0.938	0,695
19.50	1320	49.81	38400	2.55	0.68	0.66	0.65	0.63	0.064	0.491	0.015	0.016	0.925	0_787
20.00	1341	50.60	37125	2.53	0.68	0.67	0.65	0.63	0.074	0.476	0.010	0.016	0.655	0.715
20.50	1364	51.47	36000	2.51	0.68	0.67	0.65	0.64	0.083	0.460	0.008	0.015	0.538	0.722
21.00	1379	52.04	35000	2.48	0.68	0.66	0.65	0.64	0.097	0.445	0.014	0.015	0.940	0.717
21.50	1400	52.83	34000	2.46	0.68	0.67	0.65	0.64	0.106	0.430	0.009	0.015	0.613	0,728
22.00	1424	53.74	33000	2.44	0.68	0.67	0.65	0.64	0.113	0.415	0.007	0.015	0.444	0.669
22.50	1439	54.30	32000	2.41	0.68	0.67	0.65	0.64	0.127	0.400	D.014	0.015	0.904	0.770
23.00	1452	54.79	31000	2.38	0.68	0.66	0.65	9.64	0.141	0.385	0.015	0.015	1.000	0.743
23.50	1479	55-81	30000	2.37	0.68	0.67	0.66	0.64	0.145	0.370	0.004	0.015	0.241	0.769
24.00	1487	56.11	29610	2.34	0.67	0.66	0.65	0.64	0.163	0.356	0.018	0.015	1.247	0.002
24.50	1504	56.75	28740	2.32	0.67	0.66	0.65	0.64	0.174	0.341	0.011	0.015	0.753	0_800
25.00	1520	57.36	28000	2.29	0.67	0.66	0.65	0.64	0.186	0.327	0.012	0.014	0.796	0.726

* See Subsection 4.1.5.

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TABLE C-7^{*} (cont.)

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%			СТ										LNLN	DAT
Xing	BACC	% ACC	HAZ IND	PF	65 ، «۵	Q0,70	**.75	\$ * .80	log log	LOGLOG	DELLN	DELLE	DELTA	7-1 V R
25.50	1538	58.04	27000	2.28	0.67	0.66	0.65	0.64	0.196	0.312	0.010	0.014	0.675	0.816
26.00	1552	58.57	26250	2,25	0.67	0.66	0.65	0.68	0.208	0.298	D_013	0.014	0.885	0.736
26.50	1573	59.36	25500	2.24	0.67	0.66	0.65	0.64	0.215	0.284	0.007	0.018	0.087	0.863
27.00	1587	59.89	24800	2.22	0.67	0.66	0.65	0.64	0.227	0.270	0.012	0.014	0.865	0.836
27.50	1604	60.53	24000	2.20	0.67	0.66	0.65	0.64	0.237	0.255	0.010	0.014	0.688	0.900
28.00	1604	60.53	23800	2.16	0.66	0.65	0.64	0.64	0.260	0.241	0.023	0.010	1.644	0.892
28.50	1622	61,21	23000	2.15	0.66	0,65	0.64	0.64	0.269	0.227	0,009	0.016	0.608	0.907
29.00	1631	61.55	22400	2.12	0.66	0.65	0.64	0.63	0.284	0.213	0.016	0.018	1.121	0.920
29.50	1646	62,11	21800	2.11	0.65	0.65	0.64	0.63	0.295	0,199	0.011	0.018	0.763	0.961
30.00	1663	62,75	21168	2.09	0,65	0.65	0.64	0.64	0.304	0.186	0.009	0.014	0.636	0.681
30.50	1674	63.17	20800	2.07	0.65	0.65	0.64	0.63	0.317	0.172	0.014	0.014	0.981	0.941
31.00	1685	63.58	20070	2.05	0.65	0.64	0.64	0.63	0.331	0.158	0.013	0.019	0.978	0.908
31.50	1694	63.92	20000	2.03	0.64	0.64	0.64	0.63	0.346	0_184	0.015	0-014	1.098	0.855
32.00	1704	64.30	19500	2.01	0.64	0.64	0.63	0.63	0,360	0.131	0.010	0.010	1.022	0.445
32.50	1716	64.75	18986	1.99	0.54	0.64	0.63	0.63	0.372	0.117	0.012	0.014	0.893	0.900
33.00	1736	65.51	18400	1.99	0.64	0.64	0.63	0.63	0.377	0.103	0.005	0.014	0.391	0.875
33.50	1753	66.15	18000	1.97	0.64	0.69	0.64	0.63	0.385	0.089	0.008	0.014	0.567	0.814
34.00	1757	66.30	17790	1.95	0.64	0.63	0.63	0.63	0.400	0.076	0.019	0.014	1.363	0.789
34.50	1770	66.79	17200	1.94	0.63	0.63	0.63	0.63	0.415	0.062	0.011	0.014	0.799	0.754
35,00	1785	67.36	16800	1.92	0.63	0.63	0.63	0.63	0.424	0.049	0.009	0.014	0.665	0.826
35.50	1797	67.81	16316	1.91	0.63	0.63	0.63	0.63	0.435	0.035	0.011	0.014	0.845	1.821
36.00	1812	68.38	16000	1.90	0.63	0.63	0.63	0.63	0.444	0.021	0.009	0.014	0.648	1.724
36.50	1823	68.79	15600	1.88	0.63	0.63	0.63	0.63	0.456	0.008	0.012	0.014	0.893	1.740
37.00	1830	69.06	15200	1.67	0.63	0.63	0.63	0.63	0.472	0.005	0.016	0.002	7.533	1.773
37.50	1844	69.58	15000	1.86	0.63	0.63	0.63	0.63	0.481	0.019	0.009	0.014	0.687	1.719
38.00	1854	69.96	14484	1.84	0.62	0.62	0.63	0.63	0.494	0.033/	0.013	0.014	0.938	1.702
38.50	1865	70.38	14108	1.83	0.62	0.62	0.62	0.63	0.505	0.047	0.012	0.014	0.867	1.741
39.00	1882	71.02	13878	1.62	0.62	0.63	0.63	0.63	0.512	0.060	0.006	0.014	0.468	0.766
39.50	1898	71.62	13500	1.61	0.62	0.63	0.63	0.63	0.519	0.07%	0.007	0.014	0.525	0.796
40.00	1904	71.85	13120	1.BO	0.62	0.62	0.63	0.63	0.535	0.087	0.016	0.019	1.171	0.796
40.50	1917	72.34	12800	1.79	0.62	0.62	0.63	0.63	0.545	0.101	0.010	0.014	0.706	0.889
41.00	1927	72.72	12448	1.77	0.62	0.62	0.62	0.63	0.557	0.115	0.012	0.014	0.097	0.985
41.50	1965	74.15	12000	1.79	0.63	064	0.64	0.64	0.544	0.120	0.013	0.014	0.938	1.052
42.00	1965	74.15	12000	1.77	0.62	0.63	0.63	0.64	0.565	0.142) _021	0.014	1.521	1.008
42.50	1971	74.38	11712	1.75	0.62	0.62	0.63	0.63	0.580	0.156	0.016	0.014	1.134	1.068
43.00	1979	74.68	11400	1.74	0.62	0.62	0.63	0.63	0.594	0.170	0.014	0.010	0.999	1.004
43.50	1989	75.06	11000	1.73	0.61	0.62	0.63	0.63	0.606	0.183	0.012	0.014	0.862	0.932
44.00	1995	75.28	10740	1.71	0.61	0.62	0_62	0.63	0.622	0.197	0.016	0.014	1.125	0.758
44.50	2011	75.89	10450	1.71	0.61	0.62	0.63	0.63	0.620	0.211	0.006	0.014	0,446	0.800
45.00	2027	76.49	10164	1.70	0.61	0.62	0.63	0.64	0.63%	0.225	0.006	0.014	0.439	0.799
45.50	2045	77.17	10000	1.70	0.62	0.62	0.63	0.64	0.638	0.239	0.00%	0.014	0.299	0.709
46.00	2046	77.21	9852	1.68	0.61	0.62	0.63	0.63	0.658	0.253	0.020	0.014	1.428	0.694
46.50	2054	77.51	9600	1.67	0.61	0.62	0.62	0.63	0.671	0.267	0.013	D.010	0.958	0.660
47.00	2064	77.89	9360	1.66	0.61	0.61	0.62	0.63	0.683	0.281	0.011	0.014	0.817	0.747
47.50	2079	78.45	9120	1.65	0.61	0.62	0.63	0.64	0.690	0.295	0.007	0.014	0.870	0.627
48,00	2097	79.13	9000	1.65	0.61	0.62	0.63	0.64	0.693	0.309	0.004	0.015	0.,261	0.705
48,50	2104	79.40	8800	1.64	0.61	0.62	0.63	0.65	0.707	0.324	0.019	0.014	0.996	0.748
49.00	2113	79.74	8520	1.63	0.61	0.62	0.63	0.64	0.720	0.338	0.012	0.010	0.056	0.000
47.00	2126	80,23	8340	1-62	0.61	0.62	0.63	0.64	0.728	0.352	0.008	0.019	0.570	0.000
20.00	2129	80.44	8050	1.61	0.60	በፈነ	በፈን	D 60	6 786	6 367	A 610	0 0 9 0	4 368	0 868

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* See Subsection 4.1.5.

TABLE C-8. COLEMAN-STEWART VERSUS NEW HAMPSHIRE CROSSBUCKS

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	CO	LEMAN-S'	TEWART	•		N	Ем нами	PSHIRE.	СТ				_				
\$ Xine	INC CU	м	POWER	HAZARD	ENC.	CUM		POLED		HITH MATCH		e		20		4H	
	ACC #AC	C % ACC	FACTR	LNDEX	ACC	#ACC	3 ACC	FALTE	INDEX	WDIES TVAL		CUM	LESS	MATCH	LESS	MATCH	LESS MATCH
									INDEA	AUTER LANC	HICH	ATCH	#AC C	X ACC	#ACC	T ACC	XDIFF TVAL
0.50	91 91	4. 79	917	-55606	5 1			0 7 7	72000								
1_00	52 143	7 71	7 21		5 - E - T		4.13	7.17	1200.0	-9,1C -0,2	51	61	20	1.01	22	1.11	-0.17 -0.3
1.50	44 197	0 4 3	6 7 9	- 711 77		1.00		0.43	4 7 4 9 1	9.40 9.5	41	102	41	2.07	37	1.66	0.40 0.9
2.00	47 234	. 11 79	5 00	75/5/		1 5 9	9.51	5 5	24500	-0.10 -0.1	1:	134	40	2.41	51	2.57	-0.10 -0.2
2.50	50 204	14 11	6 73	70.00	20	245	12.13	617	27.11	-0.55 -0.5	(e ^)	171	56	° . ⊬?	47	3,39	-0.55 -1.0
3.00	47 334	14.31	D.17 E 20	- / 89		301	15-17	6.17	72410	-0.46 -0.7	4 ¹	520	(4	3.23	41	4.08	-0.85 -1.4
3 50	5.0 320	10.43	5.40	-61572	40	341	17.17	5.73	13599	-0.16 -^.4	10	259	67	2.38	٩'n	4.13	-9.16 -1.2
6.00	חזכ טי נוא 7ג	20 93	2+ +1	- 1494	34	175	18.00	5.40	16321	0.05 U.N	29	297	аĢ	4.40	чp	4.44	0,05 a.j
4 50	רוויף אר נאא רוב	27.82	2.2	-97191	41	416	20.47	<u>-</u> - 24	14800	-0.15 -0.1	35	322	5 L	4.59	94	4.74	-0.15 -0.2
5 00		27.33	4. 10	- 49415	45	+61	21.24	5.10	13221	-7+61 -3+6	35	357	∘€.	4.34	104	5.74	-0.01 -1.3
5 50	1/ 433	24.14	4.47	-91374	• •	404	24.00	4.SP	15001	-0_55 -0.4	л H	395	83	4.44	19	4.94	-0.55 -0.8
5.00	30 514	25.16	4.76	-93236	31	525	26.45	4.01	11200	-0.30 -0.3	37	41?	P 7	4.10	ЧÌ	4-69	-3.39 -3.4
6.10	41 550	24.23	4.70	-94972	42	567	28.53	4.16	10101	-0.35 -0.2	۶۴	469	5.7	4.64	90	4 99	-0.15 -0.5
5.50	29 589	29.69	4.51	-96432	24	591	29.70	4.58	2150	-0.10 -0.1	31	505	5.4	4.23	46	4 - 33	-1.10 -1.2
· • 00	32 6?1	31.30	4.47	- 97848	24	615	31.00	4.43	35.0.1	0.30 0.7	32	537	84	4 23	11	3.93	0.30 0.5
7.50	21 642	32.36	4.31	-95374	ני ר	638	32.14	4.29	ყეი ე	3.20 9.1	24	561	0.1	4.00	77	1 88	0.20 0.3
ศ. เกา	74 666	37.57	4.20-	100629	? 7	665	73.57	4.19	7484	3,15, 7,7	23	585	3.2	4.11	- 41	4.08	0.05 0.1
8.50	23 689	34.73	4.09-	102035	23	688	14.63	4.08	6900	0.05 0.0	1 P	607	¢ 7	6.46	.4	6 11	0.15 0.1
9.00	?? 711	35.94	1.98-	103475	22	713	35.70	3.09	640)	0.35 0.3	27	626	~ ,	2 30	0.4	4.11	0.05 0.1
9.50	20 73 1	16.84	3,98-	104734	26	746	47.10	1.90	6000	-0.25 -0.1	19	1.4.3	σá	A 4.5		4 10	2.22 0.1
10,00	21 752	37.90	3.79-	105566	و ت	150	34.26	1.44	560.)	-0.35 -1.2	7,7	5.55	17	6 16	24	4.07	- 1.29 - 0.4
10.50	22 774	39.11	3.72-	106539	31	792	39.52	7.79	5250	-(.8) -(.4	1.0	6.9.6	10	4. 67	104	4.74 5.77	
11.00	22 796	40.12	3.65-	107708	21	517	41.14	1.74	5000	-1.06 -0.5	20	713	, , , , ,	4.14	1.57	5.54	-0.81 -1.1
11.50	20 815	41.13	3.58-	106840	22	839	42.29	3.68	4901	-1.16 -0.6	20	175		4.10	1.04	7.74	-1. 16 -1.5
12.00	22 938	42.74	3.52-	109771	30	869	43.6.)	3.65	440.1	-1.56 -0.8	20	745	7 3	4.UO	1 14	3.24	-1.10 -1.7
12.50	25 863	43.50	3.48-	110853	72	891	44.91	1.59	4201	-1 41 -0 7	21	726	77	2.0	1.04	7,74	-1.56 -2.1
13.09	24 897	44.71	3.44-	111849	36	927	46.72	3.59	4777	-2 07 - 0 0	2	170	7,	2.68	1.0	5.29	~1.41 - 7.1
13.50	10 936	45.67	3.38-	112743	13	949	47.33	3.51	3800		30	- L C 	(1	2.20	111	5.54	-2.12 -3.0
14.00	8 914	46.07	3.29-	113368	17	257	48.24	7.45	3600	-2 17 -1 3		0.1	69		195	7.19	-1.(1 -2.6
14.50	21 934	47.08	3.25-	114378	B	965	48.64	1.15	3500		10	040	67	3.43	111	5.59	-2.17 -3.2
15.00	13 947	47.73	3.18-	115421	23	988	40.00	2 3 2	7760		10	P04 0.07	11	3.23	101	5.09	-1.56 -2.4
15.50	14 961	48.44	3.13-	116034	ĩś	1003	50 55	7.76	2040		20	294	6.7	3.18	104	5.24	-2.07 -3.2
16.00	13 974	49.29	1.27-	116045		1034	53 13	3 3 4	1000	2.02 1.7	1.3	3997	6.4	- 23	104	5.14	-2.12 -3.2
16.50	13 987	49.75	3.02-	117549	1	1.335	57 17	7.14	10/07	-3.02 -1.1	21	920	54	2.12	114	5.75	-1.62 -4.6
17.00	19 1005	50.66	2.98-	118183	10	1064	574L7 63 13	3.10	1000	-/.4/ -1.1	10	037	57	2.87	1.05	5.29	-2.42 -3.8
17.50	20 1025	51.66	2 95-	119049	21	1.7.74	54 10	2.12	2800	-2.47 -1.1	1.6	945	59	2.97	1.06	5.44	-2.47 -3.8
18.00	16 1041	52.47	2.91-	110534	15	1000	54.10	3.10	2000	-2.52 -1.1	14	964	61	3.07	111	5,59	-2.52 -3.9
18.50	19 1060	51.43	2.89-	120447	77	1111	54.94	2.02	2500	-2.47 -1.1		975	65	3.33	115	5,80	-2.47 -3.6
19-00	10 1020	53.01	7 86 -	120447	25		50-19	3.03	2400	-2.67 -1.1	25	10.20	60	3.02	113	5.70	-2.(7 -4.0
19-50	17 1097	56 70	2 91-	121014		1113	20.10	2.95	2400	-2,17 -0.9	9	1009	61	3.07	104	5.24	-2.17 -3.3
20.00	12 1099	55 30	2.01-	121910	1.2	1130	56.96	7.97	2243	-2.17 -0.9	21	1033	57	2.87	100	5,04	-2.17 -3.4
20.50	10 1110	1 56 35	2.11-	122622	17	1142	57.56	2.88	2130	-2.17 -0.9	10	10.40	59	2.97	102	5.14	-7.17 -3.4
21.00	13 110	57 01	2 . 1	123203	*2	1187	24-83	2.97	3000	-3,48 -1.4	34	1074	44	2.2	117	5.70	-3.48 -5.5
21.50	7 1151	57.01	2.11-	124047	g	1187	59.83	2.85	2000	-2.82 -1.2	7	1691	50	2.52	1 76	5.34	-2.82 -4.5
22.00	16 1153))(+)() 53 ///	2.01-	125210	Ű	1187	59.83	2.78	5000	-7.47 -1.0	4	1085	53	2.67	1 02	5.14	-2.47 -3.9
22.00	14 1172		7.04-	125319	8	1195	60.23	2.74	1900	-2 .17 -C.9	15	1100	52	2.62	95	4.79	-2.17 -3.5
23 00	17 1100	50.11	2.61-	126045	22	1217	61.34	2.73	1800	-2.57 -1.0	19	1119	47	2.37	96	4.94	-2.57 -4.2
23.00	10 100	27.05	2.54-	126892	11	1228	61.90	2.69	1750	-2.27 -0.9	13	1132	51	2.57	96	4.84	-2.27 -3.7
23.30	15 1201	00.53	2.58-	127554	14	1242	62.60	2.66	1620	-2.07 -0.8	12	1144	57	7.87	9.6	4:94	-2.07 -3.3
24.00	15 1216	61.29	2.55-	128421	18	1260	63.51	2.65	1600	-2.22 -0.9	20	1164	52	2.62	96	4.84	-2.22 -3.6
24.20	1228	61.90	2.53-	129052	10	1270	64.01	2.61	1520	-2.12 -0.8	13	1177	51	2.57	93	4.69	-2-12 -3-5
23.00	23 1251	63.05	2.52-	129599	25	1295	65.27	2.61	1500	-2.27 -0.9	29	1206	45	7.27	95	4.49	-2.22 -3.8

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TABLE C-8. (cont.)

		COLEMAN-S	TEWART	NF	W HAMPS	HTRE	ст					PI	n	NA	,	
	INC	CUM	POWER HAZARD	INC CUM		POWER	HAZARD	WITH	MATCH	INC	CUM	1555	MATCH	1 555	MATCH	LESS MATCH
% Xina	ACC	#ACC % ACC	FACTR INDEX	ACC #ACC	% ACC	FACTR	INDEX	%DIFF	TVAL	MTCH	MTCH	#ACC	174 2	1000	114 2	VDIEF TVAL
y													2	1000		
25.50	21	1272 64.11	2.51 -130258	1 1296	65.32	2.56	1460	-1.21	-0.5	16	1222	50	2.52	74	3.73	-1.21 -2.2
26.00	12	1284 64.72	2.49 -130984	15 131 <u>1</u>	66.08	2.54	1400	-1.36	-0.5	18	1240	44	2.22	71	3.58	-1.36 -2.5
26.50	13	1297 65.37	2.47 -131601	6 1317	66.38	2.50	1350	-1.01	-0.4	12	1252	45	2.27	65	3.28	-1 01 -1 9
27.00	12	1309 65.93	2.44 -132317	9 1326	66.83	2.48	1280	-0.86	-0.3	11	1263	46	2.32	63	3.18	-0.86 -1.6
27.50	11	1320 66.53	2.42 -132847	8 1334	67.24	2.45	1220	-0.71	-0.3	10	1273	47	2.37	61	3.07	-0 71 -1 3
28.00	7	1327 66.89	2.39 -133349	30 1364	68.75	2.46	1200	-1.86	-0.7	16	1289	38	1.92	75	3.78	-1 86 -1 5
28.50	13	1340 67.54	2.37 -134216	0 1364	68.75	2.41	1200	-1.21	-0.5	6	1295	45	2.27	69	3.48	-1.21 -2 2
29.00	17	1357 6B.40	2,36 -134843	13 1377	69.41	2.39	1120	-1.01	-0.4	20	1315	42	2.12	62	3.13	-1.01 -2.0
29.50	13	1370 69.05	2.34 -135464	13 1390	70.06	2.37	1080	-1.01	-0.4	11	1326	44	2.22	64	3.23	-1.01 -1.9
30.00	9	1379 69.51	2.32 -136307	11 1401	70.61	2.35	1020	-1.11	-0.4	7	1333	46	2.32	68	3.43	-1.11 - 2.1
30.50	5	1384 69.76	2.29 -136965	37 1438	72.48	2.38	1000	-2.72	-1.0	17	1350	34	1.71	88	4.44	-2.72 -4 9
31.00	15	1399 70.51	2.27 -137425	0 1438	72.48	2.34	1000	-1.97	-0.7	10	1360	39	1.97	78	3.93	-1.97 -3.6
31.50	12	1/11 71.12	2,25 -138010	0 1438	72.48	2.30	1000	~1.36	-0.5	9	1369	42	2.12	69	3.48	-1.36 -2.6
32.00	3	1414 71.27	2,23 -138162	19 1457	73.44	2.29	925	-2.17	-0.8	9	1378	36	1.81	79	3.98	-2.17 -4 0
32.50	8	1422 71.67	2,21 -138811	19 1476	74.40	2.29	900	-2.72	-1.0	11	1389	33	1.66	87	4.39	-2.72 -4.9
33.00	12	1434 72.23	2.19 -139459	1 1477	74.45	2.26	890	-2.17	-0.0	8	1397	37	1.86	80	4.03	-2.17 -4.0
33.50	9	1443 72.73	2.17 -139965	12 1489	75.05	2.24	840	-2.32	-0.9	13	1410	33	1.66	79	3.98	-2.32 -4 3
34.00	10	1453 73.24	2.15 -140674	27 1516	76.41	2.25	800	-3.13	-1.2	15	1425	28	1.41	91	4.59	-3.18 -5.8
34.50	33	1466 73.89	2.14 -141302	0 1516	76.41	2.21	800	-2.52	-0.9	- 9	1434	32	1.61	82	4.13	-2.52 -4.7
35.00	11	1477 74.45	2.13 -141897	3 1519	76.56	2.19	776	-2.12	-0.8	9	1443	34	1.71	76	3.83	-2 12 -4 0
35.50	10	1487 74.95	2.1] -142337	9 1528	77.02	2.17	750	-2.07	-0.7	12	1455	32	1.61	73	3.68	-2.07 -4.0
36.00	- 9	1496 75.40	2.09 -142820	17 1545	77.87	2.16	720	-2.47	-0.9	12	1467	29	1.46	78	3,93	-2 47 -4 7
36.50	9	1505 75.86	2.08 -143437	7 1552	78.23	2.14	700	-2.37	-0.9	<u> </u>	1476	29	1.46	76	3.93	-2.37 -4.6
37.00	16	1521 76.66	2.07 -144152	5 1557	78.48	2.12	660	-1.01	-0.6	14	1490	31	1.56	67	3.38	-1.81 -3.6
37.50	13	1534 77.32	2.06 -144877	7 1564	78.83	2.10	640	-1.51	-0.5	10	1500	34	1.71	64	3.23	-1.51 - 3.0
38.00	9	1543 77.77	2.05 -145267	37 1601	80.70	2.12	600	-2.92	-1.0	16	1516	27	1.36	85	4.28	-2.92 -5 5
38.50	12	1555 78.38	2.04 -145768	0 1601	80.70	2.10	600	-2.32	-0.8	11	1527	28	1.41	74	3.73	-2.32 -4.6
39.00	11	1566 78.93	2.02 -146482	0 1601	80.70	2.07	600	-1.76	-0.6	- 9	1536	30	1.51	65	3.28	-1.76 -3.6
39.50	7	1573 79.28	2.01 -146950	0 1601	80.70	2.04	594	-1.41	-0.5	7	1543	30	1.51	58	2.92	-1.41 - 3.0
40.00	14	1587 79.99	2.00 -147772	11 1612	81.25	2.03	550	-1.26	-0.4	9	1552	35	1.76	60	3.02	-1.26 -2.6
40.50	3	1590 80.14	1.98 -148351	6 1610	81.55	2.01	522	-1.41	-0.5	7	1559	31	1.56	59	2.97	-1.41 -3.0
41.00	6	1596 80.44	1.96 -149010	23 1641	82.71	2.02	500	-2.27	-0.3	7	1566	30	1.51	75	3.78	-2.27 -4.4
41.50	5	1601 80.70	1.94 -149280	0 1641	82.71	1.99	500	-2.02	-0.7	3	1569	32	1.61	72	3.63	-2.02 -3.9
42.00	2	1603 80.80	1.92 -149939	0 1641	82.71	1.97	500	-1.92	-0.7	ž	1571	32	1.61	70	3.53	-1.92 - 3.8
42.50	8	1611 81.20	1.91 -150520	0 1641	82.71	1.95	500	-1.51	-0.5	4	1575	36	1.81	66	3.33	-1.51 -3.0
43.00	6	1617 81.50	1.90 -151020	7 1648	83.06	1.93	480	-1.56	-0.5	9	1584	33	1.66	64	3.23	-1.56 -3.1
43.50	14	1631 82.21	1.89 -151577	4 1652	83.27	1.91	460	-1.06	-0.4	11	1595	36	1.01	57	2.97	-1.06 -2.2
44.00	5	1636 82.46	1.87 -152317	9 1661	83.72	1.90	450	-1.26	-0.4	5	1600	36	1.81	61	3.07	-1.26 -2.5
44.50	10	1646 82.96	1.86 -153034	7 1668	84.07	1.89	420	-1.11	-0.4	9	1609	37	1.86	59	2.97	-1.11 -2.2
45.00	- 9	1655 83.42	1.85 -153482	23 1691	85.23	1.89	400	-1.81	-0.6	17	1626	29	1.46	65	3.20	-1.81 - 3.7
45.50	11	1666 83.97	1.85 -153928	0 1691	85.23	1.87	400	-1.26	-0.4	11	1637	29	1.46	54	2.72	-1.26 -2.7
46.00	2	1668 84.07	1.83 -154170	0 1691	85.23	1.65	400	-1.16	-0.4	2	1639	29	1.46	52	2.62	-1.16 -2.6
46.50	4	1672 84.27	1.81 -154889	0 1691	85.23	1.83	400	-0.96	-0.3	4	1643	29	1.46	48	2.42	-0.96 -2.2
47.00	13	1685 84.93	1.81 -155599	1 1692	85.28	1.81	376	-0.35	-0.1	13	1656	29	1.46	36	1,81	-0.35 -0.9
47.50	2	1687 85.01	1.79 -156143	7 1699	85.64	1.80	360	-0.60	-0.2	4	1660	27	1.36	39	1.97	-0.60 -1.5
49.00	5	1692 85.28	1.78 -156751	0 1699	85.64	1.78	350	-0.35	-0.1	ŝ	1663	29	1,46	36	1.81	-0.35 -0.9
48.50	9	1700 85.69	1.77 -156918	5 1704	85.89	1.77	326	-0.20	-0.1	9	1672	28	1.41	32	1,61	-0.20 -0.5
49.00	4	1704 85.89	1.75 -157248	5 1709	96.14	1.76	320	-0.25	-0.1	6	1679	26	1,31	31	1.56	-0.25 -0.7
49.50	7	1711 86 24	1.74 -157359	14 1723	86.84	1.75	300	-0.60	-0.2	8	1686	25	1.26	37	1,86	-0.60 -1.5
50.00	3	1714 86.39	1.73 -157915	0 1723	86.04	1.74	300	-0.45	-0.2	3	1689	25	1.26	34	1.71	-0.45 -1.2

TABLE C-9. COLEMAN-STEWART VERSUS NEW HAMPSHIRE FLASHING LIGHTS

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% . Xing	INC Acc	CUM ∦ACC	% ACC	POWER Factr	HAZARD Index	INC Acc	CUM #ACC	% ACC	POWER Factr	HAZARD INDEX	WITH N %DIFF	MATCH TVAL	INC Mtch	CUM Mtch	LESS #ACC	MATCH %ACC	LESS #ACC	MATCH % ACC	LESS MATCH %DIFF TVAL
0.50	52	52	3.23	7.85	-21164	55	55	4.15	8.81	487800	-0.23	-0.3	40	40	12	0.91	15	1.13	-0.23 -0.6
1.00	38	90	6.80	6.80	-27452	34	89	6.72	6.72	337790	0.08	0.1	24	64	26	1.96	25	1,89	0.08 0.1
1.50	23	113	8.53	5.69	-31437	26	115	8.69	5.79	268380	-0.15	-0.1	20	84	29	2.19	31	2.34	-0.15 -0.3
2.00	23	136	10.27	5.14	-34861	29	144	10.83	5.44	234000	-0,60	-0.5	19	103	33	2.49	41	3.10	-0.60 -0.9
2.50	19	155	11.71	4.68	-37924	25	169	12.76	5.11	210000	-1.06	-0.8	18	121	34	2.57	48	3.63	-1.06 -1.5
3.00	21	176	13.29	4.43	-40577	29	198	14.95	4.98	183820	-1.65	-0.1	13	134	42	3.17	64	4.83	-1.66 -2.1
3.50	29	205	15.48	4.42	-43019	19	217	16.39	4.68	169400	-0.91	-0,6	15	149	56	4.23	68	5.14	-0.91 -1.1
4.00	29	225	16.99	4.25	-45179	14	231	17.45	4.36	153088	-0.45	-0.3	7	156	69	5.21	75	5.66	-0.45 -0.5
4.50	13	238	17.98	3.99	-46882	19	250	18.88	4.20	139000	-0.91	-0.5	15	171	67	5.06	79	5.97	-0.91 -1.0
5.00	11	249	18.81	3.76	-48363	23	273	20.62	4.12	128000	-1.81	-1.1	12	183	66	4.98	90	6.80	-1.81 -1.9
5.50	26	275	20.77	3.78	-50160	23	296	22.36	4.06	120800	-1.59	-0.9	18	201	74	5.59	95	7.18	-1.59 -1.6
6.00	24	299	22.58	3.76	-51/50	26	322	24.32	4.05	112720	-1.74	-0.9	22	Z23	76	5.74	99	7.48	-1./4 -1./
6.50	11	310	23.41	3.60	-53080	18	340	25.63	3.95	106981	-2.27	-1.2		232	78	5.89	108	0.10	-2.2/ -2.2
7.00	17	327	24.70	3.03	-54438	24	364	27.49	3.93	100000	-2.79	-1.4	19	251	76	5.74	113	0.00	-2.79 -2.7
7.50	14	34T	22.70	3.43		15	379	28.63	3.82	94560	-2.8/	-1.4	14	200	/0	5.74	120	0.01	-2.07 -2.0
8.00	20	308	27.04	3.30	500100	10	397	29.98	3.75	90000	-2.95	-1.4	12	277	0.7	6.14	1120	9.00 P 46	
9.00	29	307	29.29	3.39	-50210	15	407	30.74	2.62	83400	-1.51	-0.7	10	290	92	6 65	110	9 99	-2 34 -2 2
9.50	16	407	30 74	7 24	-60026	10	422	32.07	3 40	79990	-1 59	-0.7	12	305	90	6 95	113	A 53	-1 59 -1 5
10.00	16	427	31 95	3.19	-61028	12	420	32.33	3.40	75600	-1.39	-0.7	19	313	90	6 80	107	8 08	-1.39 - 1.3
10.50	20	441	33.46	3.19	-61864	14	454	34 29	3 27	72000	-0.83	-0.0	24	357	86	6.50	97	7.13	-0.83 -0.8
11.00	22	465	35.12	3.19	-62767	18	472	35.65	3.24	69225	-0.53	-0.2	22	379	86	6.50	93	7.02	-0.53 -0.5
11.50	15	480	36.25	3.15	-63470	20	492	37.16	3.21	66000	-0.91	-0.4	17	396	84	6.34	96	7.25	-0.91 -0.9
12.00	18	498	37.61	3.13	-64427	12	504	38.07	3.17	64000	-0.45	-0.2	15	411	87	6.57	93	7,02	-0.45 -0.4
12.50	19	517	39.05	3,12	-65235	11	515	38,90	3.11	61088	0.15	0.1	16	427	90	6.80	68	6.65	0.15 0.1
13.00	11	528	39.88	3.07	-65870	8	523	39.50	3.04	59800	0.38	0.2	8	435	93	7.02	88	6.65	0.38 0.4
13.50	13	541	40.86	3.03	-66646	8	531	40.11	2.97	56480	0.76	0.3	10	445	96	7.25	86	6.50	0.76 0.7
14.00	10	551	41.62	2.97	-67289	16	547	41.31	2.95	54000	0.30	0.1	11	456	95	7.18	91	6.87	0.30 0.3
14.50	13	564	42.60	2.94	-67997	15	562	42.45	2.93	51500	0.15	0.1	16	472	92	6,95	90	6.80	0.15 0.1
15.00	10	574	43.35	2.89	-68676	9	571	43.13	2.88	49940	0.23	0.1	10	482	92	6.95	89	6.72	0.23 0.2
15.50	12	586	44.26	2.86	-69330	12	583	44.03	2.84	48000	0.23	0.1	13	495	91	6.87	88	6.65	0.23 0.2
16.00	13	604	45.62	2.85	-69962	5	588	44.41	2.78	46554	1.21	0.5	13	50 8	96	7.25	80	6.04	1.21 1.2
16.50	13	617	46.60	2.82	-70808	22	610	46.07	2.79	45000	0.53	0.2	22	530	87	6.57	80	6.04	0.53 0.5
17.00	13	630	47.58	2.80	-71431	6	616	46.53	2.74	43260	1.06	0.4	11	541	89	6,72	75	5.66	1.06 1.1
17.50	6	636	48.04	2.74	-72025	.7	623	47.05	2.69	41990	0.98	0.4		550	86	6.50	13	5.51	0.98 1.0
18.00	14	64.3	48.55	2.70	-/269/	17	640	48.34	2.69	40390	0,23	0.1	14	564	79	5.97	76	5.79	0.23 0.2
10.00	19	657	49.02	2.00	-/32//	8	648	48.94	2.65	40000	0.68	0.2	14	578	/9	5.97	70	5.25	0.00 0.7
19.00	10	671	20.38	2.00	-74080	8	656	49.55	2.61	38800	0.03	6.0		585	82	0.19	71	5.44	0.03 0.9
20 00	5	676	51 06	2.00	-75417	13	669	50.53	2.59	3/500	0.15	0.1	12	597 60E	74	5.35	81	6 12	-0.76 -0.8
20.50	7	677	51 50	2.00	-75927	1/	000	51.81	2.59	36000	-0.76	-0.3		602	71	5 70	70	5 97	-0.68 -0.7
20.00	, o	691	52.39	2.92	-76477	8 0	592 700	52.27	2.55	35000	-0.68	-0.2	0	613	70	5 70	79	5.97	-0.68 -0.7
21 50	14	705	53 25	2.47	-770477	8	700	52.8/ 52 5F	2.52	34000	-0.08	-0.2	1 1 2	634	70	5 36	75	5.66	
22.00	13	718	54 21	2 46	-77730	9	715	51 00	2.49	33000	-0.30	-0.1	1.) 7	641	יי רר	5 82	74	5.59	0.23 0.2
22.50	12	730	55.14	2.45	-78196	0 7	722	54.00	2.43	2000	0.23	0.1	á	650	80	6 04	72	5.44	0.60 0 6
23.00	4	734	55.44	2.41	-78721	, 0	711	55 71	2.42	30240	0.00	0.2	ر ۲	656	79	5 89	75	5.66	0.23 0.2
23.50	5	739	55.82	2.38	-79326	л Д	735	55 51	2 36	29700	0.23	0.1	7	663	76	5.74	72	5.44	0.30 0.3
24.00	7	746	56.34	2.35	-79808	a a	744	56.19	2.34	28800	0.15	0.1	Ŕ	671	75	5.66	73	5.51	0.15 0.2
24.50	7	753	56.87	2.32	-80383	9	753	56.87	2.32	28000	0.00	0.0	7	678	25	5,66	75	5.66	0.00 0.0
25.00	11	764	57.70	2.31	-80954	ē	762	57.55	2.30	27000	0.15	0.1	12	690	74	5.59	72	5.44	0.15 0.2

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TABLE C-9. (cont.)

		COLEMA	N-STEWA	RT		NE	W HAM	PSHIRE	DT						PD	N	н		
% Xing	INC ACC	CUM #ACC % AC	POWER C Facto	HAZARD R INDEX	INC ACC	CUM ≇ACC	% ACC	POWER Factr	HAZARD INDEX	WITH M Sdiff	IATCH TVAL	INC Mtch	CUM Mtch	LESS ∦ACC	MATCH % ACC	LESS #ACC	MATCH % ACC	LESS %DIFF	MATCH TVAL
25.50	10	774 58.4	6 2.29	-81652	7	769 5	8.03	2,28	26270	0.38	0.1	8	698	76	5.74	71	5.36	0.38	0.4
26.00	9	783 59.1	4 2.27	-82346	13	782 5	9.06	2.27	25600	0.08	0.0	7	705	78	5.89	77	5.82	0.08	0.1
26.50	4	787 59.4	4 2.24	-82976	7	789 5	9.59	2.25	25000	-0.15	-0.1	8	713	74	5,59	76	5.74	-0.15	-0.2
27.00	9	796 60.1	2 2.23	-83525	6	795 6	0.05	2.22	24290	0.08	0.0	6	721	75	5.66	74	5.59	0.08	0,1
27.50	7	803 60.6	5 2.21	-84017	5	800 6	0.42	2.20	24000	0.23	0.1	8	729	74	5.59	71	5.36	0.23	0.2
28.00	11	814 61.4	8 2.20	-84541	5	805 6	0.80	2.17	23200	0.68	0.2	6	735	79	5.97	70	5.29	0.68	0.7
28.50	8	822 62.0	8 2.18	-85238	13	818 6	1.78	2.17	22500	0.30	0.1	11	746	76	5.74	72	5.44	0.30	0.3
29.00	5	827 62.4	6 2.15	-85771	7	825 6	2.31	2.15	22000	0.15	0.0	7	753	74	5.59	72	5.44	0.15	0.2
29.50	7	834 62.9	9 2.14	-86202	8	833 6	2.92	2.13	21440	0.08	0.0	9	762	72	5.44	71	5.36	0.08	0.1
30.00	5	839 63.3	7 2.11	-86722	4	837 6	3.22	2.11	20900	0.15	0.0	7	769	70	5.29	68	5.14	0.15	0.2
30,50	6	845 63.8	2 2.09	-87207	4	841 6	3.52	2.08	20400	0.30	0.1	5	774	71	5.36	67	5.06	0.30	0.3
31.00	9	854 64.5	0 2.08	-87692	5	846 6	3.90	2.06	20000	0.60	0.2	8	782	72	5.44	64	4.83	0.60	0.7
31.50	10	864 65.2	6 2.07	-88196	1	347 6	3.97	2.03	19680	1.28	0.4	3	785	79	5.97	62	4.68	1.28	1.4
32.00	5	869 65.6	3 2.05	-88771	н	855 6	4.58	2.02	19200	1.06	0.3	7	792	//	5.82	63	4.76	1.06	1.2
32.50	2	874 66.0	1 2.03	-89307	5	860 6	4.95	2.00	18792	1.06	0.3	5	797	11	5.82	63	4.76	1.06	1.2
33.00	3	8// 66.2	4 2.01	-89926	12	872 6	5.86	2.00	18040	0.38	0.1	/	804	/3	5.51	68	5.14	0.38	0.4
33.50	' '	884 66.7	1.99	-90433	0	3/8 6	5.JI	1.98	18000	0.45	0.1	9	813	71	5.30	60	4.91	0.45	0.5
34.00	2	890 67.2	2 1.98	-90941	8	885 6	5.92	1.97	1/325	0.30	0.1	0	819	/1 69	5.30	67	5.00	0.30	0.3
34.30	2	074 07.3	1 1.95 5 1 04	-7141/		092 0	/ . 3 / P 30	1.95	10346	0.00	0.0	4	023	69	5 14	74	5.21	0.00	0.0
35.00	5	007 69 1	3 1.54	-91010	11	903 0	D.20 D.50	1.90	16380	~0.45	-0.1	0 2	029	70	5 29	75	5.59	-0.45	-0.5
36.00	J	902 08.1	3 1.72	92340	4	011 6	0.00	1 01	15600	-0.38	-0.1	د ۲	222	73	5 51	74	5 50	-0.50	-0.4
36 50	6	916 69 1	8 1 90	-92570		911 6	9.01	1 80	15230	0.38	0.0	ר ר	840	76	5 74	71	5 36	0.08	-0.1
37.00	ě	925 69.8	6 1.89	-91979	Ř	919 6	9 41	1 88	15000	0.45	0.1	0 J	849	76	5.74	70	5 29	0.45	0.5
37.50	6	931 70.3	2 1 BA	-94519	2	922 6	9 64	1.86	14560	0 68	0.1	5	854	77	5.82	68	5 14	0 68	0.5
38.00	Ä	939 70.9	2 1.87	-95073	5	927 7	0.02	1.84	14220	0.91	0.3	6	860	79	5.97	67	5.06	0.91	1_0
38.50	7	946 71.4	5 1.86	-95782	Ř	935 7	0.62	1.83	14000	0.83	0.3	Ř	868	78	5.89	67	5.06	0.83	0.9
39.00	. 0	946 71.4	5 1.83	-96253	าา้	946 7	1.45	1.83	13600	0.00	0.0	6	874	72	5.44	72	5.44	0.00	0.0
39.50	9	955 72.1	3 1.83	-97065	7	953 7	1.98	1.82	13200	0.15	0.0	6	880	75	5.66	73	5.51	0.15	0.2
40.00	5	960 72.5	1 1.81	-97462	5	958 7	2.36	1.81	12800	0.15	0.0	8	888	72	5.44	70	5.29	0.15	0.2
40.50	15	975 73.6	4 1.82	-97963	6	964 7	2.81	1.80	12500	0.83	0.2	10	898	77	5.82	66	4.98	0.83	0.9
41.00	6	981 74.0	9 1.81	-98527	19	983 7	4.24	1.81	12000	-0.15	-0.0	15	913	68	5.14	70	5.29	-0.15	-0.2
41.50	4	985 74.4	0 1.79	-98875	0	983 7	4.24	1.79	12000	0.15	0.0	2	915	70	5.29	68	5.14	0.15	0.2
42.00	9	994 75.0	8 1.79	-99292	1	984 7	4.32	1.77	11792	0.76	0.2	5	920	74	5.59	64	4.83	0.76	0.9
42.50	5	999 75.4	5 1.78	-99807	4	988 7	4.62	1.76	11388	0.83	0.2	4	924	75	5.66	64	4.83	0.83	0.9
43.00	6	1005 75.9	1 1.77	-100340	1	989 7	4.70	1.74	11025	1.21	0.4	6	930	75	5.66	59	4.46	1.21	1.4
43.50	3	1008 76.1	3 1.75	-100771	4	993 7	5.00	1.72	10800	1.13	0.3	4	934	74	5.59	59	4.46	1.13	1.3
44.00	4	1012 76.4	4 1.74	-101320	8 .	1001 7	5.60	1.72	10500	0.83	0.2	6	940	72	5.44	61	4.61	0.83	1.0
44.50	8	1020 77.0	4 1.73	-101800	8 1	.009 7	5.21	1.71	10200	0.83	0.2	12	952	68	5.14	57	4.31	0.83	1.0
45.00	4	1024 77.3	4 1.72	-102352	10	019 7	5,96	1.71	10000	0.38	0.1	6	958	66	4.98	61	4.61	0.38	0.4
45.50	1	1025 77.4	2 1.70	-102695	1	1020 7	7.04	1.69	9840	0.38	0.1	0	958	67	5.06	62	4.68	0,38	0.4
46.00		1032 77.9	5 1.69	-103410	3.	1023 7	1.27	1.68	9600	0.68	0.2	6	964	68	5.14	59	4.46	0.68	0.8
40.50	7	1039 78.4	7 1.69	-103881	2	1025 7	/.42	1.66	9324	1.06	0.3	6	970	69 CC	5.Z1	55	4.15	1.06	1.3
47.00	2	1041 /8.6	3 1.67 0 1.67	-104575		1031 7	1.87	1.00	9180	0.76	0.2	5	975	60 67	4.98	56	4.23	0./6	0.9
47.50	2	1040 79.0	5 1 45 5 1 45	-102012	<u>د ۲</u>	1044 7	5.85 . 15	1.00	9000	0.15	0.0	8	963	to C	4.76	61	4.01	0.15	0.2
40.00	10	1040 79.1	⊃ 1,0⊃ 1 1,4⊑	-102236	4	1048 /	9.15 0.61	1.65	8800	0.00	0.0	2	785	64	4./0	63	4.70	0.00	0.0
40.00	10	1020 19.9	1 1 64 1 1 64	-100001	0.	1047	10.C	1.04	8500	0.30	0.1	9	374	04 41	2.03	6U 61	4.33	0.30	0.4
49.00	4	1066 80 5	1 1 604 1 1 60	106996	0. 0	1071 0) BO	1 67	8230	0.00	0.0		1012	54	4 02	59	4.01	-0.39	-0.5
50.00	1	1069 80 7	4 1 61	-107579	. و م	071 8	0.89	1 62	8000	-0.38	_0 0	11	1015	54	4 09	56	1.10	-0.15	-0.2
	-		- +-01		• •			A . V 4	0000	-0.10	- w . v	2	TO TO		4.00				

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TABLE C-10. COLEMAN-STEWART VERSUS NEW HAMPSHIRE AUTOMATIC GATES

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			C 0	LEMAN-S	TEWART			1	NEW HAMPS	HIRE	CT					PC		NH		
	% Xing	INC ACC	UU ∦AC	M C % AČC	POWER	HAZARD INDEX			4 Ρ • « ΑΓΓ Ε	OWER	HAZARD	WITH %DIF	MATCH F TVAL	INC Mtch	CUM Ntch	LESS #ACC	MATCH S ACC	HACC %	ATCH ACC	ZDIFF TVAL
					1		-	-						•••••	.,		-			
•	0,50	E	3	0.85	1.69	-5/163	4		1.98 3	.95	1272000	-1.13	-1.3	0	0	3	0.85	7	1.98	-1.13 -1.3
	1.00	6		2.34	2.34	- 20023	ź	14	3.95 3	.95	939420	-1.41	-1.0	3	Э	6	1.69	11	3,11	-1.41 -1.2
	1.50	2	14	5.90	2,04	-61713	5	17	2, 10, 3	.58	819300	~1.41	-0.9	2	5	9	2.54	14	3.95	-1.41 -1.0
	2.00	2	21	5.17	2.00	-63057	2	24	7 63 3	. 39	748000	-1.41	-0.8	4	.9	10	2.82	15	4.24	-1.41 -1.0
	2.00	1	22	6 21	2 07	-63638	9	27	7.02 3	.05	597400	-1.03	-0.9	2	11	10	2.82	16	4.52	-1.69 -1.2
	3.50	5	27	7 63	2.18	-64625	q	44	2.02 2	.30	54000	-3.6/	-1./	1	12	10	2.82	23	5.50	-3.6/ -2.3
	3.50	2	35	9.89	2.47	-65379	5	44	12.93 3	. 33	499400	-4.80	-2.0	4	10	11	3.11	28	7.91	-4.80 -2.7
	4 50	3	18	10.73	2.39	-65645	3	49	13 84 3	.25	458080	-3.11	-1.2	د د	19	10	4.52	27	7.03	-3.11 -1.7
	5.00	ี้เ	39	11.02	2.20	-65893	õ	49	13.84	.00	425000	-3.11	-1.1	2	21	19	5 00	20	7 01	-2.02 -1.5
	5.50	î	40	11.30	2.05	-66238	5	54	15.25. 2	- 77	400320	-3 95	-1 4	2	21	17	4 80	20	8 76	-3 95 -2 0
	6.00	5	45	12,71	2,12	-66745	11	65	18.36 3	.06	373800	-5 65	-1 9	5	28	17	4 80	37	10 45	-5 65 -2 7
	6,50	8	53	14.97	2,30	-67398	7	72	20.34 3	.13	350000	~5.37	-1.7	á	32	21	5.93	40	11.30	-5.37 -2.4
	7.00	2	55	15.54	2,22	-67802	5	77	21.75 3	.11	334400	-6.21	-1 9	4	36	19	5.37	41	11.58	-6.21 -2.8
	7.50	2	57	16.10	2,15	-68281	2	79	22.32 2	.98	321600	-6.21	-1.9	ż	38	19	5.37	41	11.58	-6.21 -2.9
	8.00	7	64	18.08	2,26	-68757	4	83	23.45 2	.93	308100	-5.37	-1.6	3	41	23	6.50	42	11.86	-5.37 -2.4
	8.50	3	67	18.93	2.23	-69215	6	39	25.14 2	.96	297600	-6.21	-1.8	7	48	19	5.37	41	11.58	-6.21 -2.8
	9.00	5	72	20.34	2.26	-69783	1	90	25.42 2	. 82	285200	-5.08	-1.4	5	53	19	5.37	37	10.45	-5.08 -2.4
	9.50	4	76	21.47	2,26	-70394	4	94	26.55 2	. 80	270100	-5.08	-1.4	4	57	19	5.37	37	10.45	-5,08 -2.4
	10.00	3	79	22.32	2,23	-70881	2	96	27.12 2	. 71	261000	-4.80	-1.3	2	59	20	5.65	37	10.45	-4.80 -2.3
	10.50	6	85	24.01	2.29	-71307	0	96	27.12 2	.50	254400	-3.11	-0.8	3	62	23	6,50	34	9.60	-3.11 -1.5
	11.00	5	90	25.42	2.31	-71684	0	96	27.12 2	. 47	246400	-1.69	-0.4	0	62	28	7.91	34	9.60	-1.69 -0.8
	11.50	3	93	26.27	2.28	-72049	4	100	28.25 2	.46	237000	-1.98	-0.5	2	64	29	8.19	36	10.17	-1.98 -0.9
	12.00	8	101	28.53	2,38	-72339	4	104	29.3R 2	.45	229600	-0.85	-0.2	8	72	29	8.19	32	9.04	-0,85 -0.4
	12.50	1	102	29.81	2.31	-72568		111	31.36 2	.51	220000	-2.54	-0.6	7	79	23	6.50	32	9.04	-2.54 -1.2
	13.00	0	102	23.01	2.22	- / 288 /	2	113	31.92 2	.46	210000	-3.11	-0.8	2	81	21	5.93	32	9.04	-3.11 -1.5
	13.50	5	10/	30.23	2.24	-73410	2	113	31.92 (2)	.36	202800	-1.69	-0.4	2	83	24	6.78	30	0.47	-1.69 -0.8
	14.00	4	111	31, 30	2.24	-74379	2	115	32.49.2.	. 32	197145	-1.13	-0.3	3	86	25	7.06	29	8.19	-1.13 -0.5
	14.50	0	111	37.30	2.10	-74376	2	122	33.05 2	. 28	190400	-1.69	-0.4	2	88	23	6.50	29	8.19	-1.69 -0.9
	15.00	1	122	34 46	2 22	-75630	7	122	35.30 /	. 30	183720	-2.26	-0.5	3	91	23	6.50	11	8.70	-2.26 -1.1
	16 00	2	125	35 31	2 21	-75958	2	127	35 88 2	20	170000	-0.35	-0.2	1	94	28	7.91	16	0.70	-0.85 -0.4
	16.50	, 1	122	36 16	2.19	-76207	2	120	36 44 2	. 24	160000	-0.56	-0.1	1	97	28	7.91	0L 0C	5.4/	
	17.00	4	132	37.29	2.19	-76668	3	132	37.29.2	19	161500	-0.20	-0.1	3	100	28	0.70	27	D.19 D.76	0.00 0.0
	17.50	2	134	37.85	2.16	-77111	4	136	38,42 2	20	155800	-0.56	-0.0	2	101	20	9.70	32	9 04	-0.56 -0.3
	18.00	3	137	38.70	2.15	-77389	3	139	39.27 2	.18	150000	-0.56	-0.1	2	107	30	0 47 0 47	12	9 04	-0.56 -0.3
	13.50	6	143	40.40	2.18	-77639	2	141	39.83 2	.15	147400	0.56	0.0	4	111	32	9.04	30	8.47	0.56 0.3
	19.00	ī	144	40.68	2.14	-77987	0	141	39.83 2	.10	144000	0.85	0.2	0	111	13	9 32	30	8.47	0.85 0.4
	19.50	5	149	42.09	2.16	-78414	6	147	41.53 2	.13	139950	0.56	0.1	5	116	33	9.32	31	8.76	0.56 0.3
	20.00	2	151	42.66	2.13	-78724	4	151	42.66 2.	.13	135000	0.00	0.0	. 3	119	32	9.04	32	9.04	0.00 0.0
	20.50	2	153	43,22	2.11	-79109	1	152	42,94 2.	.09	131810	0.23	0.1	1	120	33	9.32	32	9.04	0.28 0.1
	21.00	1	154	43.50	2.07	-79402	2	154	43.50 2.	.07	127500	0.00	0.0	1	121	33	9.32	33	9.32	0.00 0.0
	21.50	17	161	45.48	2.12	-79720	4	158	44.63 2.	.08	123760	0.85	0.2	5	126	35	9.89	32	9.04	0.85 0.4
	22.00	2	163	46.05	2.09	-80215	1	159	44.92 2.	.04	120000	1.13	0.2	2	128	35	9.89	31	8.76	1.13 0.5
	22.50	0	163	46.05	2.05	-80704	1	160	45.20 2.	.01	118500	0.85	0.2	1	129	34	9.60	31	8.76	0.85 0.4
	23.00	3	166	46.89	2.04	-81071	5	165	46.61 2.	.03	114504	0.28	0.1	5	134	32	9.06	31	8.76	0.20 0.1
	23.50	0	166	46.89	2.00	-81216	2	167	47.18 2.	.01	111650	-0.28	-0.1	0	134	32	9.04	33	9.32	-0.28 -0.1
	24.00	2	168	47.46	T.98	-81/28	2	169	47.74 1.	. 99	103120	-0.28	-0.1	4	138	30	8.47	31	8.76	-0.28 - 0.1
	24.50	3	171	48.31	1.9/	-82133	د	172	48.59 1.	.98	105000	-0.28	-0.1	Э	141	30	8.47	31	8.76	-0.28 - 0.1
	25,00	1	1/2	48.59	1.94	-01208	Ð	112	48.59 1,	.94	102400	0.00	0.0	0	141	31	8.76	31	8.76	U.OD U.O

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TABLE C-10. (cont.)

		COLEMAN-STE	WART			NEW HAM	PSHIRE	СТ					PD	1	N	1		
	INC	CUM	POWER	HAZARD	INC	CUM POWER	HAZARI) HAZARD	WITH	MATCH	INC	CUM	LESS	MATCH	LESS	MATCH	LESS	MATCH
% Xing	ACC	#ACC % ACC	FACTR	INDEX	ACC	\$ACC FACTR	INDEX	INDEX	%DIFF	TVAL	MTCH	MTCH	#ACC	% ACC	#ACC	% ACC	%DIFF	TVAL
						• • • • • • • • • • • • • • • • • • • •												
. 25.50	3	175 49.44	1.94	-22747	3	175 49.44	1.94	100000	0.00	0.0	2	143	32	9.04	32	9.04	0.00	0.0
26.00	3	178 50.28	1.93	-83102	4	179 50.56	1.94	98064	-0,28	0.1	Э	146	32	9.04	33	9.32	-0.28	-0.1
26.50	5	183 51.69	1.95	-83566	0	179 50.55	1.91	96000	1.13	0.2	4	150	33	9.32	29	8.19	1.13	0.5
27.00	3	186 52.54	1.95	-84045	0	179 50.56	1.87	93500	1.98	0.4	1	151	35	9.89	28	7.91	1.98	0.9
27.50	1	187 52.82	1.82	-84336	6	185 52.26	1.90	90000	0.56	0.1	6	157	30	8.4/	28	/.91	0.56	0.3
28.00	4	191 53.95	1.93	-84675	3	188 53.11	1.90	87696	0.85	0.2	5	162	29	8.19	26	7.34	0.85	0.4
28.50	1	192 54.24	1.90	-82087	1	189 53.39	1.87	86350	0.85	0.2	1	163	29	8.19	26	1.34	0.85	0.4
29.00	2	194 54.80	1.89	-85541	2	194 54.80	1.89	84000	0.00	0.0	4	167	27	1.03	27	7.63	0.00	0.0
29.50	2	190 22.37	1.00	-05982	0	194 34.60	1.00	82080	0.50	0.1	1	100	20	1.71	20	7.34	0.00	0.5
30.00	5	202 57.00	1.90	-00372		190 55.98	1.00	70200	1 40	0.2	1	172	32	9.70	27	7 74	1 40	0.5
30.50	2	204 53.65	1.09	-86844	2	200 56 50	1 97	79200	1.05	0.3	2	174	31	8 76	26	7 34	1 41	0.0
31.00	1	205 57 51	1 85	-87155	ñ	200 56 50	1 79	75400	1 69	0.2	1	175	21	8 76	25	7 06	1 69	0.7
31.50	5	211 59 60	1 86	-87488	A	204 57.63	1 80	72100	1 98	0.3	ŝ	180	31	8.76	24	6 78	1.98	0.9
32.00	,	212 59.89	1.84	-87944	7	211 59.60	1.83	71000	0.28	0.0	4	184	28	7.91	27	7.63	0.28	0.1
33.00	1	215 60.73	1.84	-88312	2	213 60.17	1.82	70000	0.56	0.1	ō	184	31	8.76	29	8.19	0.56	0.3
33.50	ō	215 60.73	1.91	-88734	2	215 60.73	1.81	68000	0.00	0.0	2.	186	29	8.19	29	8,19	0.00	0,0
34.00	2	217 61.30	1.80	-89145	0	215 60,73	1.79	66000	0.56	0.1	1	187	30	8.47	28	7.91	0.56	0.3
34.50	2	219 61.86	1.79	-89488	1	216 61.02	1.77	64880	0.85	0.1	1	188	31	8.76	28	7,91	0.85	0.4
35.00	1	220 62.15	1.78	-89723	2	218 61.58	1.76	63250	0.56	0.1	1	189	31	8.76	29	8.19	0.56	0.3
35.50	2	222 62.71	1.77	-90071	3	221 62.43	1.76	62020	0.28	0.0	4	193	29	8.19	28	7.91	0.28	0.1
36.00	2	224 63.28	1.76	-90341	1	222 62.71	1.74	60600	0.56	0.1	1	194	30	8.47	28	7.91	0.56	0.3
36.50	2	226 63.84	1.75	-90791	0	222 62.71	1.72	60000	1.13	0.2	1	195	31	8.76	27	7.63	1.13	0.5
37.00	1	227 64.12	1.73	-91142	1	223 62.99	1.70	58104	1.13	0.2	2	197	30	8.47	26	7.34	1.13	0.5
37.50	5	232 65.54	1.75	-91402	6	229 64.69	1.73	56240	0.85	0.1	0	197	35	9.89	32	9.04	0.85	0.4
38.00	1	233 65.82	1.73	-91582	3	232 65.54	1.72	55000	0.28	0.0	2	199	34	9.60	33	9.32	0.28	0.1
38.50	4	237 66.95	1.74	-92007	4	236 66.67	1.73	54000	0.28	0.0	6	205	32	9.04	31	8.76	0,28	0.1
39.00	2	239 67.51	1.73	~92187	0	236 66.67	1.71	52500	0.85	0.1	2	207	32	9.04	29	8.19	0.85	0.4
39.50	1	240 67.80	1.72	-92402	U	236 66.67	1.69	51320	1.13	0.2	0	207	10	9.32	29	8.19	1.13	0.5
40.00	0	240 67.80	1.69	-92783	2	238 67.23	1.68	50040	0.56	0.1	2	209	71	9.76	29	7 01	0,50	0.5
40.50	2	242 00,30	1.09	-93046	1	239 07.51	1.07	49000	0.00	0.1	2	211	31	8.70	20	7 91	0.03	0.4
41.00	1	243 08.04	1.07	-93530	1	240 07.00	1.05	48000	1 1 2	0.1	1	212	32	9 04	28	7.91	1,11	0.5
41.50	2	245 69 21	1 65	-93834	ō	241 68 08	1 62	46000	1 13	0.2	Ď	213	32	9.04	28	7.91	1.13	0.5
42.00	0	245 69 21	1 63	-94143	- ٦	244 68 93	1 62	45000	0 28	0.2	3	216	29	A.19	28	7.91	0.28	0.1
43.00	2	247 69.77	1.62	-94518	ő	244 68.93	1.60	44000	0.85	0.1	2	218	29	8.19	26	7.34	0.85	0.4
43.50	1	248 70.06	1.61	-94752	1	245 69.21	1.59	42650	0.85	0.1	ī	219	29	8,19	26	7.34	0.85	0.4
44.00	ī	249 70.34	1.60	-94983	2	247 69.77	1.59	41280	0.56	0.1	3	222	27	7.63	25	7.06	0.56	0.3
44.50	1	250 70.62	1.59	-94261	2	249 70.34	1.58	40000	0.28	0.0	1	223	27	7.63	26	7,34	0.20	0.1
45.00	1	251 70.90	1.58	-95404	0	249 70.34	1.56	40000	0.56	0.1	1	224	27	7.63	25	7.06	0.56	0.3
45.50	5	256 72.32	1.59	-95616	2	251 70.90	1.56	39060	1.41	0.2	6	230	26	7.34	21	5.93	1.41	0.7
46.00	3	259 73.16	1.59	-95923	4	255 73.03	1.57	37800	1.13	0.2	5	235	24	6.73	20	5.65	1.13	0.6
46.50	9	268 75.71	1.63	-96230	1	256 72.32	1.56	37200	3.39	0.5	9	244	24	6.78	12	3.39	3.39	2.0
47.00	1	269 75.99	1.62	-96429	4	260 73.45	1,56	36160	2.54	0.4	3	247	22	6.21	13	3.67	2.54	1.5
47.50	0	269 75.99	1.60	-96629	1	261 73.73	1.55	36000	2.26	0.3	1	248	21	5,93	13	3.67	2.26	1.4
48.00	1	270 76.27	1.59	-96887	1	262 74.01	1.54	34450	2.26	0.3	2	250	20	5.65	12	3.39	2.26	1.4
48.50	2	272 76.84	1.58	-97197	2	264 74.58	1.54	33600	2.26	0.3	3	253	19	5.37	11	3.11	2.26	1.5
49.00	4	276 77.97	1.59	-97514	Ű	264 74.58	1.52	32500	3.39	0.5	1	254	22	6.21	10	2.82	2.19	2.1
49,50	0	210 11.97	1,28	~ 9//13	U 1	204 /4.58	1.51	32000	3.39	0.5	U	254	22	0.21	10	2.02	3.37	2.1
50.00	2	4/0 /0.0J	1.3/	- 3 / 30 /	1	200 /4.86	T.20	10800	3.0/	υ.ь	U	204	24	υ,/Ο	11	7.11	5.07	2.2

TABLE C-11. PEABODY-DIMMICK VERSUS NEW HAMPSHIRE CROSSBUCKS

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		PE4	4800Y-		K		N	EW HAM	PSHIRE	СТ	-				n n		a ka	
% Xing	ENC	CUM		POWER	HAZARD	INC	CUM		PONER	HAZARD	NITH MATCH	LNC	CUM	1 6 5 5	MATCH	1 555	MATCH	
8	ACC	MACC	% ACC	FACTR	INDEX	ACC	#ACC	% ACC	FACTR	INDEX	XDIFF TVAL	ATCH	MICH	MACC	X ACC	HACC	X ACC	TOTEE TVAL
														0.00		-400	4 400	ADT T TAL
0.50	77	11	3.88	7.76	63891	93	83	4.18	8.37	72000	-0.30 -0.5	76	74	3	0-15	9	0.45	-0.30 -1.7
1.00	60	137	6.91	6.91	59231	52	135	6.80	6.80	47300	0.10 0.1	54	128	9	0.45	7	0.35	0.10 0.5
1.50	49	186	9.38	6.25	5676L	54	189	9.53	6.35	34500	-0.15 -0.2	50	178	8	0.40	- 11	0.55	-0.15 -0.7
2.00	51	237	11.95	5.97	54502	56	245	12.35	6.17	27200	-0.40 -0.4	53	231	6	0.30	14	0.71	-0.40 -1.8
2.50	50	287	14.47	5.79	52199	56	301	15.17	6.07	22800	-0.71 -C.6	48	279	8	0.40	22	1.11	-0.71 -2.6
3.00	38	325	16.38	5.46	51576	40	341	17.19	5.73	19200	-0.81 -0.6	42	321	4	0.20	20	1.01	-0.81 -3.3
3.50	41	366	18.45	5.27	50290	34	375	18.90	5.40	16800	-0.45 -(.3	39	360	6	0.30	15	0.76	-0.45 -2.0
4.00	32	398	20.06	5.02	49278	41	416	20.97	5.24	14800	-0.91 -0.6	30	390	8	0.40	26	1.31	-0.91 - 3.1
4.50	45	443	22.33	4 - 96	48283	45	461	23.24	5.16	13000	-0.91 -0.6	48	438	5	0.25	23	1.16	-0.91 -3.4
5.00	40	483	24.34	4.87	47467	33	494	24.90	4.98	12000	-0.55 -0.4	36	474	9	0,45	20	1.01	-0.55 -2.0
5.50	36	519	26.16	4.76	46836	31	525	26.46	4.81	11000	-0.30 -C.2	30	504	15	0.76	21	1.06	-030 -1.0
6.00	16	535	26.97	4.49	46262	42	567	28.58	4.76	10000	-1.61 -1.0	27	531	4	0.20	36	1.81	-1.61 -5.1
6.50	30	565	28.48	4.38	45597	24	591	29.79	4,58	9150	-1.31 -0.8	28	559	6	0.30	32	1.61	-1,31,-4,2
1.00	30	595	29.99	4.28	44436	24	615	31.00	4.43	8500	-1.01 -0.6	27	586	9	0.45	29	1.46	-1.01 -3.2
7.50	38	633	31.91	4.25	44296	23	638	32.16	4.29	8000	-0.25 -0.1	29	615	18	0.91	23	1.16	-0.25 -0.8
8.00	33	666	33.57	4.20	43716	27	665	33.52	4.19	7484	0.05 0.0	31	646	20	1.01	19	0.96	0.05 0.2
8.50	22	688	34.68	4.08	43238	23	688	34.68	4.08	6900	0.00 0.0	27	673	15	0.76	îs	0.76	0.00 0.0
9.00	21	709	35.74	3.97	42726	22	710	35.79	3.98	6400	-0.05 -0.0	22	695	14	0.71	15	0.76	-0.05 -0.2
9.50	21	730	36.79	3.87	42283	26	736	37.10	3.90	6000	-0.30 -0.2	20	715	15	0.76	21	1.06	-0.30 -1.0
10.00	26	756	38.10	3.81	41812	23	759	38.26	3.83	5600	-0.15 +0.I	27	742	14	0.71	17	0.86	-0.15 -0.5
10.50	19	775	39.06	3.72	41425	31	790	39.82	3.79	5250	-0.76 -0.4	20	762	13	0.66	28	1.41	-0.76 -7.3
11.00	23	798	40.22	3.66	41006	27	817	41.18	3.74	5000	-0.96 -0.5	Z 4	786	12	0.60	31	1.56	+0.96 -2.9
11.50	35	833	41.99	3.65	40611	22	839	42.29	3.68	4800	-0.30 -0.1	32	818	15	0.76	21	1.06	-0.30 -1.0
12.00	17	850	42.84	3.57	40259	30	869	43.80	3.65	4400	-0.96 -0.5	18	836	14	0.71	33	1.66	-0.96 -2.8
12.50	35	88 5	44.61	3.57	39894	22	891	44.91	3.59	4200	-0.30 -0.1	27	863	22	1.11	28	1.41	-0.30 -0.8
13.00	11	896	45.16	3.47	39515	36	927	46.72	3.59	4000	-1.56 -0.7	26	889	7	0.35	38	1.92	-1.56 -6.6
13.50	31	92 7	46.72	3.46	39205	13	940	47.38	3.51	3600	-0.66 -0.3	28	917	10	0,50	23	1.16	-0.66 -2.3
14.00	18	945	47.63	3.40	38857	17	957	48.24	3.45	3600	-0.60 -0.3	18	935	10	0.50	22	1.11	-0.60 -2.1
14.50	24	969	48.84	3.37	38493	8	965	48.64	3.75	3500	0.20 0.1	15	950	19	0.96	15	0.76	0.20 0.7
15.00	9	978	49.29	3.29	38251	23	988	49.80	3.32	3240	-0.50 -0.2	20	970	8	0.40	18	0.91	-0.50 -2.0
15.50	16	994	50.10	3.23	37990	15	1003	50.55	3.26	3070	-0.45 -(.2	13	983	11	0.55	20	1.01	-0.45 -1.6
16.00	20	1014	51.11	3.19	37657	31	1034	52.12	3.26	3000	-1.01 -0.4	23	1006	8	0.40	28	1.41	-1.01 -3.3
16.50	24	1038	52.32	3.17	37319	1	1035	52.17	3.16	2948	0.15 0.1	12	1018	20	1.01	17	0.86	0.15 0.5
17.00	15	1053	53.07	3.12	37093	19	1054	53.13	3.13	2800	-0.05 -0.0	21	1039	14	0.71	15	0.76	-0.05 -0.2
17.50	14	1067	53.78	3.07	36848	21	1075	54.18	3.10	2600	-0.40 -0.2	17	1056	11	0.55	19	0.96	-0.40 -1.5
18.00	19	1086	54.74	3.04	36561	15	1090	54.94	3.05	2500	-0.20 -0.1	18	1074	12	0.60	16	0.61	-0.20 -0.8
18.50	15	1101	55.49	3.00	36282	23	1113	56.10	3.03	2400	-0.60 -0.3	17	1091	10	0.50	22	1.11	-0.60 -2.1
19.00	13	1114	56.15	Z • 96	36097	0	1113	56.10	2.95	2400	0.05 0.0	10	1101	13	0.66	12	0.60	0.05 0.2
19.50	15	1129	56.91	2.92	35929	17	1130	56.96	2.92	2240	-0.05 -0.0	12	1113	16	0.81	17	0.86	-0.05 -0.2
20.00	4	1133	57.11	2.86	35676	12	1142	57.56	2.88	2130	-0.45 -6.2	10	1123	10	0.50	19	0.96	-0.45 -1.7
20.50	12	1145	57.71	7.82	35459	45	1187	59.83	2.92	2000	-2.12 -0.9	20	1143	2	0.10	44	2.22	-2.12 -6.2
21.00	11	1156	58.27	2.17	35229	0	1187	59.83	2.85	2000	-1.56 -0.6	9	1152	4	0.20	35	1.76	-1.56 -5.0
21.50	13	1169	58.92	2. 74	34995	0	1187	59.83	2.78	2000	-0.91 -0.4	9	1161	8	0.40	26	1.31	-0.91 -3.1
22.00	14	1183	59.63	2.71	3 48 30	8	1195	60.23	2.74	1900	-0.60 -0.2	13	1174	9	0.45	21	1.06	-0.60 -2.2
22.50	23	1206	60.79	2.70	34563	22	1217	61.34	2.73	1800	-0.55 -C.2	21	1195	11	0,55	22	1.11	-0.55 -1.9
23.00	8	1214	61.19	2 . 66	34299	11	1228	61.90	2.69	1750	-0.71 -C.3	9	1204	10	0.50	24	1.21	-0.71 -2.4
23.50	10	1224	61.69	2.63	34184	14	1242	62.60	2 • 6 6	1620	-0.91 -C.4	13	1217	7	0.35	25	1.26	-0.91 -3.2
Z4.00	22	1246	62.80	Z.62	33906	18	1260	63.51	2.65	1600	-0.71 -0.3	22	1239	7	0.35	21	1.06	-0.71 -2.6
24.50	15	1261	63.56	2.59	33693	10	1270	64.01	2.61	1520	-0.45 -(.2	14	1253	8	0.40	17	0.86	-0.45 -1.8
25.00	18	1279	54.47	2.58	33508	25	1295	65.27	2-61	1500	-0.81 -0.3	22	1275	4	0.20	20	1.01	-0.81 -3.3

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TABLE C-11. (cont.)

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		PE	480DY-1	D IMM ICH	(NE	W HAM	PSHIRE	СТ					PD	,	чн	
\$ Ying	INC	CUM		POWER	HAZARD	INC	ÇUM		POWER	HAZARD	WITH M	АЛСН	INC C	UM LES	S MATCH	LESS	MATCH	LESS MATCH
A VIUR	ACC	NACC	S ACC	FACTR	I NDE X	ACC	#ACC	X ACC	FACTR	INDEX	XDIFF	TVAL	MTCH MT	CH #AC	C & ACC	#ACC	T ACC	XDIFE TVAL
						-				· ·								
25.50	9	1288	64.92	2.55	33373	1	1296	65.32	2.56	1460	-0.40	-0.2	1 12	76 1	2 0.60	20	1.01	-0.40 - 1.4
26.00	15	1 30 3	65.68	2.53	33095	15	1311	66.08	2.54	1400	-0.40	-0.2	17 12	93 1	0.50	18	0.91	-0.40 -1.5
26.50	12	1315	66.28	2.50	32887	6	1317	66.38	2.50	1350	-0.10	-0.0	11 13	04 1	0.55	13	0.66	-0.10 -0.4
27.00	8	1323	66.68	2.47	32628	9	1326	66.83	2.48	1280	-0.15	-0.1	5 13	09 14	0.71	17	0-86	-0-15 -0-5
27.50	13	1 336	67.34	2.45	32510	8	1334	67.24	2.45	1220	0.10	0.0	7 1 3	16 21	1.01	1.8	0 91	
28.00	16	1352	68.15	2.43	32359	30	1364	68.75	2.46	1200	-0.60	-0.2	21 13	37 1	5 0.76	27	1. 36	-0.60 - 1.9
28,50	12	1 364	68.75	2.41	32135	0	1364	68.75	2.41	1200	0.00	0.0	7 1 2	44 21	1.01	20	1.0	
29.00	19	1 18 3	69.71	2.40	31935	13	1377	69.41	2.39	1120	0.30	0.1	16 17	58 2	5 1.26	19	0.96	0.30 0.9
29.50	15	1 398	70.46	2.39	31838	13	1390	70.06	2.37	1080	0.40	6.2	20 13	78 21		12	0.60	
30,00	4	1402	70.67	2.36	31544	11	1401	70.61	2.35	1020	0.05	0.0	7 13	85 1	0.86	16	0.81	0.05 0.2
30.50	13	1415	71.32	2.34	31518	37	1438	72.68	2.38	1000	-1.16	-0.4	25 14	10	5 0 25	28	1 41	
31.00	9	1424	71.77	2.32	31384	0	1438	72.48	2.34	1000	-0.71	-0.3	6 14	16	9 0.40	22	1.11	-0.71 -2.6
31,50	15	1439	72.53	2.30	31154	ō	1478	72.68	2.30	1000	0.05	0.0	4 14	20 19	0.96	18	0.91	
32.00	13	1452	73.19	2.29	30974	19	1457	73.44	2.29	925	-0.25	-0-1		38 1/	0.71	10	0.94	-0.25 -0.9
32,50	10	146.7	73.69	2.27	30747	19	1476	74.40	2.29	900	-0.71	-0.1	13 14	51 1	0.55	25	1 24	
33.00	7	1469	74.06	2.24	30653	- Îí	1477	74 45	2.26	690	-0.40	- 6 - 1	8 14	59 1		18		
33.50	18	1487	76.95	2.24	30433	12	1489	75.05	2.26	840	-0.10	-0 0	21 14	99 19	7 0 35	10	0.45	
36.00	12	1 4 9 9	75.55	2 22	30366	27	1516	76 41	2 2 2 2 2	800	- 48 0-		16 14	00		20	1 01	
34.50	7	1506	75.91	2.20	30178	21	1516	76 61	7 71	800	-0.60	-0.3	2 14	00		17		
35.00	5	1511	76.16	2.18	30057	3	1510	76 56	2 10	776	-0.50	-0.2		199 01 1/			0,00	
35.50	16	1525	76.86	2.17	29906	ر ہ	1528	77 02	2.17	750	-0.15	-01	17 16		7 0.50	10	0.41	
36-00	17	1534	77 17	2 15	29900	17	1545	77 97	2 1 4	730	-0.55	-0.1	7 1 2	26 1		10	0.50	~0.15 ~0.7
36.50	17	1547	77.97	2 14	29000	· · · · · · · · · · · · · · · · · · ·	1552	70 22	2 1 6	720	-0.35	-0.1	10.15	26 1	7 0.47	20	1.01	
37.00	16	1563	78.78	2 13	29279	5	1557	78 49	2 1 7	44.0	0.20	0.1	0 15	66 1 L		17	0.44	
37 50	20	1 545	70 00	2 10	29219	<i>,</i>	1666	70 03	2 10	660	0.30	0.1	9 1 3	1949 I.	7 U.70	13	0.00	
	10	1675	70.00	2 00	29213		1204	90.70	2 1 2	640	-1 31	-0.6	0 13	72 1		12	0.60	0.05 0.2
38 50	10	1586	70 04	2 0 7	2 9001		1401	80.10	2.10	600	-0.74	-0-2	22 15	14		27	1.30	-1.31 -4.9
30.00	11	1 500	00 24	2.00	20070		14.01	90.70	2 07	600	-0.76	- 6. 3	0 1 2	80 0	5 0.30	21	1.00	-0.70 -2.9
39 50		1400	90.24	2.00	20071	ŭ	1601	80.10	2.01	600 604	-0.45	-0-2	4 13	0.0	5 U.40	11	0.86	~0.45 ~1.8
40 00	17	1417	00.00	2.007	20302		1412	00.70	2.004	379 650	-0.05	- 6.0	0 17	190 I	0.50	11	0.55	-0.05 -0.2
40.00		1617	01. 00	2.04	20300	11	1012	01.27	2.03	550	0.25	0.1	11 16		5 U.8I	11	0.55	0.25 1.0
SU 30	0	1417	01.50	2.001	20300		1010	01.07	2.01	542	-0.05	-0.0	1 16	02 1	D U.10	10	0.81	-0.05 -0.2
41.00	7	1617	01.00	1.07	20300	23	1641	82.71	2.02	500	-1.21	-6-4	15 16		0.00	24	1. 21	-1.21 -4.9
41.00		1024	01.07	1.04	20109		10%1	02.11	1.99	500	-0.86	-0.3	0 10	23	0.05	18	0.91	-0.88 -3.9
47 50	10	1444	02.030	1 05	21900	0	1041	02011	1.05	500	-0.35	-0-1	7 10	10 1	0.20	11	0.55	-0.35 -1.8
42.00	10	1462	02.00	104	27617	7	1041	02011	1.92	500	0.15	0.1	2 10	10 10	2 0.60	9	0.45	0.15 0.7
43.00	2	1652	03.21	1.07	27607		1453	00.00	1.43	460	0.20	0.1	8 16	40 L	2 0.60		0.40	0.20 0.9
43.30	2	1034	02.02	1.92	27439		1002	03.21	1.71	400	0.10	0.0	6 16	46 8	5 0.40	6	0.30	0.10 0.5
44.00	7	1003	03.02	1.91	27329	2	1001	83.12	1.90	450	0.10	0.0	/ 16	1 22	0.50	8	0.40	0.10 0.5
44.70		1000	03.02	1.00	27529		1000	04.07	1.89	420	-0.25	-0.1	3 16	56	0.35	12	0.60	-0.25 - 1.1
43.00		1000	0%.0/	1.01	21090	23	1031	85.23	1-84	400	-1.10	-0.4	12 10	08	0.00	23	1.16	-1.16 -4.8
117 n 30	14	1002	8%,78	1.00	209/1	0	1041	85.23	1.87	400	-0.45	-0.2	13 16	81	0.05	10	0.50	-0.45 -2.7
46.00	0	1002	04.10	1.04	20971	0	1041	82.23	1.05	400	-0.45	- 0.2	0 16	81	0.05	10	0.50	-0.45 -2.1
40.00	8	1040	02.10	1.83	26/15	0	1941	85.23	1.83	400	-0.05	-0.0	6 16	87	0-15	4	0.20	-0.05 -0.4
9/0U0 /7 EC	5	1043	02.33	1+82	20018	1	1692	65.28	1.81	376	0.05	0.0	3 16	90	0.15	2	0.10	0.05 0.4
97.50	2	1048	07.70	1.80	20392	ļ	1044	02.04	1.80	1006	-0.05	-0.0	5 16	45	5 Uni5	4	0.20	-0.05 -0.4
48.00	0	1948	05.58	1.78	20287	0	1699	85.64	1.78	350	-0.05	- C - O	0 16		0.15	4	0.20	-0.05 -0.4
48.50	2	1700	85.69	1.77	26087	5	1704	85.89	1.77	326	-0.20	-0-1	4 16	49	L 0.05	5	0.25	-0.20 -1.6
04.00	4	1104	87.89	1.75	26025	5	1709	86.14	1.76	320	-0.25	-0.1	1 17	00 4	0.20	9	0.45	-0.25 -L.4
49.50	3	1707	86.04	1.74	25875	14	1723	86.84	1.75	300	-0.81	- C. 3	7 17	07 (0.00	16	0.81	-0-81 -4.0
50.00	0	1715	86.44	1.73	25684	0	1723	86.84	1.74	300	-0.40 ·	-C-I	5 17	12 :	3 0.15	11	0.55	-0.40 -2.1

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TABLE C-12. PEABODY-DIMMICK VERSUS NEW HAMPSHIRE FLASHING LIGHTS

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		PE	ABODY-I	DIMMICH	(N	ЕМ НАМ	PSHIRE	CT						הי	N	н	
t Ying	I NC	CUM		POWER	HAZARD	INC	CUM		POWER	HAZARD	STTH P	A TCH	LNC	CUM	1555	NATCH	1 65 5	MATCH	LECC MATCH
• Aing	ACC	MACC	8 ACC	FACTR	INDEX	ACC	#ACC	X ACC	FACTR	INDEX	TOIFF	TVAL	MICH	NTCH	#ACC	¥ ACC	#ACC	8 400	2DIEE TVAL
			•					١.									#		40111 1946
0.50	54	54	4.08	8.16	86919	55	55	4.15	8.31	487800	-0.08	-0.1	53	53	1	0.08	2	0.15	-0.08 -0.6
1.00	31	85	6.42	6.4Z	81747	34	89	6.72	6.72	337790	-0.30	-0.3	30	83	5	0.15	-	0-45	-0.30 -1.4
1.50	39	124	9.37	6.24	78485	26	115	8-69	5.79	268380	0.68	C.6	30	113	- 11	0.83	2	0.15	0.68 2 5
2.00	23	147	11.10	5.55	76919	29	144	10.88	5.44	234000	0.23	0.2	25	138	ģ	0.68	6	0.45	0.23 0.8
2.50	26	173	13.07	5.23	75609	25	169	12.76	5.11	210000	0.30	0.2	26	164	ġ	0.68	5	0.39	
3.00	22	195	14.73	4.91	74343	29	198	14.95	4.96	183820	-0.23	-0.2	27	191	í.	0.30	7	0.53	
3.50	26	221	16.69	4.77	73010	19	217	16.39	4.68	169400	0.30	0.2	24	215	6	0-45	, ,	0.15	0.30 1.4
4.00	13	234	17.67	4.42	72058	14	231	17.45	4.36	153088	0.23	0.1	12	227	7	0.53	4	0.30	0.23 0.9
4.50	21	255	19.26	4.28	71009	19	250	18.88	4.20	1 3 9 0 0 0	0.38	0_2	19	246	ģ	0.68	4	0-30	
5.00	20	275	20.77	4.15	70194	23	273	20.62	4.12	128000	0.15	C. 1	20	266	ģ	0.68	,	0,50	0.15 0.5
5.50	22	297	22.43	4.08	69424	23	296	22.36	4.06	120800	0.08	0.0	20	286		0.83	10	0.76	
6.00	26	323	24.40	4.07	68787	26	322	24.32	4.05	112720	0.08	C.O	29	315		. 0.60		0.53	0.00 0.2
6.50	18	341	25. 76	3.96	67983	18	340	25.68	3.95	106981	0.08	0.0	ĨŚ	330	11	0.83	10	0.76	
7.00	18	359	27.11	3.87	67418	24	364	27.49	3.93	100000	-0.38	-0.2	22	352		0.53	12	0.01	
7.50	17	376	28.40	3.79	66764	15	379	28.63	3.82	94560	-0.23	-0.1	15	367	ġ	068	12	0.91	-0 23 -0 7
8.00	11	387	29.23	3.65	66251	18	397	29.98	3.75	90000	-0.76	- C. 4	16	383	í.	0_30	14	1 06	
8.50	20	407	30.74	3.62	65644	10	407	30.74	3.62	85400	0.00	0.0	16	399	Å	0.60	 R	0 60	
9.00	14	4Z 1	31.80	3.53	65208	15	422	31.87	3.54	81700	-0.08	-0.0	15	414	ž	0.53	Ă	0.60	
9.50	11	432	32.63	3.43	64853	6	428	32.33	3.40	78880	0.30	0.1		420	12	0.91	8	0.60	0.30 0.9
10.00	18	450	33.99	3.40	64355	12	440	33.23	3.32	75600	0.76	0.3	14	434	16	1.21	6	0.00	0,76 21
10.50	10	660	34.74	3.31	63803	14	454	34.29	3.27	72000	0.45	Č. 2	12	446	14	1.06	Ä	0.60	0.45 1 3
11.00	8	468	35.35	3.21	63470	10	472	35.65	3.24	69225	-0.30	-0.1	14	460		0.60	12	0.91	-0.30 - 0.9
11.50	16	484	36.56	3.18	63026	20	492	37.16	3.23	66000	-0.60	-0.3	20	480	4	0.30	12	0.91	-0.60 -2.0
12.00	15	499	37.69	3.14	62709	12	504	38.07	3.17	64000	-0.38	-0.2	12	492	7	0.53	12	0,91	-0.38 -1.1
12.50	9	508	38.37	3.07	62254	11	515	38.90	3.11	61088	-0.53	- C. 2	8	500	a	0.60	15	1.13	-0.53 -1.5
13.00	7	515	38.90	2. 99	61912	8	523	39.50	3.04	59800	-0.60	-0.Z	10	510	5	0.38	13	0.98	-0.60 -1.9
13.50	14	529	39.95	2.96	61372	8	531	40.11	2.97	56480	-0.15	- C. 1	10	520	9	0.68	11	0.83	-0.15 -0.4
14.00	13	542	40. 94	2.92	61012	16	547	41.31	2.95	54000	-0.38	-0.2	12	532	10	0,76	15	1.13	-0.38 -1.0
14.50	20	562	42.45	2.93	60554	15	562	42.45	2.93	51500	0.00	C. O	19	551	- İ1	0.83	11	0.83	0.00 0.0
15.00	12	574	43.35	2.69	60231	9	571	43.13	2.88	49940	0.23	0.1	14	565	9	0.68	6	0.45	0.23 0.8
15.50	12	586	44.26	2.86	59859	12	583	44.03	2.84	48000	0.23	0.1	9	574	12	0.91	9	0.68	0.23 0.7
16.00	8	594	44.86	2.80	59533	5	588	44 .4 L	2.78	46554	0.45	C.2	4	578	16	1-21	10	0.76	0.45 1.2
16.50	11	605	45.69	2.77	59192	22	610	46.07	2.79	45000	-0.38	-0.1	18	596	9	0.68	14	1.06	-0.38 -1.0
17.00		611	46,15	2.71	58883	6	616	46.53	2.74	43260	-0.38	- C. I	7	603	8	0.60	13	0.98	-0.38 -1.1
17.50	12	623	47.05	2.69	58558	γ	623	47.05	2 -6 9	41990	0.00	0.0	8	611	12	0.91	12	0.91	0.00 0.0
	4	652	41.13	2.65	58278	17	640	48.34	2 .69	40390	-0.60	-0.2	12	623	9	0.68	17	1.28	-0.60 -1.6
18.50		639	48.26	2.61	58060	B	648	48.94	2.65	40000	-0.68	-0.3	8	631	8	0.60	17	1.28	-0.68 -1.8
10.00	17	026	49.55	2.61	57716	8	656	49.55	2.61	38800	0.00	0.0	9	640	16	1.21	16	1.21	0.00 0.0
19.50	14	610	50.60	2.60	57400	13	669	50.53	2.59	37500	0.08	0.0	11	651	19	1.44	18	1.36	0.08 0.2
20.00	8	678	51.21	2.56	57073	17	686	51.01	2.59	36000	-0-60	-0.2	[4	665	13	0.98	21	1. 59	-0.60 -1.4
20.50		000	21.01	2.55	56811	6	692	52.27	2.55	35000	-0.45	-0.2	9	674	12	0.91	18	136	-0.45 -l.l
21.00		DAT	52.14	2.49	20009	8	700	52.87	2.52	34000	-0-68	-0-2	5	679	12	0.91	21	1.59	-0.68 -1.6
21.50		702	53.02	2.41	56326	9	709	53.55	2 .49	33000	-0.53	-C.2	17	696	6	0.45	13	0.98	-0.53 -1. 6
26.00	Ö P	708	73041	2085	20028	6	715	54.00	2.45	32000	-0.53	-C.2	6	702	6	0.45	13	0.98	-0.53 -1.6
22.00	16	710	55 33	2.90	55112		122	24.23	2.42	30940	-0.45	- (. 2	6	710	6	0.45	12	0.91	-0.45 -1.4
23,50	4.) (1)	730	55 87	2070	22209 55217	× 4	731	77ª21	2.40	30000	0.00	0.0	12	722	9	0.68	9	0.68	0.00 0.0
24.00	<u>ہ</u>	746	56 77	2,24	55069	4	7.1	22021	2.30	29100	0.30	U.1	5	721	12	0.91	8	0.60	0.30 0.9
24.50	6	751	56.77	2 17	54770		763	54 07	2134	20000	0.08	0.0	8	(35	10	0.76		0.68	0.08 0.2
25.00	5	75 4	57,10	2.29	54505	, a	767	57 66	2.22	27000	-0.45	-0-1	2	740	11	0.83	13	0.98	-0.15 -0.4
				L0 L U	27202	~ ~	1 4 4	1 0 1 2	6.34	11000	- 11 - 40 - 1	~ /		165	10	11 16	16		-(1 65 - 1 2

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TABLE C-12. (cont.)

		PEA	ABODY-	DIMMIC	ĸ		NE	EW HAM	PSHIRE	CT						PO		чн	
\$ Vinc	1 NC	CUM		PO₩ER	HAZARD	INC	CUM		PUHER	HAZARD	HITH P	44 TCH	1 NC	CUM	LESS	MATCH	LESS	MATCH	HOTAM 2231
• Aing	ACC	#ACC	% ACC	FACTR	INDEX	ACC	#ACC	* ACC	FACTR	INDEX	8DIFF	TVAL	MTCH	нтсн	ØACC	X ACC	AACC	R ACC	ROLEE TWAL
																• • • • • •			
25.50	4	760	57.40	2.25	54262	7	769	58.08	2.28	76270	-0.68	- C. Z	6	752	8	0.60	17	1.28	-0.68 -1.8
26.00	8	768	58.01	2.23	54065	13	7 62	59.06	2.27	25600	-1.06	-0.4	11	763	š	0.30	19	1.44	-1.06 -2.9
26.50	10	778	58.76	2.22	53803	7	789	59.59	2.25	25000	-0.83	- 6. 3	7	770	Ŕ	0.60	10	1.66	-0.93 - 21
27.00	10	188	59.52	2.20	53595	6	795	60.05	2.22	24290	-0.53	-0.2	12	782	Ă	0.45	11	0.98	+0-53 -1-6
27.50	6	794	59.97	2.18	53368	5	800	60.42	2.20	24000	-0.45	-6.2	4	786	Ä	0.60	14	1.06	-0.45 -1.3
28.00	8	80 Z	60.57	2.16	53154	5	805	60.80	2.17	23200	-0.23	-0.1	5	791	11	0.83	14	1.06	-0.23 -0.4
28.50	3	805	60.80	2.13	5 3 0 0 3	13	818	61.78	2.17	22500	-0.96	-0.3	ģ	800	5	0.38	18	1_ 76	-0.90 - 7 7
29.00	8	813	61.40	2.12	52799	7	825	62.31	2.15	22000	-0.91	-0.3	9	809	4	0.30	16	1.21	-0.91 -2 7
29.50	5	818	61.7B	2.09	52625	8	833	62.92	2.13	21660	-1.13	-0.4	5	816	4	0.36	19	1-44	
30.00	7	825	62.31	2.08	52448	4	837	63.22	2.11	20900	-0.91	-0.3	6	820	5	0.38	17	1.28	-0-01 -2 -
30.50	4	829	62.61	2.05	52221	4	841	63.52	2.08	20400	-0.91	-0.3	ŝ	825	6	0.30	16	1. 21	-0.91 -2.7
31.00	7	836	63.14	2.04	51957	5	846	63.90	2.06	20000	-0.76	-C.2	6	831	5	0.38	15	1.13	-0.74 -2.2
31.50	4	840	63.44	2.01	51716	L	847	63.97	2.03	19680	-0.53	-0.2	ī	8 32	Â	0.60	15	1.13	-0.53 -1 5
32.00	5	845	63.82	1.99	51517	8	855	64.58	2.02	19200	-0.76	-0.2	4	836	g	0.68	10	1.44	-0.76 - 1.9
32.50	9	854	64.50	1.98	51278	5	860	64.95	2.00	18792	-0.45	- C.1	9	845	9	0.68	15	1.13	-0.45 -1 2
33.00	7	861	65.03	1.97	51105	12	872	65.86	2.00	18040	-0.83	-C.3	6	851	10	0.76	21	1.59	-0-83 -2.0
33.50	11	872	65.86	1.97	50869	6	878	66.31	1.98	18000	-0.45	- C. I	14	865	7	0.53	13	0. 9A	~0.45 -1.3
34.00	8	880	66.47	1.95	50648	8	886	66.92	1.97	17325	-0.45	-0.1	7	872	ß	0.60	16	1.06	-0-45 -1-3
34.5C	5	885	66.84	1.94	50476	6	892	67.37	1.95	16848	-0.53	-C.2	4	876	ģ	0.68	16	1.21	-0.53 -1.4
35.00	8	893	67.45	1.93	50250	11	903	68.20	1.95	16380	-0.76	-0.2	12	688	Ś	0.38	15	1,13	-0.76 -2.2
35.50	- 6	899	67.90	1.91	50017	4	907	68.50	1.93	16000	~0.60	-0.2	6	894	. 5	0.38	13	0.94	-0.60 -1.9
36.00	5	904	68.28	1.90	49793	4	911	68.81	£.91	15600	-0.53	-0.2	2	896	8	0.60	15	1.13	-0.53 -1.5
36.50	4	908	68.58	1.80	49600	0	911	68-81	1.89	15230	-0.23	-0.1	4	900	8	0.60	ū	0.83	-0.23 -0.7
37.00	6	914	69.03	1.87	49391	8	919	69.41	1.88	15000	-0.38	-0-1	7	907	7	0.53	12	0.91	-0.38 -L.1
37.50	10	924	69.79	1.86	49183	3	922	69.64	1.86	14560	0.15	0.0	4	911	13	0.98	11	0.83	0.15 0.4
36.00	7	931	70.32	1.85	49003	5	927	70.02	1.84	14220	0.30	0.1	5	916	15	1.13	11	0.83	0.30 0.6
38.50	5	936	70.69	1.84	48752	8	935	70.62	1.83	14000	0.08	C. 0	7	923	13	0.98	12	0.91	0.08 0.2
39.00	2	938	70.85	1 . 82	48540	11	946.	71.45	1.83	13600	-0.60	-0.2	8	931	7	0.53	15	1.13	-0.60 -1.7
39.50	7	945	71.37	1 - 81	48332	7	953	71.98	1.82	13200	-0.60	-0.2	6	937	8	0.60	16	1.21	-0.60 -1.6
40.00	11	956	72.21	1.81	48114	5	958	72.36	1.81	12800	-0.15	-0.0	10	947	9	0.68	11	0 83	-0.15 -0.4
40.50	4	96.0	72.51	1.79	48008	6	964	72.81	1.80	12500	-0.30	-0.1	5	952	8	0.60	12	0.91	-0.30 -0.9
41.00		967	73.04	1.78	47788	19	983	74.24	1.81	12000	-1.21	-0。4	11	963	4	0.30	20	1.51	-1.21 -3.3
41.50	12	979	73.94	1.78	47585	0	983	74.24	1.79	12000	-0.30	-0-1	8	971	8	0.60	12	0.91	-0.30 -0.9
42.00	8	987	(4.55	1.77	47360	1	984	74.32	1.77	11792	0.23	0.1	5	976	11	0.83	9	0.60	0.23 0.7
42.50		993	75.00	1.70	47236	4	988	74.62	1.76	11388	0.38	0.1	6	982	11	0.03	6	0.45	0.30 1.2
43.00	4	997	75.30	1.75	47082	L	989	74.70	1.74	11025	0.60	C . Z	3	9 0 5	12	0.91	4	0.30	0.60 2.0
43.50	4	1004	75 07	1.73	40403	4	993	15.00	1.72	10800	0.38	0.1	2	987	11	0.83	6	0.45	0.30 l.2
44.00	U 4	1004	12.03	1.72	40007		1001	12.00	1.12	10500	0.23	0.1	2	99Z	12	0.91	9	0.68	0.23 0.7
44.50	0	1010	76 20	1.70	40470		1009	10.21	1.71	10200	0.08	0.0	0	996	15	0.91	11	0.83	0.08 0.2
45 50	7	1010	74 01	1 40	46290	10	1019	10.90	1.1	10000	-0.68	-0.2	2	1003		0.53	16	1.21	-0.68 -1.9
45.00	÷	1024	77 34	1 407	46041	1	1020	77.09	1.69	9840	-0.23	-0.1	3	1006	11	0.83	14	1.06	-0.23 -0.6
46.50		1024	77 44	1 47	4 5 6 0 0	2	1023	77 67	1.00	9600	0.08	0.0		1011	13	0.98	12	0.91	0.00 0.2
47.00	7	1020	77.97	1.64	4550+		1027	77 07	1 4 4	9329 8182	0.23		<u>م</u>	1015	13	0.90	LO	0. 76	0.23 0.6
47.50	2	1034	78.10	1.66	45300	0 6 (1077	78 85	1.44	9000	0.00	0.0	2	1020	11	0.63	11	0-83	0.00 0.0
48-00	4	1038	78.40	1.67	45140		1044	70 15	1 7 5	8000	-0.76	-0.7	2	1023	¥	0.00	19	1.44	-0.76 -1.9
48.50	5	1041	78.78	1.62	44947	~	1054	79.61	1 600	9500	-0.83	-0-2	5	1035	•	0.99	10	1. 21	-0.70 -2.1
49.00	5	1048	79.15	1.62	44844	A	1062	80.21	1 64	8260	-0.03	-0.2	2	1041	0- 7	0.99		1.20	-0.65 -2.3
49.50	จ์	1051	79.38	1.60	44659	ŏ	1071	80.89	1.62	8000	-1.00	- 1 4	ب	1047		0.20	41	10.28	-1.51 -2.6
50.00	6	1057	79.83	1.60	66657	Ó	1071	80.89	1.62	8000	-1-06	-0.3	5	1052	9	0.30	29	1 44	-104-300

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TABLE C-13. PEABODY-DIMMICK VERSUS NEW HAMPSHIRE AUTOMATIC GATES

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		PE /	BODY-	DIMMIC	i,		NÉ	ы нач	PSHIRE	СТ					¢	חי		н	
& Ying	INC	CUM		POWE R	HAZARD,	INC	CUM		POWER	HAZARD	ыты м	ATCH	I NC	CUM	1 6 5 5	MATCH	1 6 5 5	МАТСН	LECC MATCH
* XING	ACC	# AC C	% ACC	FACIO	INDEX	ACC	MALC	# ACC	FACTR	INDEX	X D I F F	TVAI	MICH	MTCH	# 100	* ***	MACC	W ACC	TOTES THATCH
															WACC.		PACC	8 ALL	AUTER TAN
0.50	5	5	1.41	7.82	100920	7	7	1.98	3.951	272000	-0.56	-0.6	5	5	0	0 00	2	0 64	-0.64 1.4
1.00	8	13	3.67	3.67	95995	7	14	3.95	1.95	939620	-0.29	-0.2		11		0.00	<u>,</u>	0.00	-0.30 -1.4
1.50	3	16	4.52	3.01	94219	Ś	19	5.37	3.58	819460	-0.85	_0.5	2	1.5		0.00	1	0.20	-0.28 -1.0
2.00	7	23	6.50	3.25	92468	5	26	6 74	3 30	74 9000	-0.03	-0.5	,	10		0.00	د ا	0.85	-0.85 -1.7
2.50	4	27	1.63	1.05	40846	1	27	7.63	3 05	477440	-0.78	-0.1			1	0.28		0.50	-0-28 -0.6
3.00		12	9 04	3 01	80333	,	21	0 00	1.07	503(00	0.00	0.0	-	20	L	0.28		0.28	0.00 0.0
3.50	Å	40	11.30	3.23	88022	0		7.07	3.30	545400	-0.85	-0.4		29	د	0.85	6	1.69	-0.85 -1.0
6 00	5	40	17 71	3 10	84500	7		12.49	3.22	540000	-1+13	-0.4	10	34	1	0.28	5	1.41	-1.13 -1.6
4.50		51	16 61	3-10	46501	4	40	12.99	3.23	488400	-0.28	-0.1	5	44	1	0.26	2	0.56	-0.28 -0.6
5 00	2	51	17,71	3.00	0,000	2	49	13.04	3.08	428080	0.56	L. 2	5	49	2	0.56	0	0.00	0.56 1.4
5 50	2		14.97	2.99	84012	o o	49	13.84	2.11	425000	1.13	0.4	0	49	4	1-13	0	0.00	1.13 2.0
4 00			13434	2.02	03101		54	12.72	2.17	400320	0.28	0.1	4	53	2	0.56	1	0.28	0.28 0.6
6.00	~	60	10.95	2.02	82848	11	60	18.36	3.06	373800	-1.41	-0.4	7	60	0	0.00	5	1.41	-1.41 -2.2
1 00	с Б	00	10.04	2.07	62063	1	12	20.34	3.13	350000	-1.69	-0.5	6	66	0	0.00	6	1.69	-1.69 -2.4
7.00	~	11	20.06	2.87	81517	5	17	21.75	3.11	334400	-1.69	-0.5	4	70	1	0.28	7	1.98	-1.69 -2.1
7.50	6		21.75	2.90	80991	2	79	22.32	2.98	321600	-0.56	-C.2	6	76	1	0.28	3	0.65	-0.56 -1.0
8.00	5	82	23.16	2.90	80544	4	83	23.45	2.93	308100	-0.28	-0.1	3	79	3	0.85	4	1.13	-0.28 -0.4
8.50	9	90	25.42	2.99	79908	6	89	25.14	2.96	297600	0.28	C.1	8	87	3	0.85	2	0.56	0.28 0.4
9.00	1	91	25.71	2.86	79445	1	90	25.42	2.82	285200	0.28	C • 1	2	89	2	0.56	1	0.28	0.28 0.6
9.50	1	92	25.99	2.74	78941	4	54	26.55	2.80	270100	-0.56	-0.1	3	92	0	0.00	2	0.56	-0.56 -1.4
10.00	1	93	26.27	2.63	78667	2	96	27.12	2.71	261000	-0.85	-0,Z	L	93	ò	0.00	3	0.85	-0.85 -1.7
10.50	l	94	26.55	2.53	78119	0	96	27.12	2.58	254400	-0.56	-0.1	ī	94	ō	0.00	2	0.56	-0.56 -1.4
11.00	1	95	26.84	2.44	77664	0	96	27.12	2.47	246400	-0.28	- C. I	ō	94	ī	0.28	2	0.56	-0.28 -0.6
11.50	L	96	27.12	2.36	77124	4	100	28.25	2.40	237000	-1-13	-0.3	2	96		0.00	2	1.13	-1 13 -2 0
12.00	9	105	29.66	2.47	76438	4	104	29.38	2.45	228600	0.28	C.1	7	103	ž	0.56	i	0.28	0.28 0.6
12.50	4	109	30.79	7.46	76066	7	111	31.36	2.51	220000	-0.56	-0-1	4	107	2	0.56	4	1.13	-0.56 -0.8
13.00	4	113	31.92	2.46	15642	z	113	31.92	2.46	210000	0.00	0.0	5	112	ī	0 78		0.24	0.00 0.0
13.50	3	116	32.77	2.43	75103	0	113	31.92	2.36	202800	0.85	0.2	í	113	3	0.85	Ô	0.00	0.00 0.0
14.00	3	119	33.62	2.40	74681	2	115	32.49	2.32	197145	1.13	C - 4	,	115		1 13	ő	0.00	
14.50	2	121	34.18	2.36	74 05 3	2	117	33.05	2.28	190400	1.13	0.3	2	117		1 1 7	ő	0.00	1.13 2.0
15.00	1	122	34.46	2.30	73727	5	122	34 46	2 40	183720	0.00	0.0		121		1+13		0.00	1.13 2.0
15.50	i	123	34.75	2.24	13400		1.25	35.31	2.28	173600	-0.56	- (1		121		0.20	ļ	0.20	0.00 0.0
16.00	2	125	35.31	2.21	72914	5	127	35.88	2.24	1 70000	-0.56	-01	1	122	2	0.00	7	1 1 2	-0.56 -0.8
16.50	5	130	36.12	2.23	72576	2	120	36 44	2 2 2 1	168000	-0.10	-0-1	1	122	3	0.85	2	1.41	-0.56 -0.7
17.00	Ō	140	36-72	2.16	72450	2	1127	17 20	2 10	161600	-0.26	-01		120		1.13	5	0.85	0.28 0.4
17.50	4	143	37.57	2.15	71841	, ,	134	39.43	2 • 1 7	155900	-0.98	-0.1		127		0.85	~ ~	1.41	-0.56 -0.7
18.00	ő	144	37.57	2.09	71550	3	130	10.12	2.20	150000	-0.65	- 6. 2	2	152	L C	0.28	-	1.13	-0.85 -1.3
18.50	ű.	117	38 70	2 09	71253	ر د	1.59	20.01	2 . 10	167600	-1.09	-0.4	ļ	133	0	0.00	6	1.69	-1.69 -2.4
19.00	i	149	18 09	2 05	70902	2	141	20.03	2,12	14/400	-1.13	-1.2		1.37	0	0.00	4	1.13	-1.13 -2.0
19.50		146	40 49	2 00	70560	<u> </u>	1/1	27602	2.10	144000	-0.85	-0.2	L	138	0	0.00	5	0.85	-0.85 -1.7
20.00	2	144	40.00	2.07	70369	Ŷ	147	41.000	2.13	1 39950	-0.85	-0.2	6	144	0	0.00	3	0.85	-0.85 -1.7
20.00	5	14.0	41.24	2.00	10330	4	151	42.00	2.13	135000	-1-41	-0-1	2	146	0	0.00	5	1.41	-1.41 -2.2
20.00	2	140	41.01	2.04	64911	1	152	42.94	2.09	131810	-1.13	-0.Z	1	147	1	0.28	5	1.41	-1.13 -1.6
21.00	د	121	42.00	2.05	0 40 8 1	2	154	43.50	Z.07	127500	-0.85	-0.2	3	150	L	0.28	4	1.13	-0.85 -1.3
21.00	2	153	43.22	2.01	69211	4	158	44-63	2.08	123760	-1.41	-C.3	3	153	0	0.00	5	1.41	-1.41 -2.2
22.00	1	104	43.50	1.98	64038	Ļ	159	44.92	2.04	120000	-1.41	- C . 3	1	154	0	0.00	5	1.41	-L.41 -2.2
22.50	د ر	157	44.35	1.97	68746	1	160	45.20	2.01	118500	-0.85	-0-2	L	155	2	0.56	5	1-41	-0.85 -1.1
23.00	4	161	42.48	1.98	68432	5	165	46.61	2.03	114504	-1.13	- 0. 2	4	159	2	0.56	6	1.69	-1.13 -1.4
23.50	6	167	47.18	2.01	68063	2	167	47.18	2.01	111650	0.00	0.0	4	163	4	1.13	4	1.13	0.00 0.0
24.00	5	172	48.59	2.02	67730	2	169	47.74	1.99	108120	0.85	0.2	3	166	6	1.69	3	0.85	0.85 1.0
24.50	0	172	48.59	1.98	67386	3	172	48.59	1.98	105000	0.00	0.0	3	169	3	0.85	3	0.85	0.00 0.0
25.00	1	173	48.87	1.95	67137	0	172	48.59	1.94	102400	0.28	0.1	0	169	4	1.13	3	0.85	0.28 0.4

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TABLE C-13. (cont.)

		PEZ	BODY-		ς		N	EW HAMI	PSHIRE	СТ				ç	0		134	
* Ying	INC	CUM		POWER	HAZARD	1 NC	CUM		POWER	HAZARD	WITH MATCH	INC	CUM	LESS	MATCH	1555	MATCH	LESS MATCH
• AINE	ACC	#ACC	X ACC	FACTR	I ND F X	ACC	NACC	\$ ACC	FACTR	INDEX	XULFE TVAL	MICH	MTCH	#ACC	& ACC	WACC	& ACC	TOIFF IVAL
25.50	2	175	49.44	1.94	66959	3	175	49.44	1.94	100000	0.00 0.0	4	173	2	0.56	2	0.56	0.00 0.0
26.00	2	177	50.00	1.92	66697	4	179	50.56	1.94	98064	-0.56 -0.1	4	177	0	0.00	2	0.56	-0.56 -1.4
26.50	2	179	50.56	1.91	66429	0	179	50.56	1.91	96000	0.00 0.0	1	178	1	0.28	1	0.28	0.00 0.0
27.00	1	180	50.85	1.88	66178	0	179	50.56	1.87	93500	0.28 C.1	i	179	ī	0.28	ō	0.00	0.28 1.0
27.50	3	183	51.69	1.88	65731	6	185	52.26	1.90	90000	-0.56 - 0.1	2	181	2	0.56	4	1.13	-0.56 -0.8
28.00	1	184	51.98	1.86	65476	Э	188	53.11	1.90	87696	-1.13 -0.2	2	183	ī	0.28	5	1.41	-1-13 -1-6
28.50	1	185	52.26	1.83	65241	1	189	53.39	1.87	86350	-1.13 -0.2	1	184	1	0.28	5	1.41	-1.13 -1.6
29.00	4	189	53.39	1.84	65099	5	194	54.80	1.89	84000	-1.41 - 0.3	4	188	i	0.28	6	1.69	-1.41 -1.9
29.50	2	191	53.95	1.83	64804	0	194	54.80	1.86	82080	-0.85 -C.2	1	189	2	0.56	5	1.41	-0.85 -1.1
30.00	2	193	54.52	1.82	64565	4	198	55.93	1.86	80000	-1.41 -C.3	3	192	1	0.28	6	1. 69	-1.61 -1.9
30.50	2	195	55.08	1.81	64388	0	198	55.93	1.83	79200	-0.85 -0.2	2	194	ī	0.28		1.13	-0.85 -1.3
31.00	3	198	55.93	1.80	64149	2	200	56.50	1.82	78200	-0.56 -0.1	4	198	ō	0.00	2	0.56	-0.56 -1.4
31.50	2	20.0	56.50	1.79	63827	0	200	56.50	1.79	75400	0.00 0.0	0	198	2	0.56	,	0.56	0.00 0.0
32.00	1	201	56.78	1.77	63602	4	204	57.63	1.80	72100	-0.85 -0.1	3	201	ō	0.00	3	0.85	-0.85 -1.7
32.50	3	204	57.63	1.77	63370	7	211	59.60	1.83	/1000	-1.98 -C.3	ź	204	ĩ	0.28	á	2.26	-1.98 - 2.3
33.00	4	208	58.76	1.78	63153	2	213	60.17	1.82	70000	-1.41 -0.2	2	205	3	0.85	ā	2.26	
33.50	5	213	60.17	1.80	62959	2	215	.60.73	1.81	68000	-0.56 - 6.1	6	211	2	0.56	4	1.13	-0.56 -0.8
34.00	0	213	60.17	1.77	62660	0	215	60.73	1.79	66000	-0.56 - (.)	Ō	211	,	0.56	ċ	1.13	-0.56 - 0.8
34.50	1	214	60.45	1.75	62349	1	216	61.02	1.77	64880	-0.56 -0.1	ĭ	212	2	0.56	4	1.13	-0.56 -0.8
35.00	2	216	61.02	1.74	62199	2	218	61.58	1.76	6 12 50	-0.56 - 6.1	- î	215	ĩ	0.28		0.85	-0.56 -1.0
35.50	2	218	61.58	1.73	61948	3	221	62.43	1.76	62020	-0.85 -0.1	í	216	2	0.56	5	1.41	-0-85 -1.1
36.00	3	221	62.43	1.73	61663	1	222	62.71	1.74	60600	-0.28 - C.O	;	218	1	0.85	á	1 13	-0.28 -0.4
36.50	2	223	62.99	1.73	61440	ō	222	62.71	1.72	60000	0.28 0.0	5	220		0.85	2	0 56	0.20 0.4
37.00	0	223	62.99	1.70	61197	ī	223	62.99	1.70	58104	0.00 0.0	1	221	,	0.56		0.56	
37.50	4	227	64.12	\$ 1.71	60998	6	229	64.69	1.73	56240	-0.56 -0.1	ī	222	5	1.41	7	1 68	-0.56 -0.6
38.00	1	228	64.41	1.69	60652	3	232	65.54	1.72	55000	-1.13 -0.2		225	á	0.85	, 7	1 64	
38.50	0	228	64.41	1.67	60388	4	236	66.67	1.73	54000	-2.26 -0.4	2	227	í	0.28	ģ	2.54	-2.26 - 2.5
39.00	l	229	64.69	1.66	60231	Ú	236	66.67	1.71	52500	-1.98 - 0.3	ĩ	228	i	0.28	, A	2 26	-1 98 -2 3
39.50	2	Z 3 1	65.25	1.65	59970	ā	236	66.67	1.69	51320	-1.41 -0.2	2	230	i	0.28	6	1 60	
40.00	0	Z31	65.25	1.63	59828	2	238	67.23	1.68	50040	-1.98 -0.3	ì	231		0.00	7	1 00	-1 99 - 7 6
40.5C	1	232	65.54	1.62	59601	ĩ	239	67.51	1.67	49000	-1.98 -0.3		231	1	0.28	Å	2 26	-1 98 -7 3
41.00	0	232	65.54	1.60	59665	ī	240	67.80	1.65	48000	-2-26 -0.6	ň	231	i i	0 26	ő	2 56	-2 24 - 2 6
41.50	7	239	67.51	1.63	59243	i	241	68.08	1	47100	-0.56 -(.)	7	239	-	0.28	á	0.85	
42.00	2	241	68.08	1.67	59008	ō	241	68.08	1.62	46000	0.00 0.0	;	240	:	0.20	1	0.29	
42.50	0	241	68.08	1.60	58828	3	244	68.93	1.62	45000	-0.85 -0.1	ĩ	241	ò	0.00		0.85	
43.00	2	243	68.64	1.60	58528	ō	744	68.93	1.60	44000	-0.28 -0.0	,	243	0	0,00	í	0.00	
43.50	2	245	69.21	1.59	58165	ĩ	245	69.21	1.59	42650	0.00 0.0	0	243	,	0.50	2	0.56	
44.00	2	247	69.77	1.59	57940	2	247	69.77	1.59	41280	0.00 0.0	2	245	5	0.56	,	0.56	0.00 0.0
44.50	2	249	70.34	1.58	57619	2	249	70.34	1.58	40000	0.00 0.0	3	249	1	0.28	1	0.29	
45.00	0	249	70.34	1.56	57389	ō	269	70.34	1.56	40000	0.00 0.0	0	248	1	0.28	1	0,20	
45.50	ŋ	249	70.34	1.55	57197	2	251	10.90	1.56	39060	-0.56 -0.1	1	240		0.20	2	0.20	
46.00	i	250	70.62	1.54	56970		255	72.03	1.57	37800		1	250	~	0.00	2	1 41	
46.50	ī	251	70.90	1.52	56763	i	256	72.32	1.56	37200	-1.41 - 0.7		250	1	0.00	ź	1 40	
47.00	3	254	71.75	1.53	56403		260	73.45	1.56	36160	-1.69 -0.2	2	252	1	0 4 2 0	0 9	1.07	
47.50	2	256	72. 32	1.52	56214		260	73.74	1.55	36000		2	256	2	0.54	7	1 00	
48.00	ī	257	72.60	1.5	55961	1	262	74.01	1.54	34450		1	254	2	0.50		1 00	- Lovi - Lo/
48.50	2	259	73.14	1, 51	55760	,	264	74.58	1.54	33600	-1.61 -0.2	2	257	2	0.90		1 00	
49.00		262	74.01	1.51	55629	<u>د</u>	264	74.59	1.57	32500		2	260	2	0 54		1. 78	-1096 -107 -064 -00
49.50	ó	262	74.01	1.50	55422	ň	264	74.58	1.51	32000	-0.56 - 0.1	د م	260		0.00	4	1.13	-0.56 -0.8
50.00	4	266	75.14	1 50	55290		204	76 44	1 60	30900	0.78 0.0	U F	200	2,	0.20	4	1 1 1 1	
20.000		200		1.0	17200	L	200	/ 1.00	1.90	30800	Un 20 Un U	2	200	L	U₀∠B	U	0.00	0.28 1.0

APPENDIX D

REGRESSIONS

The following symbols are used as shorthand identifiers in the regressions.

- 1. AADT average daily traffic.
- 2. ACC a crossing from the accident data base.
- 3. C same as 1.
- 4. DT, DTHRU, DAY THRU number of day thru trains.
- 5 DAY SWITCH number of day switch trains.
- 6. FC, FC-ROAD the units digit of the functional classification of road over crossing.
- 7. FLGV19.DAT the file name of the regression which is the volume model for flashing lights.
- 8. GATE09 the file name of the regression which is the volume model for gates.
- 9. H, H(ofVIT16) the file name of the regression which is the volume model for crossbucks.
- 10. HWY PAVED is highway paved?
- 11. LANES the number of traffic lanes.
- 12. LOG, LOGIO refers to LOG₁₀.
- 13. LOG T**2 refers to $[LOG_{10}(T+1)]^2$.
- 14. LOG C**2 refers to $[LOG_{10}^{-1}(C+1)]^2$.
- 15. MAIN TRACKS, MAIN TRKS the number of main tracks.
- 16. MAX, MAX SPEED typical maximum speed.
- 17. MIN typical minimum speed.
- 18. N, NITE the number of night trains.
- 19. NITE SWITCH the number of night switch trains.
- 20. NITE THRU the number of night thru trains.
- 21. NOACC a crossing from the non-accident data base.
- 22. NRBY XING HWY nearby intersecting highway?
- 23. OPEN=1 NOT OPEN=2 from type of development: open is open space (1) and not open is otherwise (2, 3, 4, 5).

- 24. POP, POPULATION the tens digit of the functional classification of road over crossing.
- 25. RESID=2 NON-RESID=1 from type of development: resid is residential (2), non-resid is otherwise (1, 3, 4, 5).

26 RR ADV WARN - is railroad advance warning sign present?

- 27. SWITCH number of switch trains.
- 28. T, TRAIN number of total train movements.
- 29. TRUCKS estimated percent trucks.
- 30. TYP MAX SPEED same as 16.
- 31. TYP MIN SPEED typical minimum speed.
- 32. U=1 R=2, U=0 R=1 from highway system U is urban and R is rural.
- 33. XING ANGLES smallest crossing angle.
- 34. 1 refers to the intercept of a non-linear regression.
- 35. 10K 10,000.
- 36. * multiplication sign.
- 37. ** exponent sign.

LIST OF TABLES IN APPENDIX D

TAI	<u>BLE</u>	TITLE
D -	1	TYPICAL LINEAR REGRESSION EARLY EXPLORATORY STAGE.
D -	2	VARIABLES USED FOR VOLUME LINEAR REGRESSION MODEL.
D -	3	LINEAR REGRESSIONS COMBINING NON-VOLUME VARIABLES
		WITH BEST LINEAR VOLUME MODEL.
D -	4	SIMILAR TO D-3 DIFFERENT VARIABLES.
D -	5	ILLUSTRATION OF "MIGRATION" AND CONVERGENCE
		SUCCESSIVE ITERATIONS.
D -	6	SELECTION REGRESSION - CROSSBUCKS.
D-	7	SELECTION REGRESSION - FLASHING LIGHTS.
D-	8	SELECTION REGRESSION - AUTOMATIC GATES.
D -	9	MIGRATION - FLASHING LIGHTS - VOLUME.
D - 1	10	MIGRATION - FLASHING LIGHTS - COMPREHENSIVE.
D - 1	11	MIGRATION - GATES - VOLUME.
D - 1	12	MIGRATION - GATES - COMPREHENSIVE.
D-1	13	SELECTION REGRESSION - GATES

D-3

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TABLE D-1. TYPICAL LINEAR REGRESSION EARLY EXPLORATORY STAGE

STEP-WISE MULTIPLE REGRESSION.....BIGREG

NUMBER	0F	OBSERVATIONS	16900
NUMBER	0F	VARIABLES	28
NUMBER	0F	SELECTIONS	1

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CONSTANT TO LIMIT VARIABLES 0.00000

VARIABLE NO.	MEAN	STANDARD DEVIATION	VARIABLE
- 1	1.82444	0.54133	LANES
2	762.95160	2603.80753	AADT
3	5.62556	0.78962	FC – ROAD
4	0.81101	1.47345	POPULATION
5	1.65527	0.47530	U=1 R=2
6	2.76639	0.50006	XING ANGLE
7	1.50627	0.50021	OPEN=1 NOT OPEN=2
8	1.23834	0.42636	RESID=2 NON-RESID=1
9	1.58166	0.57752	NRBY XINGHWY Y=1 N=2
10	1.41385	0.49265	HWY PAVED Y=1 N=2
11	0.93822	0.40482	MAIN TRACKS
12	30.89574	17.71684	TYP MAX SPEED
13	14,17320	14.24937	TYP MIN SPEED
14	2.52675	3.90665	DAY THRU
15	0.87552	2.47134	DAY SWITCH
16	2.15538	3.74809	NITE THRU
17	0.43710	2.08203	NITE SWITCH
18	0.47553	0.57458	LOGIO(TRAIN + .3)
19	2.19967	0.75391	LOG10(AADT+.3)
20	35.72553	116.43390	FC * AADT
21	314.85872	1072.77081	AADT/(LANES+.3)
22	24.18595	42.33637	AADT*DTHRU **.5412
23	915.30181	2954.93851	AADT/(POP+.3)
24	29.37711	61.15177	#22 * (4-XING ANG)
25	3.72169	16.81865	MAX*DTHRU*AADT/10K
26	4.10377	20.42527	MIN*TRAIN*AADT/10K
27	1724.75166	7361.60431	AADT*DTHRU/(POP+.3)
28	0.07373	0.26134	ACC=1 NOACC=0

SELECTION....1

DEPENDE	INT	VARIABLE.	2	28
NUMBER	0F	VARIABLES	FORCED	0
NUMBER	0F	VARIABLES	DELETED	0

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VARIABLE ENTERED.....22

SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS DUCED IN THIS ST	S STEP EP	76.690 0.066		
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES RED OPORTION REDUCED	JCED	76.690 0.066	0F	1154.135
FOR 1 VARIAB MULTIPLE CO (ADJUS F-VALUE FOR STANDARD ER (ADJUS	LES ENTERED RRELATION COEFFIC TED FOR D.F.) ANALYSIS OF VAR ROR OF ESTIMATE. TED FOR D.F.)	CIENT 0.258 0.258 IANCE 1202.757 0.253 0.253			
VARIABLE <u>NUMBER</u> 22 INTERCEPT	REGRESSION COEFFICIENT 0.00159 0.03524	STD. ERROR OF REG. COEFF. 0.00005	COMPU T-VAL 34.68	TED UE 31	

STEP 2

VARIABLE ENTE	RED 10				
SUM OF SQUARE PROPORTION RE	S REDUC E D IN THIS DUCED IN THIS STE	STEP P	10.510 0.009		
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDU OPORTION REDUCED.	CED	87.200 0.076	OF	1154.135
FOR 2 VARIAE MULTIPLE CC (ADJUS F-VALUE FOR STANDARD ER (ADJUS	BLES ENTERED DRRELATION COEFFIC TED FOR D.F.) ANALYSIS OF VARI ROR OF ESTIMATE TED FOR D.F.)	IENT 0.275 0.275 ANCE 690.492 0.251 0.251			
VARIABLE <u>NUMBER</u> 22 10 INTERCEPT	REGRESSION <u>COEFFICIENT</u> 0.00145 -0.05196 0.11200	STD. ERROR OF <u>REG. COEFF.</u> 0.00005 0.00403	COMPU <u>T-VAL</u> 31.04 -12.90	TED J <u>E</u> 17)2	

VARIABLE ENTERED.....18

SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS DUCED IN THIS STE	STEP P	6.602 0.006		
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDU OPORTION REDUCED.	CED	93.802 0.001	OF	1154.135
FOR 3 VARIAB MULTIPLE CO (ADJUS F-VALUE FOR STANDARD ER (ADJUS	LES ENTERED RRELATION COEFFIC TED FOR D.F.) ANALYSIS OF VARI ROR OF ESTIMATE TED FOR D.F.)	IENT0. ANCE498. 0	.285 .285 .231 .251 .251		
VARIABLE NUMBER 22 10 18 INTERCEPT	REGRESSION <u>COEFFICIENT</u> 0.00120 -0.05616 0.03893 0.10557	STD. ERROR REG. COEFF 0.00005 0.00404 0.00380	OF COMP - T-VA -13. 10.	UTED LUE 720 916 256	

STEP 4

SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS DUCED IN THIS STE	S STEP EP		4.739 0.004		
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDU OPORTION REDUCED.	JCED		98.540 0.085	OF	1154.135
FOR 4 VARIAB MULTIPLE CO (ADJUS F-VALUE FOR STANDARD ER (ADJUS	LES ENTERED RRELATION COEFFIC TED FOR D.F.) ANALYSIS OF VARI ROR OF ESTIMATE TED FOR D.F.)	CIENT ANCE 3	0.292 0.292 94.290 0.250 0.250			
VARIABLE <u>NUMBER</u> 22 10 18 19 INTERCEPT	REGRESSION <u>COEFFICIENT</u> 0.00100 -0.03300 0.04687 0.03005 0.00790	STD. ERRO <u>REG. COF</u> 0.0000 0.0048 0.0039 0.0034	DR OF <u>EFF</u> 3 0 5	COMPU T-VAL 17.2 -6.8 12.0 8.7	TED UE 252 338 332 709	

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STEP 5	-		,		
VARIABLE ENTE	RED 4				
SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS DUCED IN THIS STR	S STEP	2.120 0.002		
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDU OPORTION REDUCED	JCED	100.661 0.087	0F	1154.135
FOR 5 VARIAB MULTIPLE CO (ADJUS F-VALUE FOR STANDARD ER (ADJUS	LES ENTERED RRELATION COEFFIC TED FOR D.F.) ANALYSIS OF VARI ROR OF ESTIMATE TED FOR D.F.)	CIENT 0. 0. CANCE 322. 0. 0.	295 295 849 250 250		
VARIABLE NUMBER 22 10 18 19 4 INTERCEPT	REGRESSION COEFFICIENT -0.03081 0.04557 0.02224 0.00874 0.01515	STD. ERROR REG. COEFF 0.00006 0.00484 0.00390 0.00370 0.00150	OF COMPI 	JTED <u>JUE</u> 507 371 590 015 331	
STEP 6					
VARIABLE ENTE	RED3				
SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS DUCED IN THIS STE	STEP P	1.379 0.001		
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDU OPORTION REDUCED.	CED	102.040 0.088	0F	1154.135
FOR 6 VARIAB MULTIPLE COI (ADJUS F-VALUE FOR STANDARD ERI (ADJUS	LES ENTERED RRELATION COEFFIC TED FOR D.F.) ANALYSIS OF VARI ROR OF ESTIMATE TED FOR D.F.)	IENT 0. 0. ANCE 273. 0. 0.	297 297 068 250 250		
VARIABLE <u>NUMBER</u> 22 10 18 19	REGRESSION <u>COEFFICIENT</u> 0.00100 -0.03058 0.04747	STD. ERROR <u>REG. COEFF</u> 0.00006 0.00483 0.00392	OF COMPL - T-VAL 17.3 -6.3 12.3	ITED <u>.UE</u> 338 328 121	

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VARIABLE ENTERED..... 11

SUM OF SQUARE PROPORTION RE	ES REDUCED IN THIS EDUCED IN THIS STE	STEP P	1.399 0.001		
CUMULATIVE SU CUMULATIVE PR	UM OF SQUARES REDU ROPORTION REDUCED.	CED	103.439 0.090	OF	1154.135
FOR 7 VARIAE MULTIPLE CC (ADJUS F-VALUE FOF STANDARD EF (ADJUS	BLES ENTERED DRRELATION COEFFIC STED FOR D.F.) R ANALYSIS OF VARI RROR OF ESTIMATE STED FOR D.F.)	IENT 0.299 0.299 ANCE 237.570 0.249 0.249)))		
VARIABLE NUMBER 22 10 18 19 4 3 11 INTERCEPT STEP 8	REGRESSION <u>COEFFICIENT</u> 0.00093 -0.03041 0.04342 0.01844 0.01094 -0.01343 0.02591 0.07526	STD. ERROR OF REG. COEFF. 0.00006 0.00483 0.00401 0.00401 0.00154 0.00278 0.00546	COMP T-VA 15. -6. 10.8 4. 7. -4.8 4.	UTED LUE 295 338 599 118 327 743	
VARIABLE ENTE	RED 9				
SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS DUCED IN THIS STE	STEP P	0.865 0.001		
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDU ROPORTION REDUCED.	CED	104.305 0.090	0F	1154.135
FOR 8 VARIAE MULTIPLE CO (ADJUS F-VALUE FOR STANDARD ER (ADJUS	BLES ENTERED RRELATION COEFFIC TED FOR D.F.) ANALYSIS OF VARI ROR OF ESTIMATE TED FOR D.F.)	IENT 0.301 0.300 ANCE 209.773 0.249 0.249			
VARIABLE <u>NUMBER</u> 22 10 18 19 4 3 11 9 INTERCEPT	REGRESSION <u>COEFFICIENT</u> 0.00093 -0.02918 0.04254 0.01800 0.01079 -0.01363 0.02616 -0.01253 0.09575	STD. ERROR OF <u>REG. COEFF.</u> 0.00006 0.00484 0.00401 0.00401 0.00154 0.00278 0.00546 0.00336	COMP T-VA 15.7 -6.0 10.6 4.4 7.0 -4.9 4.7 -3.7	UTED LUE 208 229 503 489 220 200 200 200 200 231	

D-8

0.00097

3.567

1154.135

STEP 9

17

INTERCEPT

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0.00347 0.09748

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SUM OF SQUA	RES REDUCED IN THIS REDUCED IN THIS STE	STEP P	0.790 0.001	
CUMULATIVE CUMULATIVE	SUM OF SQUARES REDU PROPORTION REDUCED.	CED	105.095 02091 OF	
FOR 9 VARI/ MULTIPLE (ADJ) F-VALUE STANDARD (ADJ)	ABLES ENTERED CORRELATION COEFFIC JSTED FOR D.F.) OR ANALYSIS OF VARI ERROR OF ESTIMATE JSTED FOR D.F.)	IENT 0.302 0.301 ANCE 188.008 0.249 0.249		
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE	ļ
22	0.00095	0.00006	15.883	
10	-0.02937	0.00484	-6.070	
18	0.03891	0.00414	9.403	
19	0.01716	0.00402	4.272	
4	0.01008	0.00155	6.505	
3	-0.01363	0.00278	-4.903	
	0.02715	0.00547	4.967	
· · · ·	-0.01254	0.00336	-3.735	
17	0.00047	0 00007	n r/7	

SUM OF SQUAP PROPORTION P	RES REDUCED IN THIS REDUCED IN THIS STE	STEP	0.732 0.001		
CUMULATIVE S CUMULATIVE F	SUM OF SQUARES REDU PROPORTION REDUCED.	ICED	105.827 0.092	OF	1154.135
FOR 10 VARIA MULTIPLE ((ADJU F-VALUE FO STANDARD E (ADJU	BLES ENTERED CORRELATION COEFFIC ISTED FOR D.F.) OR ANALYSIS OF VARI RROR OF ESTIMATE ISTED FOR D.F.)	IENT ANCE 1	0.303 0.302 70.496 0.249 0.249		
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROF REG. COEF	OF (COMPUTE	D
22 10 18 19 4 3 11 9 17 1 1 INTERCEF	0.00095 -0.02703 0.03969 0.01482 0.00968 -0.01267 0.02804 -0.01237 0.00345 0.01374 21 0.06761	0.00006 0.00488 0.00414 0.00407 0.00155 0.00279 0.00547 0.00336 0.00097 0.00400	-	15.945 -5.533 9.551 3.639 6.232 -4.536 5.125 -3.686 3.549 3.435	

TABLE D-1. (cont.)

STEP 11

VARIABLE ENTERED..... 25

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	0.635 0.001	
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	106.462 0.092	0F

1154.135

FOR 11 VARIABLES ENTEREDMULTIPLE CORRELATION COEFFICIENT...0.304(ADJUSTED FOR D.F.)....0.303F-VALUE FOR ANALYSIS OF VARIANCE...156.011STANDARD ERROR OF ESTIMATE.....0.249(ADJUSTED FOR D.F.)....0.249

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE
22	0.00116	0.00009	13.150
10	-0.02665	0.00488	-5.456
18	0.03724	0.00421	8.841
19	0.01301	0.00411	3.164
4	0.00981	0.00155	6.315
3	-0.01270	0.00279	-4.547
11	0.02700	0.00548	4.928
9	-0.01225	0.00335	-3.652
17	0.00359	0.00097	3.692
1	0.01410	0.00400	3.525
25	-0.00058	0.00018	-3.199
INTERCEPT	0.06952		

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VARIABLE ENTERED.....21

SUM OF SQUARES REDUCED IN THIS STEP	0.390
PROPORTION REDUCED IN THIS STEP	0.000
CUMULATIVE SUM OF SQUARES REDUCED	106.853
CUMULATIVE PROPORTION REDUCED	0.093

0.093 OF 1154.135

FOR 12 VARIABLES ENTEREDMULTIPLE CORRELATION COEFFICIENT...0.304(ADJUSTED FOR D.F.)....0.303F-VALUE FOR ANALYSIS OF VARIANCE...143.580OTANDARD ERROR OF ESTIMATE.....0.249(ADJUSTED FOR D.F.)....0.249

VARIABLE	REGRESSION	STD. ERROR OF	COMPUTED
NUMBER	COEFFICIENT	REG. COEFF.	T-VALUE
22	0.00116	0.00009	13.193
10	-0.02510	0.00492	-5.099
18	0.03692	0.00421	8.764
19	0.01638	0.00432	3.788
4	0.01011	0.00156	6.492
3	-0.01357	0.00281	-4.824
רו	0.02646	0.00548	4.827
9 .	-0.01246	0.00336	-3.714
17	0.00360	0.00097	3.705
ן	0.01354	0.00401	3.380
25	-0.00057	0.00018	-3.102
21	-0.00001	0.00000	-2.509
INTERCEPT	0.06808		

VARIABLE ENTERED.....20

SUM OF SQU/	ARES REDUCED IN THIS STEP	0.939
PROPORTION	REDUCED IN THIS STEP	0.001
CUMULATIVE	SUM OF SQUARES REDUCED	107.792
CUMULATIVE	PROPORTION REDUCED	0.093

0.093 OF 1154.135

FOR 13 VARIABLES ENTEREDMULTIPLE CORRELATION COEFFICIENT...0.306(ADJUSTED FOR D.F.)....0.305F-VALUE FOR ANALYSIS OF VARIANCE...133.812STANDARD ERROR OF ESTIMATE.....0.249(ADJUSTED FOR D.F.)....0.249

VARIABLE	REGRESSION	STD. ERROR OF	COMPUTED
NUMBER	COEFFICIENT	REG. COEFF.	T-VALUE
22	0.00115	0.00009	13.078
10	-0.02615	0.00493	-5.307
18	0.03682	0.00421	8.742
19	0.01469	0.00434	3.380
4	0.00996	0.00156	6.392
3	-0.01679	0.00293	-5.728
11	0.02686	0.00548	4.901
9 ·	-0.01248	0.00335	-3.723
17	0.00366	0.00097	3.767
T	0.00809	0.00424	1.906
25	-0.00056	0.00018	-3.047
21	-0.00003	0.00001	-4.549
20	0.00020	0.00005	3.893
INTERCEPT	0.10053		

VARIABLE ENTERED.....2

SUM OF SQUARES REDUCED IN THIS STEP	0.428
PROPORTION REDUCED IN THIS STEP	0.000
CUMULATIVE SUM OF SQUARES REDUCED	108.220
CUMULATIVE PROPORTION REDUCED	0.094

OF 1154.135

FOR 14 VARIABLES ENTERED	
MULTIPLE CORRELATION COEFFICIENT	0.306
(ADJUSTED FOR D.F.)	0.305
F-VALUE FOR ANALYSIS OF VARIANCE	124.791
STANDARD ERROR OF ESTIMATE	0.249
(ADJUSTED FOR D.F.)	0.249

REGRESSION	STD. ERROR OF	COMPUTED
COEFFICIENT	REG. COEFF.	T-VALUE
0.00113	0.00009	12.803
-0.02476	0.00496	-4.996
0.03742	0.00422	8.873
0.01393	0.00435	3.200
0.01006	0.00156	6.455
-0.02026	0.00321	-6.304
0.02681	0.00548	4.892
-0.01244	0.00335	-3.709
0.00361	0.00097	3.712
0.01118	0.00440	2.539
-0.00052	0.00018	-2.868
-0.00002	0.00001	-3.338
0.00036	0.00008	4.525
-0.00001	0.00000	-2.630
0.11410		
	REGRESSION <u>COEFFICIENT</u> 0.00113 -0.02476 0.03742 0.01393 0.01006 -0.02026 0.02681 -0.01244 0.00361 0.01118 -0.00052 -0.00002 0.00036 -0.00001 0.11410	REGRESSION STD. ERROR OF COEFFICIENT REG. COEFF. 0.00113 0.00009 -0.02476 0.00496 0.03742 0.00422 0.01393 0.00435 0.01006 0.00156 -0.02026 0.00321 0.02681 0.00548 -0.01244 0.00355 0.00361 0.00097 0.01118 0.00440 -0.00052 0.00018 -0.00036 0.00008 -0.0001 0.00000

STEP 15

VARIABLE ENTERED.....12

SUM OF SQUA	RES REDUCED IN THIS STEP	0.311	
PROPORTION	REDUCED IN THIS STEP	0.000	
CUMULATIVE	SUM OF SQUARES REDUCED	108.531	0F
CUMULATIVE	PROPORTION REDUCED	0.094	

FOR 15 VARIABLES ENTERED

MULTIPLE CORRELATION COEFFICIENT	0.307
(ADJUSTED FOR D.F.)	0.305
F-VALUE FOR ANALYSIS OF VARIANCE	116.834
STANDARD ERROR OF ESTIMATE	0.249
(ADJUSTED FOR D.F.)	0.249

REGRESSION	STD. ERROR OF	COMPUTED
COEFFICIENT	REG. COEFF.	T-VALUE
0.00115	0.00009	12.971
-0.02574	0.00497	-5.175
0.03148	0.00498	6.321
0.01491	0.00437	3.408
0.01090	0.00160	6.801
-0.01998	0.00322	-6.216
0.02430	0.00559	4.346
-0.01237	0.00335	-3 .6 38
0.00400	0.00099	4.051
0.01142	0.00440	2.593
-0.00059	0.00018	-3.176
-0.00002	0.00001	-2.997
0.00036	0.0008	4.535
-0.00001	0.00000	-2.659
0.00035	0.00016	2.241
0.10450		
	REGRESSION COEFFICIENT 0.00115 -0.02574 0.03148 0.01491 0.01090 -0.01998 0.02430 -0.01237 0.00400 0.01142 -0.00059 -0.00059 -0.00002 0.00036 -0.00001 0.00035 0.10450	REGRESSION STD. ERROR OF COEFFICIENT REG. COEFF. 0.00115 0.00009 -0.02574 0.00497 0.03148 0.00498 0.01491 0.00437 0.01090 0.00160 -0.02430 0.00559 -0.01237 0.00335 0.00400 0.00099 0.01142 0.00440 -0.00059 0.00018 -0.00036 0.00008 -0.0001 0.00000 0.00035 0.00016

D-15

STEP 16

VARIABLE ENTERED.....7

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	0.379 0.000	
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	108.910 0.094	OF

FOR 16 VARIABLES ENTERED

MULTIPLE CORRELATION COEFFICIENT	0.307
(ADJUSTED FOR D.F.)	0.306
F-VALUE FOR ANALYSIS OF VARIANCE	109.948
STANDARD ERROR OF ESTIMATE	0.249
(ADJUSTED FOR D.F.)	0.249

VARIABLE	REGRESSION	STD. ERROR OF	COMPUTED
NUMBER	COEFFICIENT	REG. COEFF.	T-VALUE
22	0.00115	0.00009	12.971
10	-0.02332	0.00507	-4.600
18	0.03059	0.00499	6.128
19	0.01302	0.00444	2.932
4	0.01023	0.00163	6.293
3	-0.02079	0.00323	-6.435
11	0.02482	0.00559	4.435
9 [.]	-0.01151	0.00337	-3.415
17	0.00401	0.00099	4.061
1	0.01134	0.00440	2.376
25	-0.00059	0.00018	-3.167
21	-0.00002	0.00001	-3.028
20	0.00037	0.0008	4.598
2	-0.00001	0.00000	-2.683
12	0.00040	0.00016	2.547
7	0.01151	0.00465	2.474
INTERCEPT	0.09008		

SUM OF SQUAR PROPORTION R	ES REDUCED IN THIS EDUCED IN THIS STE	STEP P	0.003 0.000	
CUMULATIVE S CUMULATIVE P	UM OF SQUARES REDU ROPORTION REDUCED.	CED	110.339 0.096 OF	1154.135
FOR 27 VARIA MULTIPLE C (ADJU F-VALUE FO STANDARD E (ADJU)	BLES ENTERED ORRELATION COEFFIC STED FOR D.F.) R ANALYSIS OF VARI RROR OF ESTIMATE STED FOR D.F.)	IENT ANCE	0.309 0.307 66.056 0.249 0.249	
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE	
22 10 18 19 4 3 11 25 21 20 2 12 7 8 26 24 14 16 6 13 15 23 27 5 INTERCEPT	0.00133 -0.2361 0.03186 0.01049 0.00981 -0.02038 0.02684 -0.1117 0.00428 0.01030 -0.00076 -0.00002 0.00036 -0.00001 0.00031 0.02157 -0.01723 0.00026 -0.00011 -0.00280 0.00252 -0.00432 0.00252 -0.00432 0.00252 -0.00432 0.00017 -0.00100 0.00000 -0.00000 -0.00138 0.11453	0.00013 0.00508 0.00590 0.00454 0.00197 0.00326 0.00580 0.00338 0.00147 0.00442 0.00025 0.00001 0.00008 0.00000 0.00597 0.00568 0.00016 0.00018 0.00597 0.00568 0.00016 0.000120 0.00120 0.00119 0.00425 0.00133 0.00000 0.00000 0.00000 0.00000	$10.163 \\ -4.649 \\ 5.397 \\ 2.310 \\ 4.976 \\ -6.243 \\ 4.629 \\ -3.308 \\ 2.906 \\ 2.330 \\ -3.044 \\ -3.114 \\ 4.540 \\ -2.696 \\ 1.726 \\ 3.615 \\ -3.031 \\ 1.634 \\ -1.415 \\ -2.341 \\ 2.110 \\ -1.016 \\ 0.916 \\ -0.754 \\ 0.729 \\ -0.662 \\ -0.229 \\ \end{array}$	

TABLE D-2. VARIABLES USED FOR VOLUME LINEAR REGRESSION MODEL

STEP-WISE MULTIPLE REGRESSION.....VOLUME

NUMBER	0F	OBSERVATIONS	22173
NUMBER	0F	VARIABLES	6
NUMBER	0F	SELECTIONS	1

CONSTANT TO LIMIT VARIABLES 0.00000

VARIABLE NO.	MEAN	STANDARD DEVIATION	VARIABLE
1	1.39356	1.08509	LOG C * LOG T
2	5.45718	3.54759	LOG C ** 2
3	0.80161	0.84164	LOG C * LOG DT
4	0.37995	0.36622	LOG DT
5	0.37084	0.38105	LOG N
6	0.08952	0.28550	ACC = 1 NOACC = 0

CORI	RELATION	MATRIX				
ROW	1 1.00000	0.36503	0.78005	0.64275	0.75581	0.26666
ROW	2 0.36503	1.00000	0.11117	-0.15990	-0.05322	0.20574
ROW	3 0.78005	0.11117	1.00000	0.90861	0.66303	0.21914
ROW	4 0.64275	-0.15990	0.90861	1.00000	0.72059	0.13931
ROW	5 0.75581	-0.05322	0.66303	0.72059	1.00000	0.16251
ROW	6 0.26666	0.20574	0.21914	0.13931	0.16251	1.00000

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TABLE D-3. LINEAR REGRESSIONS COMBINING NON-YOLUME VARIABLES WITH BEST LINEAR VOLUME MODEL

VARIABLE	DESCRIPTION
3	MAIN TRACKS
4	FC
5	NRDY XING HWY
7	AADT/(LANES + 1)
8	VOL
9	VOL*POP
10	VOL*NRBY XING HWY
20	LOG (C + 1)

where VOL is the volume equation described in B.4, Appendix B.

STEP 1

SUM OF SQUARES	REDUCED IN THIS S DUCED IN THIS STEP.	TEP	103.501 0.090	
CUMULATIVE SUM CUMULATIVE PRO	1 OF SQUARES REDUCE PPORTION REDUCED	D	103.501 0.090	OF 1154.298
FOR 1 VARIABLE MULTIPLE COR (ADJUST F-VALUE FOR STANDARD ERR (ADJUST	S ENTERED RELATION COEFFICIE ED FOR D.F.) ANALYSIS OF VARIAN COR OF ESTIMATE ED FOR D.F.)	NT 0.299 0.299 CE 1667.365 0.249 0.249		
VARIABLE NUMBER 8 INTERCEPT	REGRESSION COEFFICIENT 1.00033 -0.00004	STD. ERROR OF REG. COEFF. 0.02450	C(T-	OMPUTED -VALUE 40.833

STEP 2	2						
VARIABL	E ENTERE	D10					
SUM OF PROPORT	SQUARES	REDUCED IN TH CED IN THIS S	IS STEP TEP	•••	6.782 0.006		
CUMULAT CUMULAT	IVE SUM	DF SQUARES REI ORTION REDUCEI	DUCED	••	110.282 0.096	0F	1154.298
FOR 2 MULTI F-VAL STAND	VARIABLE PLE CORRI (ADJUSTE UE FOR AI ARD ERRO (ADJUSTE	S ENTERED ELATION COEFFI D FOR D.F.) NALYSIS OF VAF R OF ESTIMATE D FOR D.F.)	ICIENT RIANCE	0.309 0.309 894.025 0.248 0.248			
VARIA NUMB 8 10 INTER	BLE ER CEPT	REGRESSION COEFFICIENT 1.58142 -0.52980 0.00429	STD. REG 0.0	ERROR OF COEFF. D6056 D5052	COMPU T-VALU 26.T -10.48	FED JE T4 36	

VARIABLE ENTE	RED9					
SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS DUCED IN THIS STE	STEP P	2 0	.042		
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDUC OPORTION REDUCED.	CED	112 0	.324 .097	OF	1154.298
FOR 3 VARIAB MULTIPLE CO (ADJUS F-VALUE FOR STANDARD ERI (ADJUS	LES ENTERED RRELATION COEFFIC TED FOR D.F.) ANALYSIS OF VARIA ROR OF ESTIMATE TED FOR D.F.)	IENT 0 0 ANCE 608 0	.312 .312 .204 .248 .248			
VARIABLE <u>NUMBER</u> 8 10 9 INTERCEPT	REGRESSION <u>COEFFICIENT</u> 1.40695 -0.47305 0.05947 0.00632	STD. ERROR <u>REG. COEFI</u> 0.06766 0.05143 0.01033	OF (<u>F.</u>	COMPUTE <u>T-VALUE</u> 20.794 -9.198 5.759	D F 3	

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STEP 4

VARIABLE ENTERED.....3

SUM OF SQUARES	S REDUCED IN THIS DUCED IN THIS STE	STEP P	1.692 0.001	
CUMULATIVE SUN CUMULATIVE PRO	4 OF SQUARES REDU DPORTION REDUCED.	CED	114.016 0.099 OF	1154.298
FOR 4 VARIABL MULTIPLE COF (ADJUST F~VALUE FOR STANDARD ERF (ADJUST	LES ENTERED RRELATION COEFFIC FED FOR D.F.) ANALYSIS OF VARI ROR OF ESTIMATE FED FOR D.F.)	IENT 0.314 0.314 ANCE 463.750 0.248 0.248		
VARIABLE NUMBER 8 10 9 3 INTERCEPT	REGRESSION <u>COEFFICIENT</u> 1.39460 -0.49708 0.06831 0.02557 -0.01548	STD. ERROR OF <u>REG. COEFF.</u> 0.06765 0.05159 0.01045 0.00487	COMPUTED T-VALUE 20.615 -9.635 6.534 5.247	
STEP 5			-	· · · · · · · · · · · · · · · · · · ·
VARIABLE ENTER	RED7			
SUM OF SQUARES	S REDUCED IN THIS DUCED IN THIS STE	STEP P	0.859 0.001	
CUMULATIVE SUN CUMULATIVE PRO	1 OF SQUARES REDU DPORTION REDUCED.	CED	114.875 0.100 OF	1154.298
FOR 5 VARIABI MULTIPLE COP (ADJUST F-VALUE FOR STANDARD ERP (ADJUST	LES ENTERED RRELATION COEFFIC FED FOR D.F.) ANALYSIS OF VARI ROR OF ESTIMATE FED FOR D.F.)	IENT 0.315 0.315 ANCE 374.081 0.248 0.248		
VARIABLE <u>NUMBER</u> 8 10 9 3 7 INTERCEPT	REGRESSION <u>COEFFICIENT</u> 1.42520 -0.50576 0.07479 0.02241 -0.00001 -0.01234	STD. ERROR OF <u>REG. COEFF.</u> 0.06812 0.05163 0.01059 0.00494 0.00000	COMPUTED T-VALUE 20.923 -9.797 7.060 4.532 -3.739	

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D-21

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VARIABLE ENTERED..... 4

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	1.010 0.001	
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	115.885 0.100	OF

1154.298

FOR6VARIABLESENTEREDMULTIPLECORRELATIONCOEFFICIENT...0.317(ADJUSTEDFORD.F.)0.316F-VALUEFORANALYSISOFVARIANCE...STANDARDERROROFESTIMATE.....0.248(ADJUSTEDFORD.F.)0.248

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE
8	1.37206	0.06934	19.789
10	-0.48021	0.05198	-9.238
9	0.07684	0.01060	7.249
3	0.02492	0.00498	5.004
7	-0.00001	0.00000	-4.566
4	-0.01063	0.00262	-4.057
INTERCEPT.	0.04712		

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ST	ΈP	7

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SUM OF SQUARE	S REDUCED IN THIS	STEP	0.878
PROPORTION RE	DUCED IN THIS STE	P	0.001
CUMULATIVE SU	M OF SQUARES REDU	CED	116.763
CUMULATIVE PR	OPORTION REDUCED.		0.101 OF
FOR 7 VARIAB MULTIPLE CO (ADJUS F-VALUE FOR STANDARD ER (ADJUS	LES ENTERED RRELATION COEFFIC TED FOR D.F.) ANALYSIS OF VARI ROR OF ESTIMATE TED FOR D.F.)	IENT 0.3 0.3 ANCE 272.0 0.2 0.2	18 18 55 48 48
VARIABLE	REGRESSION	STD. ERROR OF	COMPUTED
NUMBER	COEFFICIENT	REG. COEFF.	T-VALUE
0	1 25120	0.00052	10 400

8	1.35128	0.06953	19.436
10	-0.47121	0.05202	-9.058
9	0.07713	0.01060	7.279
3	0.02544	0.00498	5.108
7	-0.00001	0.00000	-4.655
4	-0.01086	0.00262	-4.144
5 ·	-0.01257	0.00332	-3.785
INTERCEPT	0.06855		

STED	<u>Q</u>
	0

VARIABLE ENTERED..... 20

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	0.535 0.000	
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	117.298 0.102	0

OF 1154.298

FOR 8 VARIABLES ENTERED

MULTIPLE CORRELATION COEFFICIENT	0.319
(ADJUSTED FOR D.F.)	0.318
F-VALUE FOR ANALYSIS OF VARIANCE	239.248
STANDARD ERROR OF ESTIMATE	0.248
(ADJUSTED FOR D.F.)	0.248

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE
8	1.43432	0.07498	19.130
10	-0.48848	0.05233	-9.334
9	0,07906	0.01061	7.448
3	0.01996	0.00531	3.757
7	-0.00001	0.00000	-3,697
4	-0.01349	0.00277	-4.877
5.	-0.01283	0.00332	-3.861
20	-0.01232	0.00417	-2.955
INTERCEPT	0.11074		

TABLE D-4.LINEAR REGRESSIONS COMBINING NON-VOLUME
VARIABLES WITH BEST LINEAR VOLUME MODEL

(Different Variables from Table D-3)

VARIABLE	DESCRIPTION		
4	FC		
5	NRBY XING HWY		
9	VOL*POP		
10	VOL*NRBY XING HWY		
12	VOL/NRBY XING HWY		
18	C**2/(LANES + 1)		
19	VOL**2		
22	LOG(VOL**2)		

Where VOL is the volume equation described in B.4, Appendix B.

STEP 1

VARIABLE ENTERED.....12

SUM OF SQUARES R PROPORTION REDUC	EDUCED IN THIS S ED IN THIS STEP.	TEP	110.050 0.095		
CUMULATIVE SUM O CUMULATIVE PROPO	F SQUARES REDUCE RTION REDUCED	D	110.050 0.095	OF 1154.2	98
FOR 1 VARIABLES MULTIPLE CORRE (ADJUSTED F-VALUE FOR AN STANDARD ERROR (ADJUSTED	ENTERED LATION COEFFICIE FOR D.F.) ALYSIS OF VARIAN OF ESTIMATE FOR D.F.)	NT 0 0 CE 1783 0 0	.309 .309 .992 .248 .248		
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROF REG. COEF	R OF FF.	COMPUTED T-VALUE	
12 INTERCEPT	1.36694 0.00318	0.03236		42.237	

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2	T	E	P	2
-		-		_

VARIABLE ENTERED..... 9

SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS DUCED IN THIS STE	STEP	2.083 0.002	
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDU OPORTION REDUCED.	CED	112.134 0.097 OF	1154.298
FOR 2 VARIAB MULTIPLE CO (ADJUS F-VALUE FOR STANDARD ER (ADJUS	LES ENTERED RRELATION COEFFIC TED FOR D.F.) ANALYSIS OF VARI ROR OF ESTIMATE TED FOR D.F.)	IENT 0.3 0.3 ANCE 910.6 0.2	312 312 545 248 248	
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR O REG. COEFF.	F COMPUTED)
12 . 9 INTERCEPT	1.21243 0.06006 0.00531	0.04184 0.01032	28.976 5.817	

STEP 3

18

INTERCEPT

-0.00151

0.01248

VARIABLE ENTERED..... 18 SUM OF SQUARES REDUCED IN THIS STEP.... 1.840 PROPORTION REDUCED IN THIS STEP..... 0.002 CUMULATIVE SUM OF SQUARES REDUCED..... 113.973 CUMULATIVE PROPORTION REDUCED..... 0.099 OF 1154.298 FOR 3 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT... 0.314 (ADJUSTED FOR D.F.).... 0.314 F-VALUE FOR ANALYSIS OF VARIANCE... 618.112 STANDARD ERROR OF ESTIMATE..... 0.248 (ADJUSTED FOR D.F.)..... 0.248 VARIABLE REGRESSION STD. ERROR OF COMPUTED NUMBER COEFFICIENT REG. COEFF. T-VALUE 12 1.29576 29.121 0.04450 9 0.07423 0.01064 6.979

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0.00028

-5.471

VARIABLE ENTERED.....4

SUM OF SQUARES	S REDUCED IN THIS DUCED IN THIS STEP	STEP		1.663 0.001		
CUMULATIVE SUN CUMULATIVE PRO	1 OF SQUARES REDUC DPORTION REDUCED.	CED	1.	15.336 0.100	OF	1154.298
FOR 4 VARIABI MULTIPLE CON (ADJUST F-VALUE FOR STANDARD ERF (ADJUST	LES ENTERED RRELATION COEFFIC FED FOR D.F.) ANALYSIS OF VARIA ROR OF ESTIMATE FED FOR D.F.)	IENT	0.316 0.316 469.714 0.248 0.248			
VARIABLE NUMBER 12 9 18 4 INTERCEPT	REGRESSION COEFFICIENT 1.28031 0.07814 -0.00203 -0.01273 0.08897	STD. ERRC <u>REG. COE</u> 0.044 0.010 0.000 0.000	IR OF IFF. 59 66 30 70	COMPL <u>T-VAL</u> 28.7 7.3 -6.8 -4.7	ITED UE 14 29 35 11	

STEP 5

SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS DUCED IN THIS STE	STEP P	0.913 0.001		
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDU OPORTION REDUCED.	CED	116.249 0.101	OF	1154.298
FOR 5 VARIAB MULTIPLE CO (ADJUS) F-VALUE FOR STANDARD ER (ADJUS)	LES ENTERED RRELATION COEFFIC TED FOR D.F.) ANALYSIS OF VARI ROR OF ESTIMATE TED FOR D.F.)	IENT ANCE 37	0.317 0.317 9.056 0.248 0.248		
VARIABLE <u>NUMBER</u> 12 9 18 4 22 INTERCEPT	REGRESSION <u>COEFFICIENT</u> 1.00113 0.06582 -0.00197 -0.01325 0.73310 0.47890	STD. ERROF <u>REG. COEF</u> 0.08499 0.01112 0.00030 0.00271 0.19002	? OF C F <u>F. T</u>	OMPUTED -VALUE 11.780 5.917 -6.616 -4.900 3.858	

1154.298

STEP 6

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VARIABLE ENTERED..... 19

SUM OF SQUARES	5 REDUCED IN THIS	S STEP	0.968
	DUCED IN THIS STE	2P	0.001
CUMULATIVE SUN	I OF SQUARES REDU	JCED	117.217
CUMULATIVE PRO	PORTION REDUCED.		0.102 OF
FOR 6 VARIABL MULTIPLE COF (ADJUST F-VALUE FOR STANDARD ERF (ADJUST	ES ENTERED RELATION COEFFIC ED FOR D.F.) ANALYSIS OF VARI ROR OF ESTIMATE ED FOR D.F.)	CIENT 0.3 0.3 ANCE 318.7 0.2 0.2	19 18 89 48 48
VARIABLE	REGRESSION	STD. ERROR OF	COMPUTED
NUMBER	COEFFICIENT	REG. COEFF.	T-VALUE
12 9 18 4 22 19 INTERCEPT.	0.75094 0.06989 -0.00195 -0.01364 5.57820 -5.48395 3.01490	0.10573 0.01117 0.00030 0.00271 1.23380 1.37983	7.102 6.259 -6.562 -5.040 4.521 -3.974

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VARIABLE ENTERED.....5

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	0.937 0.001	
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	118.154 0.102	0F

1154.298

FOR 7 VARIABLES ENTERED	
MULTIPLE CORRELATION COEFFICIENT	0.320
(ADJUSTED FOR D.F.)	0.319
F-VALUE FOR ANALYSIS OF VARIANCE	275.665
STANDARD ERROR OF ESTIMATE	0.247
(ADJUSTED FOR D.F.)	0.247

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE
12	0.72686	0.10587	6.866
9	0.06971	0.01116	6.245
18	-0.00199	0.00030	-6.692
4	-0.01391	0.00271	-5.141
22	5.60566	1.23330	4.545
19	-5.48273	1.37925	-3.975
5	-0.01299	0.00332	-3.912
INTERCEPT	3.05248		

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STEP 5

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	0.599 0.001		
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	118.753 0.103	OF	1154.298
FOR 8 VARIABLES ENTERED			

MULTIPLE CORRELATION COEFFICIENT	0.321
(ADJUSTED FOR D.F.)	0.320
F-VALUE FOR ANALYSIS OF VARIANCE	242.555
STANDARD ERROR OF ESTIMATE	0.247
(ADJUSTED FOR D.F.)	0.247

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE
12	0.75952	0.10635	7.141
9	0.06678	0.01120	5.963
18	-0.00194	0.00030	-6.525
4	-0.01327	0.00271	-4.892
22	6.80342	1.29107	5.270
19	-6.63985	1.42765	-4.651
5,	-0.01282	0.00332	-3.861
10	-0.11629	0.03718	-3.128
INTERCEPT	3.67821		

TABLE D-5. ILLUSTRATION OF "MIGRATION" AND CONVERGENCE-SUCCESSIVE ITERATIONS

STEP 4

VARIABLE ENTERED..... 4 2.94402 SUM OF SQUARES REDUCED IN THIS STEP.... PROPORTION REDUCED IN THIS STEP..... 0.00032 5204.70984 CUMULATIVE SUM OF SQUARES REDUCED..... 0.57455 OF 9058.696 CUMULATIVE PROPORTION REDUCED..... FOR 4 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT... 0.75799 0.75796 (ADJUSTED FOR D.F.).... 7484.328 F-VALUE FOR ANALYSIS OF VARIANCE... 0.417 STANDARD ERROR OF ESTIMATE..... 0.41699 (ADJUSTED FOR D.F.).... VARIABLE REGRESSION STD. ERROR OF COMPUTED NUMBER COEFFICIENT REG. COEFF. T-VALUE 2 0.09136 5.890 H (OF VIT16) 0.53804E+00 0.09847 -4.718H**2 3 -0.46455E+00 -4.358 1 -0.11535E+00 0.02647 1 -0.12980E+00 0.03154 -4.115 H**3 Δ INTERCEPT -0.00008 TOLERANCE = .29440E+01STEP 4 VARIABLE ENTERED..... 1 4.28890 SUM OF SOUARES REDUCED IN THIS STEP.... 0.00047 PROPORTION REDUCED IN THIS STEP..... 5165.23438 CUMULATIVE SUM OF SQUARES REDUCED..... 0.57147 OF 9038.489 CUMULATIVE PROPORTION REDUCED..... FOR 4 VARIABLES ENTERED 0.75596 MULTIPLE CORRELATION COEFFICIENT... 0.75592 (ADJUSTED FOR D.F.).... 7390.614 F-VALUE FOR ANALYSIS OF VARIANCE... 0.418 STANDARD ERROR OF ESTIMATE..... 0.41803 (ADJUSTED FOR D.F.).... VARIABLE STD. ERROR OF REGRESSION COMPUTED REG. COEFF. NUMBER COEFFICIENT T-VALUE 2 0.08906 4.955 H (OF VIT16)0.44127E+00 -6.158 H**2 3 -0.58231E+00 0.09457 H**3 0.02983 -5.541 4 -0.16530E+00 1 -4.954 1 -0.13027E+00 0.02629 INTERCEPT 0.00005

TOLERANCE = .42889E+01

VARIABLE ENTERED..... 1 SUM OF SQUARES REDUCED IN THIS STEP.... 4.55172 PROPORTION REDUCED IN THIS STEP..... 0.00051 CUMULATIVE SUM OF SQUARES REDUCED..... 5125.97827 CUMULATIVE PROPORTION REDUCED..... 0.56965 OF 8998.48193 FOR 4 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT... 0.75475 0.75471 (ADJUSTED FOR D.F.).... 0.41799 VARIABLE REGRESSION STD. ERROR OF COMPUTED NUMBER COEFFICIENT REG. COEFF. T-VALUE 2 0.08861 4.602 H (OF VIT16) 0.40780E+00 -6.714 H**2 3 -0.62819E+00 0.09356 0.02933 -6.123 H**3 4 -0.17961E+00 -0.13417E+00 -5.105 0.02629 1 1 INTERCEPT 0.00005 TOLERANCE = .45517E+01STEP 4 VARIABLE ENTERED..... 1 SUM OF SQUARES REDUCED IN THIS STEP.... 4.64733 0.00052 PROPORTION REDUCED IN THIS STEP..... 5097.06598 CUMULATIVE SUM OF SQUARES REDUCED..... 0.56859 OF 8964.42395 CUMULATIVE PROPORTION REDUCED..... FOR 4 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT... 0.75405 0.75401 STANDARD ERROR OF ESTIMATE..... 0.41771 STD. ERROR OF VARIABLE REGRESSION COMPUTED REG. COEFF. NUMBER COEFFICIENT T-VALUE 0.08838 4.457 H (OF VIT16) 2 0.39386E+00 -6.968 H**2 0.09313 3 -0.64890E+00 0.02912 -6.389 H**3 -0.18607E+00 4 0.02626 -5.161 1 1 -0.13555E+00 INTERCEPT 0.00004

TOLERANCE = .46473E+01

D-32

TABLE D-5. (cont.)

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STEP 4

VARIABLE ENTERED..... 1

SUM OF SQUAF PROPORTION F	RES REDUCED IN THIS REDUCED IN THIS STE	STEP 4 P 0	.72331 .00053	
CUMULATIVE S CUMULATIVE F	SUM OF SQUARES REDUCED.	CED 5076	5.95288 0.56792 OF	8939.52539
FOR 4 VARIA MULTIPLE ((ADJL F-VALUE FC STANDARD E (ADJL	ABLES ENTERED CORRELATION COEFFIC JSTED FOR D.F.) OR ANALYSIS OF VARIA ERROR OF ESTIMATE JSTED FOR D.F.)	IENT 0.753 0.753 ANCE 7284.387 0.417 0.417	861 857 7 45	
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE	
2 3 4 1 INTERCEPT	0.38510E+00 -0.66143E+00 -0.18983E+00 -0.13666E+00 0.00004	0.08829 0.09299 0.02906 0.02625	4.362 -7.113 -6.533 -5.207	H (OF VIT16) H**2 H**3 1

TOLERANCE = .47233E+01

D-33

VARIABLE ENTERED..... 1

SUM OF SQUAR PROPORTION R	RES REDUCED IN THIS S REDUCED IN THIS STEP	STEP	4.75342 0.00053	
CUMULATIVE S CUMULATIVE P	UM OF SQUARES REDUCT ROPORTION REDUCED	ED	5063.00427 0.56746	OF 8982-2288
FOR 4 VARIA MULTIPLE C (ADJU F-VALUE FO STANDARD E (ADJU	BLES ENTERED ORRELATION COEFFICIE STED FOR D.F.) R ANALYSIS OF VARIAN RROR OF ESTIMATE STED FOR D.F.)	ENT 0.753 0.753 NCE 7270.675 0.417 0.417	30 26 27	
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE	
2 3 4 1 INTERCEPT	0.38069E+00 =0.66800E+00 -0.19171E+00 -0.13711E+00 0.00004	0.08826 0.09295 0.02904 0.02624	4.313 H -7.187 H* -6.602 H* -5.225 1	(OF VIT16) *2 *3

TOLERANCE = .47534E+01

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STEP-WISE MULTIPLE REGRESSION.....XBKSEL

NUMBER	0F	OBSERVATIONS	22173
NUMBER	0F	VARIABLES	19
NUMBER	0F	SELECTIONS	1

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CONSTANT TO LIMIT VARIABLES 0.00000

VARIABLE	MEAN	STANDARD	VARIABLE
NO.		DEVIATION	
1	0.33273	0.33367	LOG T
2	1.05615	0.82668	LOG C
3	0.33657	0.48075	LOG T **2
4	2.85857	3.07002	LOG C **2
5	0.69996	4.06768	T*C/5000
6	0.22695	0.27164	LOG NITE
7	0.19928	0.25083	LOG DAY THRU
8	0.12158	0.22585	LOG SWITCH
9	0.24149	0.23395	U=0 R=1
10	13.71944	12.01191	MAX SPEED
11	0.29779	0.30094	HWY PAVED
12	0.47112	0.94952	POP
13	2.34700	1.20217	FC
14	0.18081	0.31276	NRBY XING HWY
15	0.17040	0.26085	RR ADV WARN
16	0.60526	0.55643	LANES
17	0.41130	0.33122	MAIN TRKS
18	0.42599	0.22083	1
19	-0.00028	0.41708	ACC=1 NOACC=-1

D-35

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VARIABLE ENTERED.....17

SUM OF SQUARES	S REDUCED IN THIS DUCED IN THIS STEP	STEP		4.45096 0.00115		
CUMULATIVE SUN CUMULATIVE PRO	1 OF SQUARES REDUC PORTION REDUCED	ED		4.45096 0.00115	OF	3856.9758
FOR VARIABL MULTIPLE COR (ADJUST F-VALUE FOR STANDARD ERR (ADJUST	ES ENTERED RELATION COEFFICI ED FOR D.F.) ANALYSIS OF VARIA OR OF ESTIMATE ED FOR D.F.)	ENT NCE	0.03397 0.03397 25.615 0.417 0.41685			
VARIABLE NUMBER 17 INTERCEPT	REGRESSION COEFFICIENT 0.26930E-01 -0.01131	STD. E <u>REG.</u> 0.005	RROR OF COEFF. 530	COMPUTED T-VALUE 5.061	MAIN	TRKS

TOLERANCE = .44510E+01

STEP 2

VARIABLE ENTERED.....13

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	32.87232 0.00852		
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	37.32328 0.00968	0F 385	6.9758
FOR VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT (ADJUSTED FOR D.F.) F-VALUE FOR ANALYSIS OF VARIANCE 10 STANDARD ERROR OF ESTIMATE (ADJUSTED FOR D.F.)	0.09837 0.09814 08.316 0.415 0.41509		

VARIABLE <u>NUMBER</u> 17 13	REGRESSION <u>COEFFICIENT</u> 0.17329E+00 -0.32763E-01	STD. ERROR OF <u>REG. COEFF.</u> 0.01184 0.00237	COMPUTED <u>T-VALUE</u> 14.631 -13.813	MAIN FC	TRKS
INTERCEPT	0.00535	0.00237	-12.012	гс	

TOLERANCE = .32872E+02

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TABLE D-6 (cont.)

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VARIABLE ENTERED..... 11

SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS DUCED IN THIS STEP	STEP	11.00600 0.00285	
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDUC OPORTION REDUCED	ED	48.33018 0.01253 OF 3	856.9758
FOR 3 VARIAB MULTIPLE CO (ADJUS F-VALUE FOR STANDARD ER (ADJUS	LES ENTERED RRELATION COEFFICI TED FOR D.F.) ANALYSIS OF VARIAN ROR OF ESTIMATE TED FOR D.F.)	ENT 0. NCE 93. 0.4	11194 11154 7222 414 41451	
VARIABLE NUMBER 17 13 11 INTERCEPT TOLERANCE	REGRESSION <u>COEFFICIENT</u> 0.17078E+00 -0.45982E-01 0.10009E+00 0.00759 = .11007E+02	STD. ERROR REG. COEFF 0.01183 0.00289 0.01250	OF COMPUT <u>. T-VALL</u> 14.43 -15.92 8.00	TED J <u>E</u> 5 MAIN TRKS 5 FC 4 HWY PAVED
STEP 4				
VARIABLE ENTER	RED 8			
SUM OF SQUARES	S REDUCED IN THIS S DUCED IN THIS STEP	STEP	3.40733 0.00088	

CUMULATIVE	SUM OF	SOUARES	REDUCED	51.73750		
CUMULATIVE	PROPORT	TIÔN REDU	JCED	0.01341	of	3856.9758

FOR 4 VARIABLES ENTERED

OR 4 VARIABLES ENTERED	
MULTIPLE CORRELATION COEFFICIENT	0.11528
(ADJUSTED FOR D.F.)	0.11524
F-VALUE FOR ANALYSIS OF VARIANCE	75.351
STANDARD FRROR OF ESTIMATE	0.414
(ADJUSTED FOR D.F.)	0.41434
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VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE	
17	0.16764E+00	0.01185	14.150	MAIN TRKS
13	0.11275E+00	0.01282	-14.492 8.796	FC HWY PAVED
8 INTERCEPT	-0.65676E-01 0.00602	0.01474	-4.455	LOG SWITCH

TOLERANCE = .34073E+01

VARIABLE ENTERED..... 12

SUM OF SQUARES REDUCED IN THIS STEP	6.17361
PROPORTION REDUCED IN THIS STEP	0.00160
CUMULATIVE SUM OF SQUARES REDUCED	57.91112
CUMULATIVE PROPORTION REDUCED	0.01501 OF 3856.9758

FOR 5 VARIABLES ENTERED

STANDARD ERROR OF ESTIMATE

VARIABLE	REGRESSION	STD. ERROR OF	COMPUTED	
NUMBER	COEFFICIENT	REG. COEFF.	T-VALUE	
17 13 11 8 12 INTERCEPT	0.17908E+00 -0.45240E-01 0.89396E-01 -0.96059E-01 0.22343E-01 0.00678	0.01199 0.00299 0.01339 0.01557 0.00372	14.935 -15.148 6.678 -6.168 6.002	MAIN TRKS FC HWY PAVED LOG SWITCH POP

TOLERANCE = .61736E+01

VARIABLE ENTERED..... 4

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	4.30935 0.00112	
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	62.22067 0.01613 OF 3	856.9758
FOR 6 VARIABLES ENTERED		

MULTIPLE CORRELATION COEFFICIENT	0.12701
(ADJUSTED FOR D.F.)	0.12614
F-VALUE FOR ANALYSIS OF VARIANCE	60.574
STANDARD ERROR OF ESTIMATE	0.414
(ADJUSTED FOR D.F.)	0.41381

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE	-
17	0.17969E+00	0.01198	14.993	MAIN TRKS
13	-0.41808E-01	0.00306	-13.653	FC
11	0.12989E+00	0.01562	8.313	HWY PAVED
. 8	-0.87759E-01	0.01565	-5,606	LOG SWITCH
12	0.31880E-01	0.00411	7.566	POP
4	-0.87416E-02	0.00174	-5.017	LOG C **2
INTERCEPT	0.00628			

TOLERANCE = .43096E+01

D-39

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VARIABLE ENTERED..... 2

SUM OF SQUARES REDUCED IN THIS STEP	7.63773
PROPORTION REDUCED IN THIS STEP	0.00198
CUMULATIVE SUM OF SQUARES REDUCED	69.85840
CUMULATIVE PROPORTION REDUCED	0.01811 OF 3856.9758

FOR 7 VARIABLES ENTERED

(ADJUSTED FUR D.F.)	MULTIPLE CORRELATION COEFFICIENT (ADJUSTED FOR D.F.) F-VALUE FOR ANALYSIS OF VARIANCE STANDARD ERROR OF ESTIMATE (ADJUSTED FOR D.F.)	0.13458 0.13359 58.409 0.413 0.41341
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VARIABLE NUMBER	REGRESSION	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE	
17 13 11 8 12 4 2 INTERCEPT	0.15879E+00 -0.79423E-01 0.94676E-01 -0.89715E-01 0.35015E-01 -0.57484E-01 0.23373E+00 0.00451	0.01237 0.00640 0.01647 0.01564 0.00415 0.00750 0.03496	12.832 -12.402 5.747 -5.736 8.445 -7.669 6.686	MAIN TRKS FC HWY PAVED LOG SWITCH POP LOG C **2 LOG C

TOLERANCE = .76377E+01

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VARIABLE ENTERED..... 16

SUM OF SQUARES REDUCED IN THIS STEP	3.50487
PROPORTION REDUCED IN THIS STEP	0.00091
CUMULATIVE SUM OF SQUARES REDUCED	73.36326
CUMULATIVE PROPORTION REDUCED	0.01902 OF 3856.9758

FOR 8 VARIABLES ENTERED	
MULTIPLE CORRELATION COEFFICIENT	0.13792
(ADJUSTED FOR D.F.)	0.13679
F-VALUE FOR ANALYSIS OF VARIANCE	53.719
STANDARD ERROR OF ESTIMATE	0.413
(ADJUSTED FOR D.F.)	0.41324

VARIABLE NUMBER	REGRESSION	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE	
17 13 11 8 12 4 2 16 INTERCEPT	0.16021E+00 -0.81553E-01 0.89155E-01 -0.98097E-01 0.34436E-01 -0.53764E-01 0.19454E+00 0.47427E-01 0.00330	0.01237 0.00642 0.01651 0.01563 0.00415 0.00754 0.03600 0.01047	12.949 -12.706 5.400 -5.763 8.306 -7.134 5.404 4.531	MAIN TRKS FC HWY PAVED LOG SWITCH POP LOG C **2 LOG C LANES

TOLERANCE = .35049E+01

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VARIABLE ENTERED..... 9

SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS DUCED IN THIS STEP	STEP	1. 0.	17265 00030	v	
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDUC OPORTION REDUCED	ED	74. 0.	53591 01932	0 F [°]	3856.9758
FOR ⁹ VARIAB MULTIPLE CO (ADJUS F-VALUE FOR STANDARD ER (ADJUS	LES ENTERED RRELATION COEFFICI TED FOR D.F.) ANALYSIS OF VARIA ROR OF ESTIMATE TED FOR D.F.)	ENT NCE	0.13901 0.13774 48.527 0.413 0.41319			
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. EF REG. (ROR OF	COMP T-VA	UTED	
17 13 11 8 12 4 2 16 9 INTERCEPT	0.16275E+00 -0.81143E-01 0.84513E-01 -0.94786E-01 0.28006E-01 -0.58471E-01 0.21748E+00 0.49469E-01 -0.48751E-01 0.00370	0. 0. 0. 0. 0. 0. 0.	01241 00642 01660 01573 00482 00775 03704 01049 01555	13 -12 -6 5 -7 5 4 -2	.115 .640 .090 .024 .814 .548 .871 .714 .621	MAIN TRKS FC HWY PAVED LOG SWITCH POP LOG C **2 LOG C LANES U=0 R=1

TOLERANCE = .11726E+01

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STEP 16

VARIABLE ENTERED..... 10

SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS S DUCED IN THIS STEP.	STEP	1.37028 0.00036	
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDUCE	ED	75.90619 0.01968 C	DF 3856.9758
FOR 10 VARIAB MULTIPLE CO (ADJUS F-VALUE FOR STANDARD ER (ADJUS	LES ENTERED RRELATION COEFFICIE TED FOR D.F.) ANALYSIS OF VARIAN ROR OF ESTIMATE TED FOR D.F.)	ENT 0.1 0.1 ICE 44.4 0.4 0.4	4029 3886 91 13 1313	
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR O REG. COEFF.	F COMPUTE T-VALUE	D _
17 13 11 8 12 4 2 16 9 10 INTERCEPT	0.14735E+00 -0.83219E-01 0.89239E-01 -0.90911E-01 0.29163E-01 -0.55222E-01 0.20528E+00 0.48786E-01 -0.48807E-01 0.12786E-01 0.00461	0.01354 0.00646 0.01669 0.01579 0.00483 0.00782 0.03728 0.03728 0.01550 0.01580	10.879 -12.882 5.348 -5.757 6.037 -7.104 5.506 4.648 -3.083	MAIN TRKS FC HWY PAVED LOG SWITCH POP LOG C **2 LOG C LANES U=0 R-1 MAX SPEED

TOLERANCE = .13703E+01

D-43

TABLE D-7. SELECTION REGRESSION -- FLASHING LIGHTS

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STEP-WISE MULTIPLE REGRESSION.....FLGSEL

NUMBER	0F	OBSERVATIONS	14576
NUMBER	0F	VARIABLES	20
NUMBER	0F	SELECTIONS	ſ

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CONSTANT TO LIMIT VARIABLES 0.00000

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VARIABLE NUMBER	MEAN	STANDARD DEVIATION	VARIABLE
ſ	0.43189	0.35112	LOG T
2	1.49170	0.87041	LOG C
3	0.50061	0.58055	LOG T **2
4	5.21477	3.79962	LOG C **2
5 .	5.62443	17.84223	T*C/5000
6	0.29310	0.29438	LOG NITE
7	0.25977	0.26758	LOG DAY THRU
8	0.19191	0.27771	LOG SWITCH
9	0.14859	0.19996	U=0 R=1
10	13.93295	10.88430	MAX SPEED
נו	0.43079	0.20659	HWY PAVED
12	0.91970	1.08612	POP
13	1.97984	0.96749	FC
14	0.22522	0.31598	NRBY XING HWY
15	0.26576	0.25872	RR ADV WARN
16	3.69170	3.46088	TRUCKS
17	1.06265	0.77095	LANES
18	0.48267	0.36660	MAIN TRKS
19	0.44061	0.19437	1
20	0.00022	0.41248	ACC=1 NOACC=-1

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STEP 1

VARIABLE ENTERED.....18

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	1.42707 0.00058
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	1.42707 0.00058 OF 2479.802
FOR 1VARIABLES ENTEREDMULTIPLE CORRELATION COEFFICIENT0.02(ADJUSTED FOR D.F.)0.02F-VALUE FOR ANALYSIS OF VARIANCE8.39STANDARD ERROR OF ESTIMATE0.41(ADJUSTED FOR D.F.)0.41	399 399 2 2 238
VARIABLE NUMBERREGRESSION COEFFICIENTSTD. ERROR OF REG. COEFF.180.16325E-01 -0.007660.00564	COMPUTED T-VALUE 2.897 MAIN TRKS
TOLERANCE = $.14271E+01$	
STEP 2	
VARIABLE ENTERED13	
SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	7.17642 0.00289
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	8.60349 0.00347 OF 2479.802
FOR 2 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT0.058 	890 832 8 2 181
VARIABLE REGRESSION STD. ERROR NUMBER COEFFICIENT REG. COEFF. 18 0.89588E-01 0.01259 13 -0.22527E-01 0.00346 INTERCEPT 0.00158 0.00346	OF COMPUTED <u>T-VALUE</u> 7.116 MAIN TRKS -6.505 FC

TOLERANCE = .71764E+01

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VARIABLE ENTERED 3		
SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	2.09810 0.00085	
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	10.70159 0.00432	OF 2479.802
FOR 3 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT (ADJUSTED FOR D.F.) F-VALUE FOR ANALYSIS OF VARIANCE STANDARD ERROR OF ESTIMATE (ADJUSTED FOR D.F.)	0.06569 0.06464 21.053 0.412 0.41166	
VARIABLE REGRESSION STD. ERR NUMBER COEFFICIENT REG. COM 18 0.11192E+00 0.0140 13 -0.18419E-01 0.0036 3 -0.32696E-01 0.0092 INTERCEPT -0.00096 -0.00096	DR OF COMPUTER EFF. T-VALUE 19 7.941 15 -5.042 29 -3.519) - MAIN TRKS FC LOG T **2
TOLERANCE = $.20981E+01$		
STEP 4		· <u> </u>
VARIABLE ENTERED 12		
SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	1.84598 0.00074	
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	12.54757 0.00506 OF	2479.802
FOR 4 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT 0. (ADJUSTED FOR D.F.) 0. F-VALUE FOR ANALYSIS OF VARIANCE 18. STANDARD ERROR OF ESTIMATE 0. (ADJUSTED FOR D.F.) 0.	07113 06968 526 411 41154	
VARIABLEREGRESSIONSTD. ERRONUMBERCOEFFICIENTREG. COE180.10825E+000.0141	OR OF COMPUTED <u>SFF.</u> <u>T-VALUE</u> 3 7.660	MAIN TRKS

13	-0.23011E-01
3	-0.35850E-01
12	0.12776E-01
INTERCEPT	-0.00027

TOLERANCE = .18460E+01

D-46

0.00391 0.00934

0.00387

-5.889 -3.838

3.302

FC

POP

LOG T **2

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STEP 5

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VARIABLE ENTERED..... 4

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	2.12715 0.00086		
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	14.67472 0.00592	0F	2479.802
FOR 5 VARIABLES ENTERED			

MULTIPLE CORRELATION COEFFICIENT	0.07693
(ADJUSTED FOR D.F.)	0.07513
F-VALUE FOR ANALYSIS OF VARIANCE	17.347
STANDARD ERROR OF ESTIMATE	0.411
(ADJUSTED FOR D.F.)	0.41139

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED	
18 13 3 12 4	0,11973E+00 -0,16611E-01 -0.37212E-01 0.21117E-01 -0,49930E-02	0.01449 0.00430 0.00934 0.00453 0.00141	8.261 -3.860 -3.983 4.665 -3.546	MAIN TRKS FC LOG T **2 POP LOG C **2
INTERCEPT	0.00056			

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TOLERANCE = .21271E+01

VARIABLE ENTERED..... 17

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	5.05679 0.00204		
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	19.73151 0.00796	OF	2479.802

FOR 6 VARIABLES ENTERED	
MULTIPLE CORRELATION COEFFICIENT	0.08920
F-VALUE FOR ANALYSIS OF VARIANCE	19.476
(ADJUSTED FOR D.F.)	0.411 0.41099

VARIABLE	REGRESSION	STD. ERROR OF	COMPUTED	
NUMBER	COEFFICIENT	REG. COEFF.	T-VALUE	
18	0.12161E+00	0.01448	8.397	MAIN TRKS
13	-0.19399E-01	0.00433	-4.482	FC
3	-0.37895E-01	0.00933	-4.060	LOG T **2
12	0.19636E-01	0.00453	4.334	POP
4	-0.12266E-01	0.00194	-6.338	LOG C **2
17	0.43298E-01	0.00791	5.472	LANES
INTERCEPT.	-0.00120			

TOLERANCE = .50568E+01

VARIABLE ENTERED..... 7

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	1.87448 0.00076	
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	21.60598 0.00871 OF	2479.802
FOR 7 VARIABLES ENTERED		

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MULTIPLE CORRELATION COEFFICIENT	D.09334
(ADJUSTED FOR D.F.)	0.09113
F-VALUE FOR ANALYSIS OF VARIANCE	18.292
STANDARD ERROR OF ESTIMATE	0.411
(ADJUSTED FOR D.F.)	0.41086
•	

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED	
18 13 3 12 4 17 7 INTERCEPT	0.10540E+00 -0.22089E-01 -0.62118E-01 0.22907E-01 -0.12454E-01 0.44678E-01 0.83675E-01 -0.00116	0.01527 0.00440 0.01183 0.00463 0.00194 0.00792 0.02511	6.901 -5.018 -5.252 4.943 -6.435 5.641 3.333	MAIN TRKS FC LOG T **2 POP LOG C **2 LANES LOG DAY THRU

TOLERANCE = .1845E+01

TABLE D-8. SELECTION REGRESSION -- AUTOMATIC GATES

STEP - WISE MULTIPLE REGRESSION.....GATSEL

NUMBER	0F	OBSERVATIONS	3888
NUMBER	0F	VARIABLES	20
NUMBER	0 F	SELECTIONS	1

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CONSTANT TO LIMIT VARIABLES 0.00000

VARIABLE	MEAN	STANDARD	VARIABLE
NO.		DEVIATION	
1	0.62589	0.35024	LOG T
2	1.59565	0.69239	LOG C
3	0.98100	0.71782	LOG T **2
4	5.61657	3.20156	LOG C **2
5	14.80179	35.31585	T*C/500
6	0.46913	0.29430	LOG NITE
7	0.44779	0.31509	LOG DAY THRU
8	0.24501	0.27815	LOG SWITCH
9	0.11621	0.19108	U=0 R=1
10	20.61505	12.61789	MAX SPEED
11	0.46242	0.15788	HWY PAVED
12	1.22522	1.01110	POP
13	2.16860	0.80058	FC
14	0.28519	0.32094	NRBY XING HWY
15	0.29479	0.25227	RR ADV WARN
16	4.31561	3.60533	TRUCKS
17	1.17982	0.67308	LANES
18	0.73989	0.47198	MAIN TRKS
19	0.46930	0.14133	1
20	-0.00118	0.41014	ACC=1 NOACC=-1

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VARIABLE ENTERED..... 9

SUM OF SQUARES REDUCED IN THIS S PROPORTION REDUCED IN THIS STEP.	TEP 1 0	.21426 .00186	
CUMULATIVE SUM OF SQUARES REDUCE CUMULATIVE PROPORTION REDUCED	D 1 0	.21426 .00186 OF	653.87129
FOR 1 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIE (ADJUSTED FOR D.F.) F-VALUE FOR ANALYSIS OF VARIAN STANDARD ERROR OF ESTIMATE (ADJUSTED FOR D.F.)	NT 0.043 0.043 CE 7.230 0.410 0.409	309 309 0 982	
VARIABLE REGRESSION NUMBER COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE	
9 -0.79029E-01 0.00000 INTERCEPT	0.02939	-2.689	U=0 R-1
TOLERANCE = .12143E+01			
STEP 2			
VARIABLE ENTERED 18			
SUM OF SQUARES REDUCED IN THIS STPROPORTION REDUCED IN THIS STEP.	TEP 1. 0.	.35991 .00208	
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	D 2. 0.	.57417 .00394 OF	653.87129
FOR 2 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIEN	NT 0.062	274	
(ADJUSTED FOR D.F.) F-VALUE FOR ANALYSIS OF VARIAN STANDARD ERROR OF ESTIMATE (ADJUSTED FOR D.F.)	0.060 CE 7.677 0.409 0.409)67))50	
(ADJUSTED FOR D.F.) F-VALUE FOR ANALYSIS OF VARIAND STANDARD ERROR OF ESTIMATE (ADJUSTED FOR D.F.) VARIABLE REGRESSION <u>NUMBER COEFFICIENT</u>	0.060 CE 7.677 0.409 0.409 0.409 STD. ERROR OF <u>REG. COEFF.</u>	067 950 COMPUTED T-VALUE	

TOLERANCE = $.13599E \div 01$

VARIABLE ENTERED..... 13

SUM OF SQUARES PROPORTION REE	REDUCED IN THIS S DUCED IN THIS STEP	STEP	2.76406 0.00423		
CUMULATIVE SUN CUMULATIVE PRO	I OF SQUARES REDUCE	D	5.33823 0.00816	0F	653.87129
FOR 3 VARIABL MULTIPLE COF (ADJUST F-VALUE FOR STANDARD ERF (ADJUST	ES ENTERED RELATION COEFFICIE ED FOR D.F.) ANALYSIS OF VARIAN OR OF ESTIMATE ED FOR D.F.)	ENT 0 0 10 0	.09036 .08748 .657 .409 .40873		
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR REG. COEFF	OF COMP	UTED LUE	
9 18 13 INTERCEPT	-0.81648E-01 0.86826E-01 -0.28421E-01 0.00483	0.032 0.017 0.006	61 -2 63 4 89 -4	.504 924 .069	U=0 R=1 MAIN TRKS FC
TOLERANC	E = .27641+01				
STEP 4					
VARIABLE ENTER	ED 5				
SUM OF SQUARES	REDUCED IN THIS S UCED IN THIS STEP.	STEP	1.19755 0.00183		
CUMULATIVE SUM CUMULATIVE PRO	OF SQUARES REDUCE	D	6.53578 0.01000	0 F	653.87129
FOR 4 VARIABL MULTIPLE COR (ADJUST F-VALUE FOR STANDARD ERR (ADJUST	ES ENTERED RELATION COEFFICIE ED FOR D.F.) ANALYSIS OF VARIAN OR OF ESTIMATE ED FOR D.F.)	ICE 0 0 9 0	.09998 .09608 .801 .408 .40846		
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR REG. COEFF	OF COMPL	JTED LUE	
9 18 13 5 INTERCEPT	-0.98514E-01 0.10667E+00 -0.29550E-01 -0.57706E-03 0.00397	0.033 0.019 0.006 0.000	19 [°] -2. 11 5. 91 -4. 22 -2.	969 581 280 680	U=0 R=1 MAIN TRKS FC T*C/5000
TOLERAN	CE = .11975E+01				

VARIABLE ENTERED.....17

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	1.60401 0.00245		
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	8.13979 0.01245	0F	653.87129

FOR VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT... 0.11157

(ADJUSTED FOR D.F.)	0.10692
F-VALUE FOR ANALYSIS OF VARIANCE	9.787
STANDARD ERROR OF ESTIMATE	0.406
(ADJUSTED FOR D.F.)	0.40806

VARIABLE	REGRESSION	STD. ERROR OF	COMPUTED
NUMBER	COEFFICIENT	REG. COEFF.	T-VALUE
9 18 13 5 17 INTERCEPT	-0.99573E-01 0.97624E-01 -0.41718E-01 -0.85345E-03 0.34314E-01 0.00077	0.03315 0.01931 0.00793 0.00023 0.01105	-3.004 UFO R=1 5.055 MAIN TRKS -5.259 FC -3.667 T*C/5000 3.105 LANES

TOLERANCE = .16040E+01

VARIABLE ENTERED.....4

SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS DUCED IN THIS STEP	STEP	2.04109 0.00312		
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDUC OPORTION REDUCED	ED	10.18088 0.01557	0F	653.87129
FOR 6 VARIAB MULTIPLE CO (ADJUS F-VALUE FOR STANDARD ER (ADJUS	LES ENTERED RRELATION COEFFICI TED FOR D.F.) ANALYSIS OF VARIA ROR OF ESTIMATE TED FOR D.F.)	ENT 0.12478 0.11959 NCE 10.231 0.407 0.40752	2		
VARIABLE NUMBER 9	REGRESSION COEFFICIENT -0.10398E+00	STD. ERROR OF REG. COEFF. 0.03313	COMPU T-VALU -3.13	TED JE U=() R=1

	-0.10398E+00	0.03313	-3.139	⁻ U≠O R=1
18	0.10995E+00	0.01960	5.609	MAIN TRKS
13	-0.32974E-01	0.00830	-3.971	FC
5	-0.51495E-03	0.00025	-2.046	T*C/5000
17	0.68184E-01	0.01466	4.650	LANES
4	-0.13199E-01	0.00376	-3.508	LOG C **2
INTERCEPT	0.00237			

TOLERANCE = .20411E+01

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VARIABLE ENTERED..... 15

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	0.53573 0.00082		
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	10.71661 0.01639	0F	653.87129
FOR VARIABLES ENTERED			

UK VARIADELS ENTERED	
MULTIPLE CORRELATION COEFFICIENT	0.12802
(ADJUSTED FOR D.F.)	. 0.12194
F-VALUE FOR ANALYSIS OF VARIANCE	9.236
STANDARD ERROR OF ESTIMATE	0.407
(ADJUSTED FOR D.F.)	0.40745

VARIABLE	REGRESSION	STD. ERROR OF	COMPUTED	
NUMBER	COEFFICIENT	REG. COEFF.	T-VALUE	
<u>g</u>	<u>-0.85840E-01</u>	0.03462	-2.480	U=0 R=1
18	0.10615E+00	0.01971	5.386	MAIN TRKS
13	-0.32027E-01	0.00832	-3.850	FC
5	-0.53103E-03	0.00025	-2.113	T*C/5000
17	0.71871E-01	0.01480	4.856	LANÉS
4	-0.11528E-01	0.00387	-2.975	LOG C **2
15	-0.51453E-01	0.02862	-1 <i>.</i> 798	RR ADV WARN
INTERCEPT	0.00270			

TOLERANCE = .53573E+00

VARIABLE ENTERED..... 19 SUM OF SQUARES REDUCED IN THIS STEP..... 0.55937 PROPORTION REDUCED IN THIS STEP..... 0.00086 CUMULATIVE SUM OF SQUARES REDUCED..... 11.27598 CUMULATIVE PROPORTION REDUCED..... 0.11.27598 0.01724 OF 653.87129 FOR 8 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT... 0.13132 (ADJUSTED FOR D.F.)..... 0.12439 E-VALUE FOR ANALYSIS OF VARIANCE.... 8.508

I TYALUL I UK AN	ALIJIJ UF VANIANGE	0.000
STANDARD ERROR	OF ESTIMATE	0.407
(ADJUSTED	FOR D.F.)	0.40738

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE	
9	-0.11291E+00	0.03761	-3.002	U=0 R=1
18	0.93084E-01	0.02095	4.444	MAIN IRKS
13	-0.51266E-01	0.01337	-3.834	FC
5	-0.39595E-03	0.00026	-1.509	T*C/5000
17	0.64794E-01	0.01529	4.238	LANES
4	-0.17893E-01	0.00520	-3.444	LOG C **2
15 .	-0.61931E-01	0.02917	-2.123	RR ADV WARN
19	0.21691E+00	0.11804	1.838	1
INTERCEPT	0.00061			

1

TOLERANCE = .55937E+00

TABLE D-9. MIGRATION -- FLASHING LIGHTS -- VOLUME STEP 3 VARIABLE ENTERED..... 1 305.00261 SUM OF SQUARES REDUCED IN THIS STEP.... 0.04983 PROPORTION REDUCED IN THIS STEP..... 3632.59143 CUMULATIVE SUM OF SQUARES REDUCED..... 0.59347 OF 6120,92609 CUMULATIVE PROPORTION REDUCED..... FOR 3 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT... 0.77037 (ADJUSTED FOR D.F.).... 0.77033 F-VALUE FOR ANALYSIS OF VARIANCE... 7090.970 STANDARD ERROR OF ESTIMATE...... 0.413 (ADJUSTED FOR D.F.).... 0.41326 STD. ERROR OF COMPUTED VARIABLE REGRESSION REG. COEFF. NUMBER COEFFICIENT T-VALUE -83.293 1 0.03395 -0.28280E+01 3 LOG C **2 43.517 0.00190 0.82839E-01 2 LOG T 42.263 0.01778 0.75130E+00 1

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INTERCEPT

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TOLERANCE = .30500E+03

0.00043

D-57

VARIABLE ENTERED..... 1

SUM OF SQUARE PROPORTION RE	S REDUCED IN THIS DUCED IN THIS STEP	STEP	305.09638 0.04998		
CUMULATIVE SU CUMULATIVE PR	M OF SQUARES REDUC	CED	3620.20273 0.59301	OF 6104.8	30 37
FOR 3 VARIAB MULTIPLE CO (ADJUS F-VALUE FOR STANDARD ER (ADJUS	LES ENTERED RRELATION COEFFICI TED FOR D.F.) ANALYSIS OF VARIA ROR OF ESTIMATE TED FOR D.F.)	ENT 0.7 0.7 NCE 7077.4 0.4	7007 7003 07 13 1295		
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR REG. COEFI	OF COMP	UTED LUE	
3	-0.28336E+01 0.83061E-01	0.03403 0.00191	-83.2	269 1 566 LOG C	**2

0.01780

LOG T

42.301

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2	0.83061E-01
1	0 75300E+00
INTERCEPT	0.00040

TOLERANCE = .30510E+03

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VARIABLE ENTERED..... 1

SUM OF SQUA PROPORTION	RES REDUCED IN THIS ST REDUCED IN THIS STEP	EP 3	05.16557 0.05007	
CUMULATIVE CUMULATIVE	SUM OF SQUARES REDUCED PROPORTION REDUCED	36	0.59271	OF 6094.8584
FOR 3 VARI MULTIPLE (ADJ F-VALUE F STANDARD (ADJ	ABLES ENTERED CORRELATION COEFFICIEN DUSTED FOR D.F.) FOR ANALYSIS OF VARIANC ERROR OF ESTIMATE DUSTED FOR D.F.)	T 0,769 0.769 E 7068.771 0.413 0.412	988 984 3 2 76	
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPU T-VAL	JTED LUE
3 2 1 INTERCEPT	-0.28372E+01 0.83201E-01 0.75407E+00 0.00039	0.03408 0.00191 0.01782	-83.2 43.5 42.3	253 1 597 LOG C **2 325 LOG T

TOLERANCE = .30517E+03

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D-59

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STEP 3

VARIABLE ENTERED..... 1

SUM OF SQUA PROPORTION	RES REDUCED IN THIS S REDUCED IN THIS STEP	STEP	305.18782 0.05013	
CUMULATIVE CUMULATIVE	SUM OF SQUARES REDUCI PROPORTION REDUCED	ED	360 7.359 22 0.59253 C	F 6088.0764
FOR 3 VARI MULTIPLE (ADJ F-VALUE F STANDARD (ADJ	ABLES ENTERED CORRELATION COEFFICIN USTED FOR D.F.) OR ANALYSIS OF VARIAN ERROR OF ESTIMATE USTED FOR D.F.)	ENT 0.70 0.70 NCE 7063.3 0.4 0.4	6976 6972 39 13 1263	
VARIABLE NUMBER 3 2 1 INTERCEPT	REGRESSION COEFFICIENT -0.28395E+01 0.83292E-01 0.75477E+00 0.00038	STD. ERROR OF REG. COEFF. 0.03411 0.00191 0.01783	F COMPUTED T-VALUE -83.245 43.616 42.340) LOG C **2 LOG T

TOLERANCE = .30519E+03

TABLE D-10. MIGRATION -- FLASHING LIGHTS -- COMPREHENSIVE

STEP 7

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VARIABLE ENTERED..... 4

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	2.77453 0.00045		
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	3628.91696 0.59029	OF	6147.647
FOR 7 VARIABLES ENTERED			

UR / VARIABLES ENTERED	
MULTIPLE CORRELATION COEFFICIENT	0.76831
(ADJUSTED FOR D.F.)	0.76820
F-VALUE FOR ANALYSIS OF VARIANCE	2998.453
STANDARD ERROR OF ESTIMATE	0.416
(ADJUSTED FOR D.F.)	0,41589

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED) -
1	0.10197E+01	0.02098	48.603	FLGV19.DAT
2	0.13305E+00	0.01520	8.756	MAIN TRKS
5	-0.94634E-01	0.01469	-6.444	LOG T**2
6	0.16536E-01	0.00589	2.806	LANES
3	-0.36433E-01	0.00675	-5.394	FC
7	0.11623E+00	0.02614	4.446	LOG DT
4.	0.18265E-01	0.00456	4.006	POP
INTERCEPT	0.00061			

TOLERANCE = .27745E+01

VARIABLE ENTERED..... 4

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	2.80684 0.00046		
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	3566.24307 0.58810	0F	6064.050
FOR 7 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT (ADJUSTED FOR D.F.) F-VALUE FOR ANALYSIS OF VARIANCE	0.76687 0.76676 2971.351		

STANDARD ERROR OF ESTIMATE..... 0.414 (ADJUSTED FOR D.F.).... 0.41416

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE
<u> </u>	0.10305E+01	0.02112	48.782 FLGV19.DAT
2	0.13531E+00	0.01523	8.887 MAIN TRKS
5	-0.96009E-01	0.01472	-6.524 LOG T**2
6	0.17313E-01	0.00590	2.934 LANES
3	-0.36450E-01	0.00679	-5.370 FC
7	0.11870E+00	0.02622	4.528 LOG DT
4	0.18552E-01	0.00459	4.046 POP
INTERCEPT	0.00053		

TOLERANCE = .28068E+01

VARIABLE ENTERED..... 4

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	2.84340 0.00047		
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	3526.48691 0.58665	0F	6011.264

FOR 7 VARIABLES ENTERED

MULTIPLE CORRELATION COEFFICIENT	0.76593
(ADJUSTED FOR D.F.)	0.76582
F-VALUE FOR ANALYSIS OF VARIANCE	2953,634
STANDARD ERROR OF ESTIMATE	0.413
(ADJUSTED FOR D.F.)	0,41308

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTE T-VALUE	D
1	0.10376E+01	0.02122	48.901	FLGV19.DAT
2	0.13661E+00	0.01524	8,961	MAIN TRKS
5	-0.96960E-01	0.01473	-6.580	LOG T**2
6	0.17777E-01	0.00590	3.011	LANES
3	-0.36364E-01	0.00681	-5.341	FC
7	0.12034E+00	0.02626	4.583	LOG DT
4.	0.18788E-01	0.00460	4.083	POP
INTERCEPT	0.00046			

TOLERANCE = .28434E+01

VARIABLE ENTERED..... 4

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	2.86852 0.00048		
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	3501.38068 0.58570	0F	5978.068

FOR7VARIABLESENTEREDMULTIPLECORRELATIONCOEFFICIENT...0.76531(ADJUSTEDFORD.F.)0.76520F-VALUEFORANALYSISOFVARIANCE...2942.185STANDARDERROROFESTIMATE.....0.412

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTI T-VALUI	ED E
ו	0.10422E+01	0.02128	48,979	FLGV19.DAT
2	0.13737E+00	0.01526	9.004	MAIN TRKS
5	-0.97584E-01	0.01475	-6.618	LOG T**2
6	0.18064E-01	0.00591	3.058	LANES
3	-0.36259E-01	0,00682	-5.315	FC
7	0.12137E+00	0.02629	4.617	LOG DT
4 .	0.18944E-01	0.00461	4.108	POP
INTERCEPT	0.00042			

TOLERANCE = .28685E+01

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VARIABLE ENTERED.....]

FROPURITUN REL	S REDUCED IN THIS DUCED IN THIS STEP	STEP	74.79257 0.04221	
CUMULATIVE SUN CUMULATIVE PRO	1 OF SQUARES REDUC DPORTION REDUCED	ED	1110.82744 0.62684	OF 1772.1187
FOR ² VARIABI MULTIPLE COF (ADJUS F-VALUE FOR STANDARD ERF (ADJUST	LES ENTERED RRELATION COEFFICI FOR D.F.) ANALYSIS OF VARIA ROR OF ESTIMATE FED FOR D.F.)	ENT 0.7917 0.7916 NCE 3262.982 0.413 0.4126	73 57 53	
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE	
2 1 INTERCEPT	-0.19453E+01 0.18296E+00 -0.00072	0.04518 0.00873	-43.056 1 20.962 LC	OG T * LOG C
TOLERANCE	= .74793E+02			
STEP 2				
VARIABLE ENTER	RED 1			
VARIABLE ENTER SUM OF SQUARES PROPORTION RED	RED 1 5 REDUCED IN THIS DUCED IN THIS STEP	STEP	75.14251 0.04274	
VARIABLE ENTER SUM OF SQUARES PROPORTION REE CUMULATIVE SUM CUMULATIVE PRO	RED 1 S REDUCED IN THIS DUCED IN THIS STEP 1 OF SQUARES REDUC DPORTION REDUCED	STEP ED	75.14251 0.04274 1100.28178 0.62585	OF 1758.0507
VARIABLE ENTER SUM OF SQUARES PROPORTION REE CUMULATIVE SUM CUMULATIVE PRO FOR 2 VARIABL MULTIPLE COF (ADJUST F-VALUE FOR STANDARD ERF (ADJUST	RED 1 S REDUCED IN THIS DUCED IN THIS STEP M OF SQUARES REDUC PORTION REDUCED LES ENTERED RELATION COEFFICI FED FOR D.F.) ANALYSIS OF VARIA ROR OF ESTIMATE	STEP ED ED ENT 0.7911 0.7910 NCE 3249.313 0.411 0.4115	75.14251 0.04274 1100.28178 0.62585	OF 1758.0507
VARIABLE ENTER SUM OF SQUARES PROPORTION REE CUMULATIVE SUM CUMULATIVE PRO FOR 2 VARIABL MULTIPLE COF (ADJUST F-VALUE FOR STANDARD ERF (ADJUST VARIABLE NUMBER	RED 1 S REDUCED IN THIS DUCED IN THIS STEP M OF SQUARES REDUC PORTION REDUCED LES ENTERED RELATION COEFFICI FED FOR D.F.) ANALYSIS OF VARIA ROR OF ESTIMATE REGRESSION <u>COEFFICIENT</u>	STEP ED ED ENT 0.7911 0.7910 NCE 3249.313 0.411 0.4115 STD. ERROR OF <u>REG. COEFF.</u>	75.14251 0.04274 1100.28178 0.62585 11 25 53 COMPUTED <u>T-VALUE</u>	OF 1758.0507

TOLERANCE = .75143E+02

VARIABLE ENTERED..... 1

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	75.35812 0.04306
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	1094.18811 0.62527 OF 1749.9519
FOR 2VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT0.79074 0.79068F-VALUE FOR ANALYSIS OF VARIANCE3241.198 0.411 0.41090	
VARIABLE NUMBERREGRESSION COEFFICIENTSTD. ERROR OF REG. COEFF.2-0.19634E+01 0.18563E+00 INTERCEPT0.04566 0.00879	COMPUTED <u>T-VALUE</u> -43.002 1 21.129 LOG T * LOG C
TOLERANCE = .75358E+02	
STEP 2	
VARIABLE ENTERED]	
SUM OF SQUARES REDUCED IN THIS STEP	75 48342
PROPORTION REDUCED IN THIS STEP	0.04325
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	0.04325 1090.63530 0.62492 OF 1745.2375
PROPORTION REDUCED IN THIS STEP CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED FOR 2 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT 0.79052 (ADJUSTED FOR D.F.) STANDARD ERROR OF ESTIMATE 0.410 (ADJUSTED FOR D.F.) 0.41053	0.04325 1090.63530 0.62492 OF 1745.2375

TOLERANCE = .75483E+02

TABLE D-12. MIGRATION -- GATES -- COMPREHENSIVE STEP 4 2 VARIABLE ENTERED..... 9.30615 SUM OF SOUARES REDUCED IN THIS STEP.... 0.00535 PROPORTION REDUCED IN THIS STEP..... 1081,47261 CUMULATIVE SUM OF SQUARES REDUCED..... 0.62149 OF 1740.12949 CUMULATIVE PROPORTION REDUCED..... FOR 4 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT... 0.78835 0.78816 STANDARD ERROR OF ESTIMATE...... 0.412 (ADJUSTED FOR D.F.)..... 0.412 0.41202 STD. ERROR OF VARIABLE REGRESSION COMPUTED NUMBER REG. COEFF. COEFFICIENT T-VALUE 14.184 0.05205 0.73828E+00 GATE09 1 9.698 0.01958 TRACKS 3 0.18987E+00 0.09137 -9.150 1 -0.83597E+00 4 7.407 LANES 2 0.92437E-01 0.01248

TOLERANCE = .93061E+01

INTERCEPT

-0.00044

VARIABLE ENTERED..... 2

SUM OF SQUA PROPORTION	RES REDUCED IN THIS S REDUCED IN THIS STEP.	TEP	9.39554 0.00543	
CUMULATIVE CUMULATIVE	SUM OF SQUARES REDUCE PROPORTION REDUCED	D	1073.34786 0.62069 OF 172	29.2753
FOR 4 VARI MULTIPLE (ADJ F-VALUE F STANDARD (ADJ	ABLES ENTERED CORRELATION COEFFICIE USTED FOR D.F.) OR ANALYSIS OF VARIAN ERROR OF ESTIMATE USTED FOR D.F.)	NT 0.78784 0.78765 CE 1588.518 0.411 0.41116		
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE	
1 3 4 2 INTERCEPT	0.74646E+00 0.19067E+00 -0.83634E+00 0.93149E-01 -0.00048	0.05219 0.01960 0.09151 0.01249	14.244 GATEO9 9.728 TRACKS -9.139 1 7.458 LANES	

TOLERANCE = .93955E+01

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STEP 4

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VARIABLE ENTERED..... 2

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	•	9.44909 0.00548		
CUMULATIVE SUM OF SQUARES REDUCED	•	1068.51498 0.62021	0F	1722.81969
FOR 4 VARIABLES ENTERED MULTIPLE CORRELATION COEFFICIENT (ADJUSTED FOR D.F.) F-VALUE FOR ANALYSIS OF VARIANCE STANDARD ERROR OF ESTIMATE (ADJUSTED FOR D.F.)	0.78754 0.78735 1585.287 0.410 0.41065			

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTE T-VALUE	D
1.	0.74660E+00	0.05228	14.281	GATE09
3	0.19113E+00	0.01961	9.745	TRACKS
4	-0.83649E+00	0.09160	-9.132	1
2	0.93574E-01	0.01250	7,488	LANES
INTERCEPT	-0.00050			

TOLERANCE = .94491E+01

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VARIABLE ENTERED..... 2

SUM OF SQUA PROPORTION	RES REDUCED IN THIS REDUCED IN THIS STEP	STEP	9.48120 0.00552		
CUMULATIVE CUMULATIVE	SUM OF SQUARES REDUC PROPORTION REDUCED	ED	1065.62402 0.16993	OF	1718.9497
FOR 4 VARI MULTIPLE (ADJ F-VALUE F STANDARD (ADJ	ABLES ENTERED CORRELATION COEFFICI USTED FOR D.F.) OR ANALYSIS OF VARIA ERROR OF ESTIMATE USTED FOR D.F.)	ENT 0.78 0.78 NCE 1583.36 0.41 0.41	3735 3717 57 10 1034		
VARIABLE NUMBER 1 3 4 2 INTERCEPT	REGRESSION COEFFICIENT 0.74949E+00 0.19139E+00 -0.83656E+00 0.93829E+01 -0.00051	STD. ERROR O REG. COEFF. 0.05233 0.01962 0.09165 0.01250	F COMPUT <u>T-VALL</u> 14.303 9.755 -9.123 7.507		GATEO9 TRACKS 1 LANES

TOLERANCE = .94812E+01

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TABLE D-13. SELECTION REGRESSION -- GATES

STEP-WISE MULTIPLE REGRESSION.....GAT098

NUMBER	0F	OBSERVATIONS	3888
NUMBER	0F	VARIABLES	20
NUMBER	0F	SELECTIONS	1

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CONSTANT TO LIMIT VARIABLES 0.00000

VARIABLE	MEAN	STANDARD	VARIABLE
NO.		DEVIATION	
ł	0.62589	0.35024	LOG T
2	1.59555	0.69239	LOG C
3	0.90100	0.71782	LOG T **2
4	5.61657	3.20156	LOG C **2
5	14.80179	35.31585	T*C/5000
6	0.46916	0.29430	LOG NITE
7	0.44779	0.31509	LOG DAY THRU
8	0.24501	0.27815	LOG SWITCH
9	0.11621	0.19108	U=O R=1 Forced out
10	20.61505	12.61789	MAX SPEED
11	0.46042	0.15788	HWY PAVED
12	1.20522	1.01110	POP
13	2.16860	0.80058	FC
14	0.28519	0.32094	NRBY XING HWY
15	0.29478	0.25227	RR ADV WARN
16	4.31561	3.60533	TRUCKS
17	1.17982	0.67300	LANES
18	0.73989	0.47198	TRACKS
19	0.46930	0.14133	1
20	-0.00118	0.41014	ACC=1 NOACC=-1

VARIABLE ENTERED..... 18

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	1.87037 0.00286	
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	6.61961 0.01012 OF	653.87129
FOR 4 VARIABLES ENTERED		

MULTIPLE CORRELATION COEFFICIENT	0.10062
(ADJUSTED FOR D.F.)	0.09674
F-VALUE FOR ANALYSIS OF VARIANCE	9.928
STANDARD ERROR OF ESTIMATE	0.408
(ADJUSTED FOR D.F.)	0,40843

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE	
15	-0.84977E-01	0.02703	-3.144 RR ADV WARN	ł
17	0.63007E-01	0.01467	4.295 LANES	
4	-0.15191E-01	0.00351	-4.332 LOG C **2	
18	0.53622E-01	0.01601	3.350 TRACKS	
INTERCEPT	-0.00482			

TOLERANCE = .18704E+01

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VARIABLE ENTERED..... 12

SUM OF SQUARES PROPORTION REDU	REDUCED IN THIS S JCED IN THIS STEP.	TEP	1.13532 0.00174	
CUMULATIVE SUM CUMULATIVE PROF	OF SQUARES REDUCE PORTION REDUCED	D	10.29687 0.01575 ()F 653.87129
FOR 6 VARIABLE MULTIPLE COR (ADJUST) F-VALUE FOR A STANDARD ERR (ADJUST)	ES ENTERED RELATION COEFFICIE ED FOR D.F.) ANALYSIS OF VARIAN DR OF ESTIMATE ED FOR D.F.)	NT 0.1254 0.1203 CE 10.349 0.407 0.4074	9 3 8	
VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTI T-VALUE	ED
15 17 4 18 13 12 INTERCEPT	-0.54843E-01 0.67561E-01 -0.16917E-01 0.91518E-1 -0.32339E-01 0.24023E-01 0.00376	0.02777 0.01488 0.00379 0.01940 0.00770 0.00918	-1.975 4.539 -4.461 4.718 -4.201 2.617	RR ADV WARN LANES LOG C **2 TRACKS FC POP

TOLERANCE = .11353E+01

VARIABLE ENTERED..... 19

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	0.21245 0.00032		
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	11.15425 0.01706	OF	653.87129

FOR 8 VARIABLES ENTERED

MULTIPLE CORRELATION COEFFICIENT	0.13061
(ADJUSTED FOR D.F.)	0.12363
F-VALUE FOR ANALYSIS OF VARIANCE	8.415
STANDARD ERROR OF ESTIMATE	0.407
(ADJUSTED FOR D.F.)	0.40742

VARIABLE	REGRESSION	STD. ERROR OF	COMPUTED	
NUMBER	COEFFICIENT	REG. COEFF.	T-VALUE	
15	-0.67447E-01	0.02891	-2.333	RR ADV WARN
17	0.63681E-01	0.01535	4.150	LANES
4	-0.17725E-01	0.00520	-3.407	LOG C **2
18	0.90256E-01	0.02105	4.287	TRACKS
13	-0.50211E-01	0.01332	-3.769	FC
12	0.26743E-01	0.00930	2.877	POP
5	-0.40299E-03	0.00026	-1.536	T*C/5000
19 .	0.12433E+00	0.10980	1.132	1
INTERCEPT	0.00061			

TOLERANCE = .21245E+00

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VARIABLE ENTERED..... 6

SUM OF SQUARES REDUCED IN THIS STEP PROPORTION REDUCED IN THIS STEP	0.94393 0.00144		
CUMULATIVE SUM OF SQUARES REDUCED CUMULATIVE PROPORTION REDUCED	12.33853 0.01887	OF	653,87129

FOR 10 VARIABLES ENTERED

MULTIPLE CORRELATION COEFFICIENT	0.13737
(ADJUSTED FOR D.F.)	0.12881
F-VALUE FOR ANALYSIS OF VARIANCE	7.457
STANDARD ERROR OF ESTIMATE	0.407
(ADJUSTED FOR D.F.)	0.40725

VARIABLE NUMBER	REGRESSION COEFFICIENT	STD. ERROR OF REG. COEFF.	COMPUTED T-VALUE
15 17 4 18 13 12 5 19 3 6 INTERCEPT	-0.68688E-01 0.61075E-01 -0.20093E-01 0.10328E+00 -0.48599E-01 0.30711E-01 -0.87022E-04 0.85167E-01 -0.96446E-01 0.21333E+00 -0.00004	0.02892 0.01541 0.00547 0.02290 0.01333 0.00950 0.00032 0.12062 0.03625 0.08932	-2.375 RR ADV WARN 3.964 LANES -3.673 LOG C **2 4.510 TRACKS -3.646 FC 3.234 POP -0.270 T*C/5000 0.706 1 -2.661 LOG T **2 2.388 LOG NITE

TOLERANCE = .94303E+00

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APPENDIX E

EXPECTED ACCIDENT FREQUENCY PLOTS AND EOC PLOTS

Figures E-1 through E-8 give expected accident frequency per year versus percent of all crossings which are more hazardous (for a given hazard index and warning device class). (See Section 4.3.) Figures E-9 through E-11 give the percent of all accidents versus percent of all crossings (for a given hazard index and warning device class). The resulting plots of Figures E-9 through E-11 are EOC curves. (See Section 4.1.)



FIGURE E-1. TSC CROSSBUCKS MODEL (f_1 and f_3) (SEE SECTION 4.3)



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FIGURE E-2. TSC CROSSBUCKS MODEL (f_2 AND f_3) (SEE SECTION 4.3)

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FIGURE E-3. COLEMAN-STEWART CROSSBUCKS MODEL (f_2 AND f_4) (SEE SECTION 4.3)



FIGURE E-4. TSC FLASHING LIGHTS MODEL (f_2 AND f_3)



FIGURE E-5. TSC GATES MODEL (f_2 AND f_3)

E-6



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FIGURE E-6. COLEMAN-STEWART FLASHING LIGHTS MODEL $(f_2 \text{ AND } f_4)$

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FIGURE E-7. NEW HAMPSHIRE FLASHING LIGHTS MODEL $(f_1 \text{ AND } f_2)$



FIGURE E-8. TSC FLASHING LIGHTS MODEL (f_1 AND f_2)

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FIGURE E-9. COMPARATIVE EOC'S (CROSSBUCKS)



FIGURE E-10. COMPARATIVE EOC'S (FLASHING LIGHTS)



FIGURE E-11. COMPARATIVE EOC'S (AUTOMATIC GATES)

APPENDIX F

DATA BASES

The inventory data used in this study were derived from two sources: a tape containing the inventory characteristics of all public railroad crossings in the U.S., and a tape containing the inventory characteristics of railroad crossings which had an accident in 1975.* Those crossings which had multiple accidents are repeated for every accident they had. Thus, if a crossing had three accidents in 1975, the tape would contain three accounts of its inventory characteristics.

There were 219,162 public railroad crossings in existence in 1975. Included in the accident tape were 8,028 crossings, of which 943 were repetitive. Table F-1 depicts the format for these two tapes. Many of the fields in this table were descriptive in nature and hard to quantify. These fields were extracted from the data base. The resultant data base is shown in Table F-2, while Table F-3 depicts the accident data base and the non-accident data base broken down into warning device class. Table F-4 shows the repetitive nature of the accident data base, while Table F-5 depicts the data sets used in the iterative nonlinear regression.

^{*}Some of the 1975 accidents did not appear in the data base of the second tape because they could not be linked to crossings. See Subsection 2.3.1.

TABLE F-1. INVENTORY DATA BASE

LOC	LEN	TYPE	DESCRIPTION
1	7	СН	Crossing number (6 digits & check digit)
8	6	СН	Begin date (YYMNDD format)
14	6	СН	End date (YYMMDD format or 999999)
20	1	СН	Crossing status (1-changed, 2-new, 3-closed, 4-change in place)
21	2	ΖD	State code
23	1	СН	' C '
24	3	ZD	County code
27	2	ZD	State code
29	4	ΖD	City code
33	1	СН	Is city code for city or nearest city? (1-nearest city, 0-city)
34	4	СН	Railroad code
38	14	СН	Railroad division or region
52	14	СН	Railroad subdivision or region
66	7	СН	Highway number
73	17	CH	Street or road name
90	10	СН	Railroad ID number
100	6	ZD	Timetable station
106	15	СН	Branch or line name
121	6	ZD	Milepost (pic 9999V99)
127	10	СН	County map reference number
137	1	СН	Crossing type (l-pedestrian, 2-private, 3-public)
138	1	СН	Crossing position (1-at grade, 2-RR under, 3-rr over)
139	1	СН	Private crossing location (l-farm, 2-residential 3-recreational, 4-industrial)
140	1	СН	Private signs or signals (blank-not a private crossing l-signs, 2-signals, 3-no signs or signals, 4-both signs and signals)
141	15	СН	Private sign or signal description
156	1	СН	Form initiator (l-railroad, 2-state, 3-DOT, 4-file creation)
157	6	ΖD	Batch number

TABLE F-1. INVENTORY DATA BASE (cont.)

LOC	LEN	TYPE	DESCRIPTION
163	1	СН	User code
164	2	СН	Date updated
166	5	ΖD	Link Field
Remai grade	nder e.	of field	ls will be blank unless crossing is public at
171	2	ΖD	Number of daylight thru trains
173	2	ZD	Number of daylight switch trains
175	2	ZD	Number of night thru trains
177	2	ZD	Number of night switching trains
179	1	ZD	Less than one train per day? (0-no, 1-yes)
180	3	ΖD	Maximum timetable speed
183	3	ZD	Typical minimum speed
186	3	ZD	Typical maximum speed
189	1	Z D	Number of main tracks
190	2	ZD	Number of other tracks
192	10	СН	Description of other tracks
202	1	Z D	Does another railroad operate a separate track at crossing? (1-yes, 2-no)
203	16	СН	List of other railroads with separate track (four characters each)
219	_ 1	ZD	Does another railroad operate over your track at crossing (l-yes, 2-no)
220	16	СН	List of other railroads on same track (4 characters each)
236	1	Z D	Highest warning device class at crossing 8-gates, 7-flashing lights, 6-highway signals, wigwags, or bells, 5-special warning, 1-crossbucks, 3-stop signs, 2-other signal or signals, 1-none of the above)
237	1	ZD	Number of reflectorized crossbucks
238	1	Z D	Number of non-reflectorized crossbucks
239	1	ZD	Number of standard highway stop signs
240	1	ZD	Number of other stop signs
241	1	ΖD	Number of other signs (1)
242	10	CH	Description of other signs (1)

TABLE F-1. INVENTORY DATA BASE (cont.)

LOC	LEN	TYPE	DESCRIPTION
252	1	ZD	Number of other signs (2)
253	10	CH	Description of other signs (2)
263	1	ΖĐ	Number of red and white reflectorized gates
264	1	ΖD	Number of other colored gates
265	1	ZD	Number of cantilevered flashing lights over traffic lanes
266	1	ZD	Number of cantilevered flashing lights not over traffic lanes
267	1	ΖD	Number of mast mounted flashing lights
268	1	ΖD	Number of other flashing lights
269	9	СН	Description of other flashing lights
278	1	ZD	Number of highway traffic signals
279	1	ZD	Number of wigways
280	1	ΖD	Number of bells
281	20	СН	Description of special warning not train activated
301	1	ZD	Is track equipped with any signs or signals (1-no, 0-yes)
302	1	ΖD	Is commercial power available? (2-no, 1-yes)
303	1	ZD	Method of signalling for train operation: Is track equipped with signals? (2-no, 1-yes)
304	1	ZD	Does crossing provide speed selection (1-yes, 2-no, 3-N/A)
305	1	ZD	Type of development (1-open space, 2-residential, 3-commercial, 4-industrial, 5-institutional)
306	1	ZD	Is highway paved? (2-no, 1-yes) (Note different coding in B.5.1.3)
307	1	Z D	Does track run down a street (2-no, 1-yes)
308	1	Z D	Pavement markings (1-stopline, 2-RR Xing symbol, 3-none, 4-both stoplines and RR Xing symbols)
309	1	ΖD	Nearby intersecting highway? (2-no, 1-yes)
310	1	ZD	RR advance warning signs present (2-no, 1-yes)
311	1	ZD	Smallest crossing angle (1-0 to 29 degrees, 2-30 to 59 degrees, 3-60 to 90 degrees)
312	1	ΖD	Crossing surface

TABLE F-1. INVENTORY DATA BASE (cont.)

LOC	LEN	TYPE	DESCRIPTION
313	1	ZD	Number of traffic lanes
314	1	ZD	Are truck pullout lanes present? (2-no, l-yes)
315	1	ZD	Is crossing on state highway system (2-no, 1-yes)
316	2	ZD	Highway system
318	2	ZD	Functional classification of road over crossing (The tens digit codes population.)
320	6	ΖD	Estimated AADT
326	2	ΖD	Estimated percent trucks

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LOC	LEN	DESCRIPTION
1	7	Crossing number
8	2	State code
10	3	County Code
13	4	Railroad Code
17	7	Highway nubmer
24	2	Number of daylight tru trains
26	2	Number of daylight switch trains
28	2	Number of night thru trains
30	2	Number of night switch trains
32	1	Less than one train per day? (0-no, 1-yes)
33	3	Maximum timetable speed
36	3	Typical minimum speed
39	3	Typical maximum speed
42	1	Number of main tracks
43	2	Number of other tracks
45 [.]	1	Does another railroad operate on a separate track at crossing? (1-yes, 2-no)
46	1	Does another railroad operate over your track at crossing (1-yes, 2-no)
47	1	Highest warning device class at crossing (8-gates, 7-flashing lights, 6-highway signals, wigwags, or bells, 5-signal warning, 4-crossbucks, 3-stop signs, 2-other signs or signals, 1-none)
48	1	Number of reflectorized crossbucks
49	1	Number of non-reflectorized crossbucks
50	1	Number of standard highway stop signs
51	1	Number of other stop signs
52	1	Number of other signs (1)
53	1	Number of other signs (2)
54	1	Number of red and white reflectorized gates
55	1	Number of other colored gates
56	1	Number of cantilevered flashing lights over traffic lanes
57	1	Number of cantilevered flashing lights not over traffic lanes

LOC	LEN	DESCRIPTION
58	1	Number of most mounted flashing lights
59	1	Number of other flashing lights
60	1	Number of highway traffic signals
61	1	Number of wigwags
62	1	Number of bells
63	1	Is track equipped with any signs or signal (1-no, 0-yes)
64	1	Method of signalling for train operation: Is track equipped with signals? (2-no, 1-yes)
65	1	Does crossing provide speed selection (1-yes 2-no 3-N/A)
66	1	Type of development (1-open spose, 2-residential 3-commercial, 4-industrial, 5-institutional)
67	1	Is highway paved? (2-no, 1-yes)
68	1	Does track run down a street (2-no, 1-yes)
69	1	Pavement markings (l-stopline, 2-RR crossing symbol 3-none, 4-both stoplines and RR Xing symbol)
70	1	Nearby intersecting highway? (2-no, 1-yes)
71	1	RR advance warning signs present (2-no, 1-yes)
72	1	Smallest crossing angle (1-0 to 29 degrees, 2-30 to 59 degrees, 3-60 to 90 degrees)
73	1	Crossing Surface
74	1	Number of traffic lanes
75	1	Are truck pullout lanes present (2-no, l-yes)
76	1	Is crossing on state highway system (2-no, 1-yes)
77	2	Highway System
79	2	Functional classification of road over crossing*
81	6	Estimated AADT
87	2	Estimated percent trucks

*The tens digit of LOC 79 codes population. See Reference 4.

Warning Device Class	Non-Accident	Accident
Gates	11,983	707
Flashing Lights	33,969	2,650
Highway signals, wigwags, bells	3,395	169
Special warning	8,418	216
Crossbucks	141,477	3,969
Standard highway stop signs	3,525	109
Other signs	1,079	15
None	15,316	193

TABLE F-3. BREAKDOWN BY WARNING DEVICE CLASS

219,162 8,028

TABLE F-4. BREAKDOWN BY MULTI-ACCIDENTS

# Accidents	# Crossings	<pre># Repetitions</pre>
1	6,344	0
2	609	609
3	96	192
4	21	63
5	9	36
6	0	0
7	2	12
8	2	14
9	1	8
10	1	9
	7,085	943

Total sample size = 8,028 Total number of repetitions = 943

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TABLE F-5. COMPOSITION OF DISJOINT DATA BASES USED FOR MODEL CONSTRUCTION AND TESTING

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	Gates	Flashing Lights	Wigwags Bclls	Special	Cross- bucks	Stop Sign	Other Sign	None	Total
Accident	354	1,324	85	108	1,984	55	8	96	4,014
Non-accident	3,535	13,250	485	1,177	20,188	493	147	2,203	41,478
Total	3,889	14,574	570	1,285	22,172	548	155	2,299	45,492

TEST DATA BASE (SUBDATA BASE B OF FIGURE 2-1)

VALIDATION DATA BASE (SUBDATA BASE A OF FIGURE 2-1)

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Accident	353	1,326	84	108	1,985	54	7	97	4,014
Non-accident	3,535	13,250	487	1,178	20,188	493	147	2,204	41,482
Total	3,888	14,576	571	1,286	22,173	547	154	2,301	45,4 96

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APPENDIX G

NONLINEAR (LOGISTIC) VERSUS LINEAR CONSTRUCTION

AN EXPERIMENT TO COMPARE THE LINEAR MODELS WITH THE LOGISTIC MODELS

Considerable time had elapsed (and corresponding experience gained) between the development of the best comprehensive linear models 8-C and 8-D and the corresponding comprehensive logistic models which are reported as the best TSC models in this report. Consequently, it seemed desirable to obtain some comparison of the earlier linear models with the later logistic models. The earlier models had been thought to be good, and had even outperformed the New Hampshire and Coleman-Stewart models. The trouble was that the data base on which they were tested was not disjoint with the data base on which they were constructed (this was inadvertent). It was expected that the later logistic models would perform better than the earlier linear models. This was indeed the case. The surprising feature was how poorly the earlier linear models did perform. This is explained later in the Appendix, but first, a description is given of the experiment for comparing the models.

Since the data bases were reconstructed after it was discovered that they were not disjoint, the new data bases (which are purely disjoint in themselves) partially overlap the old data bases. Thus, to test linear model 8-C it was necessary to retune it on one data base and run its EOCs -- power factors on the disjoint data base. This was done, with the variables in the volume part of 8-C being run through a new linear volume regression so that the same variables appeared but the coefficients were retuned. The full regression was also completed in the same manner. Thus, a model identical with 8-C (crossbucks) in its variables but with its coefficients tuned to the latest construction data base was constructed. The EOC was run against the TSC comprehensive model; the results are shown in Table G-1. (See Appendix C and Section 2 for information on how to read EOCs.)

G-1

The test result indicated that the linear models were unsatisfactory. They were not significantly worse than the New Hampshire model, but no better. It seems that linear regression technique is inadequate to produce hazard functions, which are essentially non-linear -- probably an approximate function of a product of car and train variables. Any function can be built up out of linear terms. However, the straight linear approach was evidently not powerful enough for this purpose. As a result, the new models developed of primary interest are those of nonlinear construction which are reported on in Section 4. TABLE G-1. EOC BEST LINEAR VERSUS BEST NON-LINEAR MODEL (CROSSBUCKS)

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ATON						1.1	212		2				1	4 8										ļ					0		~	n 1 10 1				0				
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TABLE G-1. (CONT.)

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APPENDIX H

HAZARD INDEXES BASED ON ACCIDENT HISTORY

H.1 BASIC METHOD

The hazard indexes, whose development and testing is reported on in this report, are deficient in one notable respect: they do not base the hazard on accident history. Thus, although accident history is used in the development and testing of hazard indexes based on other characteristics, the hazard function itself does not have accident history as a component. Although there is not sufficient time to develop such hazard indexes for this report, a method has been developed to do so, and the necessary calculations will be presented in this appendix. The techniques here are being used in a current effort to develop accident history dependent hazard indexes.

The basic idea is simple: one of the variables determining hazard will be the number of accidents actually observed at the crossing during the data period (the year 1975 in this case). A function f(h) has already been developed which gives expected accident frequency in terms of a hazard index h (f=ce^{2h}, for example). It is now necessary to develop a function $f_c(h,y)$ which gives the expected number of accidents in a future year, given that the hazard index is h and that y accidents have been observed in a specified prior period (of a specified length in years).

First, the problem is simplified as follows: let $F_c(h,y)$ be the expected number of accidents at a crossing having hazard index h and having had y or more accidents in the year 1975. In particular, $F(h) = F_c(h,1)$ is the expected number of accidents at a crossing whose hazard index is h and which had <u>at least one</u> accident in the data year (1975). It should be remembered that h is a function of crossing characteristics other than actual accident history.

The computation of F(h) will be based on the following lemma: (Y = number of accidents in a year)

 $F_{c}(h,y) = E(Y | Y \ge y+1, h) Pr(Y \ge y+1 | h) / Pr(Y \ge y | h).$

(see end of this Appendix for derivation and definitions) and in particular:

$$F(h) = E(Y|Y>2,h) Pr(Y>2|h) / Pr(Y>1|h).$$

The quantity $E(Y|Y \ge 2,h) Pr(Y \ge 2|h) / Pr(Y \ge 1|h)$ can also be written:

(H.1)
$$F(h) = E(Y|Y \ge 2, h) Pr(Y > 2|Y > 1, h).$$

The latter quantity can be estimated on every sample which contains crossings for which $Y \ge 1$ (i.e., had an accident). Thus, it may be evaluated for every crossing which had an accident. Let $\eta_i = 0$ if $Y_i = 1$; $\eta_i = Y_i$ if $Y_i \ge 2$; and η_i be undefined if $Y_i = 0$. Then $F(h_i)$ is estimated by η_i at sample i. Suppose that:

(H.2)
$$F(h) = \lambda_1 + \lambda_2 f.$$

Then ℓ_1 and ℓ_2 can be determined by simple linear regression:

(H.3)
$$\hat{n}_i = l_1 + l_2 f$$
, where

 ℓ_1 and ℓ_2 are determined by minimizing $\sum_{i=1}^{j} (\eta_i - \eta_i)^2$. The sum over i is carried only over crossings which had at least one accident, i.e., over the crossings in the accident data base. $F(h) = \ell_1 + \ell_2 f + \ell_3 f^2$ could be similarly optimized.

The quantity F is calculated here as though the data base had no missing accidents. To correct for the missing accidents it is not sufficient to multiply by r, as was done for f (Subsection 2.3.1). However, in the next subsection we use F to make some more calculations internal to the (incomplete) data base, and then, at the end, it is shown how to correct the results for the missing accidents. In the following material notation is switched from F to F* for a reminder that it is calculated on a data base with missing accidents, and that correction for the missing accidents has not been made. One should remember that f (Subsection 2.3.1)

has been calculated on the data base with missing accidents, but has been corrected for the missing accidents. The designation f* is now introduced to denote f uncorrected for the missing accidents; then $f^* = \frac{f}{r}$ (by definition).

It is useful to mention some partial results which will illustrate the magnitudes involved. The numbers given here for ℓ_1 and ℓ_2 are to be taken as tentative, since the limits of accuracy have not been adequately assessed. A preliminary regression of the form

(H.4) $F^* = \ell_1 + \ell_2 f^*$

was run for crossbucks, flashing lights, and automatic gates. This amounts to a strong limitation on the form of F*, as it is forced to be a linear function of f alone (and not directly of C, T, number of tracks, etc.). The results are shown in Table H-1.

Table H-1 F* $\approx \ell_1 + \ell_2 f^*$

Warning Device Class	l l	^l 2
Crossbucks	0.044	1.78
Flashing lights	0.085	1.68
Automatic gates	.0.13	1.52
(Results are preliminary.	Automatic ga	tes case is
subject to larger errors	than the othe	r two cases.)

That F* is valid internal to the data base and has not been corrected for missing (unlinked) accidents. Similarly,

 $f^* = \frac{f}{r}$.

In Table H-1 the results for crossbucks and flashing lights are more reliable than those for automatic gates, which were calculated for 614 points only (total number of gate crossings with repeat accidents). Note that one equation covers both flashing lights and crossbucks in a very approximate manner:

 $F^* \approx f^*_{(.25)} + 1.7f^*$

where $f_{(.25)}^*$ represents the 25th percentile from the top of f* for the particular warning device class, i.e., for that warning device class one-fourth of the crossings have f* greater than $f_{(.25)}^*$.

H.2 ACCIDENT HISTORY DEPENDENT HAZARD INDEX

The approach just discussed uses equation H.1 (Section H.1) which is based only on the assumption that inherent hazard does not change over time. It is necessary to make such an assumption in any application of accident history.

A procedure is now outlined (using the one just developed) which provides a much more complete means of incorporating accident history into the calculation of hazard indexes (expected frequency of accident during a future time interval -- thus, absolute indexes). For a future year the expected number of accidents at a particular crossing will be calculated having given inventory characteristics and a past history of a given number of accidents in a specified time period. The goal of this calculation can be stated even more concisely: Given a crossing with specified characteristics, and given the fact that n accidents occurred in T years, find the expected number of accidents for that crossing for next year.

For this analysis the above assumption of constancy of inherent hazard with time and some other assumptions as well, are necessary. One draws on the techniques of empirical Bayseian analysis, References 9 and 10. (See Reference 9 for an application to determine insurance premium penalties for drivers based on their accident records as well as on other characteristics.)

In describing accident proneness, one customarily assumes that the specific hazard per unit time, ϕ , is not the same for all individuals (in this case crossings), but instead, has a gamma distribution with parameters a and b as indicated in H.5:

(H.5)
$$\Pr(\phi < \phi_0) = \frac{b^a}{r(a)} \int_0^{\phi_0} \lambda^{a-1} e^{-\lambda b} d\lambda$$

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(The gamma distribution with its two parameters is a "natural conjugate prior" distribution to the Poisson distribution which is introduced presently, and is required on theoretical grounds, given the time homogeneity assumption. See Reference 10.) In the present treatment, a and b are to be functions of the crossing characteristics, as recorded in the crossing inventory, or alternatively, as reflected in f^* and F^* , so that a and b are functions of f^* and F^* .

Based on the time homogeneity assumption, the probability of r accidents in T years at a crossing with a specific hazard, ϕ , has the Poisson distribution of Equation H.6:

(H.6)
$$\Pr(r=n) = \frac{(\phi T)^n e^{-\phi T}}{n!}$$

This, together with equation H.5, yields a probability distribution for the number of accidents in a year at a crossing, with ϕ unknown but a and b known. The distribution is a negative binomial distribution with parameters a and b:

(H.7)
$$P(n) = \begin{bmatrix} a+n-1 \\ a-1 \end{bmatrix} \left(\frac{b}{1+b} \right)^a \left(\frac{1}{1+b} \right)^n$$

(Reference 9 notes the custom of using the negative binomial to describe "accident proneness", and notes other references on this topic. Note that it is characteristic of the empirical Bayes procedure that a specific hazard, ϕ , is postulated but not known, and not even directly estimated.)

The negative binomial distribution results in a mean number of accidents of $\frac{a}{b}$, which, in turn, is the unconditional expected frequency of accidents, f* (for the data base at hand):

(H.8) $f^* = \frac{a}{b}$

Similarly, from the expression for F, which is here to be interpreted as F* in Equation (H.1), one derives.

$$F^* = \frac{f^* - P(1)}{1 - P(0)} \qquad \begin{cases} (P(1) = \text{probability of one accident} \\ (P(0) = \text{probability of zero accidents} \end{cases}$$

or, using Equation (H.7),

(H.9)
$$F^* = \frac{f^* - \left(\frac{b}{1+b}\right)^a \left(\frac{a}{1+b}\right)}{1 - \left(\frac{b}{1+b}\right)^a}$$

In the above equations f* and F* are used instead of f and F as a reminder that these quantities are not corrected for the missing accidents, but are calculated for the data base at hand. Thus, $f^* = \frac{f}{r}$, as noted in Section H.1 (see also Subsection 2.3.1). If one is given f* and F* for each crossing, one can determine $a(f^*, F^*)$ and $b(f^*, F^*)$ by solving equations (H.8) and (H.9).

It then follows, from Bayesian analysis, that if n accidents are observed in T years with a crossing whose characteristics yield the values f* and F*, then the expected number of accidents in any future year is given by $\overline{\phi}*(f^*,F^*,n,T)$:

(H.10)
$$\overline{\phi}^* = \frac{a}{b+T} + \frac{n}{b+T}$$

Note that from equation (H.10), with T=0 and n=0, one gets the value f^* which is the expected accident frequency conditioned on no accident history. With T=1 and n>1, and with some algebraic operations, one can also derive equation (H.1) from equation (H.10). This is a reassuring check. [Equation (H.1) is, as noted, based on fewer assumptions than is equation (H.10).] As f^* and F^* have been calculated on a data base with a fraction $1 - \frac{1}{r}$ of the accidents missing (unlinked), then this is corrected for simply by dividing b in equation (H.10) by r (from Subsection 2.3.1, r=1.41):

(H.10a)
$$\overline{\phi} = \frac{a}{\frac{b}{r}+T} + \frac{n}{\frac{b}{r}+T}$$

Equation (H.10a) now solves the problem of finding a hazard index with full dependence on accident history (for any period of time). It can even be used to rank together crossings for which the accident history is known for different numbers of years (since $\overline{\phi}$ is an absolute hazard index and provides an expected frequency of accidents). However, it could be expected to work best if all crossings had an accident history over the same time period.

Now the procedure for calculating a and b as functions of f^* and F^* will be recapitulated and expanded on. To solve equations (H.8) and (H.9), they are transformed as follows:

$$(H.11)$$
 a = bf

(H.12)
$$b_2 = \frac{\log(1 + \frac{f^*}{(1+b_1)(F^*-f^*)})}{f^*\log(1+\frac{1}{b_1})}$$

If a reasonable approximation for b is substituted for b_1 on the right-hand side of equation (H.12), then this equation yields a better approximation as b_2 . This new approximation can be put back in equation (H.12) as b_1 , resulting in a still better approximation as b_2 . This process can be iterated several times. It happens, however, that when $f^* \ge 1$, a fairly good initial approximation for b is available [equation (H.13)]:

(H.13) b
$$\approx \frac{1}{(F^* - f^*)(1 + \frac{F^*}{2})}$$

This is surprisingly accurate when F^* and f^* are not too large, as can be seen by substituting in equation H.12. For example, if $f^{*=0.33}$ and $F^{*=0.69}$, the error in equation (H.13) is less than

4 percent. In general, equation (H.13) will provide an initial estimate to be used as b_1 in equation (H.12). The resulting b_2 , if changed but slightly, is to be used for b_1 or else b_2 is substituted for b_1 to iterate the process.

It has now been shown how to calculate a and b (analytically) as functions of F* and f*. There are not empirical data involved in such a calculation, since a and b are determined as implicit functions of F* and f* by equations (H.8) and (H.9), or by equations (H.11) and (H.12).

It has already been shown how F^* could be found as a simple function of f^* [e.g., equations (H.3) and (H.4)]. F^* could be found as a function of other crossing characteristics as well (e.g., as its own function of volume variables, etc.). However, for simplicity, if F^* is found only as a function of f^* , then a and b become functions of f^* alone. Thus, when this analysis is carried out in full, a table will probably be given in the following form:

% Crossings	f*	F(f*)	a(f*)	b(f*)
0.5	•	•	•	3
1.0	•	•	o	
1.5		•	٠	•
2.0	•	•	•	•
•	•	•	•	٠
	3	•	•	•
•	•		•	•
49.0	•	٠		•
49.5	٠	•	٥	•
50.0	•	•	•	•

Some preliminary results on F* as a linear function of f* have been given in this Appendix [equation (H.4) and Table H-1]. Much more precise statements can be made about the functional dependence of F* on f*. The rough guide is just an indication; further regressions should yield fairly accurate functional relationships.

Having a and b as functions of f*, equation (H.10) now yields $\overline{\phi}$ as a function of f*, T, and n, thus completing the solution for the hazard index which depends on accident history.

When the detailed computations and the results for a and b as functions of f are given, an evaluation in the manner of the EOCs given for the ordinary hazard indexes should be included. (It will be informative to see how much accident history may enhance a hazard index.)

It is instructive to illustrate the various formulas of this section by a simple example. Suppose one is given a crossbuck crossing for which accident history has been collected for five years, during which two accidents occurred (T=5, n=2). Suppose, also, that $h_{TSC} = -0.186$ (from formula B.5.1.3). Then, from the formula in Subsection 2.3.1, f=0.230, and

$$f^* = \frac{f}{r} = \frac{0.230}{1.41} = 0.163$$

From equation (H.4),

 $F^* = 0.044 + (1.78)(0.163) = 0.334$

Then, from equation (H.13), $b \approx 5.01$. Substituting $b_1 = 5.01$ in equation (H.12), one gets $b_2 = 4.96$. One now has available the following values for the parameters appearing in equation (H.10a): a = (0.163)(4.96) = 0.8085

b = 4.96

n = 2

r = 1.41

T = 5

Substituting in equation (H.10a), one obtains

$$\overline{\phi} = \frac{0.8085}{4.96} + \frac{2}{4.96} = 0.33$$

Thus, the expected accident frequency per year, given two accidents in five years (of observation), is 0.33. This value lies between the unconditional frequency, 0.23, and the observed frequency, 0.4, i.e., $\frac{2}{5}$, as it must.

Another look at equation (H.10a) provides an additional approach in assessing the significance of $\overline{\phi}$. Letting $T_o = \frac{b}{r}$, and recalling the relation f=rf* = r. $\frac{a}{b}$, one applies simple algebra to arrive at the following transformation of equation (H.10a).



(H.14)
$$\overline{\phi} = f \cdot \frac{T_o}{T + T_o} + \frac{n}{T} \cdot \frac{T}{T + T_o}$$

Equation (H.14) shows that $\overline{\phi}$ is a weighted average of f (the unconditional expected accident frequency) and $\frac{n}{T}$ (the observed accident frequency). When T = T_o, the two terms have equal weight:

 $\frac{T_0}{T+T_0} = \frac{T}{T+T_0} = \frac{1}{2}$. One can estimate T_0 from the formulas of this subsection. In the example given above, $T_0 = \frac{4.96}{1.41} = 3.5$ years. General orders of magnitude may be noted: for crossings more hazardous than the least hazardous 70 percent and less hazardous than the most hazardous 5 percent, i.e., between the 5th and 30th percentiles from the top, one has:

Crossbucks : $5.4 < T_0 < 10.4$ years Flashing lights: $2.4 < T_0 < 4.5$ years

In both cases, $0.4 \leq fT_0 \leq 0.8$ in this limited range. In general, as f gets larger, T₀ gets smaller. Apparently, as f gets larger, fT_0 gets larger.
H.3 DEFINITIONS:

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E(Y|h) is "the expected value of the number of accidents Y, given that the hazard index has the value h". $E(Y|Y \ge y,h)$ is "the expected number of accidents in a year, Y, given the hazard index and that Y is greater than or equal to y." $Pr(Y \ge y+1|h, Y \ge y)$ means "the probability that the number of accidents in a year equals or exceeds y+1, given that it equals or exceeds y and given h."

H.4 DERIVATION OF LEMMA

Let ε be a very small fraction ($\varepsilon < 1$). Y_{ε} = number of accidents in ε of a year; $Y_{1-\varepsilon}$ = number of accidents in remaining 1- ε of that year; Y_{1} = number of accidents in a certain year; Y_{2} = number of accidents in another year Y_{2} = $Y_{\varepsilon} + Y_{1-\varepsilon}$ (All at crossing with hazard index value h) There

Then:

$$\Pr(Y_{\varepsilon} = 1 | Y_{1-\varepsilon} \ge y, h) = \frac{\Pr(Y_{\varepsilon} = 1, |Y_{1-\varepsilon} \ge y|h)}{\Pr(Y_{1-\varepsilon} \ge y|h)}$$
$$= \sum_{x=y+1}^{\infty} \frac{\Pr(Y_{\varepsilon} = 1, |Y_{1-\varepsilon} = x-1|Y-x)\Pr(Y-x|h)}{\Pr(Y_{1-\varepsilon} \ge y|h)}$$
$$\approx \sum_{x=y+1}^{\infty} \varepsilon |x| \Pr(Y=x|h) / \Pr(Y \ge y|h).$$

But,

 $\Pr(\mathbf{Y}_{\varepsilon} = 1 \mid \mathbf{Y}_{1-\varepsilon} \geq \mathbf{y}, \mathbf{h}) \approx \varepsilon E(\mathbf{Y}_{2} \mid \mathbf{Y}_{1} \geq \mathbf{y}, \mathbf{h})$

Therefore, letting $\varepsilon \rightarrow 0$ E(Y₂|Y₁ \geq y,h) = E(Y|Y \geq y + 1,h) Pr(Y \geq y+1|h)/Pr(Y \geq y|h)

= $E(Y | Y \ge y = 1, h)$ $Pr(Y \ge y+1 | Y \ge y, h)$

This formula is exact, and based only on very weak assumptions of homogeneity in time.

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